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Export Diversification Effects of the WTO Trade Facilitation Agreement

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JEL : F13, F14, F17

Keywords: Trade facilitation, Export diversification, International trade agreements, WTO

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This draft: 16 December 2014

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1 Introduction

Trade economists have for some time now emphasized the need to bring down trade costs, which by many estimates remain quite sizeable. Even for a “representative rich country”, Anderson and van Wincoop (2004) have estimated that the ad valorem equivalent of trade costs could be as high as 170%. As shown by Arvis et al. (2013), customs formalities and trade procedures that result in unnecessary delays or complexities to traders constitute an important component of trade costs. Recognizing this, the WTO 1996 Ministerial Conference in Singapore agreed “to undertake exploratory and analytical work” on this issue. The simplification of the trade procedures has been part of the WTO’s negotiating agenda since August 2004. In December 2013, WTO members concluded negotiations on a Trade Facilitation Agreement at the Bali Ministerial Conference. The new agreement will enter into force and become an integral part of the WTO Agreement once two-thirds of WTO members complete their domestic ratification process.

An example of how trade facilitation can simplify trade procedures and make them more transparent can be taken from a country which became a WTO Member in 2013 – the Lao People’s Democratic Republic. An online portal for trade has been operative since 2012.¹ On this website, all trade-related laws, regulations, measures, restrictions, licensing requirements and tariffs are indexed, cross-referenced, and made searchable by commodity code. The website also includes detailed process maps of business procedures for importing and exporting; full listings of national standards for products; procedures for clearing goods at the border; downloadable forms; and e-alerts which traders can customize to receive information.

The importance of achieving success in the WTO negotiations on trade facilitation has been underlined by a fair amount of empirical work. Various approaches for measuring the benefit of a multilateral agreement on trade facilitation have been pursued, including how much it will reduce trade costs, how much trade will increase, as well as the positive impact on jobs and on GDP. One effect that seems not to have been explored in sufficient depth is the effect on export diversification. To the extent that trade and customs procedures act like fixed costs, they prevent exporters from entering new markets or selling a wider array of products. The benefit of export diversification over selling more of the same product or selling more to the same market is the resulting reduction in risk from shocks to international trade. Exporters with diversified export baskets

¹See <http://www.laotradeportal.gov.la/index.php?r=site/index>.

or destinations are likely to be better insulated from shocks to specific markets or sectors than others.

There are various approaches taken in the literature to measure, more or less directly, trade facilitation. Several studies use the World Bank's Logistics Performance Index (LPI) and Doing Business indicators as proxies. The LPI is based on a worldwide survey of operators on the ground, providing feedback on the logistics 'friendliness' of the countries in which they operate and those with which they trade. In addition, survey data is supplemented with quantitative data on the performance of key components of the logistics chain in a given country. This includes the quality of trade and transport infrastructure. Doing Business indicators use data on the time and cost (excluding tariffs) associated with exporting and importing a standardized cargo of goods by sea transport. The time and cost necessary to complete every official procedure for exporting and importing the goods are included as well.

The OECD has developed indicators on import, export and transit trade that are closely related, and can be readily mapped on, to the families of measures included in the WTO Trade Facilitation Agreement – Information availability, Involvement of the trade community, Advance Rulings, Appeal Procedures, Fees and charges, Formalities, Cooperation, Consularization, Governance and Impartiality and Transit proceedings – see Appendix Table A-1. As explained in Moisé et al. (2011) and Moisé and Sorescu (2013), sixteen indicators were constructed based on the relevant provisions of the WTO Trade Facilitation Agreement. The indicators are in turn broken down further into some ninety-seven variables, whose values are drawn from questionnaire replies as well as publicly available data. The variables seek to reflect not only the regulatory framework in the surveyed country but also the state of implementation of the trade facilitation measures. Each of these variables is scored following a 'multiple binary scheme' where the score of two corresponds to the best performance, zero corresponds to the worst performance and a score of one to performance that lies in between.

This paper makes use of the OECD TFIs to estimate the impact of trade facilitation on export diversification, as measured by extensive margins of trade. In the baseline estimations, we consider two types of extensive margins: the number of products (HS sub-headings) by export destination, and the number of export destinations by product. We also consider theory-based extensive margins: the bilateral extensive margin suggested by Hummels and Klenow (2005), and an exporter-product extensive margin that, to the

best of our knowledge, has not previously been explored in the literature.

While we are not the first to study the extensive-margin effects of trade facilitation, we are the first to do so using the OECD TFIs. Moreover, we add to the existing literature by considering an exporter-product, not only a bilateral dimension of trade margins. A third novel contribution of this paper is the quantification of the effect of implementing trade facilitation under two realistic scenarios: (i) trade facilitation reform that moves countries that are below the median of their region to that benchmark; and (ii) reform that moves countries that are below the global median to that level.

Throughout this study, we focus on trade facilitation in the exporting country. When analyzing the number of destinations by product, this is the only viable option. When analyzing the number of products by destination, we are aware that increases in importers' trade facilitation are likely to have a positive effect. The empirical question of interest, however, concerns the effects of own's trade facilitation for given levels of trade facilitation in destination markets.

The remainder of this paper is organized as follows. The next section provides an overview of the literature on trade facilitation. Section 3 discusses the empirical methodology to estimate the effect of trade facilitation on trade margins. We first define the indicators for the different trade margins used in the empirical analysis. Next, we specify the econometric model. In Section 4, we present the empirical results. Section 5 presents estimations that use alternative measurements of trade margins and of trade facilitation. It also discusses various methodologies employed to test whether the effects are heterogeneous across counties and sectors. Section 6 includes the results of simulations under the two scenarios of convergence to the regional median and convergence to the global median. Section 7 concludes.

2 Literature

Trade facilitation has a significant potential to reduce trade costs. This effect has been quantified by a series of empirical studies that follow the methodology of Novy (2013) to infer trade costs from the observed pattern of production and trade across countries. Chen and Novy (2009) estimate that technical barriers to trade, taken as a whole, explain 4.5% of the variation in trade costs across 11 European Union member countries

between 1999 and 2003.² Arvis et al. (2013) estimate trade costs in agriculture and manufactured goods in 178 countries for the 1995-2010 period. They find that a one standard deviation improvement in the World Bank's LPI is associated with a trade cost reduction of 0.2 – 0.5 standard deviations. Using the OECD TFIs as a measure of trade facilitation, Moisé et al. (2011) estimate a cost reduction potential of around 10% of overall trade costs. In a follow-up study, Moisé and Sorescu (2013) disaggregate the cost-reduction potential across income groups. They estimate this potential to be 14.5% in low income countries, 15.5% in lower middle income countries and 13.2% in upper middle income countries.

Trade facilitation is likely to impact both variable and fixed trade costs. The formalities and requirements of a country's customs have to be met each time a shipment crosses a border. There are, however, also one-time costs such as those incurred by a firm to acquire information on border procedures. The number and complexity of the documents required for clearance can also be seen as a fixed cost. Traders have a one-time cost of learning how to fill in the forms. As the WTO Trade Facilitation Agreement contains provisions requiring countries to publish and make available information on border procedures as well as to decrease and simplify documentation requirements, it should reduce fixed costs and create new trading opportunities. Firms that did not export before may be able to do so now, since their revenues could cover the lower fixed costs of exporting (Melitz, 2003). Trade facilitation can, therefore, both expand existing trade flows (intensive margin effect) and create new trade flows (extensive margin effect).

Empirical evidence on the intensive margin effects is provided by several authors. Iwanow and Kirkpatrick (2009) find that trade facilitation positively contributes to bilateral exports of manufactured products. More recently, Moisé and Sorescu (2013) estimate a positive effect on bilateral trade flows of bilateral measures of trade facilitation constructed from the OECD TFIs. A related literature highlights the importance of time for trade. Since trade facilitation is likely to reduce the time it takes for products to cross borders, this literature is also relevant in this context. In a recent contribution, Zaki (2014) shows that the time to import (export) is equivalent to a mean ad valorem tax of 34.2% (17.6%) for developing countries. A study by Hummels and Schaur (2013) shows that each day in transit is worth 0.6%-2% of the value of the

²Their preferred specification explains 80.8% of the variation in trade costs. 42.8% is attributable to the 3-digit industry fixed effects. Of the 38% that the remaining regressors explain, geography and transport costs alone are responsible for about 25%; policy variables explain 7.6%, with technical barriers to trade (TBTs) being the most important policy factor (4.5%). TBTs therefore explain 11.8% of the variation in trade costs not accounted for by unobservable industry characteristics.

good and that time is particularly important for intermediate goods. However, Freund and Rocha (2011) find that when comparing the effects of transit, documentation, and ports and customs delays on trade, the most significant effect comes from inland transit delays. Each additional day that a product is delayed prior to being shipped reduces trade by at least 1 per cent, as found by Djankov et al. (2010). A result which combines the effects of time and costs is obtained by Hausman et al. (2013). In their study, a 1% reduction in processing costs/time leads to 0.49%-0.37% of increased bilateral trade. There is also firm-level evidence showing the adverse effect of customs delays on trade. Using a sample of Uruguayan firms, Volpe Martincus et al. (2013) show that an increase by two days in the duration of export inspections reduces exports by 16.4%. Moreover, exports would be 5.9% larger if all exports could be processed within one day.³

Some studies in this literature use econometric results from gravity equations to perform counterfactual analysis. Hoekman and Nicita (2011) simulate the effect of policy convergence by low income countries to the average of middle income countries. The percentage increase in exports (imports) of low income countries that would result from a combined convergence of the Doing Business ‘cost of trading’ indicator and of the LPI score to the average of middle income countries would be 17% (13.5%).⁴ Portugal-Perez and Wilson (2012) simulate the effects of improving trade facilitation (broadly encompassing physical infrastructure, information and communications technology, border and transport efficiency as well as business and regulatory environment). Their benchmark is an improvement half-way to the level of the top performing country in the region. The ‘ad valorem tariff-cut equivalents’ they estimate are heterogeneous across regions, with investment in physical infrastructure generally resulting in the largest trade gains. Hufbauer et al. (2013) perform a thought experiment in which countries lift their trade facilitation halfway to the region’s top performer in each category. They estimate an increase in total merchandise exports of developing countries of \$569 billion (9.9%) and an increase in total exports of developed countries of \$475 billion (4.5%).

The empirical evidence on the extensive margins effects of trade facilitation is more limited than the one on the intensive margins. Nordås et al. (2006) were among the first to show the negative effects of time to export on the probability to export. Dennis and Shepherd (2011) estimate the impact of various Doing Business indicators on the number of products that developing countries export to and import from the

³See Fernandes and Hillberry (2014) for a similar firm-level study using Albanian custom data.

⁴The LPI index alone has a higher effect than the Doing Business ‘cost of trading’ indicator. This is because improvements in the LPI also capture improvements in the quality of a country’s infrastructure.

European Union. They find that poor trade facilitation has a negative impact on developing country export diversification. Another approach is taken by Feenstra and Ma (2014). They proxy trade facilitation with port efficiency and estimate its impact on export variety, a theory-based measure of the extensive margin. They show a positive and significant effect of port efficiency on export variety. Finally, Persson (2013) distinguishes between the effects of trade facilitation (measured using the number of days needed to export from the World Bank’s Doing Business indicators) on homogenous and differentiated products. She finds that trade facilitation has a higher impact on differentiated products. Reducing export transaction costs increases the number of differentiated products by 0.7% and by 0.4% for homogenous products.

3 Empirical methodology

In this section, we define trade margins and trade facilitation indicators. Next, we outline the econometric approach.

3.1 Trade margins

We consider the relationship between trade facilitation and two indicators of trade margins: the number of exported products by destination and the number of export destinations by product.

The number of exported products by destination, npd_{ij} , counts how many products country i exports to destination j . As a baseline, we define ‘products’ as HS sub-headings (6 digit HS codes). In the HS 2002 classification that we use, there are 5224 sub-headings. For each ij pair, npd_{ij} can therefore theoretically range between 0 (no trade) and 5224 (country i exports all products to j). Panel (a) of Table 1 presents in-sample summary statistics for npd_{ij} . Overall, the variable varies between 0 and 4831 (the latter being $npd_{USA-CAN}$ – the number of HS6 sub-headings exported by the United States to Canada). Disaggregating over World Bank regions (and excluding ‘Offshore’ and ‘Industrial’ to focus on developing and emerging economies), the mean of npd_{ij} varies between 43 for Sub-Saharan Africa to 501 for East Asia and Pacific. The incidence of zeros is also highest in Sub-Saharan Africa (32% of observations) and lowest in Asia (together with Middle East and North Africa). There is, however, considerably less variation across Sub-Saharan African countries than across countries from other regions.

The number of destinations by product, ndp_{ik} , counts how many destinations are served by country i 's exports of product k . In this case, too, the baseline definition of 'product' is an HS sub-heading. Panel (b) of Table 1 presents summary statistics for ndp_{ik} . Overall, the variable varies between 0 and 169 (the latter being the number of Chinese export destinations of HS sub-heading 392690 – 'Other Articles of Plastics'; HS sub-heading 830140 – 'Other locks of Base Metal'; and HS sub-heading 940320 – 'Other Metal Furniture'). Again, the disaggregation over World Bank regions reveals relatively low scores for Sub-Saharan Africa (with an average of 1 destination served by product), and relatively high scores for Asian countries (with an average of 16 and 9 destinations served by product by East Asia and Pacific and South Asia, respectively). The incidence of zeros is also highest in Sub-Saharan Africa (68% of observations). The same incidence ranges between 36% and 51% for other regions.

Panel (c) of Table 1 presents the summary statistics for ndp_{ik} computed using HS4 trade data, that is, defining 'products' as HS headings. The number of observations and the percentage of zeros are clearly lower than for ndp_{ik} computed from HS6 trade data. Conversely, in the ij sample the sample size is determined by the number of exporting and importing countries, not by the level of sectoral disaggregation.

< Table 1 about here >

In the construction of npd_{ij} and of ndp_{ik} , we rely on mirror trade data to the extent possible because import data tend to be more complete than export data. We therefore measure exports of country i in product k using the reported imports of country j in the same product. For the few country-years for which mirror data is not available, we rely on reported export data.⁵

3.2 Trade Facilitation Indicators

The sample used for the regressions includes data for 133 countries for which OECD Trade Facilitation Indicators (TFIs) are available.⁶ Table 2 presents summary statistics for the variable TFI. This is the simple average of the country-specific indicators $TFI_i^A, TFI_i^B, \dots, TFI_i^L$.⁷ The average is unweighted because there

⁵Mirror data is not available for the years 2010, 2011 and 2012 for the following countries with TFI information: Antigua and Barbuda, Brunei Darussalam, Cuba, The Gambia, Indonesia, Iran, Kuwait, Mali, Mongolia, Papua New Guinea, Qatar and Suriname.

⁶The full list of countries by World Bank region group, with information on the date of WTO (GATT, where applicable) membership, is available in Appendix Table A-2.

⁷We only have information on indicators A-L.

is no criterion in the WTO Trade Facilitation Agreement or in its previous drafts to rank different indicators in terms of their relevance. Since each sub-indicator ranges between 0 and 2, so does TFI. Among developing and emerging economies, the scores are lowest in Sub-Saharan Africa and highest in Europe and Central Asia.⁸ There is however substantial variation within these regions, and especially within Sub-Saharan Africa (where the best-performing country, Mauritius, has a score of 1.93). The fact that the best performer in Sub-Saharan Africa (the region with the lowest average of TFI) has the highest score in the data suggests that a scenario in which all countries in the region move to the best performer's value is unlikely. We will take this into account in the simulations of Section 6.

< Table 2 about here >

3.3 Econometric model

The TFI data does not vary over time. We therefore estimate cross-sectional regressions for the year 2009. We choose this year for two reasons. First, this is suggested by Moisé and Sorescu (2013). Second, this allows us to construct measures for npd_{ij} and ndp_{ik} that are respectively based on new products and new destinations, to address endogeneity concerns (see Section 4).⁹

3.3.1 ij regressions

The ij regressions use, as dependent variable, the number of exported products, npd_{ij} . This is a bilateral measure of trade outcomes. It is therefore natural to employ a gravity framework, with the conditional mean of the dependent variable expressed as:

$$E[npd_{ij} | \mathbf{x}_{1i}, \mathbf{r}_i, \mathbf{g}_{ij}, \mathbf{m}_{ij}, \mathbf{d}_j] = g(\mathbf{x}'_{1i}\beta + \mathbf{r}'_i\theta + \mathbf{g}'_{ij}\delta + \mathbf{m}'_{ij}\eta + \mathbf{d}'_j\gamma) \quad (3.1)$$

In equation (3.1), $g(\cdot)$ is a function; \mathbf{x}_{1i} is a vector of variables that only vary across exporters i 's; \mathbf{r}_i is a vector of exporter-specific region dummies;¹⁰ \mathbf{w}_{ij} is a vector of standard bilateral gravity variables; \mathbf{m}_{ij} is a vector of multilateral resistance terms, constructed using the methodology outlined in Baier and Bergstrand

⁸It is important to note that the latter region does not include industrialized OECD countries – see Table A-2.

⁹The results with all time-varying variables averaged between 2002 and 2010 are very similar to the ones presented here and are available upon request.

¹⁰World Bank region dummies are included because in the simulations of Section 6 we average results over such regions.

(2009);¹¹ \mathbf{d}_j is a vector of importer (j)-specific effects; β , θ , δ , η and γ are vectors of coefficients to be estimated. Appendix Table A-3 provides a description of all the variables, including the data sources.¹² Table 3 presents summary statistics for all control variables.

< Table 3 about here >

As a first step, we estimate a linear fixed effects model by OLS, with the dependent variable in logs, conditioning on importer-specific effects. In this case, $g(\cdot)$ in (3.1) is the identity function. Since the dependent variable is a count variable, a model for count data is more appropriate (Dennis and Shepherd, 2011; Persson, 2013). Accordingly, we also adopt conditional Poisson and Negative Binomial (NB) Maximum Likelihood estimations, with the dependent variable in levels (always conditioning on importer-specific effects). In this case, $g(\cdot)$ in (3.1) is the exponential function.¹³

Appendix Table A-4 presents the in-sample correlations between all variables in the ij sample.

3.3.2 ik regressions

The ik regressions use, as dependent variable, the number of export destinations, ndp_{ik} . This measure of trade outcomes does not have any bilateral dimension, since it varies by exporting country i and by product k . We express the conditional mean of the dependent variable as:

$$E[ndp_{ik} | \mathbf{x}_{2i}, \mathbf{r}_i, \mathbf{h}_k] = g(\mathbf{x}'_{2i}\alpha + \mathbf{r}'_i\phi + \mathbf{h}'_k\lambda) \quad (3.2)$$

In equation (3.2), $g(\cdot)$ is a function; \mathbf{h}_k is a vector vector of product (k)-specific effects; α , ϕ and λ are vectors of coefficients to be estimated. The vectors \mathbf{x}_{2i} and \mathbf{x}_{1i} differ from one another in the following respects.

\mathbf{x}_{2i} , but not \mathbf{x}_{1i} , contains the variable $\text{Log}(\text{remoteness})$ – a weighted average of partners' (j 's) world GDP

¹¹In what Head and Mayer (2013) call the 'structural' representation of the gravity model, the multilateral resistance terms are exporter and importer specific. In cross-section estimation, it is possible to incorporate these terms as importer and exporter fixed effects *à la* Feenstra (2003). But since our measure of trade facilitation, TFI, is exporter specific, it is impossible to take this approach. Instead, given that the multilateral resistance terms enter multiplicatively in the gravity equation, we incorporate them as a pair or dyadic variable and estimate just one coefficient that measures their combined effect on the number of exported products.

¹²All variables listed in Appendix Table A-3 are self-explanatory, with the exception of the Market Access Trade Restrictiveness Index (TRI). The TRI captures the trade policy distortions imposed by the trading partners on i 's export bundle. It measures the uniform tariff equivalent of the partner country tariff and non-tariff barriers (NTB) that would generate the same level of i 's export value in a given year. The TRI index is constructed using applied tariffs.

¹³Both the Poisson and the NB fixed effect estimators are consistent even in short panels. The NB fixed effect estimator yields more efficient estimation in the presence of overdispersion (Cameron and Trivedi, 2013, p. 357). However, it does not allow for cluster-robust estimation of the variance-covariance matrix. The results are not qualitatively affected by the methodology adopted.

shares, with the inverse distance between i and j as weight (Head and Mayer, 2013). Further, \mathbf{x}_{1i} , but not \mathbf{x}_{2i} , contains the variable $\text{Log}(\text{weighted } j\text{'s TFI})$ – a weighted average of partners’ (j ’s) TFI indicators, with the inverse distance between i and j as weight.¹⁴

In this case, too, we present a fixed effect linear specification (OLS with dependent variable in logs) as a first step. $g(\cdot)$ in (3.2) is the identity function in this case. We then adopt conditional Poisson and Negative Binomial (NB) Maximum Likelihood estimations with the dependent variable in levels. $g(\cdot)$ in (3.2) is the exponential function in this case. We always condition on product-specific effects.

Appendix Table A-5 presents the in-sample correlations between all variables in the ik sample.

4 Results

4.1 ij regressions

The results of the ij regressions are in Table 4. In the OLS regressions, the dependent variable is in logs, while it is in levels in the Poisson and NB regressions. In both cases, however, the coefficients on the explanatory variables in logs can be interpreted as elasticities.¹⁵

< Table 4 about here >

The baseline results are in columns (1), (2) and (6), respectively for OLS, Poisson and NB regressions. The coefficient on the variable of interest, $\text{Log}(\text{TFI})$, is positive and statistically significant in all specifications. In the specification of column (6), the elasticity is 0.305, implying that a 1% increase in the average trade facilitation indicator is roughly associated with a 0.3% increase in the number of HS6 products exported by destination.

The coefficients on the control variables are correctly signed and statistically significant. Although the dependent variable is different, it is useful to compare the distance coefficients with the standard results from gravity studies. As reported in Table of 4 Head and Mayer (2013), the mean of the distance coefficient estimated in 159 papers ranges between -0.93 and -1.1, with a standard deviation of 0.40-0.41. The distance

¹⁴We are grateful to Richard Baldwin for suggesting the use of this variable.

¹⁵See Cameron and Trivedi (2013, p. 346). In the OLS regressions, there are fewer observations because all the observations for which $npd_{ij} = 0$ are dropped. The Poisson and NB regressions drop all observations with $npd_{ij} = 0$ across all destinations, because they do not contribute to the conditional likelihoods.

elasticity we obtain is in line with Table 4 of Head and Mayer (2013) for the OLS estimation. In the Poisson and NB model it is lower, but it is a well-established fact in the literature that the distance coefficient is lower when using count-data models. Moreover, our result is very similar to the one obtained by Persson (2013), which is the most comparable study to ours.

We see two possible concerns with the baseline estimations. First, and foremost, we cannot exclude reverse causation, that is the possibility that trade outcomes affect the incentives to invest in trade facilitation, and consequently the trade facilitation scores. In the spirit of Freund and Rocha (2011) and Portugal-Perez and Wilson (2012), we address possible reverse causality using only ‘new products’ (HS sub-headings) in the construction of the dependent variable. We proceed as follows: when computing how many products country i exported to country j in 2009, we only include the subset of products for which: (i) there were no exports from i to j (zero or missing) recorded in any of the years between 2002 and 2007; and (ii) there were positive exports from i to j recorded in at least one year between 2008 and 2010. Since npd_{ij} is, in this case, the count of new HS6 products that were not traded before 2008, it is less likely to be endogenous to trade facilitation than the indicator calculated using the set of products traded in 2009.

The use of ‘new products’ has an additional advantage. We do not necessarily exclude products that dropped from a country’s bilateral export basket during the big trade collapse of 2009. As long as a product that was not exported in any year between 2002 and 2007 started to be exported in any year before 2008 and 2010, it counts for the construction of npd_{ij} .

The results are in columns (3) and (7) of Table 4. The coefficients are slightly lower than in the baseline regressions (indicating the possibility of a small downward bias induced by reverse causality), but still positive and statistically significant.

The second possible concern with the baseline estimations of Table 4 relates to the measurement of trade facilitation. So far, we have used TFI – the unweighted average of the country-specific OECD Trade Facilitation Indicators. As an alternative, we have created a trade facilitation indicator based on Principal Component Analysis (PCA). The results of Poisson estimations are in columns (4) and (5). They are very similar to the corresponding results of columns (2) and (3).¹⁶

¹⁶To save space, we do not report NB results with TFI computed using Principal Component Analysis (PCA). They are similar to the results in columns (6) and (7) and available upon request.

4.2 *ik* regressions

The results of *ik* regressions are in Table 5. The structure of the table is similar to that of Table 4.

< Table 5 about here >

The coefficient on $\text{Log}(\text{TFI})$ is positive and significant in all specifications. The estimated elasticity in column (6) is 0.383. This implies that a 1% increase in the average trade facilitation indicator is roughly associated with a 0.38% increase in the number of destinations to which an HS6 product is exported.

In columns (3) and (7) we address possible reverse causality concerns using only ‘new destinations’ in the computation of the dependent variable. The procedure is very similar in spirit to the one described above in the case of *ij* regressions. When computing how many destination countries were served by country *i* in exporting product *k* in 2009, we only include the subset of destinations for which: (i) there were no exports of product *k* (zero or missing) recorded in any of the years between 2002 and 2007; (ii) there were positive exports of product *k* recorded in at least one year between 2008 and 2010. In this case, therefore, ndp_{ik} becomes the count of new destinations that were not served before 2008.

Also in this case, the use of ‘new destinations’ has the additional advantage that we do not necessarily exclude destinations that ceased to be served by country *i* in sector *k* during the big trade collapse of 2009. As long as a destination that was not served in any year between 2002 and 2007 started to get served in any year before 2008 and 2010, it counts for the construction of ndp_{ik} .

In the regressions with new destinations, the estimated coefficient on $\text{Log}(\text{TFI})$ is slightly larger than the baseline coefficient, both in the Poisson and in the NB specifications.

In columns (5) and (5) we present the results of the regressions that use a measure of TFI based on Principal Component Analysis, rather than the simple mean across indicators. Again, the results do not change significantly.¹⁷

¹⁷NB results are available upon request.

5 Robustness

5.1 Trade margins based on HS4 trade data

The results of ij and ik regressions using trade margins based on HS4 trade data are in Table 6. For ij regressions, Poisson and NB baseline results should be compared with columns (2) and (6) of Table 4, respectively. Results with new products (columns (2) and (4)) should be compared with columns (3) and (7) of Table 4. For ik regressions, columns (5) and (7) should be compared with columns (2) and (6) of Table 5, while columns (6) and (8) should be compared with columns (3) and (7) of Table 5. All coefficients on $\text{Log}(\text{TFI})$ are quite similar across comparable regressions. This implies that our results are not affected by the level of product disaggregation of the data.

< Table 6 about here >

5.2 Hummel-Klenow trade margins

In this section, we present econometric estimates using the theory-based ‘Hummels-Klenow extensive margins’ as dependent variables. In the regressions with country pairs, we use the following variable, directly from Hummels and Klenow (2005):

$$em_{ij} = \frac{\sum_{k \in K_{ij}} X_{wjk}}{\sum_{k \in K} X_{wjk}} \quad (5.1)$$

In equation (5.1), K_{ij} is the set of goods which country i exports to country j ; w is the reference country that has positive exports to j in all products k (in the empirical implementation, it is the rest of the world); K is the set of all products; X_{wjk} are the exports of country w to country j in product k . em_{ij} is therefore the share of those exports to j only in goods k that country i exports in total exports to country j .

In the regressions with country-product observations, we construct a similar measure (not previously used in the reviewed literature):

$$em_{ik} = \frac{\sum_{j \in J_{ik}} X_{wjk}}{\sum_{j \in J} X_{wjk}} \quad (5.2)$$

In equation (5.2), J_{ik} is the set of destinations to which country i exports product k ; w is the reference country that has positive exports of k to all destinations j (in the empirical implementation, it is the rest

of the world); J is the set of all destinations; X_{wjk} are – as in equation (5.1) – the exports of country w to country j in product k . em_{ik} is therefore the share of exports of k only to destinations j that country i exports to in total exports of product k to all destinations.¹⁸

The summary statistics for the Hummels-Klenow extensive margins em_{ij} and em_{ik} are in Appendix Table A-6. In the developing world, Hummels-Klenow extensive margins, and therefore export diversification, are lowest in Sub-Saharan Africa and highest in East Asia and Pacific. From a qualitative standpoint, these descriptive statistics are in line with the ones presented in Table 1 for npd_{ij} and ndp_{ik} . In fact, the sample correlation between npd_{ij} and em_{ij} is equal to 0.89, while the sample correlation between ndp_{ik} and em_{ik} is equal to 0.83.¹⁹

Table 7 present the results of ij and ik regressions using, as dependent variable, the Hummels-Klenow extensive margins em_{ij} and em_{ik} , respectively. The dependent variable ranges between zero and one. As suggested by Baum (2008), we use a Generalized Linear Model (GLM) with a logit transformation of the response variable and the binomial distribution. Odd-numbered columns present baseline results, in which the respective trade margin is calculated using trade data from 2009. In even-numbered columns we address concerns related to reverse causality and construct the dependent variable using only the subset of new products (in the case of em_{ij}) or new destinations (in the case of em_{ik}).²⁰

< Table 7 about here >

In the ij regressions, controlling for country characteristics, partner j dummies, region dummies and bilateral control variables, the coefficient on $\text{Log}(\text{TFI})$ is positive, but it loses statistical significance in column (2). In the ik regressions, where we control for country characteristics, product k dummies and region dummies, only the coefficient on $\text{Log}(\text{TFI})$ in column (3) is correctly signed and statistically significant. When we use the definition of the Hummels-Klenow extensive margin em_{ik} based on new destinations, this coefficient turns negative and statistically significant. There is no easy way to explain this counter-intuitive result. It should be mentioned, however, that the coefficient on $\text{Log}(\text{TFI})$ is correctly signed and statistically significant if we perform the same regression of column (4) of Table 7 using HS4 headings in the construction em_{ik} .

¹⁸We use the Stata module developed by Ansari (2013) to compute em_{ij} and em_{ik} .

¹⁹Sample correlations computed from columns (1) and (3) of Table 7, respectively.

²⁰See Section 4.1 for details on the procedure.

5.3 World Bank Doing Business indicators

Following, among others, Hoekman and Nicita (2011) and Dennis and Shepherd (2011), we also perform regressions that use, as proxies for trade facilitation, the ‘Trading across borders’ indicators of the World Bank Doing Business database. In this database, there are three indicators relevant for our purposes: number of documents to export;²¹ number of days required to export;²² and cost to export (US\$ per container).²³ To increase comparability with the results that use the OECD TFIs, we transform these variables as follows. First, we compute their inverse. Then, we rescale them between 0 (least facilitation) and 2 (most facilitation). Appendix Table A-7 presents the in-sample correlations among these variables, and the correlations between these variables and TFI.

The results are in Table 8. All the coefficients on DB docs, DB cost and DB time are consistently positive across all specifications, both in the ij and in the ik regressions. It should be noted that ij regressions use the same set of explanatory variables as Table 4 and ik regressions use the same set of explanatory variables as Table 5.

< Table 8 about here >

5.4 The elusive quest for heterogeneous effects

Beyond the central results of Section 4, we also investigate possible heterogeneity in the impact of trade facilitation on the extensive margins of trade. An important potential source of heterogeneity is between country pairs that have a PTA in place and country pairs that do not have one. There is ample evidence that most PTAs include trade facilitation provisions (see for instance Neufeld, 2014). Maur (2011) argues that in areas such as product standards and technical regulations, trade facilitation through policies such as harmonization between PTA members has the potential to introduce discrimination *vis-à-vis* excluded countries. Conversely, aspects of trade facilitation such as transparency and simplification of rules and procedures (the narrow definition of trade facilitation that we use in this paper and that is reflected in the

²¹The total number of documents required per shipment to export goods. Documents required for clearance by government ministries, customs authorities, port and container terminal authorities, health and technical control agencies and banks are taken into account.

²²The time necessary to comply with all procedures required to export goods.

²³The cost associated with all procedures required to export goods. It includes the costs for documents, administrative fees for customs clearance and technical control, customs broker fees, terminal handling charges and inland transport.

OECD TFIs), should be non-discriminatory in nature and therefore benefit all trading partners equally. If this is the case, one should not expect any heterogeneous effect of exporter’s trade facilitation on the extensive margin of bilateral trade across importers with a PTA with the exporter and importers without.

To test this prediction, we augment the ij regressions with an interaction term between the PTA dummy and $\text{Log}(\text{TFI})$. We do not obtain any consistent pattern in the results. In most regressions, the marginal effect when the PTA dummy is equal to one is not statistically different from the marginal effect when the PTA dummy is equal to zero. The evidence that an exporting country’s trade facilitation does not have a higher impact on the bilateral extensive margin with importers with which it has a PTA provides indirect support for the idea that the trade facilitation provisions captured by the OECD TFIs (and disciplined by the WTO Trade Facilitation Agreement) are non-discriminatory.

In a similar way, we investigate the effects of having a common border (without any significant result) or a common language. Although obtaining significant results in a few instances, indicating that trade facilitation is even more important if the trading partners have different languages, we do not find consistent results across all specifications.²⁴

One might be concerned that our results might be driven by developed countries only. To rule this out, we split the ij sample in four subsamples: DD (both exporter i and importer j are developed);²⁵ DG (i is developed, j is developing); GD (i is developing, j is developed); GG (both i and j are developing). As shown in Table 9, each of these subsample covers approximately one quarter of the ij dataset. Although varying in magnitude across different specifications, the results are qualitatively similar across sub-samples. We therefore reject the hypothesis that our results are driven by developed countries.

< Table 9 about here >

We further check whether the effect of trade facilitation on the extensive margins differs between final and intermediate products. Yi (2003) develops a model in which trade costs hamper vertically-specialized trade (i.e. trade along supply chains) relatively more than trade in final products.²⁶ Martinez-Zarzoso and Márquez-

²⁴The results are available upon request.

²⁵We consider ‘developed’ countries classified by the World Bank as High-income or Upper-middle-income economies. Low-income and Lower-middle-income economies are considered ‘developing’.

²⁶Yi’s model shows the magnifying trade effects of tariff reductions when vertically specialized goods cross multiple borders while they are being produced. He argues that reductions in transportation costs and trade reforms more general than tariff liberalization also have a magnifying effect on trade. Ferrantino (2012) makes the link with trade facilitation explicit. He argues

Ramos (2008) show that improvements in the Doing Business indicators ‘Number of days’ and ‘Document required’ to export/import have a relatively larger effect on technology-intensive goods and on differentiated products, as opposed to homogeneous ones. Marti et al. (2014) argue that improvements in the LPI have an effect which is larger for goods that are relatively more complex to transport. In a more direct test of Yi’s hypothesis, Saslavsky and Shepherd (2012) show that trade in parts and components – which they assume takes place largely within production networks – is more sensitive to improvements in logistics performance than trade in final goods. These papers focus on the intensive margin of trade (bilateral trade value in a gravity framework). As discussed in Section 2, Persson (2013) applies similar ideas to the extensive margins of trade. She does not explicitly consider trade in intermediate products as her focus is on product differentiation. She finds that trade facilitation has a higher extensive margin impact on trade in differentiated products than on trade in homogenous products.

In the spirit of this literature, we test for heterogeneous effects on the extensive margins of trade between intermediate and non-intermediate and products. We use two alternative definitions of intermediate products, a narrow one and a broad one. The narrow definition, adopted by WTO (2011), includes the HS sub-headings corresponding to codes 42 and 53 of the Broad Economic Categories (BEC) classification, supplemented with unfinished textile products in HS chapters 50-63. The broad definition includes the HS sub-headings corresponding to the intermediate goods of the BEC classification.

As a first step, we estimate the ij regressions in two sub-samples: one in which the dependent variable is computed across the subset of intermediate products; one in which the dependent variable is computed across the subset of all other products. We are not able to find any significant difference between estimated coefficients across these specifications. To test this result further, in the ik sample we augment the regressions with an interaction term between a dummy equal to one if the product is intermediate and the Log(TFI) variable. We do not find the coefficient of this interaction term to be significant in most specifications.²⁷

This leads us to conclude that the effect of trade facilitation on the extensive margin does not differ between intermediate and final products.

that NTMs and trade facilitation can be compared using a common metric. Efforts to reduce NTMs and efforts to increase trade facilitation should both have larger effects on trade in complex supply chains than on trade in simple supply chains. See also U.S. Chamber of Commerce (2014) and UNECA (2013) for less formal expositions, respectively by the business community and by an international organization, of the idea that trade facilitation should matter most for intermediate goods trade.

²⁷The results are available upon request.

6 Simulation results

In this section, we present the results of counterfactual analysis aimed at providing insights into the economic significance of our results. We estimate the percentage increase in the number of export destinations and in the number of exported products under two different scenarios. As shown in the ‘sd’ column of Table 2, there is wide variation in TFI scores across countries belonging to the same geographical region. This suggests that a scenario involving convergence to the top regional performer would be very unrealistic. Such a scenario is, therefore, discarded a priori. Instead we consider two realistic scenarios. The first scenario is one in which each country with a TFI score below the median of the geographical region it belongs to increases its TFI to the regional median. The second scenario considers an increase to the global median.

It is important to note that results of counterfactual analysis have to be taken cautiously. First, because they are only as good as the underlying econometric model. Although we have taken care in addressing omitted variable and reverse causality biases, we cannot control for every possible country-specific variable correlated with trade facilitation and we cannot completely exclude the endogenous co-determination of trade outcomes and trade facilitation infrastructure. Second, the counter-factual analysis does not take into account that regional (global) median values would be affected by changes in trade facilitation occurring in all countries in the region (world).

With these caveats in mind, the baseline results, grouped by region, are presented in Table 10 for ij regressions and Table 11 for ik regressions. To remain on the conservative side, we have chosen to base the simulations on the Negative Binomial (NB) results, which generally yield smaller estimated coefficients on $\text{Log}(\text{TFI})$ than the coefficients of Poisson regressions.²⁸

For ease of interpretation, it is useful to keep in mind that the entries in tables 10 and 11 represent the percentage change in the variable of interest (respectively, npd_{ij} and ndp_{ik}) that, based on the estimated regression coefficients, are predicted if country i moves from below the regional (global) median to the relevant median. The results are then averaged across regions. All countries at, or above, the relevant median are dropped from the calculation of the regional average percentage increase in the trade margin. If, say, in a given region there are 16 countries, 8 of which are below the regional median and 15 below the global median,

²⁸The simulation results using the coefficients from Poisson regressions are available upon request.

the results under the regional median scenario are averaged over the 8 bottom countries in terms of TFI, while the results under the global median scenario are averaged over all countries with the exclusion of the top regional performer.

Tables 10 and 11 have two panels each. In the upper panel, we present results based on regressions using HS6 trade data. In the lower panel, we present results based on regressions using HS4 trade data. We use both the ‘baseline’ NB and the NB specifications with new products and new destinations. Since the estimates obtained in the latter specifications address the issue of reverse causality, we take them as our preferred results. We therefore discuss only the results of even-numbered columns.

< Tables 10 and 11 about here >

As expected, the estimated gains in terms of number of products exported by destination (npd_{ij}) are quite similar in panels (a) and (b) of Table 10 (the correlation between npd_{ij} computed from HS6 trade data and npd_{ij} computed from HS4 trade data is 0.98), but also in panels (a) and (b) of Table 11. The percentage gains range from 4.1% in the case of South Asia (HS4 data, both scenarios) to 15.7% in the case of Sub-Saharan Africa (HS6 data, global median scenario). It is apparent from the table that the gains are largest in two regions, namely Sub-Saharan Africa and Latin America and the Caribbean.

In the case of the number of export destinations by product (ndp_{ik}), the estimated gains range from 7.2% for South Asia (regional median scenario, HS4 regressions) to 27.9% for Sub-Saharan Africa (global median scenario, HS6 regressions). In this case, too, the gains are largest in Sub-Saharan Africa and Latin America and the Caribbean.

7 Conclusions

This is the first paper to focus exclusively on, and to provide detailed estimates of, the prospective effect of the WTO Trade Facilitation Agreement on export diversification, as measured by the extensive margins of trade. We do so by using direct measures of trade facilitation that map into the obligations of the agreement, namely, the OECD Trade Facilitation Indicators. We explore a variety of measures of the extensive margins of trade – the number of products a country exports to a given destination (npd_{ij}), the number of destinations

to which a country exports a given product (ndp_{ik}), the Hummels-Klenow measure of the bilateral extensive margin (em_{ij}) and a similar measure of the country-product extensive margin (em_{ik}) that has not previously been explored in the literature.

The estimation results are convincing, with the coefficient on the trade facilitation variable being positive and statistically significant across almost all specifications. Using these estimates, we simulate the impact of implementing the agreement on developing countries' extensive margin of trade. Implementation of the agreement is measured using two alternative realistic scenarios – convergence to the regional median and convergence to the global median. Developing countries are likely to experience a substantial increase in the number of products exported and of destination markets. For Sub-Saharan African countries, our simulations suggest they could see an increase of up to 15.7% in the number of products exported by destination and an increase of up to 27.9% in the number of export destinations by product. For countries in Latin America and the Caribbean, our simulations suggest they could see an increase of up to 12.2% in the number of products exported by destination and an increase of up to 21.7% in the number of export destinations by product. For the reasons outlined in Section 6, these numbers have to be treated with caution. Nonetheless, they imply potentially sizeable impacts of the Trade Facilitation Agreement on extensive margins of export, and therefore on export diversification.

We see three potential avenues for further research. First, the analysis could be conducted at the level of specific TFI indicators. Such analysis would shed light on which provisions of the WTO Trade Facilitation Agreement have the most effect on the extensive margins of trade. Second, simulations could be based on actual implementation schedules which will have to be notified by WTO members. Third, research should focus on the effect of trade facilitation on trade margins computed from firm-level, as opposed to sector-level data.

Finally, we emphasize that we make no claim about the welfare effects of implementing the WTO Trade Facilitation Agreement. This would require us to estimate not only the benefits but also the costs of implementing the agreement. Notwithstanding this qualification, we know from the available literature that the costs of implementation of trade facilitation initiatives are relatively small (OECD, 2009; UNECA, 2013). At the same time, our estimations do not capture several other potential benefits of the agreement. A proper

welfare analysis would also factor in the value of locking in commitments in a multilateral agreement and other positive spillovers, such as, for instance, the reduction in the extent of rent-seeking behaviour or the environmental benefit of lower fuel consumption from shorter waiting times at the border. These topics need to be investigated further to get a more comprehensive understanding of the effects of the WTO Trade Facilitation Agreement.

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Tables

Table 1: Summary statistics of npd_{ij} and ndp_{ik} , by World Bank region

Panel (a): npd_{ij}

World Bank region	mean	median	sd	min	max	N	zeros	% zeros
Sub-Saharan Africa	43	3	200.43	0	4525	5157	1638	32%
East Asia and Pacific	501	111	793.09	0	4224	2254	382	17%
Europe and Central Asia	216	18	475.86	0	3788	3862	849	22%
Latin America and Caribbean	115	9	321.68	0	3429	3864	870	23%
Middle East and North Africa	121	20	280.25	0	3443	1772	300	17%
South Asia	283	53	548.43	0	3740	967	162	17%
Offshore	14	2	64.62	0	780	161	55	34%
Industrial	946	489	1088.38	0	4831	3220	72	2%
Whole sample	290	18	655.92	0	4831	21257	4329	20%

Descriptive statistics computed from the sample of column (6) of Table 4 and based on HS6 trade data

Panel (b): ndp_{ik} (HS6 trade data)

World Bank region	mean	median	sd	min	max	N	zeros	% zeros
Sub-Saharan Africa	1	0	4.87	0	128	167008	114129	68%
East Asia and Pacific	16	2	29.46	0	169	73066	26561	36%
Europe and Central Asia	7	1	13.49	0	135	125256	49152	39%
Latin America and Caribbean	4	1	9.12	0	137	125256	62326	50%
Middle East and North Africa	4	1	9.28	0	122	57409	28086	49%
South Asia	9	0	20.40	0	166	31314	16040	51%
Offshore	0	0	1.42	0	63	5219	3972	76%
Industrial	30	16	34.41	0	167	104380	12544	12%
Whole sample	9	1	21.16	0	169	688908	312810	45%

Descriptive statistics computed from the sample of column (6) of Table 5

Panel (c): ndp_{ik} (HS4 trade data)

World Bank region	mean	median	sd	min	max	N	zeros	% zeros
Sub-Saharan Africa	3	1	8.87	0	138	39808	18523	47%
East Asia and Pacific	28	8	40.43	0	174	17416	4039	23%
Europe and Central Asia	13	4	20.71	0	146	29856	6567	22%
Latin America and Caribbean	8	2	15.22	0	141	29856	8677	29%
Middle East and North Africa	10	3	16.74	0	137	13684	3598	26%
South Asia	17	2	31.44	0	169	7464	2519	34%
Offshore	1	0	3.00	0	67	1244	672	54%
Industrial	50	38	44.34	0	173	24880	1246	5%
Whole sample	17	3	30.41	0	174	164208	45841	28%

Descriptive statistics computed from the sample of column (7) of Table 6

Table 2: Summary statistics, TFI, by World Bank region

World Bank region	mean	median	sd	min	max	N
Sub-Saharan Africa	1.04	1.01	0.35	0.22	1.93	5157
East Asia and Pacific	1.32	1.33	0.27	0.81	1.81	2254
Europe and Central Asia	1.39	1.37	0.28	0.77	1.91	3862
Latin America and Caribbean	1.20	1.28	0.32	0.45	1.65	3864
Middle East and North Africa	1.21	1.15	0.27	0.83	1.65	1772
South Asia	1.23	1.21	0.16	1.01	1.38	967
Offshore	1.20	1.20	0.00	1.20	1.20	161
Industrial	1.50	1.54	0.18	1.13	1.86	3220
Whole sample	1.26	1.29	0.33	0.22	1.93	21257

Descriptive statistics computed from the sample of column (6) of Table 4

Table 3: Summary statistics, control variables

Variable	mean	median	sd	min	max
Log(pcGDP)	8.49	8.51	1.46	5.38	11.27
Log(market access)	-2.43	-2.38	0.76	-5.37	-1.15
Number of PTAs	40.53	41.00	25.98	0	88
Log(area)	11.90	12.01	2.11	5.76	16.65
Landlocked	0.21	0.00	0.41	0	1
Log(weighted j 's TFI)	0.26	0.26	0.06	0.01	0.35
Log(remoteness)	8.50	8.62	0.50	7.18	9.36
Log(bilateral GDP)	7.46	7.41	3.08	-2.15	18.11
PTA	0.22	0.00	0.41	0	1
Log(distance)	8.73	8.92	0.78	4.74	9.89
Common border	0.02	0.00	0.14	0	1
Common language	0.14	0.00	0.35	0	1
Colony	0.01	0.00	0.09	0	1
MR PTA	0.25	-0.05	0.81	-0.18	7.29
MR Log(distance)	10.66	-3.39	46.88	-7.59	481.69
MR Common border	0.00	-0.05	0.15	-0.05	1.56
MR Common language	0.20	-0.07	1.32	-0.10	13.54
MR Colony	0.04	-0.01	0.22	-0.01	2.51

Descriptive statistics computed from the sample of column (6) of Table 4 – except for Log(remoteness)

Descriptive statistics for Log(remoteness) computed from the sample of column (6) of Table 5

Table 4: Number of products by destination (ij regressions)

Dependent variable: $\log(npd_{ij})$ (OLS regression); npd_{ij} (Poisson and NB regressions)

	OLS			Poisson			NB		
	Baseline (1)	Baseline (2)	New HS6 (3)	PCA for TFI (4)	PCA for TFI & new HS6 (5)	Baseline (6)	New HS6 (7)		
Log(TFI)	0.290*** [0.038]	0.549*** [0.057]	0.346*** [0.078]	0.608*** [0.064]	0.319*** [0.081]	0.305*** [0.024]	0.230*** [0.028]		
Log(pc GDP)	0.113*** [0.018]	0.097*** [0.025]	0.088*** [0.030]	0.093*** [0.025]	0.090*** [0.030]	0.218*** [0.017]	0.102*** [0.019]		
Log(market access)	0.433*** [0.013]	0.401*** [0.018]	0.304*** [0.021]	0.394*** [0.018]	0.304*** [0.021]	0.320*** [0.010]	0.237*** [0.009]		
Number of PTAs	-0.002*** [0.001]	0.002*** [0.001]	0.001* [0.001]	0.002** [0.001]	0.001 [0.001]	0.003*** [0.000]	0.002*** [0.000]		
Log(area)	-0.058*** [0.007]	-0.048*** [0.009]	-0.021** [0.009]	-0.047*** [0.009]	-0.020** [0.009]	0.040*** [0.008]	0.015** [0.007]		
Landlocked	-0.248*** [0.024]	0.011 [0.028]	-0.156*** [0.027]	0.015 [0.028]	-0.155*** [0.027]	-0.078*** [0.023]	-0.200*** [0.012]		
Log(weighted j 's TFI)	4.718*** [0.291]	1.657*** [0.432]	1.082*** [0.339]	1.796*** [0.431]	1.096*** [0.343]	0.718*** [0.231]	0.708*** [0.186]		
Log(bilateral GDP)	0.788*** [0.014]	0.735*** [0.021]	0.476*** [0.027]	0.730*** [0.021]	0.475*** [0.026]	0.397*** [0.022]	0.366*** [0.015]		
PTA	0.115*** [0.041]	0.054 [0.039]	0.004 [0.033]	0.051 [0.039]	0.005 [0.033]	0.050 [0.030]	-0.002 [0.029]		
Log(distance)	-0.931*** [0.036]	-0.598*** [0.041]	-0.431*** [0.035]	-0.600*** [0.041]	-0.432*** [0.035]	-0.477*** [0.030]	-0.293*** [0.028]		
Common border	0.482*** [0.116]	-0.036 [0.086]	-0.110 [0.103]	-0.033 [0.085]	-0.109 [0.103]	-0.184*** [0.066]	-0.526*** [0.133]		
Common language	0.757*** [0.051]	0.391*** [0.055]	0.422*** [0.046]	0.383*** [0.055]	0.420*** [0.046]	0.365*** [0.040]	0.306*** [0.040]		
Colony	0.777*** [0.135]	0.582*** [0.106]	0.371*** [0.108]	0.589*** [0.106]	0.374*** [0.108]	0.337*** [0.102]	0.259** [0.126]		
Observations	16,928	21,257	21,388	21,257	21,388	21,257	21,388		
R-squared	0.740								
Log pseudolikelihood		-1.349e+06	-784586	-1.345e+06	-784897	-97952	-99089		
Number of id (j 's)	162	162	163	162	163	162	163		

Robust (clustered on partner j) standard errors in parentheses in columns (1)-(5)

Bootstrap standard errors (100 replications) in parentheses in columns (6)-(7)

* p<0.10, ** p<0.05, *** p<0.01

Partner j fixed effects, region dummies and MR controls always included

All regressions based on HS6 trade data

Table 5: Number of destinations by product (*ik* regressions)

Dependent variable: $\log(ndp_{ik})$ (OLS regression); ndp_{ik} (Poisson and NB regressions)

	OLS			Poisson			NB		
	Baseline (1)	Baseline (2)	New dest. (3)	PCA for TFI (4)	PCA for TFI & new dest. (5)	Baseline (6)	New dest. (7)		
Log(TFI)	0.345* [0.200]	0.372*** [0.007]	0.439*** [0.005]	0.484*** [0.007]	0.438*** [0.005]	0.383*** [0.006]	0.410*** [0.005]		
Log(pc GDP)	0.536*** [0.071]	0.662*** [0.005]	0.475*** [0.004]	0.646*** [0.005]	0.472*** [0.004]	0.576*** [0.003]	0.493*** [0.002]		
Log(market access)	0.370*** [0.077]	0.491*** [0.003]	0.338*** [0.003]	0.478*** [0.003]	0.336*** [0.003]	0.522*** [0.003]	0.374*** [0.003]		
Number of PTAs	0.001 [0.002]	0.004*** [0.000]	0.003*** [0.000]	0.004*** [0.000]	0.003*** [0.000]	0.002*** [0.000]	0.003*** [0.000]		
Log(area)	0.313*** [0.025]	0.380*** [0.002]	0.232*** [0.002]	0.379*** [0.002]	0.232*** [0.002]	0.349*** [0.001]	0.246*** [0.001]		
Landlocked	-0.238** [0.116]	-0.331*** [0.004]	-0.369*** [0.003]	-0.329*** [0.004]	-0.370*** [0.003]	-0.394*** [0.004]	-0.383*** [0.003]		
Log(remote)	-1.185*** [0.136]	-1.237*** [0.009]	-0.591*** [0.006]	-1.245*** [0.009]	-0.595*** [0.006]	-1.326*** [0.008]	-0.684*** [0.006]		
Observations	376,095	688,908	689,172	688,908	689,172	688,908	689,172		
R-squared	0.590								
Log pseudolikelihood		-2.621e+06	-1.422e+06	-2.614e+06	-1.421e+06	-1.432e+06	-1.235e+06		
Number of id (HS6)	5,216	5,219	5,221	5,219	5,221	5,219	5,221		

Two-way clustered (*ik*) standard errors in parentheses in column (1)

Robust (clustered on HS6 products) standard errors in parentheses in columns (2)-(5)

Bootstrap standard errors (100 replications) in parentheses in columns (6)-(7)

* p<0.10, ** p<0.05, *** p<0.01

Product (HS6) fixed effects and region dummies always included

All regressions based on HS6 trade data

Table 6: ij and ik regressions with HS4 headings

Dependent variable: npd_{ij} (ij regressions); ndp_{ik} (ik regressions)

	ij regressions				ik regressions			
	Poisson		NB		Poisson		NB	
	Baseline (1)	New HS6 (2)	Baseline (3)	New HS6 (4)	Baseline (5)	New dest. (6)	Baseline (7)	New dest. (8)
Log(TFI)	0.521*** [0.045]	0.232*** [0.059]	0.348*** [0.028]	0.194*** [0.029]	0.458*** [0.009]	0.369*** [0.008]	0.389*** [0.006]	0.365*** [0.007]
Log(pc GDP)	0.095*** [0.019]	0.078*** [0.025]	0.181*** [0.018]	0.031 [0.021]	0.598*** [0.007]	0.357*** [0.007]	0.534*** [0.005]	0.394*** [0.006]
Log(market access)	0.348*** [0.013]	0.234*** [0.018]	0.308*** [0.009]	0.193*** [0.010]	0.410*** [0.005]	0.254*** [0.004]	0.447*** [0.004]	0.291*** [0.004]
Number of PTAs	0.001*** [0.000]	0.000 [0.001]	0.003*** [0.000]	0.001*** [0.000]	0.004*** [0.000]	0.002*** [0.000]	0.003*** [0.000]	0.003*** [0.000]
Log(area)	-0.033*** [0.007]	-0.007 [0.007]	0.025*** [0.007]	-0.005 [0.008]	0.317*** [0.003]	0.164*** [0.003]	0.309*** [0.002]	0.182*** [0.003]
Landlocked	-0.036* [0.021]	-0.204*** [0.025]	-0.103*** [0.016]	-0.235*** [0.013]	-0.327*** [0.006]	-0.350*** [0.005]	-0.358*** [0.006]	-0.358*** [0.005]
Log(weighted j 's TFI)	1.450*** [0.339]	0.717*** [0.276]	0.801*** [0.241]	0.850*** [0.177]				
Log(remote)					-0.962*** [0.013]	-0.348*** [0.010]	-1.098*** [0.012]	-0.410*** [0.010]
Observations	21,386	21,388	21,386	21,388	164,208	164,208	164,208	164,208
Log pseudolikelihood	-579564	-320545	-89594	-85150	-886572	-422873	-460360	-369748
Number of id (j 's)	163	163	163	163				
Number of id (HS4)					1,244	1,244	1,244	1,244

Robust (clustered on partner j) standard errors in parentheses in columns (1)-(2)

Robust (clustered on HS4 headings) standard errors in parentheses in columns (5)-(6)

Bootstrap standard errors (100 replications) in parentheses in columns (3)-(4) and (7)-(8)

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Partner j fixed effects, region dummies, pair and MR controls included in columns (1)-(4)

Heading (HS4) fixed effects and region dummies included in columns (5)-(8)

All regressions based on HS4 trade data

Table 7: Hummels-Klenow extensive margins
 Dependent variable: em_{ij} (ij regressions); em_{ik} (ik regressions)

	<i>ij</i> regressions		<i>ik</i> regressions	
	Baseline (1)	New HS6 (2)	Baseline (3)	New dest. (4)
Log(TFI)	0.457*** [0.054]	0.118 [0.091]	0.339*** [0.015]	-0.180*** [0.017]
Log(pc GDP)	0.249*** [0.023]	0.134*** [0.041]	0.599*** [0.008]	0.586*** [0.006]
Log(market access)	0.285*** [0.019]	0.155*** [0.030]	0.504*** [0.006]	0.387*** [0.006]
Number of PTAs	-0.001* [0.000]	-0.001 [0.001]	0.004*** [0.000]	0.003*** [0.000]
Log(area)	-0.018** [0.008]	-0.009 [0.012]	0.428*** [0.002]	0.283*** [0.002]
Landlocked	-0.087*** [0.028]	-0.200*** [0.045]	-0.172*** [0.008]	-0.202*** [0.008]
Log(weighted j's TFI)	0.592 [0.386]	0.219 [0.658]		
Log(remote)			-1.608*** [0.012]	-0.932*** [0.011]
Observations	16,881	18,478	370,035	407,097
Log pseudolikelihood	-3994.8	-2127.7	-118261.8	-94469.8

Generalized Linear Model (GLM) regressions in all columns

Robust (clustered on partner j) standard errors in parentheses in columns (1)-(2)

Robust (clustered on HS6 products) standard errors in parentheses in columns (3)-(4)

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Partner j dummies, region dummies, pair and MR controls included in columns (1)-(2)

Product (HS6) and region dummies included in columns (3)-(4)

All regressions based on HS6 trade data

Table 8: Results with Doing Business variables

<i>ij regressions, dependent variable: npd_{ij}</i>						
	NB baseline			NB new HS6		
	(1)	(2)	(3)	(4)	(5)	(6)
log(DB docs)	0.183*** [0.015]			0.194*** [0.012]		
log(DB cost)		0.202*** [0.017]			0.210*** [0.009]	
log(DB time)			0.214*** [0.014]			0.167*** [0.014]
Observations	20,613	20,613	20,613	20,740	20,740	20,740
Log pseudolikelihood	-95679	-95585	-95399	-96559	-96410	-96376
Number of id (<i>j</i> 's)	162	162	162	163	163	163
<i>ik regressions, dependent variable: ndp_{ik}</i>						
	NB baseline			NB new dest.		
	(1)	(2)	(3)	(4)	(5)	(6)
log(DB docs)	0.078*** [0.003]			0.169*** [0.003]		
log(DB cost)		0.330*** [0.003]			0.212*** [0.002]	
log(DB time)			0.088*** [0.003]			0.213*** [0.002]
Observations	668,032	668,032	668,032	668,288	668,288	668,288
Log pseudolikelihood	-1.410e+06	-1.401e+06	-1.405e+06	-1.210e+06	-1.208e+06	-1.205e+06
Number of id (HS6)	5,219	5,219	5,219	5,221	5,221	5,221

Bootstrap standard errors (100 replications) in parentheses

* p<0.10, ** p<0.05, *** p<0.01

Partner *j* fixed effects, region dummies, exporter and MR controls always included in *ij* regressions

Product (HS6) fixed effects and region dummies included in *ik* regressions

All regressions based on HS6 trade data

Table 9: ij regressions with income group sub-samplesDependent variable: npd_{ij}

	Poisson baseline (1)	Poisson new HS6 (2)	NB baseline (3)	NB new HS6 (4)
<i>DD subsample: i = developed, j = developed</i>				
Log(TFI)	0.529*** [0.051]	0.589*** [0.067]	0.414*** [0.035]	0.414*** [0.043]
Observations	5,359	5,425	5,359	5,425
Log pseudolikelihood	-408360	-196406	-30534	-28399
Number of id (j 's)	81	82	81	82
<i>DG subsample: i = developed, j = developing</i>				
Log(TFI)	0.514*** [0.192]	0.122 [0.253]	0.480*** [0.055]	0.299*** [0.047]
Observations	5,427	5,427	5,427	5,427
Log pseudolikelihood	-389952	-240605	-25463	-25839
Number of id (j 's)	81	81	81	81
<i>GD subsample: i = developing, j = developed</i>				
Log(TFI)	0.208** [0.096]	0.177*** [0.059]	0.363*** [0.054]	0.200*** [0.035]
Observations	5,265	5,330	5,265	5,330
Log pseudolikelihood	-141210	-101560	-22299	-22466
Number of id (j 's)	81	82	81	82
<i>GG subsample: i = developing, j = developing</i>				
Log(TFI)	0.640*** [0.215]	0.279* [0.144]	0.185* [0.097]	0.057 [0.063]
Observations	5,206	5,206	5,206	5,206
Log pseudolikelihood	-130143	-121028	-16774	-19255
Number of id (j 's)	81	81	81	81

Robust (clustered on partner j) standard errors in parentheses in columns (1)-(2)

Bootstrap standard errors (100 replications) in parentheses in columns (3)-(4)

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$ Partner j fixed effects, region dummies, exporter and MR controls always included

All regressions based on HS6 trade data

Table 10: Simulation results based on NB ij regressionsPanel (a): Simulations based on npd_{ij} computed from HS6 trade data

	Regional median		Global median	
	Baseline (1)	New HS6 (2)	Baseline (3)	New HS6 (4)
Sub-Saharan Africa	15.5%	12.6%	20.8%	15.7%
East Asia and Pacific	6.9%	5.6%	7.3%	5.5%
Europe and Central Asia	7.5%	6.1%	6.7%	5.1%
Latin America and the Caribbean	14.2%	11.6%	16.2%	12.2%
Middle East and North Africa	5.6%	4.5%	8.7%	6.6%
South Asia	5.5%	4.5%	6.4%	4.8%

Columns (1) and (3) based on column (6) of Table 4

Columns (2) and (4) based on column (7) of Table 4

Panel (b): Simulations based on npd_{ij} computed from HS4 trade data

	Regional median		Global median	
	Baseline (1)	New HS4 (2)	Baseline (3)	New HS4 (4)
Sub-Saharan Africa	17.2%	11.6%	23.7%	13.2%
East Asia and Pacific	7.6%	5.2%	8.3%	4.6%
Europe and Central Asia	8.3%	5.6%	7.7%	4.3%
Latin America and the Caribbean	15.7%	10.6%	18.5%	10.3%
Middle East and North Africa	6.2%	4.2%	9.9%	5.5%
South Asia	6.1%	4.1%	7.3%	4.1%

Columns (1) and (3) based on column (1) of Table 6

Columns (2) and (4) based on column (2) of Table 6

Table 11: Simulation results based on NB ik regressions

Panel (a): Simulations based on ndp_{ik} computed from HS6 trade data

	Regional median		Global median	
	Baseline	New dest.	Baseline	New dest.
	(1)	(2)	(3)	(4)
Sub-Saharan Africa	20.7%	22.1%	26.1%	27.9%
East Asia and Pacific	9.5%	10.1%	9.1%	9.8%
Europe and Central Asia	10.3%	11.0%	8.4%	9.0%
Latin America and the Caribbean	19.5%	20.9%	20.3%	21.7%
Middle East and North Africa	7.7%	8.2%	10.9%	11.7%
South Asia	7.6%	8.1%	8.1%	8.6%

Columns (1) and (3) based on column (6) of Table 5

Columns (2) and (4) based on column (7) of Table 5

Panel (b): Simulations based on ndp_{ik} computed from HS4 trade data

	Regional median		Global median	
	Baseline	New dest.	Baseline	New dest.
	(1)	(2)	(3)	(4)
Sub-Saharan Africa	21.0%	19.7%	26.5%	24.9%
East Asia and Pacific	9.6%	9.0%	9.3%	8.7%
Europe and Central Asia	10.4%	9.8%	8.6%	8.0%
Latin America and the Caribbean	19.8%	18.6%	20.6%	19.4%
Middle East and North Africa	7.8%	7.3%	11.1%	10.4%
South Asia	7.7%	7.2%	8.2%	7.7%

Columns (1) and (3) based on column (3) of Table 6

Columns (2) and (4) based on column (4) of Table 6

Appendix tables

Table A-1: Mapping of OECD TFIs into DCNT and TFA provisions

Indicator	DCNT Rev. 18	TFA
A. Information availability	Articles 1 and 2	Articles 1 and 2
B. Involvement of the trade community	Article 2	Article 2
C. Advance Rulings	Article 3	Article 3
D. Appeal Procedures	Article 4	Article 4
E. Fees and charges	Article 6.1 and 6.2	Article 6.1 and 6.2
F. Formalities – Documents	Articles 7 and 10	Articles 7 and 10
G. Formalities – Automation	Articles 7 and 10	Articles 7 and 10
H. Formalities – Procedures	Articles 5, 7 and 10	Articles 5, 7 and 10
I. Cooperation – Internal	Articles 9.1 and 12	Articles 8.1 and 12
J. Cooperation – External	Articles 9.2 and 12	Articles 8.2 and 12
K. Consularization	Article 8	–
L. Governance and Impartiality	–	–
M. Transit fees and charges	Article 11	Article 11
N. Transit formalities	Article 11	Article 11
O. Transit guarantees	Article 11	Article 11
P. Transit agreements and cooperation	Article 11	Article 11

TFI's stand for "Trade Facilitation Indicators"

DCNT stands for (WTO's) "Draft Consolidated Negotiating Text"

TFA stands for (WTO's) "Trade Facilitation Agreement"

Source: Moisé et al. (2011)

Table A-2: List of countries with OECD TFI data, by World Bank region

<i>Sub-Saharan Africa</i>			
Angola (1994)	Benin (1963)	Botswana (1987)	Burkina Faso (1963)
Burundi (1965)	Cameroon (1963)	Congo (1963)	Côte d'Ivoire (1963)
Ethiopia*	Gabon (1963)	Gambia (1965)	Ghana (1957)
Kenya (1964)	Lesotho (1988)	Liberia*	Madagascar (1963)
Malawi (1964)	Mali (1993)	Mauritius (1970)	Mozambique (1992)
Namibia (1992)	Nigeria (1960)	Rwanda (1966)	Senegal (1963)
Sierra Leone (1961)	South Africa (1948)	Swaziland (1993)	Tanzania (1961)
Togo (1964)	Uganda (1962)	Zambia (1982)	Zimbabwe (1948)
<i>East Asia and Pacific</i>			
Brunei Dar. (1993)	Cambodia (2004)	China (2001)	Chinese Taipei (2002)
Fiji (1993)	Hong Kong, China (1986)	Indonesia (1950)	Korea, Rep. (1967)
Malaysia (1957)	Mongolia (1997)	Papua N. G. (1994)	Philippines (1979)
Singapore (1973)	Thailand (1982)	Viet Nam (2007)	
<i>Europe and Central Asia</i>			
Albania (2000)	Armenia (2003)	Azerbaijan*	Belarus*
Bosnia and Herzegovina*	Bulgaria (1996)	Croatia (2000)	Czech Rep. (1993)
Georgia (2000)	Hungary (1973)	Kazakhstan*	Kyrgyz Rep. (1998)
Latvia (1999)	Lithuania (2001)	Moldova (2001)	Montenegro (2012)
Poland (1967)	Romania (1971)	Russian Fed. (2012)	Serbia*
Slovak Republic (1993)	The FYROM (2003)	Turkey (1951)	Ukraine (2008)
<i>Latin America and the Caribbean</i>			
Antigua and Barb. (1987)	Argentina (1967)	Barbados (1967)	Belize (1983)
Bolivia (1990)	Brazil (1948)	Colombia (1981)	Costa Rica (1990)
Cuba (1948)	Dominican Rep. (1950)	Ecuador (1996)	El Salvador (1991)
Guatemala (1991)	Honduras (1994)	Jamaica (1963)	Mexico (1986)
Nicaragua (1950)	Panama (1997)	Paraguay (1994)	Peru (1951)
Suriname (1978)	Trinidad and Tob. (1962)	Uruguay (1953)	Venezuela (1990)
<i>Middle East and North Africa</i>			
Algeria*	Bahrain (1993)	Jordan (2000)	Kuwait (1963)
Lebanon*	Morocco (1987)	Oman (2000)	Qatar (1994)
Saudi Arabia (2005)	Tunisia (1990)	UAE (1994)	
<i>South Asia</i>			
Bangladesh (1972)	Bhutan*	India (1948)	Nepal (2004)
Pakistan (1948)	Sri Lanka (1948)		
<i>Offshore</i>			
Bahamas*			
<i>Industrial</i>			
Australia (1948)	Belgium (1948)	Canada (1948)	Cyprus (1963)
Denmark (1950)	France (1948)	Germany (1951)	Greece (1950)
Italy (1950)	Japan (1955)	Malta (1964)	Netherlands (1948)
New Zealand (1948)	Norway (1948)	Portugal (1962)	Spain (1963)
Sweden (1950)	Switzerland (1966)	United Kingdom (1948)	United States (1948)

* WTO observer government

Year of WTO (GATT, where applicable) membership in parentheses

For official country names, refer to http://www.wto.org/english/thewto_e/whatis_e/tif_e/org6_e.htm

Table A-3: Variables list

Variable	Description	Data source
<i>Dependent variables of ij regressions</i>		
$npd_{i,j}$	Number of products by destination	UN Comtrade
$em_{i,j}$	Hummels and Klenow (2005) bilateral extensive margin	UN Comtrade
<i>Dependent variables of ik regressions</i>		
$ndp_{i,k}$	Number of destinations by product	UN Comtrade
$em_{i,j}$	Product extensive margin	UN Comtrade
<i>Exporter controls (matrices \mathbf{x}_1, and \mathbf{x}_2)</i>		
TFI	Average Trade Facilitation Index	OECD TFI dataset
Log(pcGDP)	Log of GDP per capita (current US\$)	World Development Indicators (WDIs)
Log(market access)	Log of Market Access Trade Restrictiveness Index	Kee et al. (2009)
Number of PTAs	Number of Preferential Trade Agreements signed	WTO PTAs dataset
Log(area)	Log of area (sq. kms)	CEPII gravity dataset (Head et al., 2010)
Landlocked	Dummy equal one if i is landlocked	CEPII gravity dataset (Head et al., 2010)
Log(weighted j 's TFI) ^a	Inverse distance-weighted average of partners' TFI	OECD TFI and CEPII gravity datasets
Log(remoteness) ^b	Log of remoteness (Head and Mayer, 2013 definition)	CEPII gravity dataset (Head et al., 2010)
<i>Pair controls (matrix \mathbf{w})</i>		
Log(bilateral GDP)	Log of $GDP_i * GDP_j$ (current US\$)	World Development Indicators (WDIs)
PTA	Dummy equal one if i and j have a Preferential Trade Agreement	WTO PTAs dataset
Log(distance)	Weighted distance between o and d (pop-wt, km)	CEPII gravity dataset (Head et al., 2010)
Common border	Dummy equal one if o and d share a common border	– ” –
Common language	Dummy equal one if o and d share an official or primary language	– ” –
Colony	Dummy equal one if o and d were in a colonial relation post 1945	– ” –
<i>Multilateral resistance (MR) controls (matrix \mathbf{r})</i>		
MR PTA	Baier and Bergstrand (2009) MR term based on PTA dummy	WTO PTAs dataset
MR Log(distance)	Baier and Bergstrand (2009) MR term based on Log(distance)	CEPII gravity dataset (Head et al., 2010)
MR Common border	Baier and Bergstrand (2009) MR term based on Common border dummy	– ” –
MR Common language	Baier and Bergstrand (2009) MR term based on Common language dummy	– ” –
MR Colony	Baier and Bergstrand (2009) MR term based on Colony dummy	– ” –

^a Only in ij regressions^b Only in ik regressions

Table A-4: Correlations (*ij* sample)

	npd	Log(TFI)	Log(pcGDP)	Log(market access)	Number of PTAs	Log(area)	Landlocked	Log(weighted <i>j</i> 's TFI)	Log(bilateral GDP)	PTA	Log(distance)	Common border	Common language	Colony
npd	1													
Log(TFI)	0.22*	1												
Log(pcGDP)	0.35*	0.39*	1											
Log(market access)	-0.00	0.10*	-0.26*	1										
Number of PTAs	0.23*	0.26*	0.33*	0.11*	1									
Log(area)	0.19*	0.03*	-0.17*	-0.10*	-0.03*	1								
Landlocked	-0.15*	-0.05*	-0.32*	0.04*	-0.21*	0.00	1							
Log(weighted <i>j</i> 's TFI)	0.19*	0.23*	0.38*	-0.10*	0.23*	-0.06*	0.05*	1						
Log(bilateral GDP)	0.57*	0.29*	0.40*	-0.10*	0.24*	0.32*	-0.21*	0.14*	1					
PTA	0.25*	0.09*	0.09*	0.04*	0.32*	-0.00	-0.08*	0.06*	0.17*	1				
Log(distance)	-0.29*	-0.02*	-0.053	0.07*	-0.14*	0.06*	-0.05*	-0.11*	-0.07*	-0.40*	1			
Common border	0.23*	0.01	-0.01	-0.01	0.01	0.05*	0.02*	0.01	0.05*	0.21*	-0.37*	1		
Common language	0.05*	-0.06*	-0.08*	0.10*	-0.04*	-0.04*	-0.01	-0.18*	-0.13*	0.11*	-0.11*	0.12*	1	
Colony	0.12*	0.02*	0.05*	0.00	0.05*	0.02*	-0.02*	0.03*	0.08*	0.03*	-0.03*	0.05*	0.14*	1

Correlations computed from the sample of column (6) of Table 4

* p<0.05

Table A-5: Correlations (*ik* sample)

	ndp	Log(TFI)	Log(pcGDP)	Log(market access)	Number of PTAs	Log(area)	Landlocked	Log(remoteness)
ndp	1							
Log(TFI)	0.22*	1						
Log(pcGDP)	0.35*	0.38*	1					
Log(market access)	0.00	0.10*	-0.26*	1				
Number of PTAs	0.22*	0.26*	0.33*	0.11*	1			
Log(area)	0.19*	0.03*	-0.17*	-0.10*	-0.03*	1		
Landlocked	-0.15*	-0.05*	-0.32*	0.04*	-0.21*	0.00	1	
Log(remoteness)	-0.35*	-0.34*	-0.60*	0.16*	-0.47*	0.22*	0.08*	1

Correlations computed from the sample of column (6) of Table 5

* p<0.05

Table A-6: Summary statistics, Hummels-Klenow extensive margins

World Bank region	em_{ij}			em_{ik}		
	mean	sd	N	mean	sd	N
Sub-Saharan Africa	0.06	0.12	3509	0.07	0.15	50954
East Asia and Pacific	0.27	0.26	1867	0.37	0.33	46099
Europe and Central Asia	0.16	0.21	2985	0.19	0.24	75335
Latin America and Caribbean	0.10	0.17	2991	0.13	0.20	62038
Middle East and North Africa	0.12	0.16	1471	0.12	0.18	28156
South Asia	0.15	0.21	805	0.28	0.32	15147
Offshore	0.04	0.07	106	0.05	0.09	1242
Industrial	0.40	0.29	3147	0.47	0.33	91064
Whole sample	0.18	0.24	16881	0.25	0.30	370035

Descriptive statistics for em_{ij} computed from the sample of column (1) of Table 7Descriptive statistics for em_{ik} computed from the sample of column (3) of Table 7

All descriptive statistics based on HS6 trade data

Table A-7: Correlation between TFI and Doing business variables

	TFI	DB docs	DB cost	DB time
TFI	1			
DB docs	0.44*	1		
DB cost	0.30*	0.37*	1	
DB time	0.51*	0.66*	0.45*	1

Correlations computed from the sample of column (1) of Table 8

* p<0.05