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Business cycles in Bulgaria and the Baltic countries: an RBC approach

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Abstract: This paper explores the business cycle in Bulgaria and the Baltic countries: Estonia, Latvia and Lithuania during the 1993–2005 period. The paper aims at deepening the understanding of the nature of output fluctuations. The neoclassical approach will be employed, much in the spirit of the real business cycle (RBC) literature, which gives a general equilibrium picture of the transition process. The model used in this paper follows the methodology of King et al. (1988). Both the model and data series show that the major drop in output was due to productivity. In addition, the timing of the banking reforms coincides with the improvement of economic performance. This is a strong indication that banking regulations in place were crucial for the output performance throughout the period in Bulgaria and the Baltic countries, a finding that has important implications for economic policy.

Keywords: business cycles; productivity; Baltic countries; Bulgaria.

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Biographical notes: Aleksandar Vasilev is a Doctoral student at Bonn Graduate School of Economics. His primary research interests are in real business cycle theory, dynamic stochastic general equilibrium (DSGE) modelling, computational methods in macroeconomics, economic growth and development.

1 Introduction

This paper is inspired by the business cycle literature; the nature of fluctuation of macroeconomic variables is still not fully understood. Indeed, several explanations have been proposed and tested empirically, but none of them seem to provide a full account for output movements. Monetarists, led by Friedman and Schwartz (1971) argue it was mainly due to monetary shocks, while another school of economic thought sticks to real factors. Representatives of the latter include Kydland and Prescott (1982) and Fisher and Hornstein (2001), who point to productivity shocks; Prescott (2002) adds government policies, and Bergoeing et al. (2002) focus particularly on banking regulations. The effects of banks on the real side of the economy is emphasised in Bernanke (1981) and Cole and Ohanian (2001).

The reason why some economists turned to the US Great Depression episode was that it still constitutes a puzzle for the business cycle theory. There was a prolonged recession coupled with unemployment rate up to 25% over the period 1929–1939. Negative effects of a similar magnitude were experienced in many other countries all over the world at that time as well. Much later, during the '90s, all countries in Central and Eastern Europe featured a significant cyclical evolution as well. Their economies underwent a thorough transformation with a major restructuring of their previously planned economies.

This paper will explore the business cycle in Bulgaria and the Baltic countries: Estonia, Latvia and Lithuania in the 1993–2005 period. This particular time frame was used as it is supposed to provide the most reliable data. The emphasis is on these particular countries because they all kept their fixed exchange rate regimes throughout the period (with a slight exception for Bulgaria, which adopted a currency board in 1997), much like the countries during the Great Depression that were on the gold standard. Baltic countries did not exist before 1992 as independent states, while Bulgarian data is of dubious quality in the first years of the regime change. In addition, Bulgaria avoided serious reform in the early years of transition (Mihov, 2001) – the initially pursued shock therapy did not find wide public and political support, so a course towards gradualism was taken. Thus, it is put at a more or less equal footing with the Baltic countries in terms of economic reforms.

The dataset compiled by Benczur and Ratfai (2005) will be used as it contains up-to-date information on business cycle regularities in the countries from Central and Eastern Europe. Moreover, this paper builds on their study by providing a theoretical framework for much of their applied work. This study finds that productivity is the main reason for the initial drop in output in the period of interest. Thus, the paper aims at deepening the understanding of the business cycle in those countries and emphasising the importance of the financial system for the real side of the economy, where the effect is through the productivity channel.

Although transition economies in Central and Eastern Europe were thoroughly investigated throughout the previous decade, we contribute to the literature by looking at their experience through another lens. In this paper, the neoclassical approach will be employed, much in the spirit of the real business cycle (RBC) literature, which gives a general equilibrium picture of the transition process. The aim here is to use the basic neoclassical RBC model and apply it to a group of transition countries.

The model used in this paper follows the methodology of King et al. (1988, 2002) as presented also by Cerny and Lazarova (1994). Calibration parameters, taken from data on Bulgarian and Baltic economies, are used in the estimation procedure; a growth accounting exercise following Solow's (1957) approach is performed, which calculates total factor productivity (TFP) as residuals from the production function. Then, simulations are performed in order to see how the theoretical model tracks the data series of different aggregates.

A common critique is that TFP does not tell you much about the factors lying behind the Solow residuals. The way to reconcile this is to provide some episodes from the economic history of Bulgaria and the Baltic countries that could serve as possible explanations and try to judge which ones are plausible. The most difficult case is with institutional factors and structural issues since they are almost impossible to quantify. Nevertheless, we will try to determine from the data and economic history which reforms were relatively more important for the drop in the TFP.

Both the model and data series show that the major drop in output was due to productivity. In addition, the timing of the banking reforms coincides with the improvement of economic performance. This is a strong indication that banking regulations in place were crucial for the output performance throughout the period in Bulgaria and the Baltic countries.

The paper is structured as follows: the next chapter provides a review of the literature in order to place the paper in the relevant field of economic research. The stories that serve as possible explanations for the transition are discussed in the second chapter in order to illustrate the dynamics of the economies of interest. The third part describes the theoretical model and provides justification for the calibrated parameters. Model predictions are discussed and then compared against the countries' economic history as well as the existing findings in the business cycle literature. In the concluding section, the results are summarised, their importance with respect to policy is emphasised and directions for further research are provided.

2 Literature overview

The literature on the subject can be roughly divided into two clusters: papers that perform a general equilibrium analysis using a business cycle approach, and those who go for the descriptive approach. The first cluster finds TFP to be the cause for output movement, first in the US and then extended to other countries around the globe as well. They do not answer, however, what is behind this residual from estimation. Others claim it was not productivity and that the model was misspecified. Not all explanations, however, are quantifiable: some examples are structural issues such as institutional arrangements and legal framework.

Therefore, the influence of those factors is explained by another group of economists in a descriptive way only. Rostowski (1996) compares the Great Depression in the US to the transition process in Central and Eastern Europe. He notes, however, that in the initial years of transition investment fell less than GDP did compared to the depression in the US. Using a simple financial accelerator model, Rostowski (1996) concludes that in the cases of Poland and Hungary, there was a major restructuring. This was triggered by a shift in demand that followed from the trade liberalisation and that of prices, a 'real shock' to enterprises, which in the RBC literature is denoted by a shock to technology and productivity. In short, enterprises did not suffer from excess capacity but only lacked 'the right kind of productive capacity' (p.225).

Amaral and MacGee (2002) show that TFP shocks account for a significant part of the Canadian depression. However, they do not have any theory to explain TFP behaviour. Banking shocks account only for a small part of the downturn, and are insignificant in explaining the slow recovery. Fisher and Hornstein (2001) find also changes in productivity to be important factors of the Great Depression in Germany – around less than two-thirds of the fall in output. Kehoe and Ruhl (2003) suggest that movements in TFP explain the low growth period in Switzerland and New Zealand.

Hopenhayn and Neumeyer (2003) claim the fall in output per worker in Argentina in 1975–1990 was also partly due to productivity growth, together with labour reallocation and capital deepening. Their critique is that growth accounting exercises explain everything with TFP, which is exogenous in the model and there are many reasons that

lie behind that variable. In the Argentinean case, according to Hopenhayn and Neumeyer (2003), those could be commercial policies, exchange rate controls, tax structure, tariffs and non-tariff barriers, expectation of bank runs, and confiscation of deposits. These factors increased the cost of capital and relative price of investment goods. This caused a bad investment environment, which led to stagnation in capital accumulation and induced labour reallocation.

Bergoeing et al. (2002) focus their attention on Mexico and Chile in the last two decades of the 20th century. In their study, they test four stories about the different paths of economic recovery: standard monetarist story, real wage story, debt overhang story and structural reforms (trade policy, fiscal policy, privatisation, the banking system and bankruptcy procedures). Bergoeing et al. (2002) do not find these explanations to hold in the data, except for the privatisation, the banking system and bankruptcy procedures. To demonstrate this, they do a growth accounting exercise and show that the two different paths of recovery in Chile and Mexico were due to the different evolution of TFP. Then the authors look into the government policies that could account for the TFP difference. The authors conclude it was the banking system and the bankruptcy procedures that made the difference, because both countries had already undertaken privatisation before the observation period.

External shock, according to Bergoeing et al. (2002), was what triggered the TFP decrease in Mexico and Chile. In the cases of Bulgaria and the Baltic countries, this could be attributed to the movement to a market economy. This is also a technology shock, since it deals with operational efficiency and better resource management. In addition, the drop in productivity was exacerbated by the government policies (mainly banking regulations and bankruptcy procedures) in place and led to fragility in the financial sectors of the economies of interest. In a survey paper, Prescott (2002) also claims government policies can lead to prosperity or depressions.

The problem Hayashi and Prescott (2002) identify with Japan in the 1990's was the low productivity growth rate. Their growth accounting exercise, which takes TFP as exogenous, fits the data well. Japan experienced a fall in TFP productivity and increase in the capital-output ratio during the 1990's. According to them, growth theory leaves no role to financial intermediation and any frictions in it. Despite the collapse of bank loans and the 'credit crunch', firms found other ways to finance their investment – which were roughly close substitutes to bank credit.

Bernanke (1981) argues that the bankruptcy risk, especially when an economy-wide phenomenon, plays an important role in the propagation of recessions. He assumes that this problem occurs after there has been a drop in output, and financial distress magnifies that effect through the income-spending multiplier. Bernanke finds that bankruptcy risk was important in the Great Depression in 1929–1933. The limitation of his work, however, is that the authors do not look at the problem in a general equilibrium. Bernanke finds only that banking crises and loss in output are correlated but they do not measure the negative effect of the banking factor for the economy.

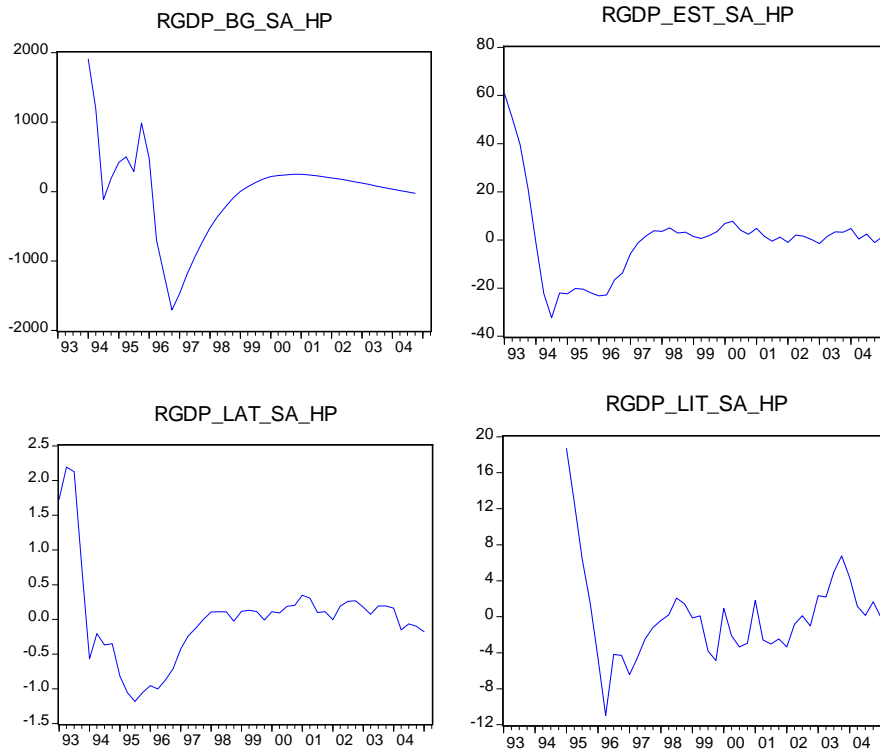
In this paper, we take the best of two worlds and combine them in pursuit of the answer to the output fall during the transition in Bulgaria and the Baltic countries. We start with a general theory in order to get the broad picture and then match our theoretical results with empirical facts. As noted, the literature in the field necessitates the use of a general equilibrium model, from which TFP will be estimated in order to see the effect of productivity on output. We will compare the movement in the TFP with the timing of reforms to see which factor could be not only a good explanation, but also to match the

time of implementation. In this way, we also show that RBC is a useful tool for studying the transition countries as well and to establish an unexploited niche in the field. In the next section, before we present the model, some background information on the Bulgarian and Baltic economies is provided.

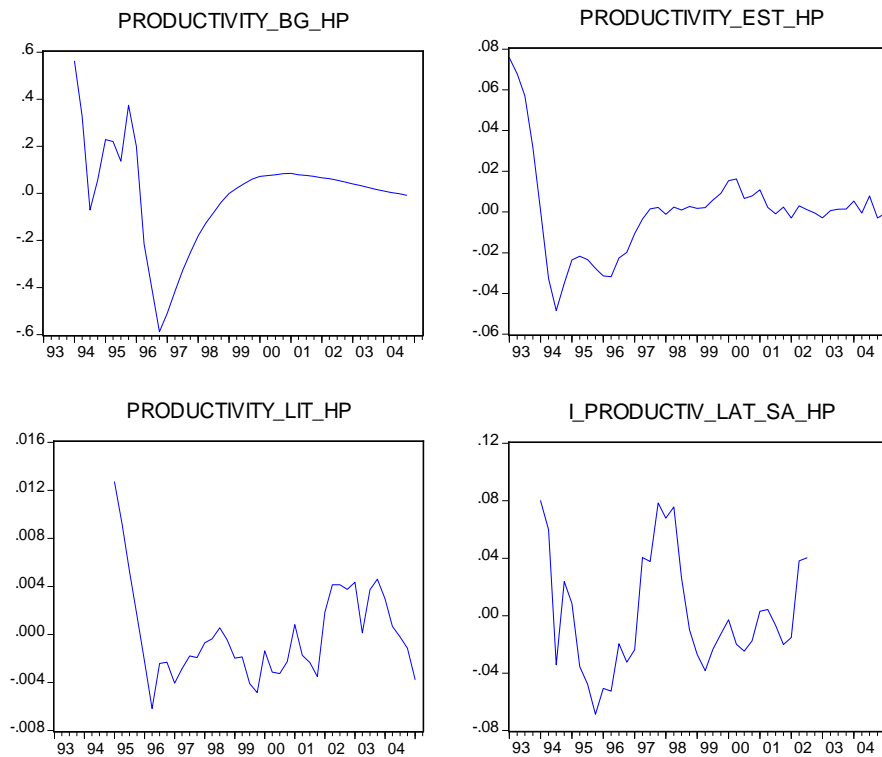
3 Economic overview of Bulgaria and the Baltic countries

Both the Baltic countries and Bulgaria were connected to the Soviet Union: Estonia, Lithuania and Latvia being part of the Russian Federation, while Bulgaria played the role of a satellite country. Heavy industry, especially that of the Bulgarian economy, was complementary to that of the Soviet Union. After the breakup of the Soviet Union in 1989, all countries faced a common external shock, which was the movement to the market economy and the world market prices of their tradables. Figure 1 below shows the rapid collapse in the beginning and then the slow recovery.

Figure 1 Cyclical fluctuations of real GDP, 1993–2005 (see online version for colours)



When we compare this to the productivity cycle (see Figure 2), we note that the cycle follows that of the real output, shown in Figure 1, very closely. The only exception is Latvia because we use industrial productivity due to data limitations. (Industrial productivity cycle, however, tracks the one of industrial production.)

Figure 2 Productivity fluctuations, 1993–2005 (see online version for colours)

Romer (2001, p.171) also notes that ‘productivity – output per worker-hour – almost always declines during recessions’ and that fluctuations are distributed very unevenly over the components of output. Our case it seems to be a productivity story as well. Thus, reforms that affect TFP directly will have a better explanatory power.

To get the whole picture, however, we should also discuss the four principle GDP components: consumption, investment, government purchases and net exports. The first factor, consumption, is highly volatile. This fact cannot be attributed solely to transformation and International Monetary Fund (IMF) stabilisation programs. Consumption spending proportional to the size of the economy is highest in Bulgaria – about 70–80% of GDP. An important component of consumer spending in all the four countries in our study is considered to be imported durables from the West. In addition, as noted by Benczur and Ratfai (2005) consumption habits are strong, and savings ‘for a rainy day’ are in hard currency.

The second component, investment, fluctuates a great deal due to the transformation in the economy: old plants are scrapped, new plants are opened, foreign direct investment (FDI) and capital inflows are attracted. During this process, government purchases are curtailed at the expense of the rapid expansion of the private sector. At the same time, fiscal balances are put in order, and budgetary items switch categories to reflect Western accounting standards. Government spending diminishes in size compared to the economy, but governments themselves do not decrease in importance as the agent setting

the rules of the market game. In addition, all the economies are small, and take international prices as given. After the fix, these countries saw their currencies appreciating in real terms. Another common feature is that all the countries have significant trade and current account (CA) deficits, mostly due to the catching-up process and the capital inflows. It is important to note as well the widespread weaknesses in the banking system (especially regulations and bankruptcy procedures) in the early '90s.

The economic overview of Bulgaria and the Baltic countries demonstrates that all four countries show a sufficient degree of similarity. This justifies a more abstract level of analysis, an approach to be pursued using a general-equilibrium model. A formal description of the model follows in the next chapter.

4 Model description

In this section, a simplified picture of the economy is presented, following the reasoning provided by King et al. (1988). Their basic RBC model will be used in order to try to explain what we observe in the data. Model predictions will then be compared to the fluctuations in the data and conclusions based on the reading from economic history will be drawn, principally that productivity and banking regulations explain a great deal of the fall in output in Bulgaria and the Baltic countries during the 1993–2005 period.

All agents in the theoretical economy are assumed to be identical, and thus can be thought of as a dynasty that is living infinitely. In addition, due to the symmetry principle applied, we can focus on just one agent, or household, and later generalise the results for the whole population. Each relationship in the economy is specified in a separate subsection. After the model is solved, a first-order Taylor approximation is used to linearise around the steady-state and examine the cyclical fluctuations.

4.1 Preferences

Preferences of an infinitely-living household are represented by the following utility function

$$U = \sum_{t=0}^{\infty} \beta^t u(C_t, L_t), \quad 0 < \beta < 1 \quad (1)$$

where

$$u(C_t, L_t) = \frac{1}{1-\sigma} C_t^{1-\sigma} v(L_t) \quad (2)$$

and C_t and L_t denote consumption and leisure, respectively, in period t . The function $u(C_t, L_t)$ is assumed to be strictly increasing, concave, twice continuously differentiable and to satisfy Inada conditions,¹ which ensure interior solution (if feasible).

4.2 Technology

There is one homogenous good in the economy, which is produced according to the following technology

$$Y_t = A_t K_t^{1-\alpha} (N_t X_t)^\alpha \quad (3)$$

The production function features constant returns to scale. In addition, it is concave, twice continuously differentiable and satisfies the Inada conditions. The labour-augmenting technology is needed for the model to achieve a steady-state: in equilibrium, the steady-state growth of the labour force is zero.

The goods produced in each period can be either invested or consumed. Physical capital depreciates at a constant rate δ , and the following equation of motion always holds true:

$$K_{t+1} = I_t + (1 - \delta) K_t \quad (4)$$

In this paper, the definition for investment by Benczur and Ratfai (2005) will be used. Those authors used only gross capital formation because the change in inventories is too small to be of importance.

4.3 Resource constraints

The time devoted to work and that used for leisure activities equal to the time endowment, here normalised to unity.

$$L_t + N_t = 1 \quad (5)$$

Consumption and investment must not exceed output

$$C_t + I_t \leq Y_t \quad (6)$$

4.4 Steady-state

Having the necessary concavity conditions, we can use the second welfare theorem and solve the central planner problem as one of a decentralised economy and find the competitive solution. In order to obtain an economy that achieves a steady-state, all variables must be written in efficiency units, e.g., $c = \frac{C}{X}$, $k = \frac{K}{X}$, $i = \frac{I}{X}$, etc. The utility function, technology and the two resource constraints are reformulated in efficiency variables. The labour-augmenting technical progress is assumed to have a constant growth rate, expressed as $X_t = e^{(\gamma-1)t}$, where

$$\gamma \geq 1 \quad (7)$$

Thus, we obtain the expressions below for the utility function, the production function, the state equation for capital and the income-expenditure constraint, respectively,

$$U = \sum_{t=0}^{\infty} (\beta^*)^t \frac{1}{1-\sigma} c_t^{1-\sigma} v(L_t) \quad (8)$$

$$y_t = A_t k_t^{1-\alpha} N_t^\alpha \quad (9)$$

$$\gamma k_{t+1} = i_t + (1 - \delta) k_t \quad (10)$$

$$y_t \geq i_t + c_t \quad (11)$$

where $\beta^* = \beta\gamma^{1-\sigma}$. The Lagrangean associated with the optimization problem of interest is

$$\mathcal{L} = \sum_{t=0}^{\infty} (\beta^*)^t \left\{ \frac{1}{1-\sigma} c_t^{1-\sigma} v(1-N_t) + \lambda_t \left[A_t k_t^{1-\alpha} N_t^\alpha - c_t + (1-\delta)k_t - \gamma k_{t+1} \right] \right\} \quad (12)$$

The representative household chooses c_t , N_t , K_{t+1} . The first-order conditions (FOCs) together with the income-expenditure constraint and the transversality condition (TVC) for capital are as follows:

$$c_t^{-\sigma} v(1-N_t) - \lambda_t = 0 \quad (13)$$

$$-\frac{1}{1-\sigma} c_t^{1-\sigma} v'(1-N_t) + \alpha \lambda_t A_t \left(\frac{k_t}{N_t} \right)^{1-\alpha} = 0 \quad (14)$$

$$-\lambda_t \gamma + \beta^* \lambda_{t+1} \left[A_{t+1} \left(\frac{N_{t+1}}{k_{t+1}} \right)^\alpha + (1-\delta) \right] = 0 \quad (15)$$

$$A_t k_t^{1-\alpha} N_t^\alpha - c_t + (1-\delta)k_t - \gamma k_{t+1} = 0 \quad (16)$$

$$\lim_{t \rightarrow \infty} (\beta^*)^t \lambda_t k_{t+1} = 0 \quad (17)$$

4.5 Steady-state dynamics

In order to obtain the quantitative impact of changes in A_t , we linearise the FOC equations (13)–(16) around the steady-state with respect to A , k , N , c , y . All those variables are expressed as percentage deviations from the steady-state, e.g.,

$$\hat{c}_t = \ln \frac{c_t}{c} = \frac{c_t - c}{c} \quad (18)$$

Using the log-linearisation technique as described by King et al. (2002), Campbell (1994) and Uhlig (1999), we obtain the following equations

$$\zeta_{cc} \hat{c}_t - \zeta_{cl} \frac{N^{ss}}{1-N^{ss}} - \hat{\lambda}_t = 0 \quad (19)$$

$$\zeta_{lc} \hat{c}_t - \zeta_{ll} \frac{N^{ss}}{1-N^{ss}} \hat{N}_t - \hat{\lambda}_t - \hat{A}_t - (1-\alpha) \hat{k}_t + (1-\alpha) \hat{N}_t = 0 \quad (20)$$

$$\lambda_{t+1} + \eta_A \hat{A}_{t+1} + \eta_k \hat{k}_{t+1} + \eta_N \hat{N}_{t+1} = \hat{\lambda}_t \quad (21)$$

$$\hat{y}_t = \hat{A}_t + \alpha \hat{N}_t + (1-\alpha) \hat{k}_t = s_c \hat{c}_t + s_i \Phi \hat{k}_{t+1} - s_i (\Phi - 1) \hat{k}_t \quad (22)$$

where $\zeta_{ab} = \frac{d \ln u_a}{d \ln b}$ is the elasticity of marginal utility of a with respect to b . For the Cobb-Douglas technology $\eta_A = \frac{\gamma - \beta(1-\delta)}{\gamma}$, $\eta_k = -\alpha\eta_A$, $s_i = (1-\alpha)\frac{\gamma + \delta - 1}{r + \delta}$, and s_c are shares of investment and consumption in output ($s_c + s_i = 1$), $\Phi = \frac{\gamma}{\gamma - (1-\delta)}$.

In order to find the dynamics of the capital formation, a system of two first-order difference equations in $\hat{\lambda}_t$ and \hat{k}_t , which are obtained from equations (19), (20) and (22):

$$\begin{pmatrix} \hat{k}_{t+1} \\ \hat{\lambda}_{t+1} \end{pmatrix} = W \begin{pmatrix} \hat{k}_t \\ \hat{\lambda}_t \end{pmatrix} + R\hat{A}_{t+1} + Q\hat{A}_t \quad (23)$$

The columns of matrix P are formed by the eigenvectors of matrix W

$$W = P \begin{pmatrix} \mu_1 & 0 \\ 0 & \mu_2 \end{pmatrix} P^{-1} \quad (24)$$

Therefore,

$$\hat{k}_{t+1} = \mu_1 \hat{k}_t + \psi_1 \hat{A}_t + \psi_2 \sum_{j=0}^{\infty} \mu_2^{-j} \hat{A}_{t+1+j} \quad (25)$$

where the expressions for ψ_1 , ψ_2 can be derived from equation (24).

4.6 Real business cycles

Now, we remove the assumption of perfect foresight and introduce a stochastic shock guiding the TFP behaviour.

The starting equation is (25), which describes the path of capital formation. Then a particular stochastic process for \hat{A}_t is specified, in this Case AR(1) with a coefficient of persistence ρ , and replace \hat{A}_{t+j} with their expected values given information at time t .

The dynamics of the state variables \hat{k}_t, \hat{A}_t is then given by the linear system

$$s_{t+1} = \begin{pmatrix} \hat{k}_{t+1} \\ \hat{A}_{t+1} \end{pmatrix} = \begin{pmatrix} \mu_1 & \pi_{kA} \\ 0 & \rho \end{pmatrix} \begin{pmatrix} \hat{k}_t \\ \hat{A}_t \end{pmatrix} + \begin{pmatrix} 0 \\ \varepsilon_{A,t+1} \end{pmatrix} = Ms_t + \varepsilon_{t+1} \quad (26)$$

The other linear equations specify how consumption, efficient labour, investment and output depend on the state-variable vector s_t . Those equations are derived from equations (19)–(22) and also using equation (30). In vector form,

$$z_t = \begin{pmatrix} \hat{c}_t \\ \hat{N}_t \\ \hat{y}_t \\ \hat{i}_t \end{pmatrix} = \begin{pmatrix} \pi_{ck} & \pi_{cA} \\ \pi_{Nk} & \pi_{NA} \\ \pi_{yk} & \pi_{yA} \\ \pi_{ik} & \pi_{iA} \end{pmatrix} \begin{pmatrix} \hat{k}_t \\ \hat{A}_t \end{pmatrix} = \Pi s_t \quad (27)$$

The formulation above facilitates the calculation of population moments once we know the variance-covariance matrix \sum_{tt} or the state vector s_t .

$$\text{var}(x_{t+j} x_t^T) = M^j \sum_{tt} \quad (28)$$

$$\text{var}(z_{t+j} x_t^T) = \Pi M^j \sum_{tt} \quad (29)$$

It is easy to verify that

$$\sum_{11} = \text{var}(\hat{A}_t) = \frac{\sigma^2}{1-\rho^2} \quad (30)$$

$$\sum_{22} = \text{var}(\hat{k}_t) = \frac{\pi_{kA}^2 \sigma^2}{(\rho - \mu_1)^2} \left(\frac{\rho^2}{1-\rho^2} + \frac{\mu_1^2}{1-\mu_1^2} - \frac{2\mu_1\rho}{1-\mu_1\rho} \right) \quad (31)$$

$$\sum_{12} = \sum_{21} = \text{cov}(\hat{A}_t, \hat{k}_t) = \frac{\pi_{kA}^2 \sigma^2}{\rho - \mu_1} \left(\frac{\rho^2}{1-\rho^2} - \frac{\mu_1}{1-\mu_1\rho} \right) \quad (32)$$

4.7 Parameterisation of the model

Calibration approach, pioneered by Kydland and Prescott (1982) will be used when solving the model. Parameters of the model were estimated using quarterly data for Bulgarian and the Baltic countries. Data series are taken from Benczur and Ratfai's (2005) dataset.

In this paper, an estimate by Ganey (2005) will be used, with standard errors in brackets:

$$\hat{\alpha}_{BG} = 0.43(0.07) \quad (37)$$

A linear depreciation scheme is applied, where

$$\delta = \frac{I_t - (K_{t+1} - K_t)}{K_t} \quad (38)$$

Ganey (2005) finds it to be 0.05 for Bulgaria when annual data is used. We set $\delta = 0.0125$, since we use quarterly data for all the four countries because we assume the same production technology being available everywhere.

The variable number of hours worked was not reported in the dataset used. As there are no observations to determine both hours and stationary hours, we set them to 0.4, as done in Cerny and Lazarova (1994) for Czechoslovakia. This corresponds to a 9-hour

working day, which is a reasonable assumption, given the fact that working hours are fixed during the period. In addition, Cooley and Prescott (1995) also do not differentiate between total hours and employment in their model.

On the balanced growth path all variables grow at a constant rate $\gamma - 1$, that is

$$\begin{aligned}\ln Y_t &= \ln y_t + (\gamma - 1)t = \ln \bar{y}_t + \hat{y}_t + (\gamma - 1)t \\ \ln C_t &= \ln c_t + (\gamma - 1)t = \ln \bar{c}_t + \hat{c}_t + (\gamma - 1)t \\ \ln I_t &= \ln i_t + (\gamma - 1)t = \ln \bar{i}_t + \hat{i}_t + (\gamma - 1)t \\ \ln K_t &= \ln k_t + (\gamma - 1)t = \ln \bar{k}_t + \hat{k}_t + (\gamma - 1)t\end{aligned}\tag{39}$$

Output, consumption, investment and capital are regressed on a common time trend and independent constant terms in order to obtain an estimate for $\gamma - 1$:

$$\begin{array}{cccc}\hat{\gamma}_{BG} - 1 = 0.005, & \hat{\gamma}_{EST} - 1 = 0.013, & \hat{\gamma}_{LAT} - 1 = 0.0133, & \hat{\gamma}_{LIT} - 1 = 0.0129 \\ (0.001) & (0.0005) & (0.0009) & (0.0009)\end{array}$$

It is not surprising that estimates show a common trend growth rate for the Baltic countries, a fact which we utilise in the calibration process later on.

TFP was calculated using the definition in Benczur and Ratfai (2005). Since there is not much variation in capital due to the deterministic nature of the estimated series, Solow residuals will follow closely the movement in labour productivity, obtained as a ratio of GDP and employment. Seasonally adjusted and detrended productivity series were then tested for stationarity using the standard Dickey-Fuller test. At 1% level of significance, we rejected the presence of a unit root. Then we regresses \hat{A}_t , HP-detrended on its lagged value, assuming productivity follows an AR (1) process.² The coefficient in front of the lagged term measures persistence (standard errors reported in brackets). In addition, the variance of the shock was calculated from the residuals of the regression, and is assumed to be normally distributed.

$$\begin{aligned}\hat{A}_t &= \alpha + \rho \hat{A}_{t-1} + u_t \\ \hat{\rho}_{BG} &= 0.76(0.08), s.d(\hat{u}_t)_{BG} = 0.11011 \\ \hat{\rho}_{EST} &= 0.56(0.05), s.d(\hat{u}_t)_{EST} = 0.00831 \\ \hat{\rho}_{LAT} &= 0.49(0.12), s.d(\hat{u}_t)_{LAT} = 0.001991 \\ \hat{\rho}_{LIT} &= 0.59(0.08), s.d(\hat{u}_t)_{LIT} = 0.027727\end{aligned}$$

Persistence measures for the productivity shock is quite similar among the Baltic countries, in the range 0.49–0.59, while Bulgaria has a much higher estimate – 0.76, together with the highest volatility in the shock process.

Below are provided the calibration parameters and steady-state values, which are divided into two groups, those for Bulgaria and another set for the Baltic countries. That particular division was necessitated by the great similarity in terms of economic performance in Estonia, Latvia and Lithuania. To solve the model, we use the Morten Ravn's (2006) GAUSS code for the model by King et al. (1988).

<i>Parameters</i>	<i>Bulgaria</i>	<i>Baltic countries</i>
Labour share α	0.43	0.43
Share of time allocated to work N	0.4	0.4
Growth rate of TFP γ	1.005	1.013
Real interest rate r	$0.065 / 4 = 0.015$	$0.065 / 4 = 0.015$
Depreciation rate δ	0.0125	0.0125
Persistence of TFP shock ρ	0.76	0.56
St. dev. of TFP shock	0.11011	0.00831
Discount factor $\beta^* = \beta\gamma^{1-\sigma}$	0.989	0.989
Elasticity of MU of cons. σ	1	1

<i>Steady-state values</i>	<i>Bulgaria</i>	<i>Baltics</i>
A^{ss}	10	10
Y^{ss}	1467.735	1467.7352
C^{ss}	733.192	672.1673
K^{ss}	12628.061	12628.0611
I^{ss}	694.543	795.5678
K^{ss} / Y^{ss}	8.6038	8.6038
$W = AP^{ss} = Y^{ss} / N$	3669.338	3669.3379

A peculiarity of the model is that it predicts too high capital-output ratio of around 8.6. That art effect, however, can be explained with the excessive capital accumulation during the central planning era and low utilisation of machines at the same time. In Bulgaria, the ratio is above 3.5, while for a typical market economy it is around 3.

4.8 Stochastic simulation

Two simulation of the model was done with 500 experiments and a horizon of 50 observations in order to correspond to the time span in the data. One of the specifications was calibrated using data for Bulgaria, and the other corresponds to the Baltic countries. The latter is named after Estonia in the paper because Estonian data for the technology shock was used in this simulation. Then the predicted results are compared to the estimates in Benczur and Ratfai (2005) in Tables 1–6 below.

Both model specifications generally underpredict the absolute and relative volatilities of the variables in the data but that is a typical feature of such simple models, e.g., Cooley and Prescott (1995). The only exclusion is the relative volatility of employment for the Estonian model. In addition, the model calibrated for Estonia tends to overpredict cyclical measures and under-predict persistence estimates, with the exception of investment persistence, which is an exact match. Despite this particular deficiency of the model, actual persistence values fall within two standard deviations from the simulated ones and thus are considered to be statistically close to that of the model.

Table 1 GDP

	<i>GDP volatility</i>	<i>Autocorrelation</i>
Bulgaria	4.2	0.66
Estonia	2.46	0.67
Latvia	1.89	0.65
Lithuania	2.53	0.56
Model EST	1.1027 (0.1434)	0.375(0.1343)
Model BG	0.2827 (0.0454)	0.5478 (0.1257)

Notes: GDP volatility is measured as the standard deviation of GDP.

Autocorrelation is measured as the AR (1) coefficient in the series.

Source: Benczur and Ratfai (2005), author's estimates

Table 2 Consumption (C)

	<i>BG</i>	<i>EST</i>	<i>LAT</i>	<i>LIT</i>	<i>Model EST</i>	<i>Model BG</i>
Abs.volatility ¹	5.48	3.41	2.38	2.45	0.1406 (0.0194)	0.0669 (0.0177)
Rel. volatility ²	1.30	1.38	1.39	0.97	0.1275 (0.0045)	0.2346 (0.0385)
Cyclicalit ³	0.78	0.60	0.22	0.26	0.9155 (0.0314)	0.6438 (0.0490)
Persistence ⁴	0.56	0.65	0.18	0.28	0.4379 (0.1337)	0.8355 (0.0725)

Notes: ¹Absolute volatility is measured as the standard deviation of the series.

²Relative volatility is measured as the ratio of the standard deviation of the variable of interest and real GDP.

³Cyclicalit is measured as the contemporaneous correlation between the variable and real GDP.

⁴Persistence is measured as the AR(1) coefficient in the series.

All data is at quarterly frequency, seasonally adjusted and detrended using HP filter.

Source: Benczur and Ratfai (2005), author's estimates

Table 3 Investment (I)

	<i>BG</i>	<i>EST</i>	<i>LAT</i>	<i>LIT</i>	<i>Model EST</i>	<i>Model BG</i>
Abs. volatility	14.3	12.3	10.45	9.39	1.9272 (0.2516)	0.553 (0.0872)
Rel. volatility	3.42	5.02	6.12	3.71	1.7475 (0.0064)	1.9568 (0.0217)
Cyclicalit	0.39	0.69	0.20	0.72	0.9997 (0.0064)	0.9945 (0.0022)
Persistence	0.45	0.37	0.09	0.62	0.3763 (0.1343)	0.5328 (0.1268)

Note: All data is at quarterly frequency, seasonally adjusted and detrended using HP filter.

Source: Benczur and Ratfai (2005), author's estimates

Table 4 Employment (N)

	<i>BG</i>	<i>EST</i>	<i>LAT</i>	<i>LIT</i>	<i>Model EST</i>	<i>Model BG</i>
Abs. volatility	4.5	1.32	3.29	2.15	0.5855 (0.0764)	0.1473 (0.0230)
Rel. volatility	1.0	0.54	1.74	0.85	0.5309 (0.0021)	0.5216 (0.0094)
Cyclicalit	-0.17	0.47	0.63	0.13	0.9983 (0.0007)	0.9773 (0.0094)
Persistence	0.9	0.75	0.84	0.68	0.3793 (0.1341)	0.5254 (0.1275)

Note: All data is at quarterly frequency, seasonally adjusted and detrended using HP filter.

Source: Benczur and Ratfai (2005), author's estimates

Table 5 Technological productivity¹ (A)

	<i>BG</i>	<i>EST</i>	<i>LAT</i>	<i>LIT</i>	<i>Model EST</i>	<i>Model BG</i>
Abs. volatility	6.7	2.18	3.08	3.10	0.8423 (0.1098)	0.2203 (0.0346)
Rel. volatility	1.6	0.89	1.74	1.22	0.7638 (0.002)	0.7797 (0.011)
Cyclicalit	0.7	0.84	0.42	0.73	0.9992 (0.0003)	0.9903 (0.0039)
Persistence	0.8	0.56	0.49	0.59	0.3776 (0.1342)	0.5295 (0.127)

Notes: ¹Productivity measure in the data corresponds more to the real wage in the model, as it is defined as output/labour hours. By no means, it is claimed that the definition for productivity used in this paper, as well as in Romer (2001) and Benczur and Ratfai (2005) captures the whole effect of the TFP. No matter which of the two measures we use, however, the results obtained are qualitatively identical.

All data is at quarterly frequency, seasonally adjusted and detrended using HP filter.

Source: Benczur and Ratfai (2005), author's estimates

Table 6 Real wages

	<i>BG</i>	<i>EST</i>	<i>LAT</i>	<i>LIT</i>	<i>Model EST</i>	<i>Model BG</i>
Abs. volatility	7.6	2.41	3.45	6.16	0.5193 (0.0674)	0.1422 (0.0248)
Rel. volatility	1.8	0.98	1.83	2.43	0.4709 (0.0021)	0.5021 (0.0166)
Cyclicalit	0.6	0	0.23	0.45	0.9978 (0.0009)	0.9762 (0.0072)
Persistence	0.6	-0.01	0.73	0.77	0.3743 (0.1343)	0.6269 (0.1222)

Note: All data is at quarterly frequency, seasonally adjusted and detrended using HP filter.

Source: Benczur and Ratfai (2005), author's estimates

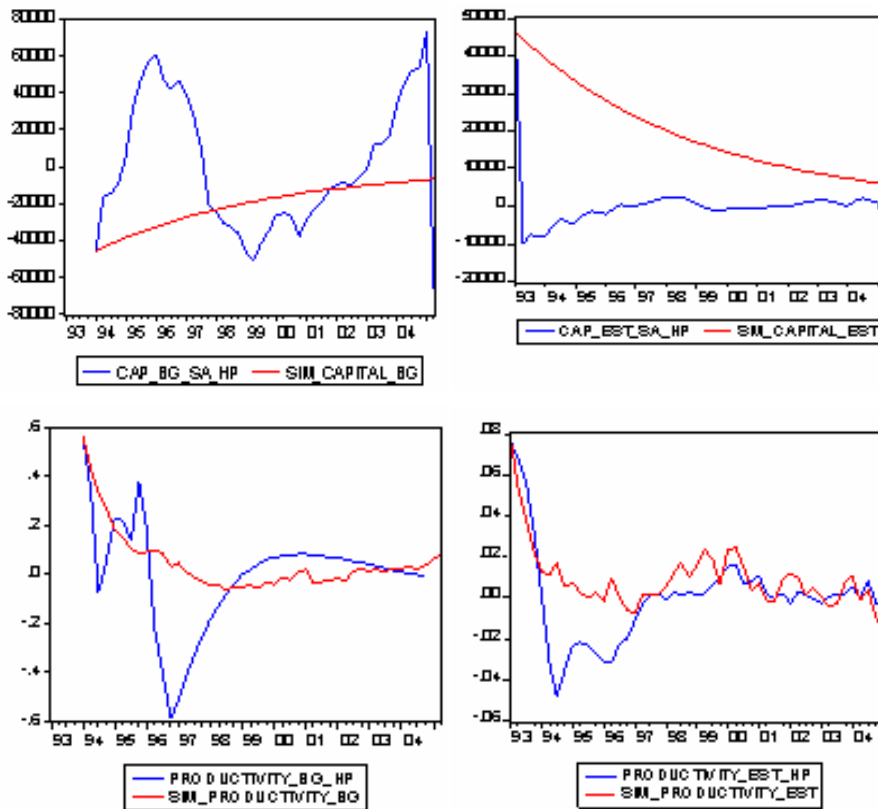
The model for Bulgaria gets closer estimates, although it overestimates the cyclicalit parts as well. The model performs much better with persistence measures: it overpredicts the measure for consumption, but gets a close match for investment and real wages, and a value in the 99% confidence interval for output autocorrelation and persistence measures for employment and technology shock.

At first sight, the model in its two specifications does not seem to be an effective one. Some important limitations of the analysis should be noted, however. The model is extremely simplistic: Our model does not have a government sector and there is no external sector or money, which is a somewhat restrictive framework. In addition, we use quarterly data for ten years and the model has a horizon of 50 observations. This may be too short a period for long-term tendencies to show up in the model. Cerny and Lazarova (1994) obtain similar results for Czechoslovakia when they use annual data for 1948–1983 period. Their explanation lies in the fact that central planning is not well described by aN RBC model. In our case, the countries undertook serious transformation throughout the period and were not considered as market economies until towards the end of the period. The transition countries are likely to be off the balanced-growth path. The assumption of the same technology and preferences may not hold: the utility function may not be logarithmic and capital and labour markets are far from perfect in the countries we investigate. Employment, for example, is still not entirely market-based during the period with a high public sector and too much administrative staff. In addition, there were privatisation deals that had clauses for preserving the levels of employment.

Data limitation is also an issue, since industrial employment was used for Latvia, which is a very crude estimate.

The model, however, performs quite well in tracking the cyclical path of capital and productivity.³ That is the great achievement of the RBC model, where all variables are driven by productivity shocks. The financial crisis episode in Bulgaria in 1996–1997 is not captured by the model, though. This, however, cannot be blamed on the model, as this crisis period was an outlier event. Figure 3 shows the match of simulated and actual series of capital and productivity cycles.

Figure 3 Cyclical fluctuations of actual vs. simulated capital and productivity (see online version for colours)



The explanation why the simulated capital series do not closely follow the cycle in the data is due to the assumption that capital stock is zero before the period. This seems to be a strict assumption, as demonstrated by Figure 3 above. However, the same problem with capital is noted by Cerny and Lazarova (1994) in their RBC study on Czechoslovakia. The conflict between theory and data may be due to the linear depreciation scheme used to estimate capital series. This is in line with Doms and Dunne's (1998) study, which shows that depreciation in US plants features important non-linearities.

In conclusion, both specifications generally capture the cyclical effects and persistence measures and this is an important contribution to the literature. In addition, the overall poor fit is a general phenomenon in the RBC literature: King et al. (1988) admit that in order for a neoclassical RBC model to match data well, the productivity shock should be highly persistent, with an autocorrelation coefficient of the order of 0.95–0.99. In this way, the dynamics of output are almost entirely determined by the persistence of the shock. Moreover, in their analysis, the authors do not use HP-filtered series. As a robustness check (available upon request), we perform a sensitivity analysis for both specifications by taking the autocorrelation coefficient to be 0.95 as in King et al. (1988). It does not provide qualitatively different results, though.

As we have demonstrated in this chapter, the basic RBC model fits the facts in Bulgaria and the Baltic countries quite well, given its oversimplified structure. This shows that neoclassical models have a great potential in researching business cycles in transition economies. In the next section, we connect model predictions with particular events in the countries of interest during 1993–2005.

5 Predictions of the theoretical model in light of the economic history in Bulgaria and the Baltic countries

After the completion of the technical exercise, we go back to the data once again in order to do a reading of economic history in Bulgaria and the Baltic countries in light of the evidence from the model. Below we provide some stories, following the structure proposed by Bergoeing et al. (2002). Following their argument, the paper focuses on monetary aggregates, trade, foreign debt and structural reforms.

5.1 Standard monetarist story

One of the frequently prescribed medicines for a country in a recession is to expand its money supply in the short- to medium-term (because in the long-run that monetary expansion is going to cause hyperinflation) in order to speed up recovery. This was a common story for the Great Depression, as noted by Bergoeing et al. (2002). There is ample evidence (Friedman and Schwarz, 1971; Bernanke, 1995; Bernanke and James, 1990) that countries that abolished the gold standard and increased their money supply recovered much of their lost output. One could argue that almost all central bankers in Central and Eastern Europe in the beginning of the transition process were monetarists, since they applied the approach advocated by this school of economic thought.

Three of the countries in our sample, however, Bulgaria (since April 1997), Estonia (June 1992) and Lithuania (April 1994) have currency board arrangements (CBA), while Latvia (1994) went for a fixed ER. Therefore, money supply in all those countries becomes an endogenous variable, a function of the capital flows into the country. In case of negative productivity shocks and/or capital outflows, CBA exacerbates the problem since the money supply will contract and make the recession at hand even bigger. The three Baltic countries went immediately for a fix after separating from Russia, while Bulgaria had a free float until 1997 due to lack of international reserves. Bulgaria also had a bad growth experience during the period 1991–1997, except for a feeble recovery

in 1994–1995 before the hyperinflation in 1996–1997. The conjuncture changed dramatically for the better after the introduction of the CBA.

In conclusion, the monetarist story is to a great extent unable to explain the recovery of the countries in the sample. Actually, a fixed regime in the Baltic countries and Bulgaria established long-run low inflation credibility and led to higher growth and FDI. Banking regulations were improved, western accounting standards were implemented, a better screening process was put in place, and monetisation of deficits was forbidden.

5.2 Structural reforms story

This story made the difference for some Latin American countries (see Bergoeing et al., 2002) and served as a good explanation for the relatively better economic performance of Chile compared to Mexico in 1980–2000. Structural reforms notion unites reform packages in the spheres of trade and fiscal policy, as well as privatisation, banking regulations and bankruptcy procedures. Of all the reforms, the ones that seem to have the biggest effect on the productivity, and their implementation to coincide with a recovery, are the banking regulations and bankruptcy procedures.

5.2.1 Trade policy

From the very beginning, the Baltic countries oriented themselves towards the EU, while Bulgaria was undecided whether to affiliate to Russia or to the Western Europe. Still, all countries went for large-scale trade liberalisation and elimination of tariffs after recommendations by the World Trade Organization (WTO) and IMF missions. In addition, the Baltic countries used to trade with Russians metals, which were quoted at below market prices. Thus, there were good profit opportunities for exporting Baltic firms and banks, and the latter entered. But Russia also moved to world prices and the arbitrage opportunities decreased. Banks had to find other profitable lending opportunities to sustain their position. Through this channel, a banking crisis had an effect on trade through trade credit arrangements and in this way put a strain on the budget.

5.2.2 Fiscal policy

In theory, tax reforms have an effect on capital accumulation and real wages. Tax rates also influence the incentive to work, and high rates do not translate automatically into a higher fiscal revenue. The reason behind this was the absence of an efficient tax administration. Thus, fiscal authorities in the transition countries decreased tax rates, which led to an increase in tax revenue. Moreover, in a fixed exchange rate regime, fiscal policy is stronger but you cannot expand it in a traditional way because it will undermine the credibility of that hard peg (which happened in Argentina in 2002). There is also an important lag between implementation of changes in government policy and their effect on productivity. In our case, the relatively short time span may be unable to catch this effect. Most of the countries are running balanced budgets, or small surpluses, as required by law in Estonia. Indeed, all four countries were running significant budget deficits during the banking crises: 5–6% in Lithuania in 1993–1994, for example, and even reaching 19% in Bulgaria in 1996.

5.2.3 Privatisation

There are some measurement issues with this story: first, the different privatisation programs in Bulgaria and the Baltic countries, which makes it hard to determine which one was the most successful. In addition, it is difficult to estimate the effect of privatisation since there was a lot of 'cherry-picking' on the investors' side: that is, *ex ante* profitable companies were privatised first. Nevertheless, there is a consensus among economists that privatisation leads to the establishment of a market relationship between firms and banks. It is also a way to deal with the bailing out problem: For the Baltics, for example, there was a common (mis)perception that banks are Russian-backed (Flemming and Talley, 1996). Thus, privatisation in the banking sector led to an opening of private banks and a closing down is inefficient state.

Banking crises have an important implication for the real side of the economy. Usually, the negative effect is in the slower recovery in production, and growth in general. That is due to people's loss of faith in the banking system, which impedes the intermediation of savings into investment. It is also the case that there was a change in the government after a banking crisis in all the Baltic countries and in Bulgaria.

Fleming and Talley (1996) note the major ways in which the banking system is exposed to stress: in all four countries, it is usually the enterprises that are the main borrowers. During the course of transition, they become subject to hard budget constraints and no longer receive government subsidies to cover their losses. In addition, inflation declines and real interest rates become positive, so it is harder on the firms' side to service their debts. Budget deficit in Bulgaria and Latvia ballooned, which decreased the demand for short-term Treasury bills by banks. As a result, the financing of the budget became problematic: the Latvian and Bulgarian Government, for example, also allocated credit denominated in foreign currency with public guarantee to the private sector. These funds were mostly wasted and now the government has to pay them back, which made the situation at hand even worse. This liability of the government led to inflation and a deterioration of banks' balance sheets, setting the ground for bank panics. Table 7 below provides detailed information on the deterioration of the quality of bank loans in Bulgaria, which inevitably happens when banks provide loans without the necessary screening process and sufficient collateral.

What the government can do is to establish an adequate legal framework for banking and financial services, develop effective bank supervision and regulation, free banks from political influence and foster effective bank management. Accounting based on Western standards is needed, and not one based on historical costs, as well as audits. The institutional demands that that rules of the game be properly specified, and a bigger governance role be given to the Central Bank and bank supervision. There is also the need to react very quickly before it becomes too late: Liquidation of banks, rehabilitation for those that are too big to fail – will have a serious negative effect on the banking system and the economy as a whole. The expansion of loans should not be so fast because it decreases the quality of the credits given, and there should be more careful screening in place.

After the introduction of the currency board in Bulgaria, there was a significant change in banks' behaviour. This was mostly due to the improvements in bank legislation and regulation: banks were required to report every large credit they gave and to better monitor the status of the loans outstanding. A loan was categorised now as

non-performing if the delay of payments was more than 30 days, which is much shorter period of time compared to 90 days previously.

Table 7 Credit portfolio of Bulgarian Commercial banks, 1993–1999 (in percent)

	1993	1994	1995	1996*
Standard	7.61	17.69	25.91	43.67
Doubtful Group A	82.75	66.88	54.55	33.89
Doubtful Group B	2.19	3.46	4.18	10.67
Loss	7.45	11.97	15.35	11.77
Provisions available	7.18	23.58	23.74	105.42
	1997*	1998*	1999*	
Standard	58.2	69.0	73.3	
Watch	8.6	10.0	9.3	
Substandard	5.9	5.6	2.2	
Doubtful	4.7	1.7	3.3	
Loss	22.7	13.6	12.0	

Note: *Banks in liquidation are excluded.

Source: Caporale et al. (2002)

5.2.4 Bankruptcy procedures⁴

In theory, the reform in bankruptcy procedures should increase the incentives for capital accumulation, as well as its efficiency, as was the case in Chile (Bergoeing et al., 2002). Higher productivity growth can be an effect of the right timing in privatisation, banking and bankruptcy procedures reforms. Before reforms, government policies distorted the functioning of the market mechanism. Resources were misallocated both within and between sectors. As a result of this, the economy goes below its production possibilities frontier (PPF), due to the fall in aggregate productivity. Moreover, it is not only those static negative effects, but also dynamic consequences dealing with the entry/exit decisions of firms.

Static effect is due to the discriminatory attitude towards government and private firms. The loss-making state enterprises are subsidised from the corporate tax revenue collected from the profit-making private enterprises. This is done in order to preserve the jobs of the employees in the public sector. In addition, government companies produce too much, while the private sector produces too little. On the aggregate level, production is lower than the one in a market environment. Resources are not used to their maximum efficiency and this is translated into lower productivity growth.

The connection with the banking story is that it was very often the case that the two sectors received credits at different interest costs. This was done either through a transfer from the budget to state companies, or simply private companies paid a higher interest cost to compensate for the below market rate loan given to the public enterprises. This credit policy also distorts capital allocation, as illustrated in the simple models in Bergoeing et al. (2002). In the short-run, capital is assumed to be fixed in quantity and its allocation among sectors is distorted by the credit policy in place. This misallocation of capital will decrease the incentive to take loans in order to invest in productive capacity

and will decrease the speed of capital accumulation. In this way, the lower level of capital in the economy will lead to a lower level of production.

The dynamic effects are connected to the firms' entry/exit decision: with government protection, high-cost producers that cannot cover their expenses are kept in the market. This has a discouraging effect on entry of privately-managed enterprises and also poses barriers to the entry of new firms. This is bad for the economy since usually private firms come with newer technology. Without innovation and learning, however, the losses in productivity over time rise in an exponential manner. Atkinson and Kehoe (quoted in Bergoeing, 2002) give an idea how big losses for the economy can be when they accumulate over time. In this sense, opening for trade exercises pressure on inefficient firms to close and leaves more room for private companies.

The countries' economic history has demonstrated that banking regulations and bankruptcy procedures have important effects on the real side of the economy, which acts through the productivity channel. More specifically, they affect the allocative efficiency of resources, both in a static and dynamic manner. The other stories either do not have a significant effect on TFP, or do not hold water in our case. This does not mean they are not worth discussing: debt and fiscal policy, for example, have important background effects.

6 Conclusions

In this paper, the cyclical fluctuations and productivity growth during transition in Bulgaria and the Baltic countries: Estonia, Latvia and Lithuania was explored. The present paper aimed at deepening the understanding of business cycles in transition countries and emphasising the importance of the financial system for the real side of the economy. This study estimated a simple RBC model, using the classical paper by King et al. (1988).

Both the model and data series show that the major drop in output was due to productivity, a result obtained also in Amaral and MacGee (2002), Hayashi and Prescott (2002), Kehoe and Ruhl (2003). In addition, the analysis performed showed that the fall in TFP may be due to the banking regulations. In addition, the timing of the banking reforms coincides with the improvement of the economic performance, as in Bernanke (1981), Bergoeing et al. (2002), Hopenhayan and Neumeier (2003).

The major limitation of the model is that it is too simplistic. Despite the simplicity, we were unable to find data on some of the variables and had to use estimates in our calibration. In addition, there is no government sector with distortionary taxation explicitly modelled, e.g., as in Baxter and King (1993) and McGrattan (1994). In future work, government shocks will be allowed for, which is expected to solve the shock persistence problem. A promising area of improvement is to go deeper into the microfoundations: the theoretical framework may incorporate, for the sake of richer analysis, additional complications such as factor underutilisation, variable labour supply and labour hoarding, wage and price adjustment, monopolistic markups, capital adjustment costs, and heterogeneity of capital.

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References

- Amaral, P. and MacGee, J.C. (2002) 'The Great Depression in Canada and the United States: a neoclassical perspective', *Review of Economic Dynamics*, Vol. 5, No. 1, pp.45–72.
- Baxter, M. and King, R.G. (1993) 'Fiscal policy in general equilibrium', *American Economic Review*, Vol. 83, pp.315–334.
- Benczur, P. and Rattai, A. (2005) 'Economic fluctuations in Central and Eastern Europe: the facts', *CEPR Discussion Paper* # 4846.
- Bergoeing, R., Soto, R., Kehoe, P.J. and Kehoe, T.J. (2002) 'Decades lost and found: Mexico and Chile since 1980', *Review of Economic Dynamics*, Vol. 5, No. 1, pp.166–205.
- Bernanke, B. (1981) 'Bankruptcy, liquidity, and recession', *American Economic Review*, Vol. 71, pp.155–159.
- Bernanke, B. and James, H. (1990) 'The gold standard, deflation, and financial crisis in the great depression: an international comparison', NBER WP 3488.
- Bernanke, B. (1995) 'The macroeconomics of the great depression: a comparative approach', *Journal of Money, Credit and Banking*, Vol. 27.
- Campbell, J.Y. (1994) 'Inspecting the mechanism: an analytical approach to the stochastic growth model', *Journal of Monetary Economics*, Vol. 33, pp.463–506.
- Caporale, G.M., Hristov, K., Miller, J.M., Nenovsky, N. and Petrov, B. (2002) 'The banking system in Bulgaria', in Zeljko Sevic (Ed.): *Banking Reforms in South-Eastern Europe*, Edward Elgar, Cheltenham, UK.
- Cerny, A. and Lazarova, S. (1994) 'Growth and business cycles in Czechoslovakia according to the basic neoclassical RBC model', *CERGE-EI Working Paper* 65.
- Cole, H.L. and Ohanian, L.E. (2000) 'Re-examining the contributions of money and banking shocks to the US Great Depression', *Federal Reserve Bank of Minneapolis Research Department Staff Report* 270.
- Cole, H.L. and Ohanian, L.E. (2001) 'New deal policies and the persistence of the Great Depression: a general equilibrium analysis', *Federal Reserve Bank of Minneapolis Research Department Working Paper* 597.
- Cooley, T.F. and Prescott, E.C. (1995) 'Economic growth and business cycles', in Thomas F. Cooley (Ed.): *Frontiers of Business Cycle Research*, Princeton University Press, Princeton, NJ.
- Doms, M.E. and Dunne, T. (1998) 'Capital adjustment patterns in manufacturing plants', *Review of Economic Dynamics*, Vol. 1, No. 2, pp.409–429.
- Fisher, J. and Hornstein, A. (2001) 'The role of real wages, productivity, and fiscal policy in Germany's great depression 1928–1937', Federal Reserve Bank of Chicago, WP 2001-07.
- Fleming, A. and Talley, S. (1996) 'The Latvian banking crisis: lessons learned', World Bank Policy Research Working Paper No. 1590.
- Friedman, M. and Schwartz, A.J. (1971) *A Monetary History of the United States, 1867–1960*, Princeton University Press.
- Ganev, K. (2005) 'Measuring total factor productivity: growth accounting for Bulgaria', *Bulgarian National Bank Discussion Paper* 52.
- Hayashi, F. and Prescott, E.C. (2002) 'The 1990s in Japan: a lost decade', *Review of Economic Dynamics*, Vol. 5, pp.206–235.

- Hopenhayn, H.A. and Neumeyer, P.A. (2003) 'The Argentine Great Depression 1975–1990', *Working Paper*.
- Kehoe, T.J. and Ruhl, K.J. (2003) 'Recent Great Depressions: aggregate growth in New Zealand and Switzerland', *Working Paper*.
- King, R.G., Plosser, C.I. and Rebelo, S.T. (1988) 'Production, growth and business cycles', *Journal of Monetary Economics*, Vol. 21, pp.195–232.
- King, R.G., Plosser, C.I. and Rebelo, S.T. (2002) 'Production, growth and business cycles: technical appendix', *Journal of Computational Economics*, Vol. 20, pp.87–116.
- Kydland, F.E. and Prescott, E.C. (1982) 'Time to build and aggregate fluctuations', *Econometrica*, Vol. 50, pp.1345–1370.
- McGrattan, E. (1994) 'The macroeconomic effects of distortionary taxation', *Journal of Monetary Economics*, Vol. 33, pp.573–601.
- Mihov, I. (2001) 'The economic transition in Bulgaria, 1989–1999', in Mario I. Blejer and Marco Screb (Eds.): *Transition: The First Decade*, The MIT Press, Cambridge, MA.
- Prescott, E.C. (1986) 'Theory ahead of business cycle measurement (1986)', *Federal Bank of Minneapolis Quarterly Review*, pp.9–22.
- Prescott, E.C. (2002) 'Prosperity and depression: 2002 Richard T. Ely Lecture', *Federal Reserve Bank of Minneapolis Research Department Working Paper* 618.
- Ravn, M. (2006) 'King-Plosser-Rebelo Gauss programmes', *Quantitative Macroeconomics and Real Business Cycles Codes, Software Component # 104*, available at <http://ideas.repec.org/c/dge/qmrbcd/104.html>.
- Romer, D. (2001) *Advanced Macroeconomics*, 2nd ed., McGraw-Hill/Irvin, New York, NY.
- Rostowski, J. (1996) 'Comparing two Great Depressions: 1929–33 and 1989–93', in S. Zecchini (Ed.): *Lessons from the Economic Transition: Central and Eastern Europe in the Nineties*, OECD and Kluwer Academic Publishers, Boston, p.590.
- Solow, R. (1957) 'Technological change and the aggregate production function', *Review of Economics and Statistics*, Vol. 39, pp.312–320.
- Uhlig, H. (1999) 'A toolkit for analyzing nonlinear dynamic stochastic models easily', in R. Marimon and A. Scott (Eds.): *Computational Methods for the Study of Dynamic Economies*, pp.30–61, Oxford University Press, Oxford.

Notes

- 1 $\lim_{C \rightarrow 0} \frac{\partial U(C_t, L_t)}{\partial C_t} = \infty, \lim_{C \rightarrow \infty} \frac{\partial U}{\partial C_t} = 0, \lim_{L \rightarrow 0} \frac{\partial U(C_t, L_t)}{\partial L} = \infty, \lim_{L \rightarrow \infty} \frac{\partial U(C_t, L_t)}{\partial L} = 0$
- 2 The regression for Latvia was run for industrial productivity only due to data limitations.
- 3 Both empirical and model data are H-P filtered, as in Kydland and Prescott (1982).
- 4 The exposition in this subsection relies heavily on the arguments provided by Bergoeing et al. (2002).

Appendix

All data were seasonally adjusted using X-11 multiplicative procedure, except for the net capital flows, for which X-11 additive was used because of the negative values in some of the observations. In addition, X-11 is a standard procedure in the literature. The trend was extracted using Hodrick-Prescott filter using the default $\lambda = 1,600$ parameter. Productivity is defined as real GDP divided by total employment, as done by Romer (2001) and Benczur and Ratfai (2005). For Latvia, GDP was divided by industrial employment due to data limitations.