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### **Tests for Normality in Linear Panel Data Models**

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# Tests for normality in linear panel data models

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**Abstract.** A new Stata command, xtsktest, is proposed to explore non-normalities in linear panel data models. The tests explore skewness and excess kurtosis allowing researchers to identify departures away from gaussianity in both error components of a standard panel regression, separately or jointly. The tests are based on recent results by Galvao, Montes-Rojas, Sosa-Escudero and Wang (2013), and can be seen as extending the classical Bera-Jarque normality test for the case of panel data.

Keywords: st0001, xtsktest, skewness, kurtosis, normality, panel data

#### 1 Introduction

The need to check for non-normal errors in regression models obeys to both methodological and conceptual reasons. From a strictly methodological point of view, lack of Gaussianity sometimes harms the reliability of simple estimation and testing procedures, and calls for either better methods under alternative distributional assumptions, or for robust alternatives whose advantages do not depend on distributional features. Alternatively, whether errors should be more appropriately captured by skewed and/or leptokurtic distributions may be a statistical relevant question *per se*.

The normality assumption also plays a crucial role in the validity of inference procedures, specification tests and forecasting. In the panel data literature, Blanchard and Mátyás (1996) examine the consequences of non-normal error components for the performance of several tests. Montes-Rojas and Sosa-Escudero (2011) show that non-normalities severely affect the performance of the panel heteroskedasticity tests by Holly and Gardiol (2000) and Baltagi et al. (2006). Despite these concerns the Gaussian framework is widely used for specification tests in the one-way error components model; see, for instance, the tests for spatial models in panel data by Baltagi et al. (2003) and Baltagi et al. (2007). Even though there is a large literature on testing for skewness and kurtosis in cross-sectional and time-series data, including Erickson and Whited (2010), Bai and Ng (2005), and Bera and Premaratne (2001), to cite a few of an extensive list results for panel data models are scarce. A natural complication is that, unlike their cross-section or time-series counterparts, in simple error-components models lack of Gaussianity may arise in more than one component. Thus, an additional problem to that of detecting departures away from normality is the identification of which component is causing it. Previous work on the subject include Gilbert (2002), who exploits cross-moments, and Meintanis (2011), who proposes an omnibus-type test for normality in both components jointly based on empirical characteristic functions.

Galvao et al. (2013) develop tests for skewness (lack of symmetry), kurtosis, and normality for panel data one-way error component models. These tests are important in practice because, in the panel data case, the standard Bera-Jarque test is not able to disentangle the departures of the individual and remainder components from non-Gaussianity. The tests are constructed based on moment conditions of the within and between transformations of the OLS residuals. These conditions are exploited to develop tests for skewness and kurtosis in the individual-specific and the remainder components, separately and jointly. The tests are particularly useful for the case where the number of individuals, N, goes to infinity, but the number of time periods, T, is fixed and might be small.

The proposed tests are implemented in practice using a bootstrap procedure. Since the tests are asymptotically normal, the bootstrap can be used to compute the corresponding variance-covariance matrices of the statistics of interest and carry out inference. In particular, the tests are implemented using a crosssectional bootstrap. We formally prove the consistency of the bootstrap method applied to our case of short panels.

The new command **xtsktest** implements a battery of tests to identify nonnormalities in standard error components panel models, based on recent results by Galvao et al. (2013). For standard regression models, the classical Bera-Jarque test (implemented in **Stata** with **sktest**), is a simple procedure that detects departures away from gaussianity in the form of skewness and excess kurtosis in the regression error term. A natural concern in the case of panel data models is to identify which error component (if not both) is the source of non-normalities. The proposed tests allows researchers to explore skewness and excess kurtosis in each component separately or jointly. In this context, the proposed procedure can be seen as extending the famous Bera-Jarque tests for the case of simple panel data models.

Section 2 reviews the results of Galvao et al. (2013) and presents the tests. Section 3 describes the **xtsktest** syntax. Next we illustrate the procedure by applying the new tests to an investment model studied by Fazzari et al. (1988). We conclude with practical suggestions on the proper use of the tests.

#### 2 Skewness and kurtosis in the one-way error components model

Consider the following standard panel data one-way error components model

$$y_{it} = x_{it}b + u_i + e_{it}, \quad i = 1, \dots N, \ t = 1, \dots, T,$$
(1)

where b is a p-vector of parameters, and  $u_i$ ,  $e_{it}$ , and  $x_{it}$  are copies of random variables u, e, and x, respectively (b does not contain a constant). As usual, the subscript *i* refers to individual, and t to time. Here  $u_i$  and  $e_{it}$  refer to the individual-specific and to the remainder error component, respectively, both of which have mean zero.

The quantities of interest are each component skewness,

$$s_u = \frac{E[u^3]}{(E[u^2])^{3/2}}, \text{ and } s_e = \frac{E[e^3]}{(E[e^2])^{3/2}},$$

and kurtosis,

$$k_u = \frac{E[u^4]}{(E[u^2])^2}$$
, and  $k_e = \frac{E[e^4]}{(E[e^2])^2}$ .

Galvao et al. (2013) construct statistics for testing for skewness and kurtosis in the individual-specific and the remainder components, separately and jointly. When the underlying distribution is normal, the null hypotheses of interest become  $H_0^{s_u}$ :  $s_u = 0$  and  $H_0^{s_e}$ :  $s_e = 0$  for skewness and  $H_0^{k_u}$ :  $k_u = 3$  and  $H_0^{k_e}$ :  $k_e = 3$  for kurtosis. Moreover, under normality, the null hypotheses for these cases are given by

$$H_0^{s_u\&k_u}: s_u = 0 \text{ and } k_u = 3,$$
  
 $H_0^{s_e\&k_e}: s_e = 0 \text{ and } k_e = 3.$ 

Following Galvao et al. (2013), the statistics for symmetry are  $\widehat{SK}_{u}^{(1)} = \widehat{E[u^3]}$  and  $\widehat{SK}_{e}^{(1)} = \widehat{E[e^3]}$  and those for kurtosis are  $\widehat{KU}_{u}^{(1)} = \widehat{E[u^4]} - 3(\widehat{E[u^2]})^2$ and  $\widehat{KU}_{u}^{(1)} = \widehat{E[e^4]} - 3(\widehat{E[e^2]})^2$ . Similar statistics can also be presented in a standarized way as  $\widehat{SK}_{u}^{(2)} = \frac{\widehat{E[u^3]}}{(\widehat{E[u^2]})^{3/2}}$  and  $\widehat{SK}_{e}^{(2)} = \frac{\widehat{E[e^3]}}{(\widehat{E[e^2]})^{3/2}}$  for symmetry and  $\widehat{KU}_{u}^{(2)} = \frac{\widehat{E[u^4]}}{(\widehat{E[u^2]})^2} - 3$  and  $\widehat{KU}_{e}^{(2)} = \frac{\widehat{E[e^4]}}{(\widehat{E[e^2]})^2} - 3$  for kurtosis. Each statistic is consistent and properly standarized follows a N(0,1) asymptotic law under the corresponding null hypothesis. However, they may differ in small samples. Moreover, tests for joint symmetry and kurtosis are constructed using  $(\widehat{SK}_{u}^{(j)})^2 + (\widehat{KU}_{u}^{(j)})^2$  and  $(\widehat{SK}_{e}^{(j)})^2 + (\widehat{KU}_{e}^{(j)})^2$ , j = 1, 2, each following a  $\chi_2^2$ asymptotic law under the corresponding null hypothesis. The tests implementation uses bootstrap to estimate the variances of the skewness and kurtosis test statistics in practice.

#### 3 The xtsktest syntax

#### 3.1 Syntax

The command syntax is:

xtsktest depvar [indepvars] [if] [in] [, reps(#) seed(#) standard

The simplest case is when it is executed only by invoking its name.

#### 3.2 Options

xtsktest supports the following options:

reps(#) specifies the number of bootstrap replications. The default is reps(50).

seed(#) specifies the seed for the random-number generator. See [R] set seed.

standard specifies if the skewness and kurtosis statistics are standardized by the estimated variance. The default is no standardization.

#### 3.3 Remarks

xtsktest can be used both as a standard command or as a post estimation command after ols or random effects model (see [R] regress and [XT] xtreg). In the first setting the command requires at least one variable in the varlist, while in the post-regression background varlist is not required. Example 1 shows the first option; examples 2 and 3 use xtsktest as a post estimation command.

#### 3.4 Saved results

xtsktest stores the following estimates saved to e():

```
Matrices
```

```
e(xtsk_text) s kewness and kurtosis test
results, one per row; first
column: point estimation,
second: standard errors,
third:p-values.
e(joint) joint skewness and kurtosis
test results, one per row;
first column: chi-squared
statistics, second:p-values.
```

#### 4 Empirical application: Investment equation

In this section, we apply the developed tests to Fazzari, Hubbard, and Petersen's (1988) investment equation model, where a firm's investment is regressed on an observed measure of investment demand (Tobin's q) and cash flow. This is one of the most well-known models in the corporate investment literature and we use this application as a way to illustrate our theoretical results. Following Fazzari et al. (1988), investment–cash-flow sensitivities became a standard metric in the literature that examines the impact of financing imperfections on corporate

investment (Stein 2003). These empirical sensitivities are also used for drawing inferences about efficiency in internal capital markets (Lamont 1999; Shin and Stulz 1998), the effect of agency on corporate spending (Hadlock 1998; Bertrand and Mullainathan 2005), the role of business groups in capital allocation (Hoshi et al. 1991), and the effect of managerial characteristics on corporate policies (Bertrand and Schoar 2003; Malmendier and Tate 2005).

Tobin's q is the ratio of the market valuation of a firm and the replacement value of its assets. Firms with a high value of q are considered attractive as investment opportunities, whereas a low value of q indicates the opposite. Investment theory is also interested in the effect of cash flow, as the theory predicts that financially constrained firms are more likely to rely on internal funds to finance investment (see e.g. Erickson and Whited (2000)). The baseline model in the literature is

$$I_{it}/K_{it} = \alpha + \beta q_{it-1} + \gamma C F_{it-1}/K_{it-1} + u_i + e_{it}, \qquad (2)$$

where I denotes investment, K capital stock, CF cash flow, q Tobin's q, u is the firm-specific effect and e is the innovation term.

We check for skewness and kurtosis in both u and e using the proposed tests. We are interested in testing for skewness and kurtosis for at least three reasons. First, testing normality plays a key role in forecasting models at the firm level. Second, asymmetry in both components is used for solving measurement error problems in Tobin's q. The operationalization of q is not clear-cut, so estimation poses a measurement error problem. Many empirical investment studies found a very disappointing performance of the q theory of investment, although this theory has a good performance when measurement error is purged as in Erickson and Whited (2000). Their method requires asymmetry in the error term to identify the effect of q on firm investment. Third, skewness and kurtosis by themselves provide information about the industry investment patterns. Skewness in u determines that a few firms either invest or disinvest considerably more than the rest, while kurtosis in u determine that a few firms locate at both sides of the investment line, that is, some invest a large amount while others disinvest large amounts too. Skewness and/or kurtosis in e show that the large values of investment correspond to firm level shocks.

We follow Almeida et al. (2010), who considered a sample of manufacturing firms (SICs 2000 to 3999) over the 2000 to 2005 period with data available from COMPUSTAT's P/S/T, full coverage. Only firms with observations in every year are used, in order to construct a balanced panel of firms for the five year period. Moreover, following those authors, we eliminate firms for which cash-holdings exceeded the value of total assets and those displaying asset or sales growth exceeding 100%. Our final sample consists of 410 firm-years and 82 firms. Because we only consider firms that report information in each of the five years, the sample consists mainly of relatively large firms.

To demonstrate the use of **xtsktest** to this application, we must first open the dataset and declare it to be panel data. See [XT] **xtset**.

<sup>.</sup> use investment.dta

```
. xtset idcode time
    panel variable: idcode (strongly balanced)
    time variable: time, 2 to 6
        delta: 1 unit
```

Consider first an OLS estimation of the effect of Tobin's q and cash flows on investment.

#### . regress investment tobinq cashflow

Source	SS	df	MS		Number of obs = $410$ F(2, 407) = $89.07$
Model   Residual    Total	.536747282 1.22632448	2 . 407 .	268373641 003013082		Prob > F = 0.0000 R-squared = 0.3044 Adj R-squared = 0.3010 Root MSE = .05489
investment	Coef.				
tobinq   cashflow   _cons	.0384663 .1117721 .0669764	.009402 .009614 .008787	22 4.09 22 11.63	0.000 0.000 0.000	.0199834 .0569492 .0928724 .1306718 .0497016 .0842512

Next, consider a one-way error components random effects model.

. xtreg investment tobinq cashflow, re

Random-effects GLS regression Group variable: idcode					of obs = of groups =	110
R-sq: within = 0.1014 between = 0.3583 overall = 0.2779				Obs per	group: min = avg = max =	5.0
corr(u_i, X)	= 0 (assumed	1)		Wald chi Prob > c	. ,	84.09 0.0000
investment	Coef.	Std. Err.	z	P> z	[95% Conf.	Interval]
cashflow	.0673706   .0824715   .0516002	.0115191	7.16	0.000	.0598944	
sigma_u sigma_e rho	.0380806 .03857635	(fraction c	of variar	nce due to	o u_i)	

The results show a positive and significant effect of both Tobin's q and cash flows on investment flows in both models. The random effects model also shows that there is considerable variation across firms in terms of unobservables. In fact, half of the variation is due to the firm specific component  $u_i$  and other half to the remainder component  $e_it$ . Note that the presence of firm specific effects determines that OLS standard errors are not correct, while the random effects are. Consider the use of **xtsktest** as an estimation command of the skewness and kurtosis of each component. The command can be implemented in the following three equivalent ways: as a standard command (example 1) or as a post estimation command after OLS (example 2) or random effects (example 3). We will consider the implementation with 500 bootstrap replications and without a random number seed (= 123) (default options have 50 bootstrap replications and no random number seed).

. \* Example 1, command mode xtsktest investment tobing cashflow, reps(500) seed(123) (running \_xtsktest\_calculations on estimation sample) Bootstrap replications (500) 50 100 150 200 250 300 350 400 450 500 Tests for skewness and kurtosis Number of obs 110

14 dillo O 1	01 000		110
Replica	ations	=	500

(Replications based on 82 clusters in idcode)

	Observed Coef.	Bootstrap Std. Err.	z	P> z		-based Interval]
Skewness_e   Kurtosis_e   Skewness_u   Kurtosis_u	.0000387 9.33e-06 .0000511 4.27e-08	.0000137 1.92e-06 .0000171 1.27e-06	2.81 4.87 2.99 0.03	0.005 0.000 0.003 0.973	.0000117 5.58e-06 .0000176 -2.44e-06	.0000656 .0000131 .0000847 2.53e-06
Joint test for Joint test for			chi2(2) = chi2(2) =			2 = 0.0000 2 = 0.0115

. \* Example 2, post-estimation command after OLS

- . regress investment tobing cashflow
   (output omitted)
- . xtsktest, reps(500) seed(123)
   (output omitted)

. \* Example 3, post-estimation command after OLS

. xtreg investment tobing cashflow, re

(output omitted)

. xtsktest, reps(500) seed(123)

#### (output omitted)

The screen output shows in the first column the observed coefficients of the four statistics (without standarization,  $\widehat{SK}_e^{(1)} = .0000387$ ,  $\widehat{KU}_e^{(1)} = 9.33e - 06$ ,  $\widehat{SK}_u^{(1)} = .0000511$  and  $\widehat{KU}_u^{(1)} = 4.27e - 08$ ) used for symmetry and kurtosis for each error component. The next columns show the standard errors computed by bootstrap replications, the z statistics, p-values and the 95% confidence intervals using the normal approximation. Finally, the lower part of the results output show the joint test for normality on each component of the error term and their p-values. The model shows that both components are asymmetric (with right symmetry), while only the remainder component e has excess kurtosis. Thus, while we expect the occurrence of large positive investment shocks ( $E[e^3] > 0$ ) these are systematic in some firms (i.e.  $E[u^3] > 0$ ). Asymmetry thus produces the rejection of the null hypothesis of normality in both error components, although the rejection is stronger for the remainder than for the firm specific component.

We also evaluate symmetry and kurtosis in each component using the standardized statistics,  $\widehat{SK}_{e}^{(2)}$ ,  $\widehat{KU}_{e}^{(2)}$ ,  $\widehat{SK}_{u}^{(2)}$  and  $\widehat{KU}_{u}^{(2)}$ . These can be implemented with the option standard.

. xtsktest investment tobinq cashflow, reps(500) seed(123) standard (running _xtsktest_calculations on estimation sample)
Bootstrap replications (500) + 1+ 2+ 3+ 4+ 5
400
450
Tests for skewness and kurtosis Number of obs = 41
Replications = 50
(Replications based on 82 clusters in idcode
Observed Bootstrap Normal-based
Coef. Std. Err. z P> z  [95% Conf. Interval

1	Coef.	Std. Err.	Z	P> z	[95% Conf.	Interval]
Skewness_e   Kurtosis_e   Skewness_u   Kurtosis_u	.6040947 3.645848 .9857612 .0220666	.1523494 .6932803 .2176725 .4963144	3.97 5.26 4.53 0.04	0.000 0.000 0.000 0.965	.3054954 2.287044 .5591309 9506917	.902694 5.004653 1.412391 .994825
Joint test for Joint test for	•		chi2(2) = chi2(2) =		Prob > chi Prob > chi	

Note: standardized coefficients

As expected, the results do not differ from those presented with the non standardized statistics. The numeric results however provide an easier interpretation of the excess kurtosis in the remainder component with a value of  $\widehat{KU}_{e}^{(2)} = 3.645848$  and firm-specific component  $\widehat{KU}_{u}^{(2)} = .0220666$ . The joint test for normality in u, however, provides a higher chi-squared value with a clearer rejection than in the previous examples using non-standardized coefficients.

#### 5 Conclusion

This paper implements tests for skewness/symmetry and kurtosis of the error components in linear panel data random effects models. The **xtsktest** procedure allows the evaluation of each error component third and fourth moments. This can be used as an alternative to the Bera-Jarque test in panel data models.

As discussed in the Introduction, checking for skewness and kurtosis in the error components plays an important role in testing and estimation procedures in linear panel data models. Deviations from symmetry and kurtosis of 3 ivalidate methods that are not robust to normality. Moreover, estimating third and fourth moments is also important for forecasting in panel data models (see Baltagi 2008, for a discussion).

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