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Did quantitative easing affect interest rates outside the US? New evidence based on interest rate differentials

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Did Quantitative Easing Affect Interest Rates Outside the US? – New Evidence Based on Interest Rate Differentials

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Ansgar Belke, Daniel Gros, and Thomas Osowski¹

Did Quantitative Easing Affect Interest Rates Outside the US? – New Evidence Based on Interest Rate Differentials

Abstract

This paper explores the effects of non-standard monetary policies on international yield relationships. Based on a descriptive analysis of international long-term yields, we find evidence that long-term rates have followed a global downward trend prior to as well as during the financial crisis. Comparing interest rate developments in the United States and the Eurozone, it appears difficult to find a distinct impact of the Fed's QE1 on US interest rates for which the global environment – the global downward trend in interest rates – does not account. Motivated by these results, we analyze the impact of the Fed's QE1 program on the stability of the US-Euro long-term interest rate relationship by using a CVAR and, in particular, recursive estimation methods. Using data between 2002 and 2014, we find limited evidence that QE1 caused a breakup or a destabilization of the transatlantic interest rate relationship. Taking global interest rate developments into account, we thus find no significant evidence that QE had an independent, distinct impact on US interest rates.

JEL Classification: C32, E43, E44, E58, F31, G01, G15

Keywords: Quantitative easing; unconventional monetary policies; time series econometrics; Cointegrated VAR (CVAR); recursive methods

January 2016

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1. Introduction

Huge adverse shocks generated by the financial crisis caused economic decline as well as turmoil on financial markets in 2008. Even after sharp reductions in (short-term) interest rates, central banks worldwide could not reduce the effects of the financial crisis substantially. With interest rates near zero, central banks lost their main policy tool because the zero lower bound proved to be a larger constraint than previously assumed by a large share of economists.¹ In response, central banks around the globe started unprecedented policy interventions, the so-called “nonstandard measures”.

Regarding the pure size of measures, the Fed has been the most active central bank by implementing several non-standard measures - most notably several rounds of Quantitative Easing (QE). The first round of QE was announced and launched in November 2008 mainly aiming at reducing the turmoil on financial markets and stabilizing the US economy. After the termination of QE1 in March 2010, QE 2 started in November 2010 followed by Operation Twist in September 2012 and an additional round of QE (QE3) in September 2012. Apart from the Fed, the Bank of England (BoE, 2009-2014) and the Bank of Japan (BoJ, since 2010) also made use of large-scale asset purchase programs in order to generate additional monetary stimulus at the lower zero bound.

While there have been noticeable differences² between the QE rounds conducted by the Fed and the central banks of other leading industrialized countries, a common aim of QE has been to put pressure on long-term yields.³ By reducing long-term yields, the Fed expected to further stimulate economic activity and prevent significant declines in inflation rates.⁴ Furthermore, another mechanism of QE runs via the exchange rate. However, long-term assets (and thereby interest rates) as well as exchange rates are often more affected by expectations about the future than by current economic conditions. Therefore, the announcement of a program can have a stronger impact than the actual implementation. With respect to the impact of QE on the nominal exchange rate, it is important to note that not each country can benefit from a nominal depreciation of their local currency, if several central banks start large-scale asset purchase programs at the same time.

¹ See Chung et al. (2012).

² For a comparison of QE designs in the USA see Rosengren (2015) and Fawley/Neely (2013).

³ Apart from Operation Twist in the 1960s, the BoJ launched a purchase program in March 2001. However, even the BoJ claimed that the policy has been largely ineffective.

⁴ However, the impact on yields is not clear a priori. If QE strongly increased expectation about future inflation and growth, yields should actually increase.

To measure the longer-term impact of QE on interest rates, interest rate relationships and exchange rates is inherently difficult, because one has to make many assumptions about how asset prices such as the exchange rate and the interest rate would have evolved in the absence of QE.⁵ Furthermore, there is evidence that the impact of QE crucially depends on the design of the program, the economic environment and a country's economic structure.⁶ For example, in the case of Japan (long-term) interest rates had been very low for a long period and seemed little affected by the increasingly aggressive asset purchases of the BoJ. However, the yen started to depreciate strongly after the asset purchase program was greatly increased in size and scope. By contrast, the effective dollar exchange rate moved little around the announcement and implementation dates of the different asset purchase programs operated over the last seven years by the Federal Reserve. In the case of the UK, one actually observed a trend-wise appreciation of the pound over the period during which the BoE bought large amounts of gilts, and there was apparently some impact, albeit only temporary, on long-term interest rates (Gros/Aldici/De Groen, 2015).

However, the real difficulties go even deeper. The majority of available studies just look at developments within the country undertaking QE and neglect the global environment. Global financial markets are highly integrated and (long-term) rates have been highly correlated across advanced economies, not only along a downwards trend, but also during cyclical ups and downs. Over most periods, rates have declined as much, sometimes more, in areas where QE was *not* undertaken. There is no sign that the fact that the ECB did not undertake bond purchases when they were undertaken by the US and the UK did in anyway prevent interest rates in the Euro area from following US rates downwards when only the US implemented QE (Gros/Aldici/De Groen, 2015).

According to the literature focusing on the national (but also international) transmission of QE shocks, authors generally pronounce two main transmission channels: the signaling channel and the portfolio-balance channel. Although both channels might explain a certain amount of the movements of financial variables in response to QE, we believe that the global comparative evidence is also and probably even more compatible with the view that QE did not “move” interest rates, but appeared to be important because major central banks undertook purchases when they realized that the recession caused by the financial crisis would be longer and more

⁵ For the counterfactual analysis in macroeconometrics with an empirical application to QE see Pesaran and Smith (2012).

⁶ See Rosengren (2015) for an assessment of how the design of the Fed's QE programs have affected their effectiveness. Fratzscher et al. (2013) for an empirical comparison of the effects of QE1 and QE2.

severe than anticipated. In this regard, it is also possible to explain the decline in interest rates before announcements of QE have been made by central banks. This observation might highlight that markets were quicker to revise their expectations and rates had thus come down before central banks started to buy assets. Therefore, market participants as well as central banks with their announcements of QE reacted to the same driving force – namely stronger adverse effects of the financial crisis than previously expected. In this regard, the prolonged weakness affected most of the developed world. Interest rates thus fell trend-wise in most advanced economies independently of whether QE was implemented by the national central bank (Gros/Aldici/De Groen, 2015).

Our view that the central banks program as well as reductions in interest rates had a common underlying source is somehow compatible with implications of the signaling channel, if one assumes that QE generated new information about the (future) state of the (global) economy for market participants. Following this argument, QE has been a signal that the crisis would be longer and more severe. Market participants reduced their expectation about future growth putting downward pressure on interest rates. However, if one follows this interpretation the fall in interest rates might have occurred anyway – at the latest when market participants would have revised their expectations about the severity of the crisis.⁷

Although several empirical studies credit QE with strong falls in US interest rates, rates fell as much in the Euro area, where QE was not undertaken (until recently). This finding might imply that several studies which neglect the global downward trend might give QE too much credit. The absence of a clear, distinct impact of QE episodes on interest rates and the exchange rate (e.g. the US), where QE was undertaken, should be puzzling. Although some (event) studies pronounce a very strong impact of QE on interest rates in the country where QE was implemented, international long-term interest rates remained to be highly correlated. Therefore, QE had little impact on interest rate differentials (USD versus Euro). This aspect also explains why asset purchases had little impact on exchange rates. If QE had had such a strong impact on interest rates as often asserted (i.e. in the order of 100 basis points according to several event studies⁸), one would have expected a strong impact on the exchange rate.

One way to test the hypothesis that large-scale asset purchases had a separate, identifiable impact on long term interest rates (in the currency area where they are undertaken) for which the

⁷ See Glick/Leduc (2011) for a similar interpretation.

⁸ See, for instance, Gagnon et al. (2011).

global downward trend in interest rates does not account is to estimate the cointegrated relationship between US and Euro area interest rates and to test whether one finds a structural break in this relationship around the time QE was undertaken in the US.

Apart from the hypothesis stated above, one has to admit that the overall effects of large asset purchase programs are still not well understood - even based on a domestic perspective. In this regard, QE shocks might differ from conventional interest rate shocks in normal times not only with regard their relative magnitude, but also by changing relationships between economic variables.⁹ It is well-known in theoretical and empirical literature that extraordinary and sustained macroeconomic policy actions can affect economic relationships and cause structural changes. When the Federal funds rate reached the zero lower bound and the Fed announced QE in November 2008, the Fed effectively changed its monetary policy variable from the Federal funds rate to its balance sheet size (Belke / Klose, 2013). In contrast to the pre-crisis era, it is now not possible anymore to measure monetary policy by simply looking at one interest rate.¹⁰ Motivated by this circumstance, several authors have argued in favor of econometric models which incorporate possible structural changes.¹¹ This aspect is further pronounced by Chen et al. (2013) who highlight that pre-crisis models could have become obsolete, as unconventional monetary policy might be transmitted in different ways compared to monetary policy actions in normal times.¹² To our knowledge, this paper is one of the first which tries to test empirically whether QE has changed economic relations on international financial markets.

The paper at hand uses a cointegration approach to analyze whether the Fed's QE1 has caused a structural change of the US-European interest rate relationship. As empirical approach, we use the Johansen procedure (CVAR) in order to estimate the long-run relationship between US and European interest rates also taking developments of the nominal exchange rate into account. We use monthly data from 2002 to 2014. After estimating the potential long-run relationship, we use recursive methods proposed by Johansen/Juselius (2006) in order to check for structural changes respectively parameter constancy. Our focus on QE1 results from two main reasons. Firstly, the general impression and empirical evidence on QE in the USA indicate that QE1 has been the program with the highest impact on financial variables. Therefore, if one assumed to

⁹ Analyzing the link between the monetary base and the money supply (defined as M1, M2 or M3) from a national perspective, it appears that the relationship has been completely broken since 2008 (Gros/Alcidi/De Groen (2015). See also McLeay et al. (2014).

¹⁰ See the growing empirical literature which tries to measure the monetary policy stance by using "shadow rates" (e.g. Lombardi/Zhu, 2014).

¹¹ See Kapetanios et al. (2012) and Baumeister/Benati (2012).

¹² Gambacorta et al. (2014) argue in a similar fashion.

find an independent effect on US interest-rates relative to Euro interest rates, QE1 might appear to be the natural choice. The second reason is related to the statistical approach of this paper. As recursive tests loose power to detect structural breaks at the end of the data sample, it would be difficult to reject the assumption of parameter constancy / structural constancy for Operation Twist, QE3 and partly QE2, even if a structural break was present.

The paper proceeds as follows. Section 2 contains a brief summary of the QE programs of the Fed and a descriptive analysis of their effects on domestic and global interest rates as well as exchange rates. Furthermore, a short review of empirical results of international QE transmission is provided. The data used as well as the estimation approach is depicted in section 3. The estimation process as well as the results are presented in section 4. Conclusively, section 5 sums up our results and provides an outlook for further research.

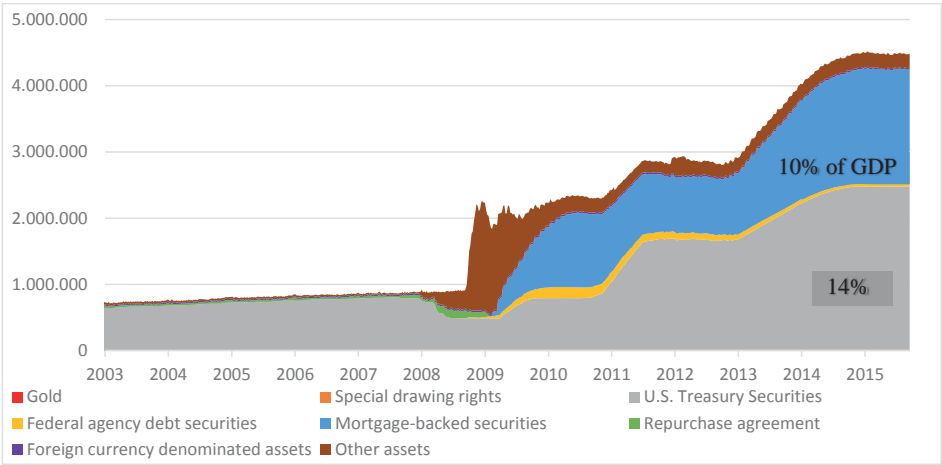
2. Quantitative Easing and global financial markets

2.1 Purchase programs in the USA

In November 2008, the Federal Reserve announced the first round of QE (QE1). These purchases included government sponsored enterprise debt (GSEs) and agency mortgage-backed securities (MBS) of up to \$600 billion. After announcing the willingness to extend the program in January 2009, the FOMC decided to purchase an additional \$750 billion in (agency) MBS, \$100 billion in agency debt and also started to purchase long-term Treasury securities worth of \$300 billion in March 2009. In total, the Fed purchased asset worth of \$1.75 trillion between November 2008 and March 2010, an amount twice the magnitude of total Federal Reserve assets prior to 2008.

In October 2010, the FOMC announced the second round of QE (QE2). It contained purchases of \$600 billion of Treasuries and was finished in June 2011. A few months later, the implementation of a maturity extension program, the so-called ‘Operation Twist’ (OT), was launched. By purchasing \$400 billion of Treasury bonds with maturities of 6 to 30 years and selling bonds with maturities less than 3 years, the FOMC intended to extend the average maturity of the Fed’s portfolio. Eventually, the third round of QE (QE3) started in September 2012. It targeted a monthly purchase of \$85 billion through the purchase of mortgage-backed securities (\$40 billion) and longer-term Treasury securities (\$45 billion). In contrast to the other programs, the continuation of QE3 was tied to improvement in the labor market. Overall, the Fed balance sheet increased by about \$3.5 trillion (roughly 20% of GDP).

Figure 1: Federal Reserve Balance Sheet (\$ mil.)



Note: Percentage refers to 2014 GDP.
Source: Federal Reserve.

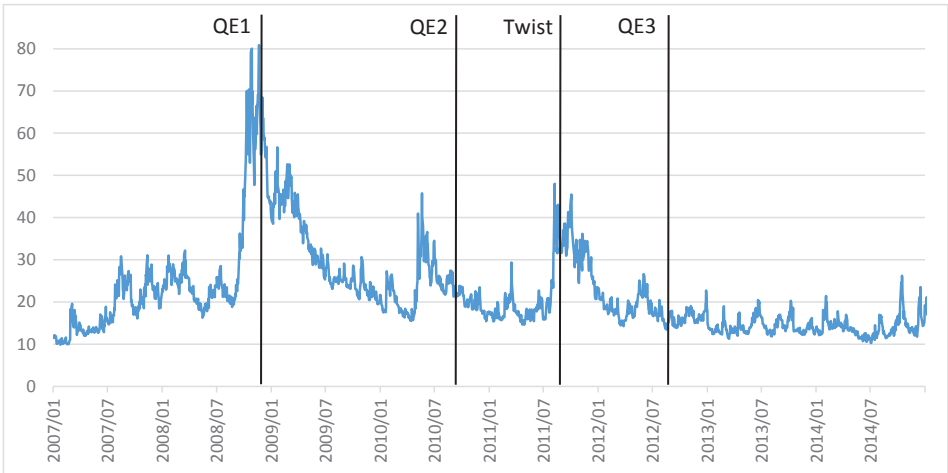
As shown in Figure 1, the balance sheet of the Federal Reserve has now reached over \$4 trillion, or close to 25% of GDP. Two assets clearly dominate on the asset side: Treasury securities and federal agency securities. The latter are all guaranteed by the Federal government of the United States. It is thus formally true that the Federal Reserve has intervened in the market for securitized mortgages, but it has bought only securities guaranteed by the government. In terms of the evolution of the balance sheet, one can clearly see the impact of QE 1, 2 and 3.

The Fed’s QE programs did not only differ in terms of their concrete design, but also with regard to the underlying economic environment during the time of their implementation. In this regard, Figure 2 presents the CBOE Volatility Index which is one of the most common measures of sentiment and systemic risk of the US stock market. Apparently, the perception of systemic risk differed strongly over time and therefore also between the starting points of the QE programs.

As expected, the highest amount of market uncertainty arose after the beginning of the financial crisis marked by the bankruptcy of Lehman in September 2008. While markets stabilized in 2009 and perception of systematic risk remained on an overall low level in the following years, two peaks of uncertainty are observable in Mai 2010 and August 2011. While the increased perception of systematic risk in Mai 2010 corresponds to the beginning of the European debt crisis, the second peak can be linked to Standard & Poor’s downgrade of the US credit rating.

Comparing the levels of uncertainty around the starting points of the QE programs, one can conclude that QE1 was implemented in a completely different environment compared to QE2 and QE3. While QE1 was implemented during the height of the crisis and therefore in an environment of huge uncertainty, QE 2 and QE3 were introduced by the Fed when financial markets had already stabilized. This aspect should be kept in mind when QE programs are compared and might also relate to the general assessment of several empirical papers that QE1 has been the most effective program of the Fed.

Figure 2: CBOE Volatility Index (VIX)



Source: Chicago Board Options Exchange (CBOE)

2.2 Literature review

While numerous empirical papers focus on domestic effects of QE, empirical evidence on international effects is still growing. In this section, we provide a survey of the literature which has attempted to quantify the effects of the Fed’s QE programs. As our paper focuses on the effects of QE on interest rate relationships, the following literature review primarily focuses on impacts on financial markets – interest rates and exchange rates.

According to current empirical studies, the general impact of large-scale asset purchase programs seems to vary considerably across countries or regions and also depends on the time and circumstances of their implementation. For the US, it looks like QE1 was the most effective in influencing financial markets, unemployment and inflation, while QE2 was far less effective. As of today, the overall effects as well as the magnitude of such shocks are highly uncertain.

Two main sources of uncertainty might explain large differences regarding the results of empirical estimations which try to estimate the impact of QE. Implemented as a direct response to the financial crisis, it appears to be extremely difficult to separate between effects of large-scale purchase programs and financial markets as well as macroeconomic conditions. Secondly, the majority of estimation methods and models rely on strong assumptions (for instance, about the transmission mechanisms of QE). Changing the assumptions might heavily influence the results. In relation to this point, Rudebusch et al. (2007) show that although there is no structural relationship between the term premium and GDP, a reduced-form empirical analysis supports the existence of an inverse relationship between the term premium and real economic activity.

The aspects mentioned appear to be especially relevant for several event studies which tend to find very large effects of QE compared to studies using different empirical frameworks. The two main drawbacks of event studies are heavy assumptions about the identification of monetary policy shocks and the focus on a very short period of time. In this regard, the standard event study methodology does not provide an estimate of the persistence of a monetary policy shock (Wright, 2011).¹³

Although the aim of QE was to support the economic development as well as the performance of the labor market in general, a large share of studies focus on its effect on long-term yields (especially Treasury bond yields). Regarding the effects of QE on domestic interest rates, the general consensus is that QE (especially QE1) had a reducing effect on US medium and long-term yields. Gagnon et al. (2011) investigate the effects of QE1 by using event study as well as time series methods. They find that the cumulative effect of LSAP announcements on yields of US Treasury bonds as well as US agency debt declined up to 150 basis points. By scaling the Fed purchases to '10-year equivalents', the authors measure the amount of duration which the Fed removed from the market. Across the three asset classes which were purchased during QE1, the purchases account for more than 20% of the total outstanding 10-year equivalents. Gagnon et al. (2011) argue that by reducing the net supply of assets with long duration, the program was successful in reducing the term premium by 30 to 100 basis points. In accordance with their results, the authors highlight the importance of the portfolio balance channel relative to the signaling channel.

¹³ See Hamilton (2011) for several critical remarks on measuring the effects of QE by using event studies.

While Christensen/Rudebusch (2012) find similar cumulative reductions using an event study, their empirical results however stress the importance of the signaling channel.¹⁴ Wright (2011) generates interesting insights using a structural VAR with daily data to identify monetary policy shocks. While he finds significant effects on long-term yields, these effects die off quite fast, with an estimated half-life of two month. These results might put the very large effects of event studies somehow into perspective.

Apart from event study methodology, further evidence is presented by Hamilton/Wu (2012) who use a term structure model to predict the effect of a change in the central bank's asset structure (short-for-long-term debt swap) and also indirectly the effect of buying \$400 billion in long-term Treasuries.¹⁵ Their results are much lower compared to event studies mentioned, as they find that such a policy would cause a reduction of the ten-year rate of (only) 13 basis points. Similar results have been obtained by Neely (2014) and Meyer/Bomfim (2010).¹⁶

Chung et al. (2011) find effects which are not negligible. Based on counterfactual model simulations, they find that the past and projected expansion of the Federal Reserve's securities holdings since late 2008 are roughly equivalent to a 300 basis point reduction in policy interest rates (since 2009 through 2012). Model simulations suggest that the additional stimulus provided by the purchases has kept the unemployment rate at a lower level (1½ percentage points by 2012) than what it would have been in the absence of the purchases and also argued that the asset purchases have probably prevented the US economy from falling into deflation.

Liu et al. (2014) find smaller effects. By using a change-point VAR model, they estimated that the Fed's asset purchase program reduced 10-year spreads by an average of 90 basis points over the crisis period. Without the asset purchase program, the unemployment rate was estimated to have been 0.7 percentage points higher and inflation, on average, 1 percentage point lower in 2010.

Regarding the effects on international financial markets, the majority of paper find cross-border effects as well as effects on exchange rates. Fratzscher et al. (2013) examine the international effects of QE1 and QE2. They find that QE1 was effective in lowering sovereign yields and

¹⁴ Further studies using event study methodology: Krishnamurthy/Vissing-Jorgensen (2011), D'Amico/King (2013).

¹⁵ The purchase amount roughly corresponds to the amount of Treasury bonds bought during QE1.

¹⁶ For further evidence, see Krishnamurthy/Vissing-Jorgensen (2011).

raising equity markets in the US and abroad. According to their results, QE1 might have generated a safe haven effect causing a strong global rebalancing of portfolios out of emerging markets and into US equity and funds and thereby putting upward pressure on the US Dollar (USD). However, regarding the effects of QE2, the authors find that this program has overall been ineffective in lowering yields worldwide and has caused sizeable capital outflows, mainly into emerging economies, and thereby marked a USD depreciation.

Neely (2013) puts more weight on the effects of QE on Treasury yields of developed countries.¹⁷ Using an event study as well as a portfolio-balance model, Neely (2013) finds substantial evidence that QE1 announcements have reduced sovereign yields in the US and abroad. Furthermore, Neely (2013) finds significant evidence that QE has generated a general depreciation of the USD. Bauer/Neely (2015) use dynamic term structure models in order to uncover whether international yields have declined as a results of signaling or portfolio-balance effects. They find that the relative importance of the signaling channel increases with an economy's sensitivity to signals from conventional US monetary policy. Consistent with the notion that Canada is highly sensitive to US monetary policy, the authors find large signaling effects for Canadian Treasury yields. For Australian and German Treasury bonds, the others find especially large portfolio balance effects.

2.3 Effects of the Fed's QE on interest rates and exchange rates and inflation

The aim of large-scale asset purchases is to lower long-term interest rates. Given that short-term rates are already at the zero lower bound, this amounts to a flattening of the yield curve. Furthermore, QE might also work by increasing inflation expectations and thereby decreasing real interest rates.

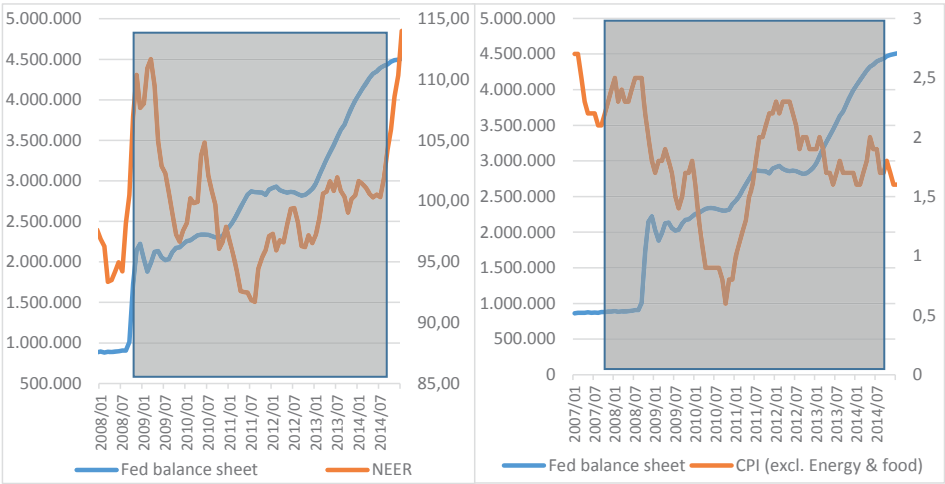
Keeping the aims of QE in mind, this section focuses on the short- and more medium-term evolution of interest, exchange and inflation rates around major QE operations. Therefore, this descriptive approach is somehow in contrast to the summarized findings of the academic literature presented in the previous chapter, which usually adopts a shorter-term view and state of the art econometric methods. We take this approach in order to check whether large-scale asset purchases had an observable impact on key variables.

Figure 3 and 4 show the evolution of the central bank balance sheets and the evolution of the nominal effective exchange rate and that of inflation for the United States and the Euro area.

¹⁷ Neely uses data from the following countries: USA, Australia, Germany, Japan and UK.

Shaded areas indicate the QE episodes considered here. In addition, two tables show the short-term impact of the announcement as well as the conduct of the Quantitative Easing programs on the long-term interest rates as well as the nominal and inflation-adjusted real exchange rates.

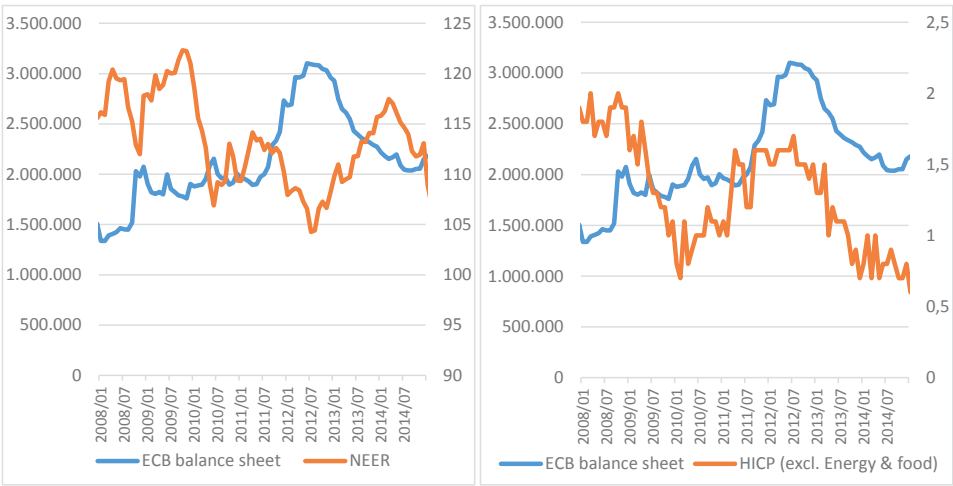
Figure 3: US: Central bank balance sheet, exchange rate and inflation



Source: Federal Reserve.

For the US, it is difficult to detect any observable impact of the QE programs on either the exchange rate or inflation. The two panels of Figure 3 show that both the exchange rate and inflation underwent large swings, which are seemingly unrelated to the various rounds of asset purchases by the Federal Reserve. The USD had been depreciating sharply before the onset of the financial crisis. Starting in April 2008, the USD started to appreciate strongly and continued to do so, even after QE1 was implemented in November 2008. The peak was reached in March 2009 (after a rough doubling of the monetary base under QE1). A phase of USD weakness followed, which only partially coincided with QE2 until 2011. Since then the USD has appreciated trend-wise despite further tremendous increases in the balance sheet of the Federal Reserve under QE3. With regard to the relationship between QE and inflation, it is difficult to find a clear impact of QE on inflation. Inflation continued to fall for about two years after the start of QE1, then reversed in coincidence with QE2, but then again trended downwards despite QE3 being implemented.

Figure 4: Euro area: Central bank balance sheet, exchange rate and inflation



Source: European Central Bank and Eurostat.

For the Euro area, the link between the central bank’s balance sheet and both the exchange rate and inflation appears to be stronger and more persistent (see the two panels of Figure 4). This is surprising since there has been no QE in the Euro area (until now), and the ECB could influence the size of its balance sheet only indirectly via its offers of long-term lending to banks at favorable rates.

Next, we focus on the impact of QE on long-term interest rates. We adopt an approach which focuses on the overall effects of the QE programs. Table 1 provides an overview of the impact of QE on long-term interest rates. The column entitled ‘change’ shows the difference between the long-term interest rate one quarter before the start of the actual asset purchases and the rate one quarter after the start of the asset purchase. This variable should thus capture both the announcement effect and the impact of the initial implementation.¹⁸ As three of the four entries in this column are negative, one can conclude that overall QE had the intended impact of reducing long-term interest rates. QE2 marks the only exception as the long-term interest rates actually increased.

¹⁸ We assume that the market had learned enough about the actual impact of the asset purchases after one quarter to anticipate correctly the rest.

Table 1: Impact of Quantitative Easing programs on interest rates

	Long-term interest rate (%)				
	Before	At start	After	Change	Compared to Euro area (core)
Euro area					
PSPP (March 2015)	1.6
United States					
QE1 (Nov 2008)	3.9	3.3	2.7	-1.1	0.1
QE2 (Nov 2010)	2.8	2.9	3.5	0.7	0.0
Twist (Sept 2011)	3.2	2.4	2.0	-1.2	0.0
QE3 (Sept 2012)	1.8	1.6	1.7	-0.1	-0.1

Note: The data before, at and after the start of the Quantitative Easing programs refer, respectively, to the quarterly averages for the quarter before the start of the programs, the quarter in which the programs started and the quarter after the start of the programs. Note that long-term interest rates refer to average government bonds maturing in about ten years published by the OECD. The Euro area core is proxied by the long-term interest rates for Germany. Source: OECD.

Table 1 also provides in the last column the evolution of the interest rate differential, i.e. the difference between US and core Euro area interest rates. This column shows entries that are mostly close to zero indicating no change of the interest rate differential around the announcement and introduction of QE programs.

Table 2: Impact of Quantitative Easing programs on exchange rates

	Nominal effective exchange rate (index 2010=100)			
	Before	At start	After	Change
Euro area				
PSPP (March 2015)	100.3
United States				
QE1 (Nov 2008)	95.8	106.0	108.8	13.0
QE2 (Nov 2010)	100.7	97.4	96.1	-4.6
Twist (Sept 2011)	93.6	94.2	97.8	4.3
QE3 (Sept 2012)	99.0	98.9	97.5	-1.5

Note: The data before, at and after the start of the Quantitative Easing programs refer, respectively, to the quarterly averages for the quarter before the start of the programs, the quarter in which the programs started and the quarter after the start of the programs. The nominal effective exchange rates (NEER) are the three-month averages of the BIS effective exchange rate indices. An increase in the NEER means that the currency has appreciated in nominal terms.

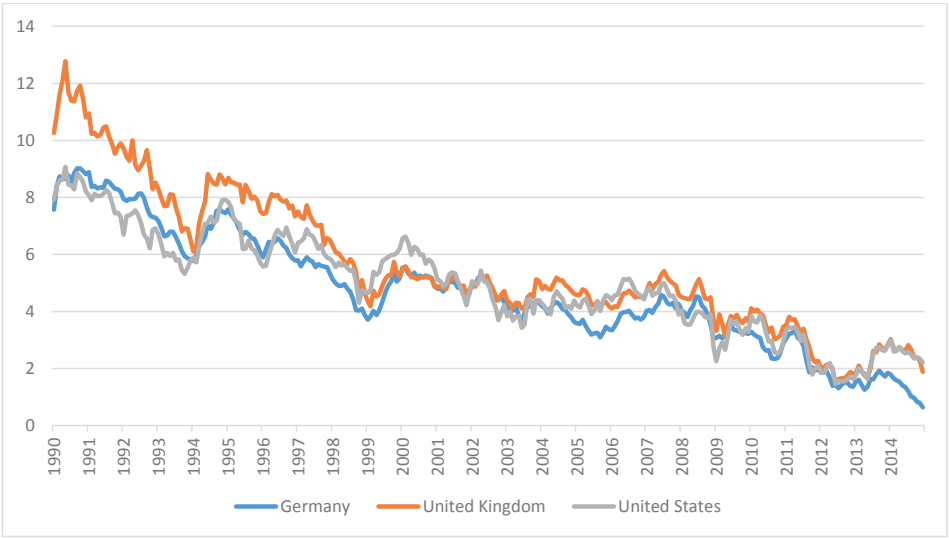
Source: Authors' calculations based on BIS.

Table 2 provides similar information on the reaction of the (effective nominal) exchange rate around major QE episodes. The column change again shows the (percentage) difference between the nominal effective exchange rate one quarter before and one quarter after the start of the asset purchases to illustrate the combined announcement and implementation effect. As a negative sign indicates a depreciation of the exchange rate, QE1 enters with the wrong sign, in the sense that the exchange rate appreciated (although one would expect QE to result in a depreciation).

2.4 Global financial markets and national QE

So far our analysis has followed the usual approach of looking for a link between asset purchases and financial market variables at the national level. In reality, however, financial markets in advanced countries are very open and highly integrated. This implies that one should not just look at US financial variables when trying to measure the impact of QE. However, disentangling the impact of QE in globally integrated financial markets is much more difficult, as one needs to adopt a comparative approach.

Figure 5: Long-term interest rates in major currency areas since 1990



Source: OECD.

A first key observation is that in reality (long-term) interest rates have followed a common long-term trend across major currency areas. Global financial markets are highly integrated and (long-term) rates have been highly correlated across advanced economies, not only along a

downward trend, but also during cyclical ups and downs, as illustrated in Figure 5. The correlation is too tight and has lasted too long to be just a coincidence (Gros/Aldici/De Groen, 2015). The most natural interpretation is that there exists a global capital market which is integrated across currency boundaries. Short-term interest rates are determined by central banks directly and can thus deviate strongly whenever the policy stance is different. Since 2009 short-term interest rates have basically been equal to zero in both the US and the Euro area, but long-term (here 10-year) rates have fluctuated, albeit around a clear common downwards trend.

The effectiveness of large asset purchases by the Federal Reserve should thus not be measured simply by the associated fall in US interest rates, *but a fall in the interest rate differential between the US and the Euro area (or other major markets)*. However, if one uses this metric, one must conclude that large asset purchases by the Fed have failed to have a differential impact on the US. Table 3 shows that over most QE periods rates have declined as much, and sometimes more, as they have in areas where QE was not undertaken. The small size of the changes in the interest-rate differentials is striking. For the US, no QE episode is associated with a change in the interest rate differential of more than 0.1%.

Table 3: Counterfactual impact of Quantitative Easing programs on long-term interest rates

Long-Term interest rate (%)		
	Change ($T_{+1q} - T_{-1q}$)	Compared to Euro area (core)
United States		
QE1 (Nov 2008)	-1.13	0.06
QE2 (Nov 2010)	0.67	-0.05
Twist (Sept 2011)	-1.16	0.00
QE3 (Sept 2012)	-0.12	-0.07

Note: The figures represent the decline/increase in the long-term interest rates in the period around the start of the Quantitative Easing programs. The changes are calibrated deducting the average interest rate in the quarter after the start minus the average of the quarter before the start. Long-term interest rates refer to average government bonds maturing in about 10 years published by the OECD. The Euro area core is proxied by the long-term interest rates for Germany.

Source: Own elaboration based on data from OECD.

Table 3 shows the movement of the transatlantic interest rate differential just around major large-scale asset purchases in the US. Figure 6 shows the evolution of the long-term (10-year) interest-rate differential between the US and the Euro area (proxied by the main riskless rate, i.e. the German rate). Since the ECB did not undertake QE until only at the very end of this

period, one would expect that the repeated round of large-asset purchases by the Federal Reserve should have resulted in a lowering of long-term US rates relative to Euro area rates (i.e. the line should have gone up). However, the opposite has been the case: US rates increased relative to Euro area rates if one compares the period just before QE1 (say May 2008) to January 2014 (i.e. long before it could be anticipated that the ECB would also eventually engage in large-scale purchases of government bonds). Over this period, the Federal Reserve has bought bonds worth in total over 20% of US GDP, but US interest rates actually increased (slightly) relative to Euro area rates.

Figure 6: Transatlantic long-term interest rate differential from 2007



Note: The difference between the German and US long-term interest rates is calculated by deducting the US rate from the German rate. Long-term interest rates refer to the monthly average government bonds maturing in about ten years. The vertical lines indicate the announcements by the Fed of the different quantitative-easing measures. Source: Own elaboration based on data from OECD.

This finding has important policy implications. The ECB in particular has been criticized for not having undertaken asset purchases earlier, and it was even argued that one reason the absence of a common fiscal agent for the Euro area is so important is that it has much delayed the decision of the ECB to undertake large purchases of public sector bonds. However, there is no sign that the fact that the ECB did not undertake bond purchases when they were undertaken by the US (and the UK) did in any way lead to higher interest rates in the Euro area. The ECB did of course undertake other ‘unconventional’ monetary policy operations, but these were confined to providing more liquidity to the banking system with the longest maturity being (until recently) 3-year operations. It is thus difficult to explain the co-movement of US and Euro area rates with similar monetary policy operations.

As we have stated in the introduction, we believe that the severity of the crisis and economic recession in developed economies led to (further) reductions in long-term interest rates across countries along the downward trend. In this regard, QE has only been a reaction to the crisis, but did not itself reduce interest rates. The small and non-persistent impact of US QE on the interest rate differential as well as the limited impact on exchange rates and inflation point in that direction as well. As we do not find an independent, separate effect of the US QE on the US economy which cannot be related to the global downward trend, the global comparative evidence thus suggests that several studies might overestimate the impact of QE. However, as analysis so far has solely focused on descriptive methods, we will analyse the relationship between US interest rates, European interest rates and the nominal exchange rate in a more sophisticated econometric approach in the chapters 3 and 4.

2.5 Fiscal policy, debt management and the portfolio balance channel

Before we turn to the main analysis of this paper, we would like to discuss one additional aspect which is closely related to the effectiveness of QE: fiscal debt management. Regarding the main channel of transmission, several academics as well as officials have highlighted the importance of the portfolio-balance channel.¹⁹ This transmission channel is based on the preferred habitat and imperfect asset substitutability theory and predicts that a reduction in the net supply of a given asset should in fact reduce its term premia and thereby its yield (D'Amico/King, 2012). In this regard, if the central bank buys large amounts of long-term (Treasury) assets, the central bank shortens the maturity structure of debt instruments that private investors have to hold, changing the relative net supplies and thereby reducing long-term interest rates. Empirical as well as theoretical papers mainly find evidence that the portfolio balance channel works.

However, the theoretical approach assumes that the (fiscal) debt management is exogenous and that it does not respond to measures taken by the central bank and therefore does not change its behavior. Greenwood et al. (2014) analyses the debt management of the US Treasury during the QE rounds and highlights that the fiscal side tended to supply the markets with longer maturity than during normal times / before the crisis. Regarding the supply of long-term government debt, the authors show that the amount of government debt with a maturity over 5 years held by the public (excluding the Fed's holding) has actually risen from 8 percent of GDP in 2007 to 15 percent in 2014. Focusing on the volume of 10-year duration equivalent debt, the stock has actually doubled from 13 percent of GDP to 26 percent over the same interval. Despite

¹⁹ See Yellen (2011) and Bernanke (2012).

massive asset-purchase programs by the Fed, the pressure to absorb (long-term) government debt has increased rather than decreased since the beginning of the crisis.

In this regard, the central bank and the fiscal side have been pushing in opposite directions with debt management policies at least partly offsetting the impact of monetary policy. Analyzing the reasons, Greenwood et al. (2014) find that roughly two thirds of the increased supply of long-term Treasury debt can be related to the tremendous increase in outstanding debt due to large deficits in the recent years, while the remaining one-third is due to the Treasury's active policy of extending the average maturity of government debt. The net result of these two opposing policies has overall been still a substantial increase in the longer-term securities held by the public: the fiscal deficits combined with the lengthening of maturity by the debt management office have increased the supply (measured in the equivalent of 10- year bonds) by close to 30% of GDP. But the various rounds of asset purchases by the Federal Reserve took about 15% of GDP from the market. Therefore, about one-half of this increase in longer-dated US federal securities has been undone by the various rounds of asset purchases of the Federal Reserve.

Greenwood et al (2014) also document that the weighted-average duration of federal debt securities issued by the Treasury has increased from about 4 years in 2008 to 4.6 years in late 2014. However, if one aggregates the Treasury and the Federal Reserve, the (weighted-average) duration has actually fallen to 2.9 years (Greenwood et al., 2014, p. 11) – a reduction of 1.7 years. This lower effective average maturity of the US public (federal) debt might now become relevant as the Federal Reserve is about to start increasing rates. The increase in rates will lead to a higher cost of debt service more quickly than if the duration of the public debt had been at the 4.6 years, which apparently was the target of the Treasury since 2008.

Greenwood et al. (2014) conclude that the common impression that the Fed asset purchases reduce long-term interest rates through the portfolio-balance effect might be wrong, as “the totality of policy has raised rather than reduced the quantity of long-term government debt held by private investors.” In this regard, it is argued that the fiscal sector's policy reaction has crowded out the portfolio-balance effects of QE which should have theoretically been the case.

We now turn to our econometric tests of the QE impact.

3. Data and empirical approach

With regard to our research objective, interest rate measures have to be chosen for the US as well as for the Euro area. The choice for the US is straightforward, as Treasury bond yields appear as the most common choice. As interest rate measure for the Euro area, German Treasury bond yields are used. We argue that German Treasury bonds are considered to be the least risky bonds in the Euro area. In this regard, we hope to avoid distortions of the cointegration relationships which might be generated by rising risk-premia of other sovereign bonds in the Euro area in presence of the European debt crisis. Furthermore, the use of Treasury bond yields is also motivated by the fact that they are often regarded as benchmark for domestic interest rates. As measures of interest rates, we use ten-year Treasury bonds yields. This implies that we focus on long-term interest rate measures.²⁰ We further include the nominal exchange rate (USD/Euro) in our estimations. We employ monthly end-of-period data between 2002:01 and 2014:12. For estimation purposes, we use the logarithm of the exchange rate variable.

Figure 7: Nominal exchange rate (\$/€)

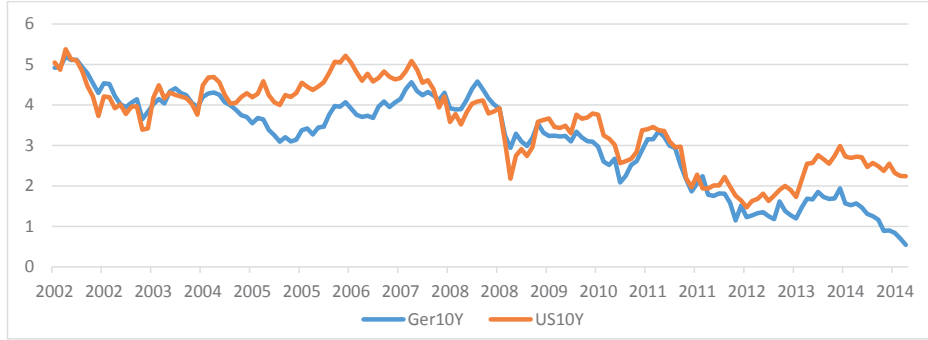


Source: Thomson Reuters Datastream

The time series are presented in Figures 7 and 8. Regarding the nominal exchange rate, it is once again difficult to observe a clear impact of the Fed's QE rounds. After November 2008, the exchange rates does show a certain amount of volatility, but appears to fluctuate around a constant level. The Treasury bonds yields presented in Figure 8 once again show a downward trend in interest rates as well as a strong correlation between both interest rates. Once again, no clear impact of QE is visible.

²⁰ We performed similar estimations using five-year and seven-year yields. Overall, we obtained almost identical results.

Figure 8: Treasury bond yields – USA and Germany



Source: Thomson Reuters Datastream

The econometric framework applied is a cointegrated vector autoregressive (CVAR) model which allows us to model the impact of domestic interest rate shocks on foreign interest rates as well as the exchange rate while taking care of the feedback between the variables. Our choice is also based on the CVAR's feature to avoid an a priori division of variables into exogenous and endogenous. As we include interest rate measures for the US and the Euro area as well as the nominal exchange rate, any ex-ante causality classification would be arbitrary. The basic representation is the p -dimensional vector autoregressive model with Gaussian errors ($\varepsilon_t \sim iid N(0, \Omega)$):

$$X_t = A_1 X_{t-1} + \dots + A_k X_{t-k} + \varphi D_t + \varepsilon_t, \quad t = 1, \dots, T \quad (1)$$

where X_t is a vector containing the variables of interest; D_t is a vector of deterministic variables containing constants, linear trends and dummy variables. Reformulating the model in an error-correction form allows to distinguish between stationarity that is created by linear combinations of the variables and stationarity created by first differencing:

$$\Delta X_t = \Pi X_{t-1} + \Gamma_1 \Delta X_{t-1} + \dots + \Gamma_{k-1} \Delta X_{t-k-1} + \varphi D_t + \varepsilon_t, \quad t = 1, \dots, T \quad (2)$$

Equation (2) presents the ECM representation of the VAR model. The VECM form of the model gives an intuitive explanation of the data, separating long and short-run effects. While Γ_i contains the short run information, Π contains the long-run relationships. Based on the insights of Stock/Watson (1988) that a cointegrating relationship represents that two or more time series share a common stochastic trend and assuming that our variables are $I(1)$, the rank (r) of matrix Π has to be reduced ($r < p$). The reduced rank matrix can be factorized into two $r \times p$ matrices α and β ($\Pi = \alpha\beta'$). The factorization provides r stationary linear combinations of the variables (cointegrating vectors) and $p - r$ common stochastic trends of the system.

As mentioned in chapter two, there is a certain probability that the Fed's introduction of Quantitative Easing might have altered the potential long-run relationship between US and Euro area interest rates. In accordance with our theoretic approach, if QE was effective in reducing long-term interest rates, it should have an effect on US interest rates for which the global downward trend in interest rates does not account for. Therefore, QE might have had a separate, identifiable impact on long-term interest rates in the USA which should show up as a break in the long-run relationship. We thus check for a potential structural break around the time of the Fed's announcements of QE1.

As our empirical approach is based on the Johansen procedure (CVAR), we use a large set of recursive techniques as proposed by Johansen/Juselius (2006) to check for structural changes in the relationship. The fundamental idea of recursive testing is to start with a baseline model estimated for a subsample and to gradually extend the end point of the sample until the full sample is covered. After every extension of the sample, the test statistics are re-estimated. The recursive methods used are:

- 1) The log likelihood test as a broad test giving us hints at the general appropriateness of the model. In this regard, the test is quite similar to the recursive Chow tests used in single equation models.
- 2) Recursive tests based on the eigenvalues allowing us to obtain detailed information about the constancy of the individual cointegration relations.
- 3) Recursive tests of the cointegration space. Because the eigenvalues are a quadratic function of α and β , the previous group of test is not able to differentiate between non-constancy related to α or β . The "max test of a constant β " and the test of " β_i equal to a known β " focus on spotting non-constancy in the β structure.
- 4) Recursively calculated prediction tests are used in order to check for systematic predictive failure of the model over a specific period of time.

Additionally, backward recursive estimation techniques are used. As there have been several rounds of QE in the USA between the end of 2008 and 2014 as well as changes in the conduction of QE, it is pretty difficult to identify potential dates of the structural breaks ex ante. In this regard, recursive test possess the advantage that the precise date of a potential structural break is not needed. However, recursive tests lose power to discriminate between structural stability and non-stability, if a potential structural break is near the end of the sample. In this regard, it appears to be even more difficult to analyze the relationship for structural breaks caused by QE2, Operation Twist and QE3. Therefore, we focus our analysis on QE1 which was announced

and started in November 2008 by the Fed, as the announcement and implementation of QE1 roughly split our sample in half. Furthermore, our choice to look at the impact of QE1 is also motivated by the impression that QE1 is generally considered to be more effective compared its successor programs.

4. Empirical Results

4.1 Unit Root Tests

According to Stock/Watson (1988), a cointegrating relationship can also be regarded as the occurrence of common stochastic trends of individual time series. As a first step of our analysis, the order of integration of every time series used in our model has to be determined. As common in empirical literature, Augmented Dickey Fuller (ADF) and Phillips Peron (PP) tests are used. In order to generate robust results, we perform different specifications regarding deterministic components, as the integration order is of essential importance for the subsequent cointegration analysis. The entire data sample is used, the maximal number of lags is 12 and the Bayesian information criterion is utilized to determine the appropriate number of lags included in the ADF test equations.

Table 4: Unit root tests

<i>Levels</i>	Augmented Dickey Fuller Test		Phillips Perron Test	
	Intercept	Intercept & Trend	Intercept	Intercept & Trend
US10Y	0.6262	0.6409	0.1579	0.5068
Ger10Y	0.6262	0.6490	0.6511	0.5750
EXR	0.9183	0.1843	0.9171	0.4475
<i>1. Differences</i>				
US10Y	0.000***	0.000***	0.000***	0.000***
Ger10Y	0.000***	0.000***	0.000***	0.000***
EXR	0.000***	0.000***	0.000***	0.002***

Note: Asterisks refer to level of significance: *10%, **5%, ***1%.

The results presented in Table 4 indicate that the levels are at least integrated of order one as the time series possess at least one stochastic trend. With regard to the results of the first differences, the tests propose that they are (trend-) stationary. Therefore, we conclude that the time series in levels are integrated of order one ($I(1)$).

As mentioned above, the introduction of QE can be regarded as an unparalleled event in the recent history of monetary policy. Therefore, not only relationships between variables might have changed, but also the behavior of the time series themselves. As depicted by Perron (1989), structural breaks might have strong effects on the results of unit root test, sometimes leading to wrong implications generated by tests. In order to further strengthen the robustness of our unit root tests, Zivot-Andrews tests are used. The Zivot-Andrews tests allow for a single break in the intercept, the trend or both (Zivot/Andrews, 1992).

Table 5: Unit root test allowing for structural breaks - Zivot-Andrews tests

<i>Levels</i>	Break (Intercept)	Break (Trend)	Break (Intercept + Trend)
US10Y	-3.68381	-3.58673	-4.01533
Ger10Y	-3.96228	-3.25455	-3.93304
EXR	-3.04287	-3.53209	-3.98429
<i>1. Differences</i>			
US10Y	-12.498***	-12.368***	-12.509***
Ger10Y	-12.369***	-12.171***	-12.353***
EXR	-12.678***	-12.513***	-12.696***

Note: Asterisks refer to level of significance: *10%, **5%, ***1%.

The Bayesian information criteria is used for the purpose of lag length selection.

Table 5 contains the results of the Zivot-Andrews tests. The results support the findings of the ADF and PP tests. With regard to the levels, the null hypothesis of a unit root cannot be rejected even after we allow for various types of structural breaks. As the tests rejected the null hypothesis for the first differences, the variables can be considered to be $I(1)$ in levels. Therefore, we feel legitimized to use cointegration approaches.

4.2 Cointegrated VAR estimations

4.2.1 Estimation of the long-run relationship

We focus on the relationship between long-term (10-year) bond yields and neglect further real variables such as real GDP since we are interested mainly in the impact of QE on financial markets. Furthermore, if a cointegrating relationship can be detected using a sub-system, the long-run relationship should also be present in a larger model which includes additional variables. In this regard, our model M_t^{10Y} contains the following variables:

$$M_t^{10Y} = (US10Y, Ger10Y, LEX)_t'$$

Regarding the model, we have to choose a specification regarding the lag length. Furthermore, we have to decide which deterministic components are to be included into the VAR as well as into the cointegrating space. Regarding the deterministic components, an intercept into the unrestricted VAR and the cointegrating space is included. Because the model resembles the Uncovered Interest Parity (UIP), there appears to be no theoretical reason for a linear trend in the cointegrating space.²¹ With regard to the lag length, we include three lags into the unrestricted VAR. Although the information criteria suggest two lags, the results of the residual analysis tremendously improve if an additional lag is included. As the number of the degrees of freedoms is still high, we decide to include three lags. For the following dates, dummy variables are included into the VAR-equations: 2004:04, 2008:10 and 2008:12. However, no dummy variable enters the cointegrating vector(s).

Table 6: Residual analysis –diagnostic testing on the unrestricted VAR (3)-Model; M_t^{10Y}

Multivariate Test				
Residual autocorrelation:				
LM(1):	ChiSqr(9)	=	17.274	[0.045]
LM(2):	ChiSqr(9)	=	12.794	[0.172]
LM(3):	ChiSqr(9)	=	9.984	[0.352]
LM(4):	ChiSqr(9)	=	8.641	[0.471]
Test for ARCH:				
LM(1):	ChiSqr(36)	=	39.863	[0.302]
LM(2):	ChiSqr(72)	=	72.418	[0.464]
Univariate Tests				
	ARCH(3)	Normality	Skewness	Kurtosis
Δ Ger10Y	1.322 [0.724]	0.769 [0.681]	-0.145	2.750
Δ US10Y	3.299 [0.348]	3.069 [0.216]	-0.022	3.508
Δ LEX	4.914 [0.178]	0.422 [0.810]	-0.028	3.070

Note: *p*-values in brackets

Diagnostic tests of the VAR model are presented in Table 6. In order to avoid bias of the trace test, the model has to be well specified especially regarding residual autocorrelation and the normality of the residuals. Overall, the assumptions of the CVAR appear to be satisfied. The

²¹ While the results of the LR-Test of Exclusion do not recommend to exclude a deterministic trend from the cointegrating space. However, its inclusion does not fundamentally change the results of our analysis. After imposing over-identifying restrictions, the deterministic trend also becomes insignificant.

tests do not indicate any issues regarding the assumption of normality. Furthermore, there is only small evidence that autocorrelation and ARCH-effects are present. Therefore, we conclude that the model is well-specified.

Table 7: LR Trace Test for the unrestricted VAR (3)-model, M_t^{10Y}

r	$p - r$	Eigenvalue	Trace	95% crit. Value	p -value
3	0	0.148	40.359	34.565	0.011
2	1	0.084	15.773	19.932	0.177
1	2	0.016	2.455	9.219	0.682

Note: p -values for testing the null hypotheses of $H_0: r = 0$ and $H_0: r \leq i + 1$, $i = 0, 1, 2$, respectively, for different ranks r .

The results of the trace tests are presented in Table 7.²² The results clearly indicate the presence of a single cointegrating relationship. Therefore, the rank of the Π -matrix is restricted to one.

Table 8: The just-identified long-run cointegration relations for $r = 1$, M_t^{10Y}

	US10Y	Ger10Y	EXR	Constant
$\hat{\beta}_1$	0.116 [1.755]	-0.129 [-2.011]	1 [NA]	-0.353 [-2.900]
	Δ US10Y	Δ Ger10Y	Δ EXR	
$\hat{\alpha}_1$	0.163 [1.424]	0.189 [2.000]	-0.051 [-4.114]	

Note: The first column reports long-run coefficients β . The second column shows the adjustment coefficients α . t -values in brackets. The line above provides the test statistic for over-identifying restrictions, which is an LR-test [p-value].

As the rank of Π is chosen to be one, normalizing on one of the variables is sufficient in order to generate the just-identified long-run relation. The results are presented in Table 8. The long-run relationship is in line with theoretical expectations. An interest rate increase in one country leads to an appreciation of the own currency. Regarding the adjustment process, we see the expected reaction of the exchange rate as it contributes to reduce deviations from the estimated steady state relationship. While the adjustment of the German yield also contributes in reducing

²² We simulate the asymptotic distribution of the trace test. The following settings are used: length of random walk: 400, number of replications: 2500.

equilibrium errors, the US yield shows no (significant) sign of adjustment. All estimated β -coefficients are significant at least at the 10% level.

After estimating the just-identified long-run cointegration relations, placing restrictions on the β -vector can be used to test for specific relationships suggested by economic theory. Firstly, we test for proportionality between the interest rate measures by restricting the coefficient of the exchange rate variable to zero. However, the test of restricted model clearly rejects the imposed restriction (p -value = 0.002). Therefore, the exchange rate appears to be an important component of the estimated long-run relationship.

Secondly, we restrict the interest rates to have the same magnitude, but different signs. Therefore, we check for a relation among the interest rate differential between German and US yields and the exchange rate which is one common assumption of the UIP. The model is accepted (p -value = 0.785) at the 10-percent significance level. The results are presented in Table 9.

Table 9: The over-identified long-run cointegration relations for $r = 1$, M_t^{10Y}

	US10Y	Ger10Y	EXR	Constant
$\hat{\beta}_1$	0.128 [2.041]	-0.128 [-2.041]	1 [NA]	-0.400 [-9.691]
	Δ US10Y	Δ Ger10Y	Δ EXR	
$\hat{\alpha}_1$	0.145 [1.238]	0.193 [2.002]	-0.052 [-4.021]	

Test of restricted model: CHISQR(1) = 0.074 [0.785]

Note: The first column reports long-run coefficients β . The second column shows the adjustment coefficients α . t -values in brackets. The line above provides the test statistic for over-identifying restrictions, which is an LR-test [p-value].

According to our estimations, we obtain the following long-run relationship:

$$0.400 + 0.128 * (Ger10Y - US10Y) = EXR$$

The estimated relationship indicates that a reduction of the US (German) yield should lead to a depreciation (appreciation) of the US-Dollar vis-à-vis the Euro which is in line with economic theory. Movements of the exchange rate and the German yields thus reduce deviations from the long-run relationships.

4.2.2 Did QE1 cause a structural break?

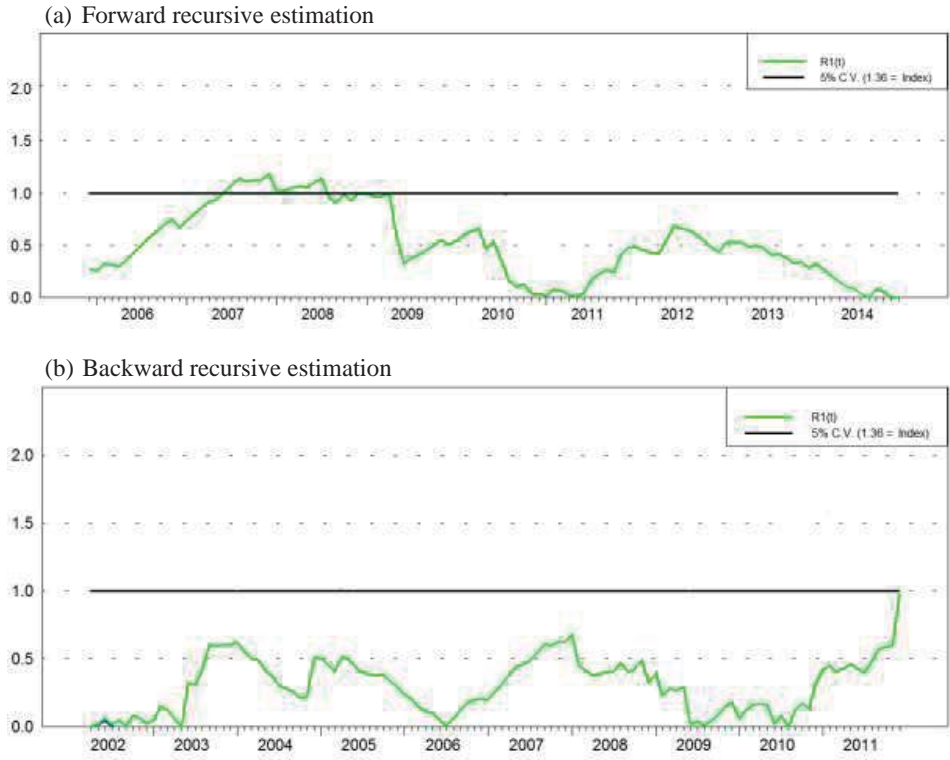
Even though, the trace tests highly recommend one cointegrating relationship and we have found a significant and theoretically correct cointegration relationship, this does not exclude the possibility that the model suffers from parameter non-constancy. The purpose of recursive tests is to identify whether we have a constant parameter regime and, if this is not the case, to identify where in the sample period the data strongly suggests a change in the structure (Johansen/Juselius, 2006). Regarding the research question of the paper, the time period between 2008:11 and 2010:03 is of importance. In particular, two time dates might be of particular interest regarding the possibility of a structural break caused by QE1: Firstly, the announcement of QE1 in 2008:11 and secondly, the FOMC's decision to buy Treasury bonds as well as an quantitative expansion of the QE program in 2009:03.

With regard to our recursive tests, the baseline sample contains data from 2002:04 to 2005:12.²³ Besides the use of forward recursive test, we are also making use of backward recursive tests. In order to keep the amount of figures manageable, we only include the backward recursive estimation if we find contradicting evidence compared to the forward procedure. Regarding the backward recursive estimations, the base sample contains data from 2011:12 to 2014:12.

Because the log-likelihood test is similar to the Chow test used in single equation estimations, the test is quite useful in identifying the timing of a structural break. With regard to the results presented in Figure 10(a), the test indicates some instability starting around the beginning of the subprime mortgage crisis in 2007. However, the evidence is limited. The test statistic slightly exceeds the black line which represents the critical value at the 5% significance level and only for a brief period of time. After mid-2008, the test statistic stays below the critical value. Surprisingly, we find no evidence of a structural break around the breakdown of Lehman-Brothers and the subsequent beginning of the global financial crisis. In this regard, we find no clear impact of QE on the test statistics and therefore no evidence of significant parameter non-constancy around the introduction of QE1 in 2008:11. Further evidence of a structural break starts to appear in the mid of 2010 – at the beginning of the European debt crisis. The empirical realization of the test statistic quickly increases, but fails to reject the null hypothesis of parameter constancy. Checking the robustness of this results by using backward recursive testing, we once again find no effect of QE on the structure of the relationship.

²³ In order to check for robustness, we also used the following baseline samples: 2002:04 to 2004:12 and 200:04 to 2006:12. We obtained similar results which are available on request.

Figure 10: Test for Constancy of the Log-Likelihood, M_t^{10Y}



Next, we focus on the stability of the cointegrating relationship by checking the constancy of the eigenvalues. Figure 11 depicts the recursively estimated trace test statistics. Both models – $X(t)$ and $R1(t)$ – show some signs of instability in the estimated cointegration relation (blue graph).²⁴ A first sign of instability appears in mid-2008, as the blue graph decreases strongly. Afterwards, the relationship stabilizes as the blue graph increases and remains significant at the 5% significance level. Around the announcement and implementation of QE1, the cointegration relationship shows an almost linear development indicating no signs of structural changes caused by QE1. The largest amount of instability is found around mid- 2010 which timely corresponds to the beginning of the European Debt Crisis, as the cointegration relationship gradually declines and even becomes insignificant at the 5% level. Eventually in mid-2011 the relationship once again starts so increase till the end of the sample. Additional evidence is generated

²⁴ While the $X(t)$ model contains short-run and long-run information of the data, the $R_1(t)$ only contains long-run information.

in Figure 12 which illustrates the fluctuations of the eigenvalues. The graph stays below one indicating no sign of instabilities.

Figure 11: Recursively estimated Trace Test statistics, M_t^{10Y}

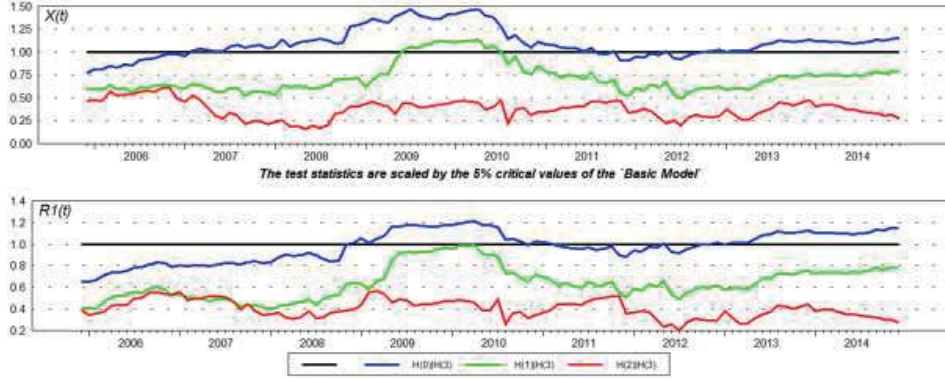
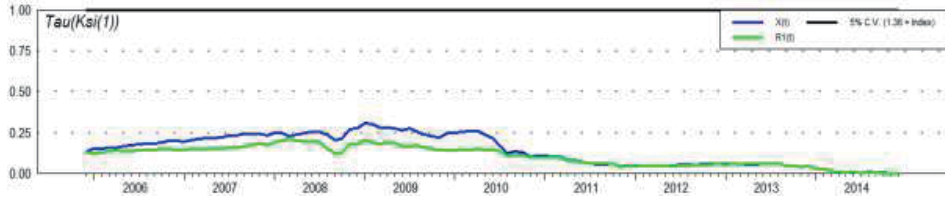


Figure 12: Fluctuations of the Eigenvalue, M_t^{10Y}



Because the eigenvalue can be expressed as a quadratic term of α and β , those tests do not differentiate between instability caused by α and β . As a next step, we focus on the components of our long-run relationship, β . We start by presenting the results of the “max test of constant beta” test which can be considered as a rather conservative test. In this regard, rejection of the null hypothesis would indicate strong evidence of non-constancy in the long-run relationship. The results are presented in Figure 13 and show no significant evidence of a possible structural break in the long-run relationship.

Next, the test of “ β_t equals a known β ” is used to check for the constancy of β . The main idea of the test is to obtain an estimate $\tilde{\beta}$ based on a chosen reference period. The recursive testing will then extend the sample beyond the reference sample checking whether the parameters remain constant over time. In this regard, the results can be sensitive to the chosen reference

sample.²⁵ We chose the reference sample as 2002:04 to 2006:12. Different reference samples were employed, but the selection of the reference sample did not change the results fundamentally. The results are presented in Figure 14. A first period of instability can be established around the beginning of 2008 and therefore around the meltdown of Bear Stearns. However, stability is not rejected. Surprisingly, we once again find no evidence of a structural meltdown around the peak of the financial crisis in late 2008. Correspondingly, we also find no evidence of an impact of QE1. The most striking result is the large amount of instability which begins to develop around the beginning of the European debt crisis in mid-2010. From this data point onwards, the test results indicate that long-run relationship has changed in comparison with the pre-crisis era as indicated by the rejection of the null hypothesis.

Figure 13: Test of Beta Constancy, M_t^{10Y}

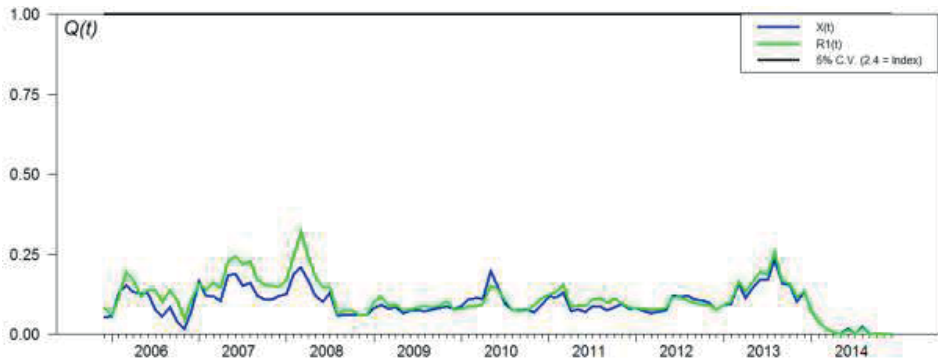
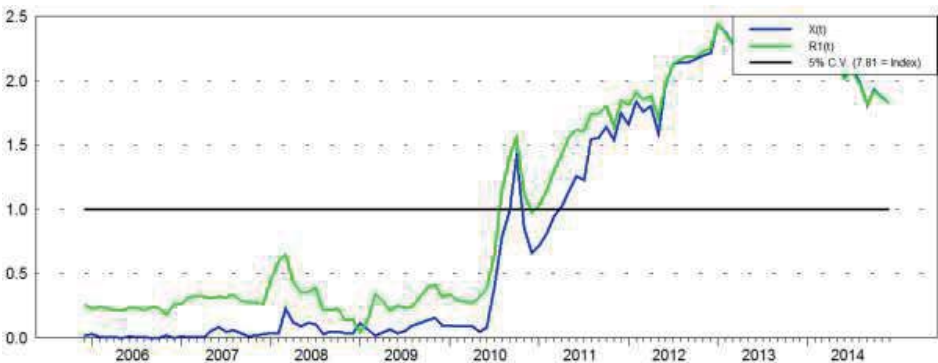


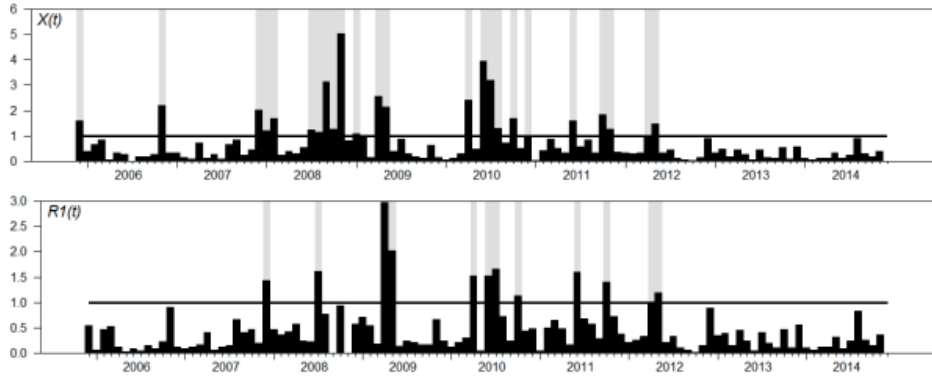
Figure 14: Test of known beta, M_t^{10Y}



²⁵ In order to obtain robust results, we performed several tests using the following reference sample periods: 2002:04 to 2004:12, 2002:04 to 2005:12, 2002:04 to 2007:06 and 2002:04 to 2007:12. However, the test does not appear to be sensitive to the reference sample.

As a final check, we look at the one step-ahead predictions of the system. The results depicted in Figure 15 generate further empirical evidence that there is a certain amount of instability between mid-2007 to mid-2008. With regard to the introduction of QE1, the evidence of a break is rather limited. However, we find prediction errors in 2009:03 and 2009:04 which might correspond to the announcement of the FOMC to start buying US Treasury bonds. The second bulk of prediction errors occurs during the onset of the European debt crisis in mid-2010.

Figure 14: One step-ahead predictions of the full system, M_t^{10Y}



To sum up our results, we detect some empirical evidence of structural changes in the model. First of all, we come up with limited evidence that there is some degree of instability prior to the outbreak of the global financial crisis (between mid-2007 to mid-2008). Most importantly with regard to our research question, we do not find evidence that the announcement and conduct of QE1 generated a change in the relationship of “risk-free” bonds of the USA and Europe. With regard to our results, the highest amount of instability can be found in mid to late 2010. Based on the timing, we believe that the beginning of the European debt crisis might be held responsible for this pattern. In this regard, uncertainty about the future of the Euro and the probability of a breakup of the Euro area might have destabilized the relationship between the nominal exchange rate and the interest rate differential. The crisis also resulted in large capital flows into Germany – especially from the periphery states of the Euro area – driving down German Treasury yields. This aspect can be characterized as a “safe haven effect” and might have further destabilized the relationship between German and US yields.

In order to verify the robustness of results presented, additional estimations were conducted using five and seven-year yields. The estimations generated almost identical results. Therefore, our results do not appear to be sensitive regarding the choice of yields. Furthermore, we checked the robustness by including the CBOE Volatility Index (VIX) as an exogenous variable in order

to correct for systematic risk perception of US financial markets. Again, the trace test recommends the presence of one cointegrating relationship ($r = 1$). While the VIX significantly enters the short-term dynamics of the model, the variable is not significant in the long-run relationship. Once again, we find no evidence that QE generated a structural change in the transatlantic interest-rate relationship.²⁶

5. Conclusions

Did Quantitative Easing generate a structural break in the relationship between European and US long-term interest rates which, in turn, may be considered as an individual effect of QE on US interest rates and financial markets? The estimation results of our CVAR analysis generated theoretically expected relationships between US and European yields as well as the nominal exchange rate. By using recursive estimation methods, we did not find significant evidence of a structural break in the interest rate relationship caused by QE1.

Following our hypothesis, we therefore conclude that there is no evidence that QE had an effect on US interest rates which cannot also be explained by the global downward trend in (long-term) interest rates. While we cannot reject QE1 having no impact on the transatlantic interest rate relationship, we find evidence that the beginning of the European debt crisis had a more destabilizing impact on the relationship.

As our analysis focused on the effects of QE1, it is possible that QE2 or QE3 had different effects on the interest rate relationship. As the asset purchases during QE2 comprised solely of US Treasury bonds, there may have been a strong(er) effect on the transatlantic relationship of “risk-free” bonds. On the other hand, the effects of QE2 and also QE3 are generally considered to be even smaller compared to QE1. In this regard, we do not believe that QE2 and QE3 had a more pronounced effect on the interest rate relationship than QE1.

References

- Bauer, M.D. / Neely, C.J. (2014): International channels of the Fed's unconventional monetary policy, *Journal of International Money and Finance*, Vol. 44, pp. 24-46.
- Baumeister, C. / Benati, L. (2012): Macroeconomic effects of a spread compression at the zero lower bound, Bank of Canada Working Paper, 2012-12.

²⁶ Results of the robustness checks are available on request.

Belke, A. / Klose, J. (2013): Modifying Taylor reaction functions in presence of the zero-lower-bound – Evidence for the ECB and the Federal Reserve, *Economic Modelling*, Vol. 35, pp. 515-527.

Chen Q. / Filardo, A. / He, D. / Zhu, F. (2013): International spillovers of central bank balance sheet policies, *BIS Working Paper*, No. 66, Bank for International Settlements, Basle.

Christensen, J. / Rudebusch, G. (2012): The response of interest rates to US and UK Quantitative Easing, *Economic Journal*, Vol. 122(564), pp. F385-F414.

Chung, H. / Laforte, J.-P. / Reifschneider, D. / Williams, J.C. (2012): Have we underestimated the likelihood and severity of zero lower bound events?, *Journal of Money, Credit and Banking*, Vol. 44(1), pp. 47-82.

D'Amico, S. / King, T.B. (2012): Flow and stock effects of large-scale Treasury purchases: Evidence on the importance of local supply, *Journal of Financial Economics*, Vol. 108 (2), pp. 425-448.

Fawley, B / Neely, C. (2013): Four stories of Quantitative Easing, *Federal Reserve Bank of St. Louis Review*, Vol. 95(1), pp. 51-88.

Fratzscher, M. / Lo Duca, M. / Straub, R. (2013): On the international spillovers of US Quantitative Easing, *Discussion Papers, DIW Berlin*, No. 1304, German Institute for Economic Research, Berlin.

Gagnon, J. / Raskin, M. / Remache, J. / Sack, B. (2011): The financial market effects of the Federal Reserve's large-scale asset purchases, *International Journal of Central Banking*, Vol. 7(1), pp. 3-43.

Gambacorta, L. / Hofmann, B. / Peersman, G. (2014): The effectiveness of unconventional monetary policy at the zero lower bound: a cross-country analysis, *Journal of Money, Credit, Banking*, Vol.46, pp. 615-642

Glick, R. / Leduc, S. (2013): Central bank announcements of asset purchases and the impact on global financial and commodity markets, *Journal of International Money and Finance*, Vol. 31 (8), pp. 2078-2101.

Greenwood, R. / Hanson, S.G. / Rudolph, J.S. / Summers, L.H. (2014): Government debt management at the zero lower bound, *Hutchins Center on Fiscal & Monetary Policy, Working Paper*, No. 5, Brookings.

Gros, D. / Aldici, C. / De Groen, W.P. (2015): Lessons from Quantitative Easing: Much ado about so little?, CEPS Policy Brief, No. 330, Center for European Policy Studies, Brussels.

Hamilton, J.D. (2011): Evaluating Quantitative Easing using event studies, URL: http://econbrowser.com/archives/2011/07/evaluating_quan.

Hamilton, J.D. / Wu, J.C. (2012): The effectiveness of alternative monetary policy tools in a zero lower bound environment, *Journal of Money, Credit and Banking*, Vol. 44(1), pp. 3-46.

Johansen, S. / Juselius, K. (2006): The Cointegrated VAR model: methodology and applications, Oxford University Press, Oxford, 2nd ed.

Kapetanios, G. / Mumtaz, H. / Stevens, I. / Theodoridis, K. (2012): Assessing the economy-wide effects of Quantitative Easing, Bank of England Working Paper, No. 443, London.

Krishnamurthy, A. / Vissing-Jorgensen, A. (2011): The effects of Quantitative Easing on interest rates: channels and implications for policy, National Bureau of Economic Research Working Paper, No. 17555, Cambridge/MA.

Liu, P. / Mumtaz, H. / Theodoridis, K. / Zanetti, F. (2014): Changing macroeconomic dynamics at the zero lower bound, Bank of England, London, mimeo.

Lombardi, M. / Zhu, F. (2014): A shadow policy rate to calibrate US monetary policy at the zero lower bound, BIS Working Paper, No. 452, Bank for International Settlements, Basle.

Meyer, L.H. / Bomfim, A.N. (2010): Quantifying the effects of Fed asset purchases on Treasury yields. *Macroeconomic Advisers Monetary Policy Insights: Fixed Income Focus*, June 17.

McLeay, M. / Radia, A. / Thomas, R. (2014): Money creation in the modern economy, Bank of England Quarterly Bulletin 2014/1, London, pp.1-14.

Neely, C. J. (2010). The large-scale asset purchases had large international effects, Federal Reserve Bank of St. Louis Working Paper 2010 - 018G, St. Louis, July.

Perron, P. (1989): The great crash, the oil price shock, and the unit root hypothesis, *Econometrica*, Vol. 57(6), pp. 1361-1401.

Pesaran, H.M. / Smith, R. (2012): Counterfactual analysis in macroeconometrics: an empirical investigation into the effects of Quantitative Easing, Institute for the Study of Labor, Bonn.

Rosengren, E.S. (2015): Lessons from the U.S. Experience with Quantitative Easing, President & Chief Executive Officer Federal Reserve Bank of Boston, The Peterson Institute for International Economics and Moody's Investors Service's 8th Joint Event on Sovereign Risk and Macroeconomics, Frankfurt, Germany February.

Rudebusch, G.D. / Sack, B.P. / Swanson, E.T. (2007): Macroeconomic implications of changes in the term premium, Federal Reserve Bank of St. Louis Review, Vol. 89(4), pp. 241-69.

Stock, J. / Watson, M. (1988): Testing for common trends, Journal of the American Statistical Association, Vol. 83 (404), pp. 1097-1107.

Wright, J.H. (2012): What does monetary policy do to long-term interest rates at the zero lower bound?, The Economic Journal, Vol. 122 (564), pp. F447-F466.

Zivot, E. / Andrews, D.W. (1992): Further evidence on the Great Crash, the oil-price shock, and the unit-root, Journal of Business & Economic Statistics, Vol. 10(3), pp. 251-270.