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## Minimum Wages and Firm Value

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IZA DP No. 9914

## **Minimum Wages and Firm Value**

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April 2016

# Minimum Wages and Firm Value

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## **ABSTRACT**

### **Minimum Wages and Firm Value\***

How does the value of a firm change in response to a minimum wage hike? The evidence we have to date is not well-suited to answer this question, principally because events that have been studied are not completely unknown to the stock market or have uncertainty associated with them. This paper exploits the announcement of a sizeable change in the minimum wage in the UK that was both totally unanticipated and free of uncertainty. The stock market response of employers of minimum wage workers is examined in an event study setting, looking at minute-by-minute changes surrounding the announcement and at cumulative abnormal returns on a daily basis before and after the announcement. The analysis uncovers significant falls in the stock market value of low wage firms. The size of the fall in value is compared to the fall in profitability in response to the wage cost shock that will be induced by the announcement and is seen to be of a comparable magnitude.

JEL Classification: J23, L25

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“We strongly support the National Minimum Wage and want to see further real-terms increases in the next Parliament. We accept the recommendations of the Low Pay Commission that the National Minimum Wage should rise to £6.70 this autumn, on course for a Minimum Wage that will be over £8 by the end of the decade.”

Conservative Party Manifesto, April 14<sup>th</sup> 2015.

“I am today introducing a new national living wage. We will set it to reach £9 an hour by 2020. The new national living wage will be compulsory. Working people aged 25 and over will receive it. It will start next April at the rate of £7.20. The Low Pay Commission will recommend future rises that achieve the Government’s objective of reaching 60 percent of median earnings by 2020.”

Budget Speech, July 8<sup>th</sup> 2015.

“I’ve talked to several chief executives and been surprised by the impact on their profits. In one [big] company, it would wipe out all of their profits”

Paul Drechsler, President, Confederation of British Industry, September 2015.

## **1. Introduction**

Ever since minimum wage floors were first introduced to labour markets around the world, a perennial research question of high relevance for labour market policy has been how firms adjust to wage cost increases brought about by increases in the minimum wage. The first port of call for much of the literature has been to study the labour demand response of firms, and this has at various points in time generated research and policy controversies about what minimum wage increases do to employment or unemployment.<sup>1</sup> As evidence of employment losses from minimum

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<sup>1</sup> Surveying the (mostly US time series) literature that studied data up to the late 1970s, Brown, Gilroy and Kohen (1982) concluded that minimum wages reduced teenage employment, but had less effect on adult employment. The next phase of research were the more micro-based studies of the 1990s, spearheaded by the Card and Krueger (1994) paper on fast food restaurants and Card and Krueger’s (1995) book, which both questioned the earlier findings and found no evidence of disemployment effects. The introduction of the UK National Minimum Wage in April 1999 also generated a number of pieces of research failing to find significant employment effects (see, inter alia, Machin, Manning and Rahman, 2003, Stewart, 2002, 2004, or Dolton, Bondibene and Stops, 2015). In the US there has recently been another revival of research on minimum wage effects, with some focus on geography and differences across state borders. As before this is proving controversial on whether or not minimum wages reduce employment (see, inter alia, Dube, Lester and Reich, 2010 and 2016, Meer and West, 2015, or Neumark, Salas and Wascher, 2014).

wage hikes has proven elusive in a number of settings, a smaller body of research has placed a focus on looking for other margins of firm adjustment. Whilst there are many such margins, which may differ for firms operating in different sectors, some areas considered in research have been the scope to pass minimum wages on in terms of higher prices (e.g. Aaronson and French, 2007; Lemos, 2008), whether firms cut back on wage increases for other higher paid workers and so reduce wage inequality (e.g. Lee, 1999; DiNardo, Fortin and Lemieux, 1996; Dickens and Manning, 2004) and on whether minimum wage increases squeeze firm profitability (e.g. Draca, Machin and Van Reenen, 2011).

In this paper we study minimum wage effects on firm profitability in a different way from the direct before/after analysis of changes in accounting profitability that result from minimum wage changes. Instead we study the impact of the announcement of a minimum wage change on the stock market value of firms. This approach has been adopted in a couple of studies before, first by Card and Krueger (1995) who studied twenty three events in the US between 1987 and 1989 that eventually led to minimum wage increases in 1990 and 1991, and by Pacheco and Nalker (2006) who undertook an event study looking at changes in shareholder values following a significant reform to the youth minimum wage in New Zealand. Neither of these studies delivers very clear results, primarily because the nature of the ‘events’ that were examined do not allow for a clean event-study. Such a study would require a completely unexpected minimum wage change that had no uncertainty attached to its introduction. To take one example from Card and Krueger, on June 13, 1989 President Bush vetoed a minimum wage rise. The stock market reaction to this event shows no significant effect on firm value. But as Card and Krueger note, it is difficult to know whether this veto conveyed new information to the market, since the White House had promised to veto the bill

when it was first passed by the House three months earlier. And if it did contain new information, how did it change the probability of a minimum wage change?

The event studied in this paper is able to significantly improve upon such concerns. A completely unanticipated and sizable change in the UK minimum wage system was announced in the newly elected Conservative government's emergency budget that was called after its election to power in May 2015. As the quote from the budget speech of July 8 2015 reproduced at the top of the first page of this paper testifies, the Chancellor George Osborne announced that the UK government would introduce a new National Living Wage (NLW) of £7.20 per hour for workers aged 25 and over. Not only was this announcement from a right of centre government that has traditionally been against minimum wages, it was also totally unexpected, with other government ministers and the body which advises the government on minimum wages (the Low Pay Commission) not knowing that it would occur.<sup>2</sup> Thus the major advantage of our study compared to the other stock market reaction research in this area is that the announcement we study was completely unanticipated.

The event study approach has been very widely used in finance (see Kothari and Warner, 2008, or MacKinlay, 1997), but it has also been used by labour economists in several settings, most notably as a means of studying the effects of unions on firm performance.<sup>3</sup> The seminal paper studying union effects on stock market values was by Ruback and Zimmerman (1984) whose event study of union representation elections uncovered evidence of a negative effect of union wins on the equity value of US firms. Subsequently, Bronars and Deere (1990) uncovered similar effects while Lee and Mas

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<sup>2</sup> The BBC News reported that day as follows: "In a surprise announcement at the end of his speech, he said workers aged over 25 would be entitled to a "national living wage" from next April, to soften the impact of in-work benefit cuts." <http://www.bbc.co.uk/news/uk-politics-33437115>

<sup>3</sup> Another example of event studies in labour economics is Farber and Hallock's (2009) analysis of the stock market value effects of job loss announcements.

(2012) used a much larger sample and wider time window to find a longer run impact of unions on firm value. These union papers usefully inform the research design we adopt in our event study, but in the different setting of minimum wage changes.

In this paper, the differential stock market response of employers of minimum wage workers is compared to that of employers of higher wage workers in an event study setting looking at minute-by-minute changes surrounding the announcement and at cumulative abnormal returns in the days before and after the announcement. There is evidence of significant falls in the stock market value of low wage firms. Within a day of the budget, firm values were around 1.2 percent lower for the employers of minimum wage labour and ended up stabilizing around 2 to 3 percent lower after five days. Much of this adjustment was rapid and had happened within a couple of days. The size of the fall in firm value is compared to the fall in profitability in response to the wage cost shock that will be induced by the announcement and is seen to be of a comparable magnitude.

The remainder of the paper is structured as follows. In Section 2, the relationship between minimum wages and profitability is first considered, followed by a discussion of the system of minimum wages that operates in the UK and then how this has altered following the introduction of the new National Living Wage. Section 3 describes the data and event study methodology. The results are discussed in Section 4, and Section 5 offers some concluding remarks.

## **2. Minimum Wages, Profitability and the New National Living Wage**

### *Minimum Wages and Profitability*

For a competitive profit maximizing firm employing  $L$  workers at wage rate  $W$ , using other factors of production at price  $R$  and selling its output at price  $P$ , profits are



maximized at  $\Pi(W, R, P)$ . For such a firm, the derivative of the profit function with respect to the wage rate is  $\partial\Pi/\partial W = -L(W, R, P)$ , the negative of the demand for labour and the second derivative is  $\partial^2\Pi/\partial W^2 = -\partial L/\partial W$ . The introduction of a minimum wage at a level  $M$ , above the prevailing wage  $W$ , reduces firm profits by  $\Delta\Pi = \Pi(W, R, P) - \Pi(M, R, P)$ .

Following Ashenfelter and Smith (1979), this can be approximated as:

$$\Delta\Pi \cong -L\Delta W + \frac{1}{2} \frac{\partial L}{\partial W} (\Delta W)^2 \quad (1)$$

where  $\Delta W = M - W$ .

The first term on the right-hand side of (1) is the wage bill effect on profits ( $-L\Delta W$ ) and the second can be thought of as the labour demand ( $\frac{1}{2} \frac{\partial L}{\partial W} (\Delta W)^2$ ) effect on profits. Equation (1) can be rewritten as:

$$\Delta\Pi = -WL \left( \frac{\Delta W}{W} + \frac{\eta}{2} \left( \frac{\Delta W}{W} \right)^2 \right) \quad (2)$$

where  $\eta = \frac{W}{L} \frac{\partial L}{\partial W} < 0$  is the elasticity of labour demand.

Equation (2) offers one way of thinking about the profit response of a firm to a minimum wage hike. If there is “no behavioural response”, which in this setting means no impact on labour demand, the second order effect in (2),  $(\frac{\eta}{2} \left( \frac{\Delta W}{W} \right)^2)$ , is zero. The fall of profits that would result from the imposition of a minimum wage  $M$  is then equal to the proportionate change in the wage multiplied by the wage bill.

If adjustment can occur, then the labour demand effect in the second term is non-zero. This can offset the profit loss to the extent that firms can substitute away from low-wage workers into other factors (e.g. capital). One interesting question is the speed at which such adjustment could occur. In the event study setting of the empirical work in this paper, this is particularly interesting when one attempts to gauge the size of the

profit reduction from a change in market value (which is the present discounted value of firm profits).<sup>4</sup>

Other models focus on the particular mode of adjustment to the wage cost shock from a higher minimum wage. In Aaronson and French (2007), for example, firms have constant marginal costs and thus a horizontal supply curve and so, irrespective of the nature of competition in the product market, the minimum wage increase is passed on entirely to consumers in the form of higher prices. In putty-clay models (like Aaronson, French and Sorkin, 2015) the same is true.<sup>5</sup> Some other alternatives to adjusting prices that allow the firm to negate the negative impact on profits include compression of the internal wage structure, efficiency wages/productivity improvements and conventional labour demand responses (see, for example, Hirsch, Kaufman and Zelenska, 2015).

Equation (2) also usefully illustrates the inverse relationship between a firm's initial wage and the profit change. It shows that, the lower the initial wage, then the greater the fall in profits associated with the imposition of a minimum wage. This logic underpins what we do in our empirical work where we focus on the stock market response of employers of minimum wage workers in an event study setting. The means by which we define firms that employ minimum wage workers is considered in Section 3 of the paper where we also describe the data that we use.

### *Minimum Wage Setting in the UK*

A National Minimum Wage (NMW) was introduced to the UK labour market in April 1999. Prior to that, minimum wages did not play a role in wage determination

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<sup>4</sup> See also Abowd's (1989) classic study of union wage increases and firm performance. Abowd estimates a version of equation (2) examining the effects of unanticipated increases in the wage bill (which he defines as union wealth) on the present discounted value of profits as reflected in changes in stock market values (or shareholder wealth). Interestingly, the findings are unable to reject the simple no behavioural response model where the second order effect is zero.

<sup>5</sup> See Sorkin (2015) who emphasises the distinction between modes of adjustment in the short and long run. Clemens and Wither (2014) also document that minimum wages take some time to be reflected in employment levels

as the system that used to operate (the Wages Councils who set sectoral minima in low wage sectors, only covering about 10 percent of UK workers) had been abolished by the Conservative government in 1993 (Dickens, Machin and Manning, 1999).

The rate at which the National Minimum Wage was introduced was determined by a body set up by the Labour government which was elected in May 1997. The Low Pay Commission (LPC) was instituted as an advisory body by the National Minimum Wage Act of 1998. The LPC has nine Commissioners, three of which are from business, three of which are from employee representation groups, and three are members who are independent from the social partners. These last three are the Chair and two academics who are experts in labour economics and industrial relations.

The LPC remit is set by the government each spring, with a main focus on coming up with evidence-based recommendations on the main adult minimum wage rate and the associated age-specific minima. The LPC assesses evidence from a wide range of sources (e.g. academic research, site visits, an annual consultation procedure with oral evidence taken from a wide range of stakeholders). It then makes recommendations in a report submitted to government in February, to which the government responds on acceptance or rejection of the recommendations, and then if accepted (as the main adult rate always has been since introduction) the NMW is updated on October 1<sup>st</sup>.<sup>6</sup>

Figure 1 shows the rates from 1999 to 2015. In April 1999 the National Minimum Wage was first introduced at a rate of £3.60 per hour for workers aged over 21, together with a youth development rate at £3.00 per hour for 18-21 year olds. Through time more rates have been introduced: in 2004 a minimum wage for 16-17

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<sup>6</sup> For more detail on the functioning of the LPC see Brown (2002), Butcher (2012) and Metcalf (1999, 2002).

year olds was introduced, and an apprentice minimum wage in 2010. Also in 2010 the adult rate was extended to 21 year olds, so that by 2015 there were four rates in place: the adult minimum (now for those aged 21 and over) which had reached £6.70 by October 2015; the youth development rate for 18-20 year olds of £5.30; the rate for 16-17 year olds of £3.87; and the apprentice rate of £3.30.<sup>7</sup>

In many quarters, the operation of the LPC has been deemed a success. The Institute of Government's 2010 polling of 159 members of the Political Studies Association rated the NMW as the most successful government policy of the previous thirty years.<sup>8</sup> This reflects the evidence-based approach leading to little in the way of employment losses from the NMW, and the independence of the LPC in being able to make its deliberations largely free from political intervention.

### *The New National Living Wage*

After being in a coalition government with the Liberal Democrats as the UK government in power between May 2010 and 2015, the Conservative party was elected outright in the May 2015 election. It called an emergency budget for July 8 2015 and in this budget the Chancellor George Osborne made the completely unexpected announcement of introducing a new National Living Wage that would raise the NMW for age 25 year olds and older workers by 50 pence from April 2016. The main reason for this was to offset the tax credit cuts that the Chancellor introduced in the budget in his strong programme of austerity cuts. In the Appendix we outline all the measures announced in the budget statement. There were a number of changes to personal

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<sup>7</sup> The UK government has almost always accepted the LPC's recommendations on rates. This has always been true for the recommendations adult, development and 16-17 year old rates. The recommendation on apprentice rates has twice been met with a government instituted change: first in 2011 when the LPC recommended a freeze of the rate but government intervened to increase it by 5 pence; then more markedly in 2015 when the LPC recommended raising the rate by 7 pence from £2.73 to £2.80 as the business secretary pushed the rate a further 50 pence up to £3.30.

<sup>8</sup> See: <http://www.bbc.co.uk/news/uk-politics-11896971>.

taxation and to the taxation of dividends and corporate profits. The only changes other than the NLW that were focused on particular businesses were a rise in insurance premium tax and reform of the taxation of banks – neither of which are relevant for our study. This gives us confidence that no other announcement in the budget had a significant impact on low-wage firms in particular.

From a political economy perspective, this is a striking and radical intervention. It comes from a political party that has traditionally been strongly against minimum wages and, indeed, which strongly opposed the introduction of the NMW in the first place. The very poor real wage performance of most workers in the UK labour market since 2008 (when median real wages have fallen by around 10 percent, but such falls are seen across most of the wage distribution as well) has altered this standpoint to some extent.<sup>9</sup> It is true that all of the main UK political parties (including the Conservatives) have recognised that minimum wages are both popular amongst the general public and that they can play a role in raising wages (and by association living standards).<sup>10</sup>

The surprise of the budget announcement and the size of the wage shock is what we exploit in our event study of the impact on the stock market value of firms.<sup>11</sup> The 50 pence NLW supplement on adult minimum wages came as a complete surprise to the market. The Chancellor also introduced a target level for the adult minimum wage of £9 per hour to be reached by 2020. This was also news to the stock market, as the

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<sup>9</sup> See Blundell, Crawford and Jin (2014) and Gregg, Machin and Fernandez-Salgado (2014) for more detail on the nature of real wage falls in the UK labour market.

<sup>10</sup> On the popularity of the UK minimum wage, a 2014 Gallup poll reported that 66 percent of those polled were in favour of increasing the minimum wage.

<sup>11</sup> In addition to being unanticipated, we noted earlier that a key additional requirement for a successful event study, particularly when trying to evaluate the size of any estimated effect, is that there be no uncertainty over the introduction of the new minimum wage. The announcement considered in this paper satisfies that requirement because the 1998 National Minimum Wage Act gives government ministers the power to set the minimum wage without reference to the LPC. The government confirmed in writing to the LPC on budget day that such an order would be made.

Conservative Manifesto before the May 2015 election (quoted at the start of the paper) was clear in the aspiration of reaching £8 per hour by that date.

The introduction of the NLW also alters the role of the LPC in its future deliberations, as it now has a target to work to. In practical terms the government intervention has also effectively introduced a new age band into the structure of minimum wages that operate for low wage workers in the UK. This is shown in Table 2 where the new structure of minimum wage rates that will apply from 2016 is compared to those of 2014 and 2015.

The new NLW also offers an ‘experiment’ not made possible by previous increases in the minimum wage coming from the LPC recommendations. In due course, it will be interesting to study the employment and other economic effects of this big increase, of 10.8 percent compared to the £6.50 rate at the time of the announcement, or of 7.5 percent compared to the already accepted LPC rate of £6.70 that was made effective after the budget in October 2015. By 2020, the targeted minimum wage of £9 is 12.5 percent higher than the £8 level that had previously been suggested.

As a result of the minimum wage changes and 2020 target, the number of workers in the UK labour market who are covered by the minimum is expected to rise significantly. Figure 2 shows actual coverage from 1999 to 2014 and expected coverage from 2015 onwards (defined as the number of workers paid at or below the relevant minimum and up to 5 pence above). There is a significant increase resulting from the change. In 2015 before the change the number of covered workers had risen gradually to reach 1.6 million by 2015. Afterwards, due to the new NLW of April 2016 and the 2020 target, there is a sharp increase, straight away jumping to over 2.5 million, and reaching 3.8 million by 2020.

### **3. Data and Modelling Approach**

#### *Data*

The equity price data come from Datastream and Bloomberg. The principal sample frame is made up of the constituents of the FTSE All-Share Index. This index comprises eligible companies listed on the London Stock Exchange main market that pass screenings for liquidity and investability. The index captures 98 percent of the UK's market capitalization. There are 643 constituents of the index, with a mean (median) market capitalization of £3.1bn (£599mn).

We exclude from our analysis all investment trusts and private equity funds, giving a final sample of 442 firms. We have daily prices, total returns and volume. We also extract daily data on market capitalization, dividend yield and the price-book ratio. Trade-level data for the announcement date are taken from Bloomberg. These data contain the price, volume and exact time of every trade during the official trading day (8am to 4.30pm). We use this data to compute the volume-weighted price for each stock by minute of the trading day.

#### *Treatment Firms*

To estimate the effect of the increase of the minimum wage on firm market values, we need to construct a sample of firms that are exposed to the treatment. Our principal results follow the approach of Draca, Machin and Van Reenen (2011) by using accounting information on the average wage of the firm to sort firms. The annual accounts report the total wage and salary costs for the firm and the average number of employees. This gives us a measure of the average annual wage per employee. Since the new minimum wage is set at £7.20, a worker who is employed full-year full-time would earn £14,976.

We therefore identify the quoted firms in our sample who are expected to be affected by the minimum wage announcement on the basis of their average wage per employee being less than £15,000 per year. The strength of this identification approach depends on the extent to which minimum wage workers are concentrated in firms at the lower end of the wage distribution. Unfortunately, firms are not required to report any information on the distribution of wages within the firm, only the average wage.

To therefore assess the usefulness of the approach we adopt, we study a different data source looking at the segregation of wages across firms in the UK using the 2013 cross-section of the Annual Survey of Hours and Earnings (ASHE) and the Annual Business Survey (ABS). These are matched worker-firm level data that allows us to look at within-firm wage distributions and explore the association between average wages and the intensity of low-wage workers. We have a sample of 62,594 workers employed in the private sector who can be matched to firm-level data (7,795 firms) that includes the average wage per employee.

We follow the Low Pay Commission procedure of defining a minimum wage worker as any worker who receives an hourly wage (excluding overtime) that is up to five pence above the minimum wage effective at the time of the survey (April 2013). Overall, 6.2 percent of our sample are minimum wage workers. This is very similar to the 7 percent figure reported for all private sector workers in the UK in the 2014 Low Pay Commission report (page 22). Our final restriction is that firms must employ at least 100 workers and have an average wage of at least £5,000. These two sample restrictions reduce our sample to 59,535 workers (and 5,853 firms), but the share of minimum-wage jobs remains at 6.2 percent.

In Figure 3, we plot the proportion of minimum-wage workers in a firm against the firm's average annual wage for those firms with an average wage below £30,000



(40,046 workers in 3,684 firms). We split the sample into vintiles of the average wage, with a sample of 186 firms in each bin. There is clearly a strong decline in the proportion of minimum-wage workers as the average wage rises. Only for those bins to the left of the £15,000 cutoff are at least 10 percent of the workforce on the minimum wage. On average, 21 percent of workers in firms with an average wage of less than £15,000 are minimum-wage workers. Alternatively, one can note that 75 percent of minimum-wage workers work in firms that have average annual wages of less than £15,000.

These figures underestimate the impact of the minimum wage on our equity sample, because they reflect the average share of minimum-wage workers across all industries. Our equity sample however is heavily focused on two particular industries (see the list of firms and their sectors in the Appendix), namely non-specialised Retail Trade (47190) and Beverage Serving activities (56302). As Table 2 shows, if we further restrict the ASHE/ABS sample to only include these two industries, we have a sample of 2,924 workers (and 84 firms). In these industries, 30.3 percent of workers in firms with an average wage of less than £15,000 are minimum-wage workers.

We also require that the firm has a majority of its employees based in the UK. Some of the low average wage firms are predominantly operating in low-wage economies, but have chosen to list on the London stock exchange. Clearly these firms are not affected by the UK minimum wage. Our final sample of NMW firms comprises the twenty companies that are listed (grouped by sector) in the Appendix. Together, these 20 firms employ over 600,000 workers.

### *Event Study Methodology*

We follow the by now standard approach in the finance literature to estimate the effect of the minimum wage change on a firm's equity value. We compute the

abnormal return as the difference between a stock's actual return and the expected return. For firm  $i$  at time  $t$ , the abnormal return is simply:

$$AR_{it} = r_{it} - E[r_{it}|X_t]$$

where  $r_{it}$  is the actual realized return and  $E[r_{it}|X_t]$  is the expected return, with  $X_t$  denoting the information set at time  $t$ .

We consider a number of alternative specifications for  $E[r_{it}|X_t]$ . Perhaps the most common approach is to use the Capital Asset Pricing Model (CAPM) to estimate the sensitivity of the  $i$ th firm's return to a market index (i.e.  $X_t$  is the market return) and use the predicted values as an estimate of the expected return. This approach is adopted by Ruback and Zimmerman (1984), Card and Krueger (1995) and Lee and Mas (2012). To implement this, we estimate a daily return model of the form:

$$r_{it} = \alpha_i + \beta_i r_{mt} + \epsilon_{it}$$

where  $r_{mt}$  is the return on the equal-weighted FTSE All Share index. The market model is estimated over a twelve-month period up to the 15<sup>th</sup> April 2015, which is 60 days prior to the minimum wage announcement. The abnormal return from the CAPM model is then simply:

$$AR_{it} = r_{it} - (\hat{\alpha}_i + \hat{\beta}_i r_{mt})$$

It is well-known however that the cross-section of stock returns can be predicted by more than simply the market return (Fama and French, 1992). We therefore also present results using a four-factor model for expected returns that includes the market return, a size return based on market capitalization, a value return based on dividend yield and a momentum return.<sup>12</sup>

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<sup>12</sup> For the four-factor model, we estimate each factor return by allocating all stocks in the FTSE All-Share index into (a) a large, medium and small-cap portfolio based on the 30<sup>th</sup> and 70<sup>th</sup> percentile rank on the previous trading day, (b) a high, medium and low dividend yield portfolio based on the 30<sup>th</sup> and 70<sup>th</sup> percentile rank on the previous trading day and (c) a high, medium and low momentum portfolio based on the 30<sup>th</sup> and 70<sup>th</sup> percentile rank on the previous trading day of returns over the period 126-21 days prior to the ranking. We then generate daily factor returns by taking the difference between the two

One further factor that could affect our results is that firms with sizeable minimum wage exposure tend to be heavily distributed in certain industries that employ more low wage workers such as retail, hotels, restaurants and bars. Evidence suggests that stock returns have an important industry component (Moskowitz and Grinblatt, 1999; Fama and French, 1997) and if, by chance, these particular industries experienced abnormal returns relative to the market since the announcement date for reasons unrelated to the minimum wage, we would ascribe those returns to the minimum wage announcement. The expected return definition used above will not account for this. We therefore also construct two-digit SIC industry returns (excluding the minimum wage firms themselves) and use these as our measure of expected returns. So in our main results we consider estimates of abnormal returns from three alternative models: (1) the CAPM, (2) a four-factor model and (3) a 2-digit industry model.

#### **4. Results**

In this section we discuss the results of our analysis. We begin by presenting minute-by-minute evidence to demonstrate the strong reaction of our minimum wage sample to the exact announcement time of the minimum wage increase. We then examine the subsequent daily abnormal returns. Finally we present evidence from a regression model that suggests a significant ability of the market to isolate minimum wage firms from other, arguably similar, firms.

##### *Intra-day Announcement Effects*

The Chancellor of the Exchequer began the budget statement at 12:33 on July 8<sup>th</sup>. At 13:35, he announced the decision to raise the minimum wage, one hour and two

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extreme portfolios for each factor. See Dube, Kaplan and Naidu (2011) for another example of an event study using factor models.

minutes into the speech. He concluded the speech four minutes later, at 13:39. We can therefore exploit the intra-day price change in our sample to examine whether there was any difference in returns between the NMW firms and the non-NMW firms prior to the announcement time (13:35) and whether there was a subsequent divergence. We have minute-by-minute data on the trade price and volume traded of each stock. We present results for both the equal-weighted index of NMW firms and a value-weighted index that accounts for the very different trading volumes that are present in the intra-day data.

To motivate the analysis, Figure 4 shows minute-by-minute share price moves for three low wage firms from the market open on budget day to 24 hours after the NLW announcement. The three firms are a retail firm (Home Retail Group), a pub group (JD Wetherspoon) and a hospitality firm (The Restaurant Group). In all three cases, there is a marked dramatic drop in their share prices at the precise time that the minimum wage announcement was made. The initial drop within half an hour was of the order of 2-3 percent and was broadly sustained for the rest of the trading day. On market opening at 8:00 AM the next day, the pub JD Wetherspoon took another hit dropping a further 2 percent, presumably as the market had more time overnight to absorb and assess the information contained in the announcement.

Analysis of the whole sample of NMW firms is provided in Table 3. Panel A gives the minute-by-minute cumulative abnormal returns ( $CAR(X, Y)$ ) for the NMW stocks from the time of announcement ( $X = A$ ) over the subsequent  $Y$  minutes. In Panel B we report the pre-announcement returns. The abnormal returns are calculated for a market model (i.e. adjusting returns for the overall market return over the same period) and for the two-digit industry model. The first two columns of results use equal-weighting, whilst the final two columns are value-weighted. Figure 5 displays the

cumulative raw returns to the NMW and non-NMW stocks from the beginning of the trading day to 24-hours after the announcement, while Figure 6 shows the equal-weighted abnormal return for the NMW stocks. The 95 percent confidence interval around the mean abnormal return is also shown for the latter.

For the pre-announcement effects, there are two key points. First, there was essentially no trend in the market in the hours prior to the budget and no significant difference between the NMW and non-NMW stocks. Second, all of the budget announcements prior to the minimum wage announcement had very little effect on market prices and almost no effect on the relative performance of the two groups of firms. This is consistent with the evidence in Section 2 that other announcements in the budget were not directed at low-firm wage in particular.

From 13:35 onward, the picture is very different. There was a sharp fall in the price of NMW stocks. One hour on from the announcement, the NMW firms on average experienced a fall of as much as 69bp relative to the market on an equal-weighted basis, and 176bp on a value-weighted basis.<sup>13</sup> Although these effects weaken as the trading day finished, all observations on the average abnormal returns of the NMW firms are negative from the announcement time. Moreover, on market opening the day after budget day, the gap significantly widens again as Figures 5 and 6 make very clear, so that 24 hours after the NLW announcement the fall in cumulative abnormal returns is around 1.2 percent. Overall, we take this as strong evidence that the decline in NMW prices relative to non-NMW prices occurred as a direct result of the minimum wage announcement and was not as a consequence of other Budget changes (or indeed other events in the market).

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<sup>13</sup> Note that the smaller equal-weighted as compared to value-weighted returns reflect the fact that two very small firms (Game Digital and Card Factory) had price gains during the afternoon on very small volume.

Before we move on to consider the results on the evolution of daily cumulative returns, it is interesting to look at the intra-day pattern of abnormal returns for the NMW firms in the weeks before the announcement. We do this so as to both undertake a placebo-type experiment and to see whether the estimated abnormal returns in the 24-hours after the announcement were indeed significantly different from zero on the basis of the empirical distribution of abnormal returns for this portfolio of stocks. Figure 7 plots the equal-weighted abnormal returns for the NMW firms for every 24-hour period (measured from 13:35 to 13:34 on the subsequent day) from 1<sup>st</sup> May up to and including the budget day (47 observations). In the Figure, the budget day observation is the black bar, whilst the 46 pre-budget observations are denoted by gray bars. The difference between the black bar and the gray bars is striking and two points are clear. First, pre-budget abnormal returns are generally within the range of  $\pm 0.5\%$  for this portfolio over a 24-hour period (for 90% of observations). Second, the decline observed after the NLW announcement at  $-1.28\%$  was by far the largest abnormal return observed across all 47 observations.

#### *Cumulative Abnormal Returns*

Having demonstrated an intra-day response to the minimum wage announcement, we now turn to consider daily cumulative abnormal returns from the announcement day onward. We compute these equal-weighted cumulative abnormal returns from day  $X$  to day  $Y$  as  $CAR(X, Y) = \sum_{t=X}^{t=Y} AR_{it}$ . So for example denoting  $X=A$  as announcement time,  $CAR(A, 10)$  measures the cumulative abnormal return on the NMW stocks for the post-announcement part of Budget day and the first ten trading days from then. We use three expected return models: (1) the CAPM model, (2) a 4-factor model and (3) a two-digit industry-matched model.

Table 4 presents the key results for the mean CARs, whilst Figure 8 plots the CAR (using the industry-matched model for returns) with associated 95 percent confidence intervals for the 10 days either side of the announcement. Panel A of the table gives post-announcement returns for a one-, two-, three-, four-, five-, and ten-day horizon, while Panel B reports the pre-announcement returns for corresponding horizons in the pre-Budget period. We study daily returns for a 10-day window since the announcement is clear and public, so we would expect the market to re-price reasonably fast. Longer horizons are increasingly likely to bring in other events not related to the NLW that may shift abnormal returns differentially for treatment and control firms. We do however comment below on estimated CARs for up to three months after the announcement.

The decline over the announcement day and the following day,  $CAR(A, 1)$ , is in the range -1.4 to -1.7 percent, depending on the expected return model we adopt. Over the five-day period, the decline is between -2.1 percent and -3.0 percent. At this point, the decline seems to have stabilized, since there is at most an additional 20bp decline over the subsequent five days. This suggests that the market reacted quickly to the announcement, with around one-half to three-quarters of the adjustment occurring by the end of the first-day after the Budget. All the estimated negative cumulative returns are statistically significant at the 5 percent level or better.

Panel B shows weak evidence that the NMW firms outperformed over the pre-announcement period, though none of the CARs are statistically significant. This suggests that if anything the results may slightly underestimate the effect of the minimum wage announcement, since we know that stocks tend to have momentum over these holding-period horizons (see Jegadeesh and Titman, 1993).

As already noted, other events may start to impact the estimated CARs as the time window is extended further. That said, when we did widen the window and estimate a twenty- and sixty-day CAR, we obtained estimates (and associated standard errors) of -2.919 (1.884) and -3.968 (3.755) for the industry-matched model. The significantly reduced precision of these estimates is not at all surprising, but the magnitude does show that the significant fall experienced within the first few days of the announcement appears to be permanent.

#### *Possible Small Sample Issues*

As with other event studies, small sample issues may lead to possible biases. The first concerns the relatively small number of firms in the treatment group and the volatility of returns for some of the less traded firms. To consider this, we have therefore also estimated CARs for the median firm at each horizon and by trimming the two highest and lowest CARs within the sample of twenty firms in the same way as the mean regressions in Table 4. The results for the five-day horizon are reported in Table A1 in the Appendix. All these alternative approaches continue to show significantly negative effects – though the point estimates are somewhat less negative in general. We also find that 15 of the 20 firms in the NMW portfolio have a negative abnormal return five-days after the announcement (regardless of expected return model). Using 18 months of pre-announcement data, we find that for this sample of 20 firms, such a large number of negative returns across the portfolio lies outside the 95 percent confidence interval.

A second issue is that the distribution of cumulative abnormal returns may not be normally distributed as assumed. We therefore also undertook a non-parametric analysis where we estimated CARs in an estimation window that predates the NLW announcement. As one example of this, we estimated five-day industry-matched CARs



based on daily data from September 2013 through April 2015 and considered their distributions. The 5<sup>th</sup> percentile CAR was -1.59 and the 95<sup>th</sup> percentile CAR was +1.73; the 1<sup>st</sup> percentile was -2.67 and the 99<sup>th</sup> percentile +2.34. We can then ask, given how this specific portfolio of stocks moved over the last two years, how likely is it that over a five-day window they would decline by the size of our estimated CAR of -3.047. This has an empirical  $p$ -value of 0.002. In general, these empirical confidence intervals broadly matched the standard confidence intervals used in Table 4.

### *Regression Based Analysis*

In a final analysis of the returns data, we examine the cross-section of all post-announcement abnormal returns for the full sample of 442 firms. We are interested primarily here in whether there is evidence to suggest that the market was able to accurately identify the firms most at risk from the minimum wage rise and whether our identification approach to minimum-wage firms is plausible given the market response.

Table 5 reports some regression results for this cross-section of returns for the five-days following the announcement (results that prove qualitatively similar for ten-day returns are given in Appendix Table A2). We use the CAPM model to estimate expected returns, but the results are robust to using any of the alternative return models considered before. In column (1) we simply report the coefficient on a dummy if the firm was in our NMW sample. The coefficient of -2.131 on this dummy is of course equal to the mean abnormal return after five-days reported in the first column of Table 4. In column (2) we additionally add controls for 2-digit industry, market-capitalization size quintile and pre-announcement returns. These controls marginally increase the size of the estimated negative effect for the NMW firms to -2.331.

In column (3) we consider whether the negative cumulative returns for the NMW sample are merely a result of a more general decline for lower-wage firms. Since

we use the cutoff of less than £15,000 average wage to identify the NMW firms, we consider whether firms in the £15-20k, £20-25k and £25-30k average wage bracket experienced any similar pattern (with £30k+ being the omitted group). There is no evidence to support this idea, suggesting that the market focused closely on the lowest-wage firms.

In column (4) we divide the NMW sample into two equal-sized groups of 10 firms. One group, NMW High  $\pi$  Impact, are those firms in which the mechanical percentage reduction in pretax profits from a 4 percent rise in the wage bill (the estimated effect of the minimum wage hike for the average firm in our sample – see below) is above the median, and the second group, NMW Low  $\pi$  Impact, are those for which it is below the median. We would expect the cumulative abnormal returns to be more negative for the NMW High  $\pi$  Impact group, and this is exactly what we find – the magnitudes are quite different in numerical terms though, because of the inevitably small sample involved, the standard errors do not allow us to formally reject equality between the two groups.

Table 6 further refines the analysis, looking at a minimum wage exposure variable that we were able to match to our firms from the ASHE and ABS data described above. This measures the actual proportion of workers paid the National Minimum Wage in April 2013, and is defined for 275 of the original 442 firms (and 19,441 workers). We lose firms for two key reasons. First, some firms are listed on the London Stock Exchange but the holding company is not UK-domiciled. These firms do not appear in the ABS as the sampling frame is only UK-registered companies. We lose 99 companies as a result. Second, as the ASHE wage data is only a 1 percent sample, some firms do not have any workers recorded in the data. For each firm that we can match, we compute the minimum wage exposure (NMW Exposure) by calculating the

percentage of workers that had an hourly wage no more than 5p higher than the October 2012 minimum wage rate of £6.19. The advantage of this measure is that it captures the full range of minimum-wage exposure for all firms, rather than just focusing on 20 firms.

Columns (1) and (2) of Table 6 report the regression of 5-day abnormal returns after the national living wage announcement on this exposure variable (Table A3 in the Appendix reports results for 10-day returns). It should be noted that the results in Table 5 are replicated almost exactly if we re-estimated on the reduced sample of 275 firms. The exposure variable is significantly negative, so that firms with a higher proportion of minimum wage workers experience a larger decline in value following the announcement. Taking the coefficient in column (2), the results suggest that firms with a 10 percentage point higher fraction of the workforce employed on the minimum wage experienced a 62bp additional decline in stock price following the announcement.

We also define a dummy variable equal to 1 if the firm is one of the 20 firms with the largest estimated minimum wage exposure. Columns (3) and (4) show regression results for this dummy variable.<sup>14</sup> The estimates are very similar to those in Table 5, which suggests that the results we have established are robust to reasonable alternative definitions of the particular constituents of the portfolio of NMW stocks.

#### *Gauging the Impact on Profitability*

Finally, we turn to an analysis of the magnitude of the estimates that we have obtained. Consider the sample of firms that have an average wage of less than £15,000 in the matched worker-firm data discussed in Section 3. We noted above that for the two principal industries in our equity sample, 30 percent of workers in these firms are

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<sup>14</sup> This sample of 20 firms therefore differs somewhat from that used in the results in Table 5 which are based on the average wage in the firm (from the company annual report).

minimum-wage workers. To evaluate the impact of a minimum-wage hike for these firms, we start by noting that for the low-wage firms in these industries, minimum-wage workers account for 21 percent of the wage bill. This of course is lower than their share of employment since they are by definition the lowest paid workers in the firm. So the simple direct effect of increasing the minimum wage by 10 percent for these workers would generate a wage bill rise of 2.1 percent.

But this calculation ignores two additional effects. First, all workers currently above the minimum wage but who would fall below the new minimum wage, must have their wages raised to at least the new minimum. 27 percent of workers are in this group, and to raise their wages at least to the new minimum adds an additional 1.3 percent to the wage bill, giving an overall increase of 3.4 percent. Second, it is usually assumed that workers seek to protect their relativities following a minimum wage increase<sup>15</sup>. A simple method of capturing this is to assume that workers on the old minimum wage get the full 10 percent increase and that all workers within 20 percent of the old minimum receive some wage increase on a smoothly-sliding scale that maintains wage rank order, places everyone at least at the new minimum and tapers the minimum wage effect to vanish for all workers with wages already above 10 percent of the new minimum. This gives a total wage bill increase of 4.1 percent. We would argue that this suggests a fairly tight bound on the wage bill effect of the increase, for the average firm in our sample, to be between 3 and 4 percent.

How reasonable is this calculation? Perhaps the best evidence comes directly from one of the firms in our NMW sample. Next plc (a large clothes retailer) released

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<sup>15</sup> This indirect effect is often termed a spillover or ‘ripple’ effect. Dube, Galiano and Leonard (2015) use payroll data from an individual firm in the US to suggest that this ripple effect is observed for workers earning up to 15 percent above a newly implemented Federal minimum wage. For other work on spillovers based on individual-level microdata see Autor, Manning and Smith (2016) for the US or Dickens and Manning (2004) for the UK.

its half-yearly report in September 2015. They provided a detailed calculation of the impact of the NLW on their wage bill. By 2020, the firm estimated that the wage bill would be £27m higher as a result of paying the NLW (including the associated costs of maintaining relativities) on a total wage bill of £720m, or 3.8 percent.

Table 7 calibrates the impact of this scale of wage bill shock on firm profits. We use data from the last three (pre-NLW announcement) annual reports of the 20 firms in our equity sample. All these reports predate the minimum wage announcement. We normalize sales to be 100 and compute the average for each firm of the wage bill and pre-tax profits. The figures in column (1) report the baseline. On average, the wage bill is equal to 18.6 percent of total sales and the pre-tax profit equals 6.0 percent. This is a relatively low profit margin, though is common for firms in these sectors. We focus on pre-tax profits because we assume that any hit to the wage bill feeds through all the profit measures e.g. firms cannot for example alter their financing costs to offset the wage bill rise. With a real interest rate of 3 percent, a pre-tax profit of 6 gives a present value of 206.

Now consider a rise in the minimum wage of 10 percent (the average rise over the next four years) that raises the wage bill by 4 percent (column (2)). Assuming no offsetting effects, pre-tax profits fall by 12 percent. If this rise in the wage bill is permanent (i.e. no subsequent reduction in the minimum wage) and there are no offsetting effects, the present value of the firm drops by 12.4 percent to 180.5. Alternatively, suppose that the firm takes the hit in full for five years but then successfully generates fully offsetting effects elsewhere on the income statement. Then firm value declines by 1.6 percent.

We can also evaluate the size of the response in terms of variable exposure to the minimum wage. We reported in Table 6, that a firm with a 10 percent higher share

of minimum-wage workers experienced an additional estimated 62bp fall in value. In the final column of Table 7, we show that a firm employing 40 percent minimum wages workers (10 percentage points above the baseline), would experience a decline in present value of 17 percent. Again assuming full offset in the medium-run, the drop in firm value is predicted to be 2.3 percent, 70bp more than for the baseline firm. This very closely matches the estimate from our empirical analysis.

Whilst these numbers are inevitably somewhat back-of-the-envelope, they do suggest that the size of the effect we have estimated in our empirical work (a 2-3 percent decline) is consistent with the market believing that firms will be able to substantially offset the costs of the minimum wage, at least in the medium term. Crucially, our estimates are inconsistent with there being a large, permanent effect on firm profits where employers are unable to adjust to respond to the additional wage costs induced by the minimum wage increase.

How will firms achieve this offset in practice? Again, we can look at corporate reports issued since the announcement that have commented on the NLW introduction. At the time of writing, nine firms in our sample have released reports that discuss the NLW. The text of statements pertaining to the NLW are reported in section C of the Appendix. Consideration of the statements make clear two things. First of all, to varying degrees, they all expect to be able to significantly mitigate the effects of the rise in the wage bill. Second, the margin of adjustment seems highly unlikely to be employment. In fact, none of the nine say they anticipate employment reductions in response to the NLW.

Instead, the nine company statements focus on other margins of adjustment to the wage shock that the NLW will induce. Factors mentioned include increasing prices, raising productivity and increasing cost efficiency. Most mentioned is the idea that the

cost increases will be passed on in terms of higher prices (or termed inflationary pressure in some of the reports). For example, Next plc (whose half-yearly report we have referred to above) say that a 1 percent increase in product price over the next four years would completely offset their additional wage costs from the National Living Wage.

## **5. Conclusions**

Based on a stock market event study, the empirical research presented in this paper describes what happened to the market value of firms employing minimum wage workers when a completely unexpected announcement of raising the minimum wage occurred. The setting is an emergency budget that was called promptly after the new election of a right of centre government in the UK. The UK Chancellor of the Exchequer made a surprise announcement on budget day (July 8 2015) that he would introduce a new National Living Wage some 11 percent higher than the prevailing National Minimum Wage for workers aged 25 and above.

The impact of the announcement is studied both intra-day and for up to ten days either side of the announcement. Unlike the work on stock market responses to minimum wages in other settings, where the unanticipated nature and certainty of the announced rises are much less precise than in our setting, we find there to be a strong and lasting impact on stock market values. The size of this reduction in firm value resulting from the NLW announcement is compared to the fall in profitability in response to the wage cost shock that will be induced by the announcement and is seen to be of a comparable magnitude, assuming that firms can adjust over time. Thus the announcement of NLW introduction seems to have had the impact of significantly

reducing the expected profits of UK firms for when it is to be implemented in April 2016.



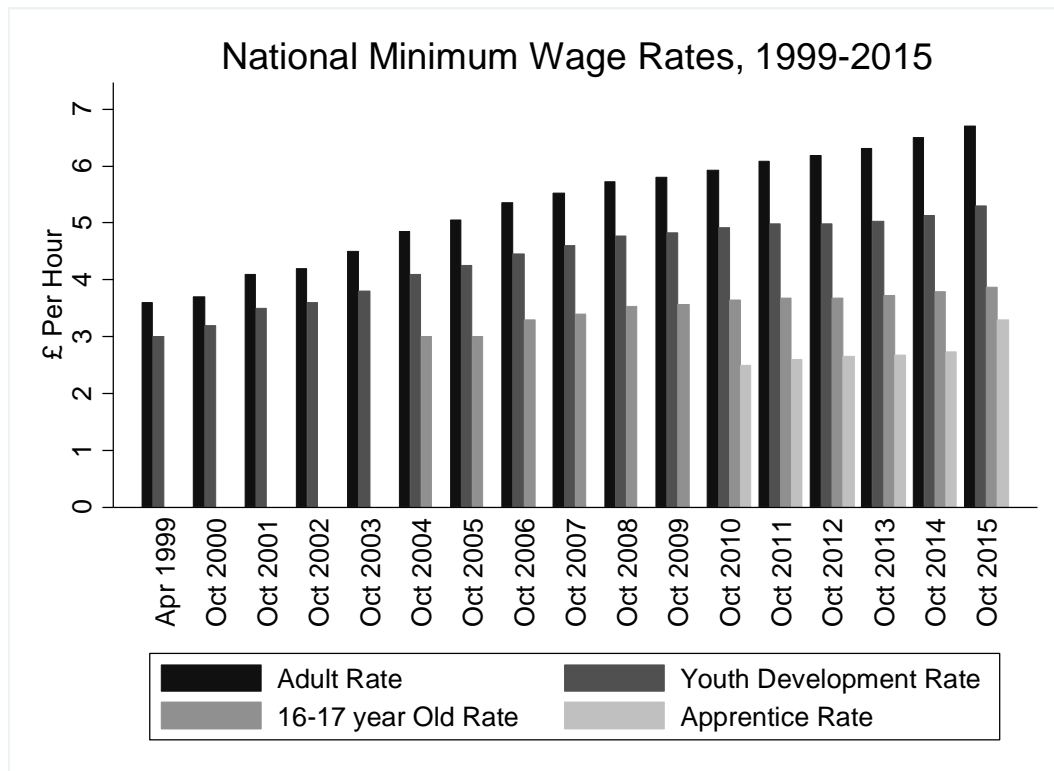
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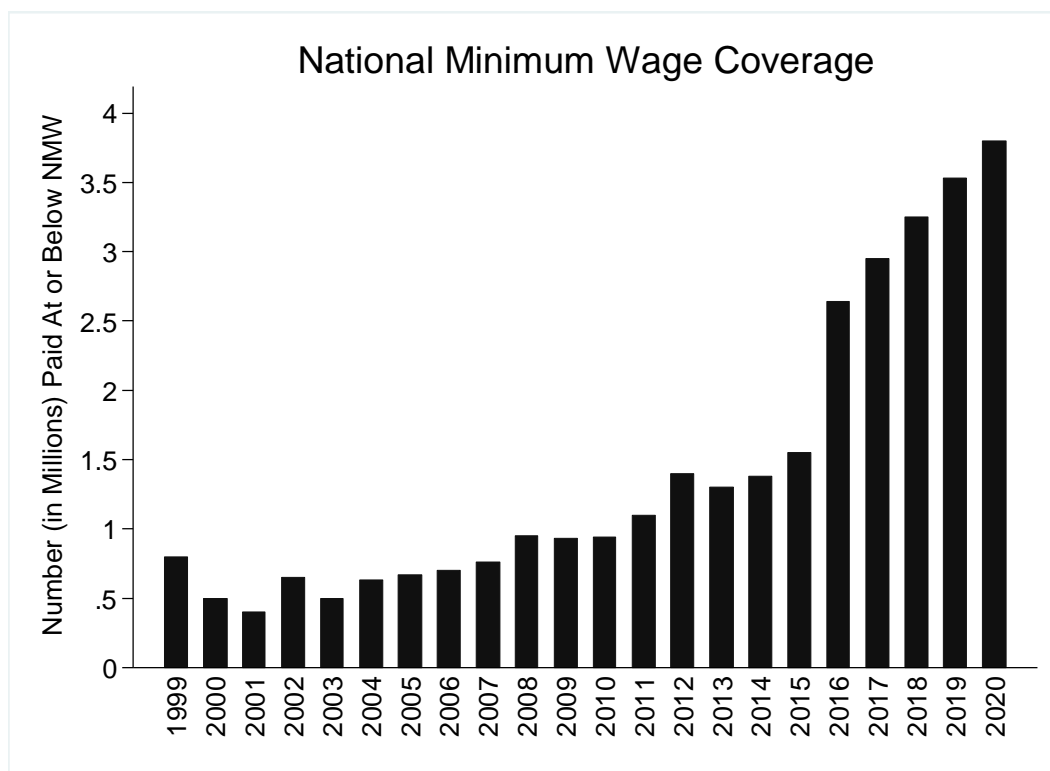
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**Figure 1:**  
**UK National Minimum Wages, 1999-2015**



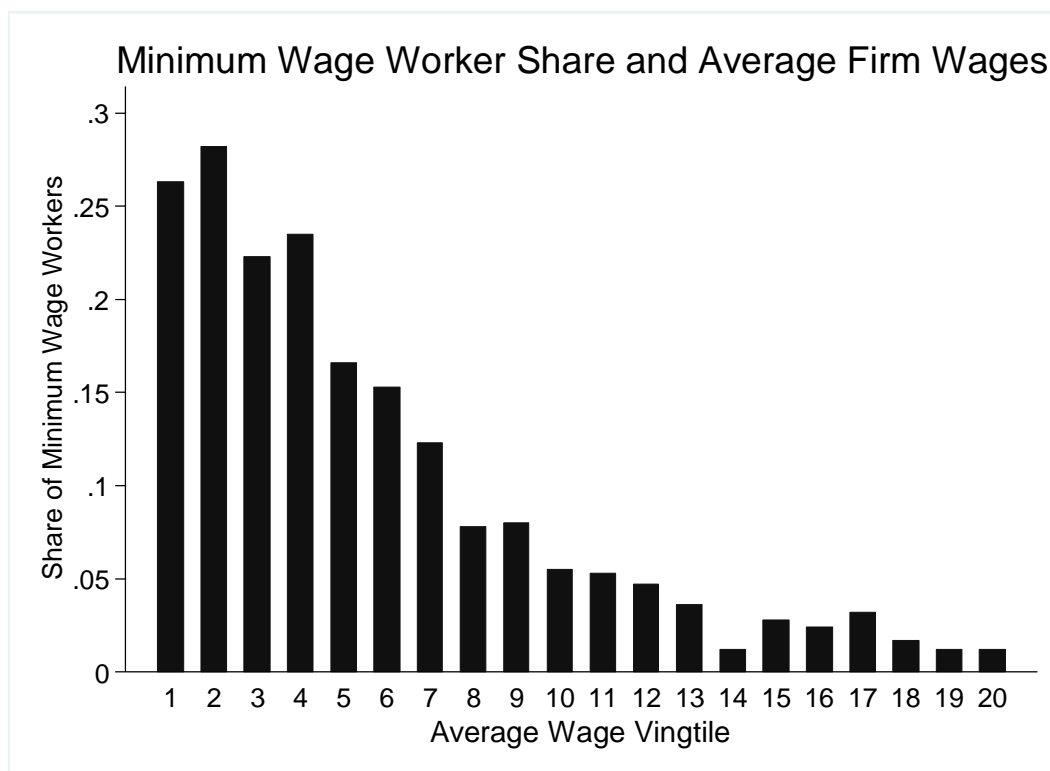
Notes: From Low Pay Commission annual reports.

**Figure 2:**  
**Actual and Estimated Minimum Wage Coverage, 1999-2020**



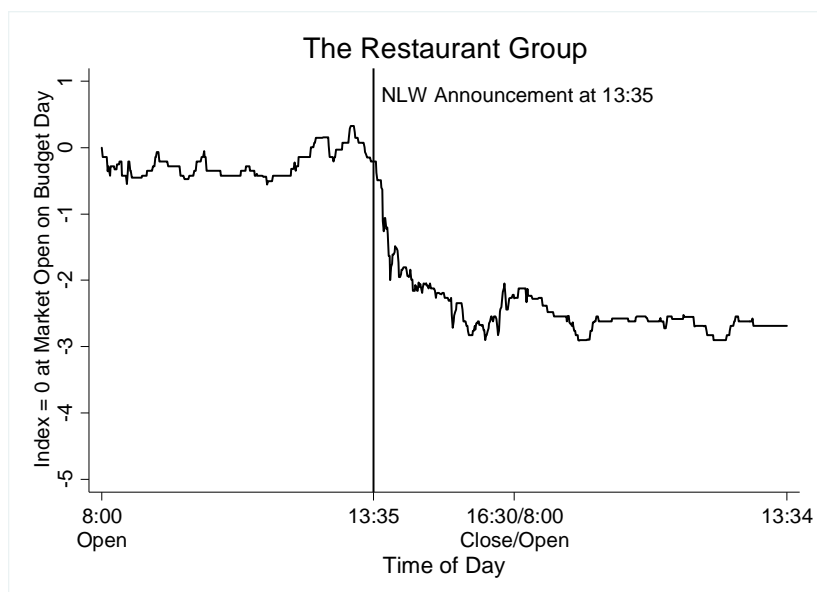
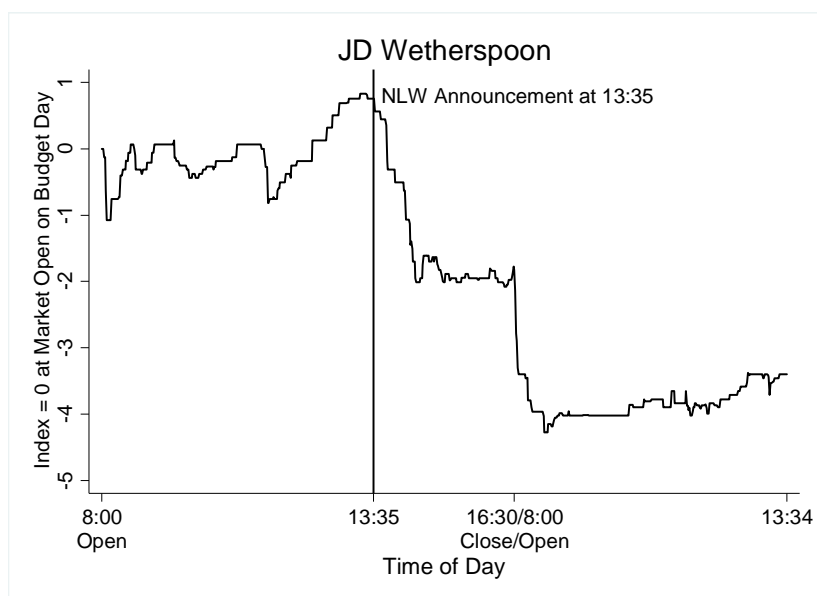
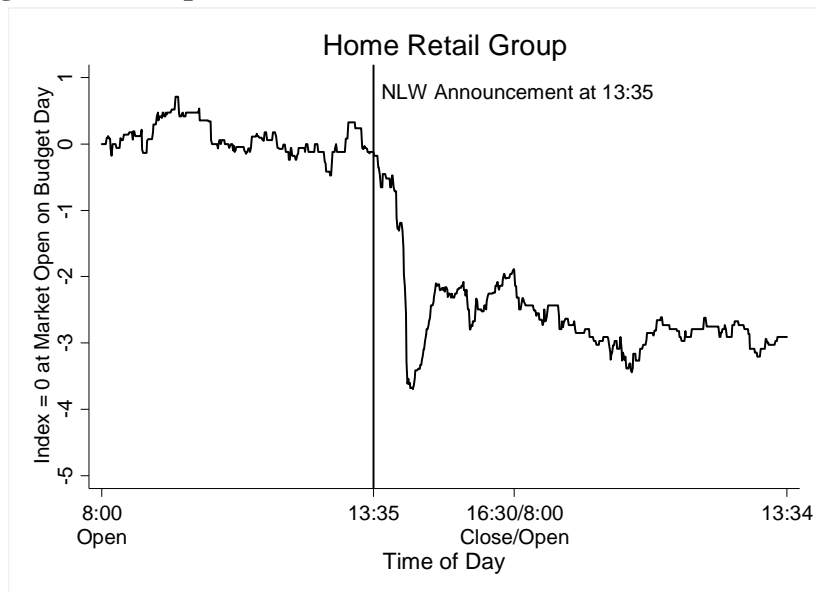
Notes: Low Pay Commission calculations from Annual Survey of Hours and Earnings (ASHE) from 1999 to 2014 and from 2015 onwards using wage forecasts from ASHE 2014. Paid at or below the minimum is based upon the Low Pay Commission procedure of defining a minimum-wage worker as any worker who receives an hourly wage (excluding overtime) that is up to five pence above the minimum wage effective at the time of the survey

**Figure 3:**  
**Minimum Wage Shares and Firm Average Wages**

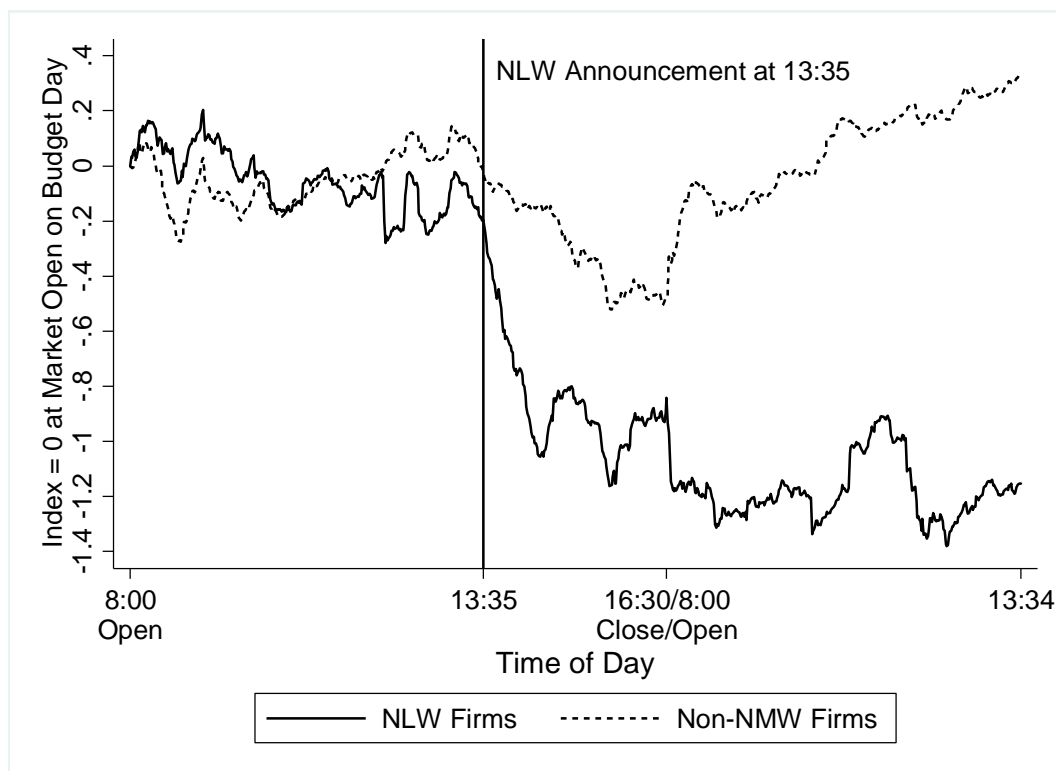


Notes: The y-axis shows the proportion of minimum-wage workers in the firm. The x-axis shows the average annual wage in the firm divided into vingtile bins of 5 percentiles from lowest (left) to highest (right) – a total of 20 bins for annual wages from £5,000 to £30,000. Derived from matched worker-firm data (40,046 workers in 3,684 firms) from the 2013 Annual Survey of Hours and Earnings (ASHE) and the Annual Business Survey (ABS).

**Figure 4: Examples of Share Price Movements of NMW Firms**



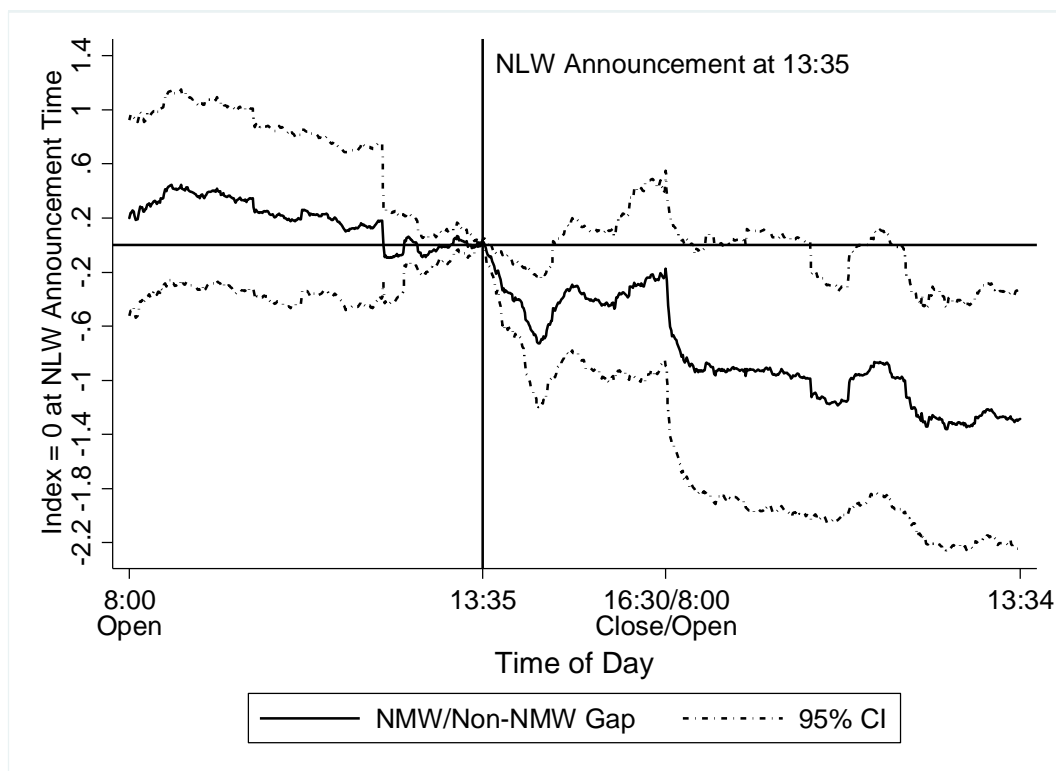
**Figure 5:**  
**Returns On Budget Day and 24 Hours After The NLW Announcement**



Notes: Based on 442 FTSE All-Share Index quoted firms, comprising 20 NMW firms and 422 Non-NMW firms. The 20 NMW firms are listed in the Appendix.

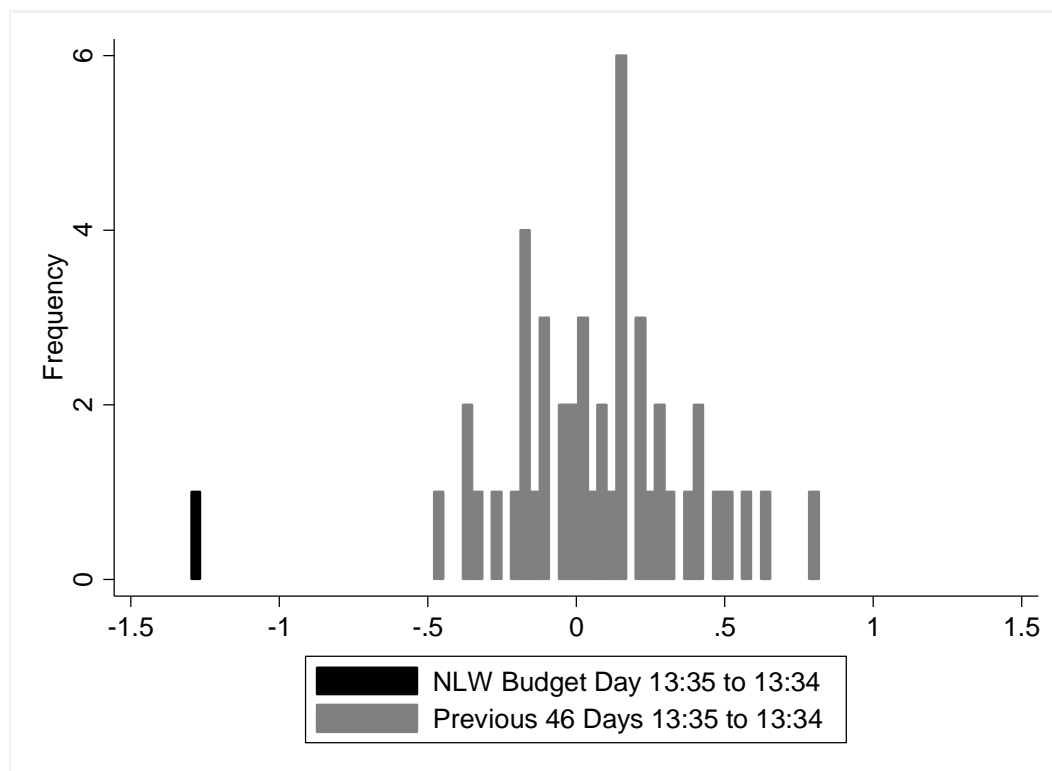


**Figure 6:**  
**Gap in Returns On Budget Day and 24 Hours After The NLW Announcement**



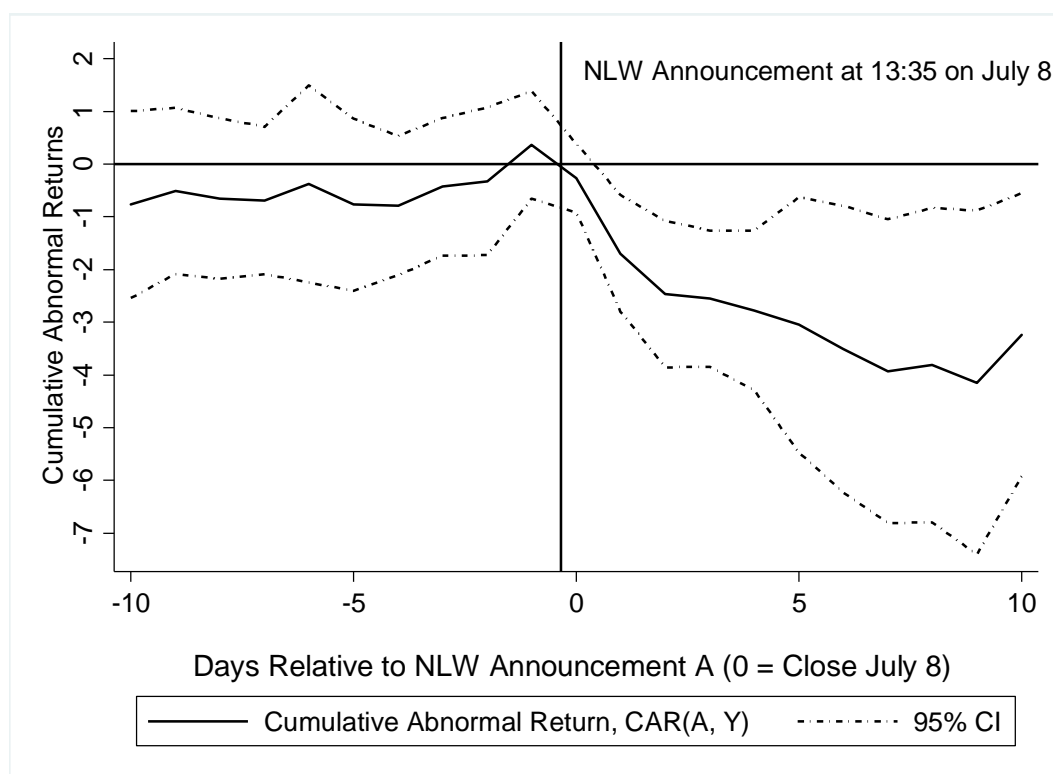
Notes: Based on 442 FTSE All-Share Index quoted firms, comprising 20 NMW firms and 422 Non-NMW firms. The 20 NMW firms are listed in the Appendix.

**Figure 7:**  
**Frequency Distribution of Abnormal Returns From May 1 to Budget Day**



Notes: Based on 442 FTSE All-Share Index quoted firms, comprising 20 NMW firms and 422 Non-NMW firms. The 20 NMW firms are listed in the Appendix.

**Figure 8:**  
**Daily Cumulative Abnormal Returns**



Notes: Based on 442 FTSE All-Share Index quoted firms, comprising 20 NMW firms and 422 Non-NMW firms. The 20 NMW firms are listed in the Appendix. Cumulative abnormal returns for the NMW firms are measured using a 2-digit industry return index as the expected return measure. The announcement date (Day 0) was budget day on July 8 2015, with the announcement A occurring at 13:35, two-thirds of the way through the trading day.

**TABLE 1. AGE VARIATIONS IN MINIMUM WAGES, 2014-2016**

Uprating Date	October 2014	October 2015	April 2016
Decision Process and Date	Government Accepted Low Pay Commission Recommendations in March 2014	Government Accepted Low Pay Commission Recommendations in March 2015	Chancellor Introduced National Living Wage For 25+ Workers in July 2015 Budget
Adult NMW, Age 21+	6.50	6.70	
Adult NMW, Age 21-24			6.70
Adult NLW, Age 25+			7.20
Youth NMW, Age 18-20	5.13	5.30	5.30
Youth NMW, Age 16-17	3.79	3.87	3.87

Notes: From Low Pay Commission annual reports and July 2015 budget.

**TABLE 2. MINIMUM WAGE SHARES AND FIRM AVERAGE WAGES**

	All Firms			Firms With Average Wage < £15,000		
	Number of Firms	Mean Wage	Share of Minimum Wage Workers	Number of Firms	Mean Wage	Share of Minimum Wage Workers
All	3684	16943	0.096	1259	11878	0.210
NMW Firms	84	14046	0.247	58	11736	0.303
Non-NMW Firms	3600	17170	0.093	1201	11894	0.206

Notes: Derived from matched worker-firm data (40,046 workers in 3,684 firms) from the 2013 Annual Survey of Hours and Earnings (ASHE) and the Annual Business Survey (ABS). The NMW firms are those in the two particular industries that dominate our equity sample, non-specialised Retail Trade and Beverage Serving activities.

**TABLE 3. ESTIMATES OF INTRA-DAY CUMULATIVE ABNORMAL RETURNS**

	Equal-Weighted		Value-Weighted	
	(1) Market Model	(2) Industry- Matched Model	(3) Market Model	(4) Industry- Matched Model
<b>Panel A: NMW Firms – Post-Announcement</b>				
CAR(A,20)	-0.342** (0.131)	-0.271* (0.125)	-1.419** (0.551)	-1.246** (0.537)
CAR(A,60)	-0.694** (0.238)	-0.638** (0.242)	-1.759** (0.681)	-1.564** (0.662)
CAR(A,120)	-0.469 (0.288)	-0.575 (0.312)	-1.495** (0.447)	-1.408** (0.425)
CAR(A, Market Close)	-0.280 (0.323)	-0.262 (0.334)	-0.996* (0.495)	-0.815 (0.491)
CAR(A, 24 Hours)	-1.300** (0.491)	-1.420** (0.519)	-2.019** (0.196)	-1.997** (0.209)
<b>Panel B: NMW Firms – Pre-Announcement</b>				
CAR(-20,A)	-0.040 (0.045)	-0.121* (0.054)	0.034 (0.038)	-0.080 (0.040)
CAR(Budget Start, A)	-0.023 (0.089)	-0.054 (0.093)	0.149 (0.086)	0.086 (0.075)
CAR(Market Open, A)	-0.183 (0.368)	0.038 (0.363)	-0.164 (0.159)	0.298 (0.139)

Notes: CAR(A, Y) denotes the cumulative abnormal return from announcement t time A (13:35) to minute Y relative to the announcement time. There are 20 firms in the NMW sample. Panel A reports results for the post-announcement period and Panel B reports results for the pre-announcement period. The market closed 175 minutes after the announcement, it opened 334 minutes before the announcement and the budget began 62 minutes before the announcement. The cumulative abnormal return for each firm is equal-weighted in columns (1) and (2) and weighted by their share of the total value of all trades in NMW stocks over the relevant period in columns (3) and (4). \*\* denotes significance at 1 percent level; \* denotes significance at 5 percent level.

**TABLE 4. ESTIMATES OF MEAN DAILY CUMULATIVE ABNORMAL RETURNS**

	(1) CAPM Model	(2) 4-Factor Model	(3) Industry-Matched Model
<b>Panel A: NMW Firms – Post-Announcement</b>			
CAR(A,1)	-1.428** (0.523)	-1.516** (0.542)	-1.693** (0.564)
CAR(A,2)	-1.799** (0.608)	-1.967** (0.620)	-2.465** (0.709)
CAR(A,3)	-1.661** (0.617)	-1.823** (0.612)	-2.553** (0.660)
CAR(A,4)	-1.533* (0.656)	-1.639* (0.656)	-2.774** (0.773)
CAR(A,5)	-2.131* (0.917)	-2.231* (0.910)	-3.047** (1.235)
CAR(A,10)	-2.268* (1.128)	-2.464* (1.165)	-3.241** (1.370)
<b>Panel B: NMW Firms – Pre-Announcement</b>			
CAR(-1,A)	-0.327 (0.512)	-0.357 (0.533)	-0.368 (0.520)
CAR(-5,A)	0.500 (0.800)	0.402 (0.805)	0.767 (0.832)
CAR(-10,A)	1.647 (0.934)	1.523 (0.963)	0.768 (0.904)

Notes: CAR(A, Y) denotes the cumulative abnormal return from announcement time on July 8<sup>th</sup> (A) to close on day Y, where Y is relative to 8<sup>th</sup> July. There are 20 firms in the NMW sample. Panel A reports results for the post-announcement period and Panel B reports results for the pre-announcement period. \*\* denotes significance at 1 percent level; \* denotes significance at 5 percent level.

**TABLE 5. REGRESSION ESTIMATES OF FIVE-DAY ABNORMAL RETURNS**

	(1)	(2)	(3)	(4)
NMW	-2.131* (0.894)	-2.331* (0.943)	-2.266* (1.032)	
NMW High $\pi$ Impact				-3.318* (1.607)
NMW Low $\pi$ Impact				-1.071 (1.044)
Average Wage £15-20k			-0.196 (0.909)	-0.146 (0.912)
Average Wage £20-25k			0.262 (0.568)	0.295 (0.565)
Average Wage £25-30k			0.291 (0.899)	0.297 (0.902)
Sample Size	442	442	442	442
2-Digit Industry	N	Y	Y	Y
5-Day Prior Return	N	Y	Y	Y
Size Quintiles	N	Y	Y	Y

Notes: The dependent variable is the five-day cumulative abnormal return,  $CAR(A,5)$ , using the CAPM model. NMW are our sample of 20 firms with average wage less than £15,000. NMW High  $\pi$  Impact are the 10 NMW firms for which a 4 percent rise in the wage bill would generate the largest percentage decline in pre-tax profits. NMW Low  $\pi$  Impact are the other ten firms in the NMW sample. Robust standard errors in parentheses. We also control for foreign exposure with a dummy equal to one if the majority of firm employment (or sales) is outside the UK. \*\* denotes significance at 1 percent level; \* denotes significance at 5 percent level.

**TABLE 6. REGRESSION ESTIMATES OF FIVE-DAY ABNORMAL RETURN  
USING ACTUAL EXPOSURE**

	(1)	(2)	(3)	(4)
NMW Exposure	-5.020* (2.422)	-6.208* (2.964)		
NMW Sample			-2.123* (0.920)	-2.185* (1.041)
Sample Size	275	275	275	275
2-Digit Industry	N	Y	N	Y
5-Day Prior Return	N	Y	N	Y
Size Quintiles	N	Y	N	Y

Notes: The dependent variable is the five-day excess return using the CAPM model. NMW Exposure is the percentage of workers in the firm that earn within 5p on the National Minimum Wage in the ASHE sample for the firm. NMW Sample is a dummy variable equal to 1 if the firm is one of the 20 firms with the largest value for NMW Exposure and where the measure of exposure is based on at least 30 observations in ASHE. We also control for foreign exposure with a dummy equal to one if the majority of firm employment (or sales) is outside the UK. Robust standard errors in parentheses.

\*\* denotes significance at 1 percent level; \* denotes significance at 5 percent level.



**TABLE 7. THE NATIONAL LIVING WAGE EFFECT ON FIRM PROFITS**

	Baseline Firm	10 Percent Increase in Minimum Wage (30 Percent Minimum Wage Workers)	10 Percent Increase in Minimum Wage (40 Percent Minimum Wage Workers)
Sales	100	100	100
Wage Bill	18.6	19.3	19.6
Other Costs	53.9	53.9	53.9
Gross Profit	27.5	26.8	26.5
Operating Profit	8.7	8.0	7.7
Pre-Tax Profit	6.0	5.3	5.0
Present Value of Pre-Tax Profits	206	180.5	170.9
Percent Decline in Firm Value (Permanent)		-12.4	-17.0
Decline in Firm Value (Hit to 2020, Thereafter No Effect)		-1.6	-2.3

Notes: Column 2 assumes a 3 percent real interest rate, 4 percent rise in wage bill (resulting from a 10 percent minimum wage increase, assuming 30 percent of workers are NMW workers). Column 3 assumes a 5.5 percent rise in the wage bill (resulting from a 10 percent minimum wage increase, assuming 40 percent of workers are NMW workers). If, rather a 10 percent increase, the minimum wage increase is graduated as a path of 7 percent, 9 percent, 11 percent and 12 percent increases to get to the £9.00 per hour target in 2020 the calculation turns out to be almost identical.

## Appendix

This Appendix gives details on: the named NMW firms we study; the 2015 budget measures; statements about the NLW made by companies in our sample in their company accounts; tables containing additional results.

### A. List of National Minimum Wage Firms By Sector

#### Bars and Restaurants

Greene King plc  
J D Wetherspoon plc  
Marston's plc  
Mitchells & Butlers plc  
The Restaurant Group plc

#### Retail

*Apparel*  
Next plc

*Broadline*  
B&M European Value Retail  
Debenhams plc  
Marks and Spencer Group plc  
Home Retail Group plc

*Food*  
McColl's Retail Group plc  
WM Morrison Supermarkets plc  
Greggs plc

*Home Improvement*  
Dunelm Group plc

*Specialty*  
Card Factory plc  
Game Digital plc  
Poundland Group plc  
WH Smith plc

#### Services

*Business Support*  
Mitie Group plc

*Recreational*  
Cineworld Group plc

## **B. Budget 2015 Measures**

In addition to the National Living Wage announcement, the budget contained the following measures (HM Treasury, 2015):

1. A rise in the tax-free personal allowance from £10,600 (2015/16) to £11,000 (2016/17).
2. Raising the threshold for higher-rate marginal income tax from £42,385 (2015/16) to £43,000 (2016/17).
3. Removing the family home from inheritance tax liability.
4. Reducing the tax-free amount of pension contributions for those earning more than £150,000.
5. Eliminating the ability of individuals to permanently claim non-domicile status for tax purposes.
6. Reforming the welfare system by (a) freezing many benefits in nominal terms for four years, (b) reducing the cap on household total benefit claims to £20,000 per year, (c) requiring 18-21 year olds claiming benefit to move on to an apprenticeship or start work within 6 months of claiming.
7. A real terms increase in the Defence budget for the remainder of the Parliament.
8. Reforming dividend tax credit to favour those with smaller dividend income.
9. The main rate of corporation tax will be cut to 19% in 2017 and 18% in 2020.
10. Annual investment allowance confirmed at £200,000 per year.
11. A rise in employment allowance (from £2,000 to £3,000), which provides an offset to employer national insurance contributions.
12. Insurance premium tax rises from 6% to 9.5%.
13. Restricting tax relief for wealthier landlords.
14. Reforming the taxation of banks with a gradual move toward taxing profits rather than balance-sheet size.
15. A levy of 0.5% (from 2018) on all firms paybill to fund 3 million new apprenticeships.
16. 30 hours of free childcare for 3 and 4 year olds from 2017 – increased from 15 hours.
17. Student maintenance grants to be replaced with loans.
18. Reform of the road tax system.
19. Public sector pay increases limited to 1% per annum for the next four years.
20. A crackdown on tax evasion and non-compliance.

### C. Firm Responses to NLW

Up to April 2016, the following nine companies in the sample we study have published half-year or full-year company accounts, including statements about the National Living Wage:

#### 1). Next plc

Half-year results, 10 September 2015

“NEXT’s total UK wages, at circa £600m, represent 15% of our sales. Of these, about £100m relates to those who would qualify for the NLW. Our adult starter rate will already be £7.04, so the increase to £7.20 in 2016 is manageable. We estimate the cost of implementing the NLW next year will be £2m.

Looking beyond 2016, we estimate that the cost of the further rises in the NLW to 2020, *over and above the rate of general wage inflation*, will be £27m per annum. Of this, £11m relates to the wages of those who will be paid the NLW. The remaining £16m is the knock-on effect of maintaining wage differentials for supervisors, junior managers and other more skilled or demanding roles within the business (such as specialist call centre work).

The £27m additional NLW cost is not immaterial but, in the context of NEXT’s wider cost base, is not transformative. We estimate that we would need to increase prices by around 1% to compensate for this cost which, taken over four years, is unlikely to have a material effect on the trading performance of the business. It should be noted that this is probably a pessimistic view of the required price rise, as we have assumed no improvements in productivity. In reality, we hope to be able to compensate for some wage inflation through increased productivity measures throughout the business.”

#### 2). Debenhams plc

Full-year results, 22 October 2015.

Cash impact £3m in FY2016, £8m in FY2017 (total wage bill = £345m). Both estimates include the cost of maintaining pay differentials.

“We anticipate that in the context of our continuing investment in systems and infrastructure and our focus on operational effectiveness, productivity gains will mitigate the majority of this impact on our P&L in these years”

#### 3). Marks and Spencer Group plc

Half-year results, 4 November 2015

“The new NLW rate will be £7.20 per hour effective from 1 April 2016, and as our permanent customer assistant rate is already £7.41, we do not expect any short term impact on our staffing costs”.

4). Home Retail Group plc

Half-year results, 21 October 2015

“It is anticipated that, without any actions to mitigate the increases in wages attributable to the new NLW, the associated FY17 group cost increase would be around £15m.” (total wage bill = £581m). “The Group will look at all opportunities to mitigate an element of this NLW cost increase through its existing operational excellence cost efficiency programmes.”

5). WM Morrison Supermarkets plc

Full-year results, 10 March 2016.

“In September we announced an increase in hourly pay for in-store colleagues, to £8.20 from a previous minimum of £6.83. We are recognizing the contribution of our excellent and dedicated colleagues, who are fundamental to the Morrisons turnaround. As announced, the extra investment will be £40m, and more than 90,000 staff will benefit across all age brackets. The new rate is £1 per hour above the National Living Wage that is to be introduced from April 2016” (total wage bill = £1.728bn).

6). Greggs plc

Half-year results, 29 July 2015

“The proposed increases to the minimum wage are likely to drive inflationary pressure in the broader sector over the coming years. We have consistently paid rates of pay above this level, with our standard rate for hourly-paid shop staff at £7.11, currently nine per cent higher than the national minimum wage. We are assessing the medium-term impact of further increases on our business”

7). Poundland Group plc

Half-year results, 19 November 2015.

“Although Poundland is not a minimum wage employer, we need to maintain pay differentials when new rates are introduced. Without mitigation, the increase in wages attributable to the new NLW is estimated at £4.3m in the 2017 financial year”

“We are mitigating this additional cost. We have accelerated investment in the current year in productivity including the trial of hybrid self-scanning checkouts, new sales based ordering systems, shelf ready packaging, the installation of LED lighting and other operational measures. We will spend an additional £2m this year in these areas, as we seek to fully mitigate the additional costs of the NLW.”

8). WH Smith plc

Full-year results, 15 October 2015.

“In July, the government announced the introduction of the NLW from April 2016. We anticipate the impact of it in the 2015/16 financial year to be slightly over £1m for the Group, with around two-thirds of the cost impacting the High Street business. While identifying the impact is complex, we have calculated that without any actions to offset the impact, the additional annual cost across the Group would be around £2-3m, or around 0.5% of the Group’s total cost base each year. As with all inflationary pressures, we will look to offset these increases through our existing pipeline of initiatives, as well as identifying further opportunities for improved productivity and efficiency.” (total wage bill = £189m).

9). Mitie Group plc

Half-year results, 23 November 2015

“We are positive about this move [introduction of NLW], which ensures that those of our people who are affected, are better rewarded and feel more motivated to do the jobs they do. It will also improve retention rates across our business.”

“The areas of our business that are impacted by this change are single service contracts in cleaning, a small area of security, catering and healthcare. One third of the contracts in these areas of the business come up for re-tender each year, and all work bid or re-bid since July 2015 onwards has factored in the NLW increases. Many contracts in the group have wage rates above this level”.

“We also believe that this is an important move for the outsourcing industry, as it affects all competitors equally and creates a level playing field in terms of bidding. We do expect outsourcing prices to increase in the future and will not be advocating job cuts in reaction to this change. However we will continue to work with clients to identify cost efficiencies in other areas.”

## D. Additional Tables

**TABLE A1. ALTERNATIVE ESTIMATES OF DAILY CUMULATIVE ABNORMAL RETURNS**

	(1) CAPM Model	(2) 4-Factor Model	(3) Industry-matched Model
Mean	-2.131* (0.917)	-2.231* (0.910)	-3.047** (1.235)
Median	-1.148* (0.563)	-1.358** (0.510)	-2.213* (1.125)
Trimmed Mean	-1.343** (0.394)	-1.392** (0.395)	-2.444** (0.658)
Negative Return Count	15* (5, 14)	15* (5, 14)	15* (4, 14)

Notes: Estimates are for the five-day cumulative abnormal return,  $CAR(A,5)$  for the NMW stocks. The trimmed mean excludes the two largest and two smallest returns. Negative Return Count is the number of stocks (out of 20) that have a negative five-day CAR (with the empirical 95 percent confidence interval below). \*\* denotes significance at 1 percent level; \* denotes significance at 5 percent level.

**TABLE A2. REGRESSION ESTIMATES OF TEN-DAY ABNORMAL RETURNS**

	(1)	(2)	(3)	(4)
NMW	-2.268* (1.101)	-2.540* (1.136)	-2.676* (1.220)	
NMW High $\pi$ Impact				-4.089* (1.765)
NMW Low $\pi$ Impact				-1.080 (1.383)
Average Wage £15-20k			-0.321 (1.012)	-0.259 (1.107)
Average Wage £20-25k			-0.934 (0.751)	-0.894 (0.749)
Average Wage £25-30k			0.615 (1.115)	0.619 (1.117)
Sample Size	442	442	442	442
2-Digit Industry	N	Y	Y	Y
10-Day Prior Return	N	Y	Y	Y
Size Quintiles	N	Y	Y	Y

Notes: The dependent variable is the ten-day cumulative abnormal return,  $CAR(A,10)$ , using the CAPM model. NMW are our sample of 20 firms with average wage less than £15,000. NMW High  $\pi$  Impact are the 10 NMW firms for which a 4 percent rise in the wage bill would generate the largest percentage decline in pre-tax profits. NMW Low  $\pi$  Impact are the other ten firms in the NMW sample. We also control for foreign exposure with a dummy equal to one if the majority of firm employment (or sales) is outside the UK. Robust standard errors in parentheses. \*\* denotes significance at 1 percent level; \* denotes significance at 5 percent level.



**TABLE A3. REGRESSION ESTIMATES OF TEN-DAY ABNORMAL RETURN  
USING ACTUAL EXPOSURE**

	(1)	(2)	(3)	(4)
NMW Exposure	-6.850* (2.920)	-7.816* (3.319)		
NMW Sample			-2.537* (1.025)	-2.428* (1.149)
Sample Size	275	275	275	275
2-Digit Industry	N	Y	N	Y
10-Day Prior Return	N	Y	N	Y
Size Quintiles	N	Y	N	Y

Notes: The dependent variable is the ten-day excess return using the CAPM model. NMW Exposure is the percentage of workers in the firm that earn within 5p on the National Minimum Wage in the ASHE sample for the firm. NMW Sample is a dummy variable equal to 1 if the firm is one of the 20 firms with the largest value for NMW Exposure and where the measure of exposure is based on at least 30 observations in ASHE. We also control for foreign exposure with a dummy equal to one if the majority of firm employment (or sales) is outside the UK. Robust standard errors in parentheses. \*\* denotes significance at 1 percent level; \* denotes significance at 5 percent level.