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CDS and equity market reactions to stock issuances in the U.S. financial industry: Evidence from the 2002-13 period

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CDS and Equity Market Reactions to Stock Issuances in the U.S. Financial Industry: Evidence from the 2002-13 Period

Marcia Millon Cornett, Hamid Mehran, Kevin Pan, Minh Phan, and Chenyang Wei *Federal Reserve Bank of New York Staff Reports*, no. 697 November 2014

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Abstract

We study market reactions to seasoned equity issuances that were announced by financial companies between 2002 and 2013. To assess the risk and valuation implications of these seasoned equity issuances, we conduct an event analysis using daily credit default swap (CDS) and stock market pricing data. The major findings of the paper are that CDS prices respond quickly to new, default-relevant information. Over the full sample period, cumulative abnormal CDS spreads drop in response to equity issuance announcements. The reactions are significantly stronger during the financial crisis. At that time, the federal government injected equity into financial institutions to ensure their viability. The market reacted to the equity issue announcements by assessing significantly lower costs for default protection via credit default swaps. The evidence indicates that single-name CDS based on financial firms' default probabilities are potentially useful for private investors and regulators.

Key words: financial institutions, stock issuance, credit default swaps, financial crisis

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CDS and Equity Market Reactions to Stock Issuances in the U.S. Financial Industry: Evidence from the 2002 – 2013 Period

1. Introduction

Financial theory and empirical analysis has shown that, because security values are determined by the distribution of the market value of a firm's assets, market prices of different securities issued by a firm are closely linked. Specifically, market prices of financial securities issued by firms should reflect the default risk of the firms. Less explored is the relation between market prices of various credit related securities for the same firm. For example, the link between the relatively new, but fast growing credit derivatives market and equity markets has only begun to be explored. One such market is the single name credit default swap (CDS) market. A single name CDS is a contract that provides protection against a default event involving a single issuer, or 'name.' Therefore, single name CDS contracts should reflect the pure issuer default risk of a firm irrespective of any issue specific risk. Thus, single name CDS contracts are of particular interest in that they may serve as a benchmark for measuring and pricing credit risk of a firm.

In this paper, we study CDS and equity market reactions to 129 seasoned equity issuances that were announced by 56 financial companies between 2002 and 2013. We split the sample period, analyzing market reactions before and during the financial crisis. We take a unique approach to measuring a change in the market's perception of the probability of firm default. Specifically, we collect and analyze CDS data to assess the risk and valuation implications of these seasoned equity issuances. We focus on financial institutions because they were heavily involved in, and some would argue at the center of, the financial crisis. Further, unlike any other industry, financial institutions were required to increase capital during and after the crisis. Thus,

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¹ For example, Norden and Weber (2009) analyze the relationship between credit default swap (CDS), bond, and stock markets between 2000 and 2002. Focusing on the intertemporal co-movement, they find that stock returns lead CDS and bond spread changes. Further, they find that the CDS market is more sensitive to the stock market than the bond market and the strength of the co-movement increases the lower the credit quality and the larger the bond issues.

this industry represents a unique opportunity to examine the relation between single name CDS securities and equity securities issued by the firms. To further highlight the uniqueness of this relation in the financial industry, we compare results for financial firms to those for 356 seasoned equity issuances that were announced by 172 nonfinancial companies between 2002 and 2013.

Consistent with previous empirical findings, we find that the stock market generally reacts negatively to announcements of common stock issuances by financial firms: the three-day cumulative abnormal stock returns (CARs) average -1.43% percent over the sample period. We also find that reactions are more negative during the financial crisis, albeit not significantly so: the average CARs range from -1.58% to -2.02% in these years. The major findings of the paper are that CDS prices respond quickly to new, default-relevant information. Over the full sample period, cumulative abnormal CDS spreads (CAS) drop significantly in response to equity issuance announcements by an average of 10.12 bps. Further, the reactions are significantly stronger during the financial crisis. The mean CASs range from -16.15 bps to -18.84 bps. These patterns are not found for nonfinancial firms that issue equity over the same period. Cross-sectional analysis suggests that riskiness of the institution is one of the key determinants of both abnormal stock returns and abnormal CDS spreads. Equity issuances by riskier institutions (rated speculative grade) are received more favorably in the CDS market and by stock investors. However, the relations fail to hold during the financial crisis. During the financial crisis, the federal government injected equity (through the TARP program) into financial institutions. The equity was made senior to common stock, but not to other capital providers. Thus, the market reacted to this by assessing lower costs for default protection via credit default swaps and insignificant changes in returns on equity issue announcements. Further, creditors obtained an additional buffer against losses in the event of default. As a result, the certification effect of an equity issue announcement

prior to the financial crisis was not necessarily the dominant effect during the crisis. During the financial crisis, issuances of equity by higher risk financial firms did not provide the same degree of default-relevant information as they did before the crisis. Thus, the evidence indicates that single named CDSs based on financial firms' default probabilities are potentially useful for private investors and regulators.

The remainder of the paper is organized as follows. Section 2 summarizes recent research on CDS contracts as measures of default risk. Section 3 describes the data and methodology used in the analysis. Section 4 discusses the results. Finally, Section 5 concludes the paper.

2. CDS contracts as a measure of default risk

In this paper, we take a unique approach to measuring a change in the market's perception of the probability of a debt default. Specifically, we collect and analyze single name credit default swap data. In the event of a default, the CDS writer compensates the CDS buyer for losses on a given face value of the underlying debt. The spread on the CDS contract is the price paid to the writer of the CDS for selling this insurance contract to the CDS buyer. When the market perceives that the probability of a debt default decreases (increases), the spread charged on the CDS decreases (increases). Thus, we use changes in the CDS spread to measure changes in the market's perception of the probability of a debt default.

Several papers have examined the ability of CDS spreads to measure risk. Hull et al. (2004) explore the extent to which credit rating announcements by Moody's are anticipated by participants in the credit default swap market. They conclude that either credit spread changes or credit spread levels provide helpful information in estimating the probability of negative credit rating changes. Jorion and Zhang (2007) examine the intra-industry information transfer effect of

Chapter 11 bankruptcies, Chapter 7 bankruptcies, and jump events, as captured in the credit default swaps (CDS) and stock markets. They find strong evidence of contagion effects (positive correlations across CDS spreads) for Chapter 11 bankruptcies and competition effects (negative correlations across CDS spreads) for Chapter 7 bankruptcies. They also find that a large unanticipated jump in a company's CDS spread leads to the strongest evidence of credit contagion across the industry. Ismailescu and Kazemi (2010) examine the effect of sovereign credit rating change announcements on the CDS spreads of the event countries, and their spillover effects on other emerging economies' CDS premiums. They find that positive events have a greater impact on CDS markets in the two-day period surrounding the event, and are more likely to spill over to other emerging countries. Alternatively, CDS markets anticipate negative events, and previous changes in CDS premiums can be used to estimate the probability of a negative credit event.

Most recently, papers have focused on CDS spread movements around the financial crises in the U.S. and Europe. For example, Fontana and Scheicher (2010) study the relative pricing of euro area sovereign CDS and the underlying government bonds for ten euro area countries from January 2006 to June 2010. They find that, since September 2008, CDS spreads have on average exceeded bond spreads, which may have been due to 'flight to liquidity' effects and limits to arbitrage. Arce et al. (2012) analyze the extent to which prices in the sovereign CDS and bond markets reflect the same information on credit risk in the context of the European Monetary Union. They find that there are persistent deviations between both spreads during the crisis. Levels of counterparty and global risk, funding costs, market liquidity, volume of debt purchases by the European Central Bank in the secondary market, and the banks' willingness to accept losses on their holdings of Greek bonds are found to be significant factors in determining which market leads price discovery. Andenmatten and Brill (2011) test whether the co-movement of sovereign

debt CDS premia increased significantly after the Greek debt crisis started in October 2009. Their results indicate that during the Greek debt crisis there were not only periods of interdependence, but also periods characterized by a significant increase in the co-movement of sovereign credit risk as measured in CDS premia.

Di Cesare and Guazzarotti (2010) examine the determinants of credit default swap spread changes for a large sample of U.S. nonfinancial companies from January 2002 and March 2009. They use variables that have previously been found to have an impact on CDS spreads. Since the onset of the crisis, they find that CDS spreads have become much more sensitive to the level of leverage, while volatility has lost its importance. They also show that since the beginning of the crisis CDS spread changes have been increasingly driven by a common factor, which cannot be explained by indicators of economic activity, uncertainty, and risk aversion. Alexopoulou et al. (2009) confirm the existence of a long-run relationship between CDS and corporate bond markets, and the tendency for CDS markets to lead corporate bond markets in terms of price discovery. They find that the outbreak of the financial crisis induced a substantial increase in risk aversion and a shift in the pricing of credit risk, with CDS markets becoming more sensitive to systematic risk, while bond markets priced in more information about liquidity and idiosyncratic risk. Moreover, the crisis also brought about a systematic disconnection between the two markets caused by the significant change in the lead-lag relationship, with CDS markets always leading bond markets. Finally, Galil et al. (2014) use a cross-sectional analysis to investigate the determinants of CDS spreads around the financial crisis. Fundamental variables such as historical stock returns, historical stock volatility, and leverage explain CDS spreads after controlling for ratings. During the crisis, while fundamental variables maintained their explanatory power, the explanatory power of ratings decreased to almost zero.

A few recent studies have focused on single name CDS spreads for financial firms. For example, Kallestrup et al. (2012) show that financial linkages across borders are priced in the CDS markets beyond what can be explained by exposure to common factors. They construct a measure which takes into account the relative size and riskiness of bank exposures to domestic government bonds and other domestic residents. This measure helps explain the dynamics of bank CDS premia after controlling for country specific and global risk factors. Demirgue-Kunt et al. (2010) investigate the impact of government indebtedness and deficits on bank stock prices and CDS spreads for a sample of international banks. They find that the change in bank CDS spreads in 2008 relative to 2007 reflects countries' deterioration of public deficits. Ballester et al. (2013) use bank CDS spreads to evaluate contagion among banks and banking sectors in different countries during the financial crisis. They find the Eurozone troubles barely affected U.S. banks. De Bruyckere et al. (2013) use correlations in CDS spreads at the bank and at the sovereign level to document significant contagion between bank and sovereign credit risk during the European sovereign debt crisis.² Possibly most relevant for our study, Flannery et al. (2010) evaluate the viability of credit default swaps (CDS) spreads as substitutes for credit ratings for large financial institutions that were prominently involved in the financial crisis. They show that CDS spreads incorporate new information about as quickly as equity prices and significantly more quickly than credit ratings.

² Eichengreen et al. (2009) identify common factors in the movement of banks' credit default swap spreads. They find that fortunes of international banks rise and fall together even in normal times along with short-term global economic prospects. But the importance of common factors rose to exceptional levels from the outbreak of the financial crisis to past the rescue of Bear Stearns, reflecting a diffuse sense that funding and credit risk was increasing. Following the failure of Lehman Brothers, the interdependencies briefly increased to a new high, before they fell back to the pre-Lehman elevated levels.

3. Data and methodology

3.a. Data

The initial sample examined in this paper includes all U.S. financial institutions that issued equity between 2002 and 2013.³ This sample is collected from S&P's Capital IQ database. We then limit the sample to those financial firms having CDS pricing data reported by Markit.⁴ The final sample we examine includes data on 129 equity issuances by 56 financial institutions.⁵ Table 1 lists the year-by-year distribution of the announcements and shows a fairly balanced flow in financial institutions' equity issuance activities between the first and the second half of the sample period. Figures 1 and 2 illustrate the general movements of the CDS spreads and stock market values during the sample period. Figure 1 plots a daily series of the mean CDS spreads of our sample firms and Figure 2 plots the daily evolution of the combined equity market capitalization⁶ of the sample firms. Note the significant increase in CDS spreads and decrease in financial firm values during the financial crisis.

We examine stock returns and CDS spreads for the whole sample period, as well as two sub-periods, termed in the paper, the "pre-crisis" and the "crisis" periods. We experiment with two cutoff dates to delineate the pre-crisis and crisis periods using two prominent events in 2008: 3/24/2008 (the collapse of Bear Stearns) and 10/14/2008 (the TARP Capital Purchase Program announcement). Signs of significant problems in the U.S. economy first arose in late 2006 and the first half of 2007 when home prices plummeted and defaults by subprime mortgage borrowers began to affect the mortgage lending industry as a whole, as well as other parts of the economy

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³ The sample firms include all U.S. financial companies except for REITs, due to their distinct organizational structure.

⁴ Markit is a private company headquartered in London. The firm provides CDS end-of-day quotes for approximately 450 of the most liquid CDS contracts, including G20 sovereigns and large financial corporations.

⁵ Issuance transactions with public offering or shelf registration features are excluded as are non-secondary issue related transactions.

⁶ Equity market capitalization is calculated as the daily stock price times the number of shares of common stock outstanding (from CRSP data tapes).

noticeably. As mortgage borrowers defaulted on their mortgages, financial institutions that held these mortgages and mortgage backed securities started announcing huge losses on them. A prime example of the losses incurred is that of Bear Stearns. In the summer of 2007, two Bear Stearns hedge funds suffered heavy losses on investments in the subprime mortgage market. The two funds filed for bankruptcy in the fall of 2007. Bear Stearns' market value was hurt badly from these losses. The losses became so great that in March 2008 J.P. Morgan Chase and the Federal Reserve stepped in to rescue the then fifth largest investment bank in the United States before it failed or was sold piecemeal to various financial institutions. On March 24, 2008, J.P. Morgan and Bear Stearns announced a merger agreement. J.P. Morgan Chase purchased Bear Stearns for \$236 million, or \$2 per share. The stock was selling for \$30 per share three days prior to the purchase and \$170 less than a year earlier. The collapse of Bear Stearns and its sale to J.P. Morgan marked the beginning of the mortgage crisis for major financial institutions.

On October 14, 2008, then-Secretary of the Treasury Paulson announced a revision in TARP implementation in which the Treasury directly injected up to \$250 billion⁸ of TARP funds (through the Capital Purchase Program (CPP)) into the U.S. banking system through the purchase of senior preferred stock and warrants in qualifying financial institutions (QFIs).⁹ The CPP was intended to inject equity into financial institutions that were suffering from temporary liquidity and other financial problems due to the financial crisis, but were otherwise in decent shape. By the

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⁷ Using principal components analysis to identify common factors in the movement of banks' credit default swap spreads, Eichengreen et al. (2012) find that the importance of common factors rose steadily to exceptional levels from the outbreak of the Subprime Crisis to past the rescue of Bear Stearns, reflecting a diffuse sense that funding and credit risk was increasing.

⁸ This amount was eventually lowered to \$204.9 billion.

⁹ Qualifying financial institutions (QFIs) include bank holding companies, financial holding companies, insured depository institutions, and savings and loan holding companies that are established and operating in the United States and that are not controlled by a foreign bank or company.

time it closed on December 31, 2009, 707 applications were approved and funded by the Treasury through the CPP.¹⁰

3.b. Seasoned equity issues of nonfinancial firms

In order to examine whether any observed relation between CDS spreads and equity market reactions is unique to financial firms, we also collect information on equity issuances by nonfinancial firms between 2002 and 2013. The initial sample of nonfinancial firm equity issuances is also collected from S&P's Capital IQ database. We then limit the sample to those nonfinancial firms having CDS pricing data reported by Markit. The final sample of nonfinancial firms we examine includes data on 356 equity issuances by 172 firms. Table 1 lists the year-by-year distribution of the announcements.

3.c. Calculation of cumulative abnormal stock returns (CARs)

We use standard event study techniques to calculate unexpected changes in stock prices in response to equity issue announcements by the sample firms. Stock return data are collected from the Center for Research in Security Prices (CRSP) data tapes. Stock return data are identified for 54 of the 56 financial firms (123 of the 129 equity issue announcements) and 170 of the 172 nonfinancial firms (337 of the 356 events). We calculate normal stock return performance using a

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¹⁰ Cornett et al. (2013) look at how the pre-crisis health of banks is related to the probability of receiving and repaying TARP capital. They find that financial performance characteristics that are related to the probability of receiving TARP funds differ for the healthiest ("over-achiever") versus the least healthy ("under-achiever") banks. Ng et al. (2011) find that publicly traded TARP banks experienced significantly lower equity returns relative to non-TARP banks during TARP's initiation period. They also find that equity markets adjusted the values of the TARP banks upward in the quarters following TARP injections. Veronesi and Zingales (2010) also provide an analysis of the valuation effects of TARP, but for just the initial financial institutions that received TARP infusions on October 14, 2008. They find that, while the primary effect of TARP was to benefit bondholders, the program created little value for shareholders. While they find that valuation benefits for banks exceeded the costs imposed on taxpayers, they argue that other potential rescue strategies would have yielded larger net benefits.

125 trading day window, ending five trading days before the announcement date. A Fama-French three factor model is used to predict the expected excess return:¹¹

$$Return_{excess i,t} = \alpha + \beta_1 (Market - R_f)_t + \beta_2 (SMB)_t + \beta_3 (HML)_t$$
 (1)

where R_f is the risk-free return rate, SMB stands for "small (market capitalization) minus big," and HML stands for "high (book-to-market ratio) minus low."

The abnormal return, $AR_{i,t}$, is then calculated as the actual, $Return_{excess}$, less the predicted, $Return_{excess}$. Summation over the three-day window (the trading day before the announcement, the announcement day, and the following trading day) gives the three-day announcement period cumulative abnormal stock returns (CARs):

$$CAR (-1,1)_i = AR_{i,1} + AR_{i,0} + AR_{i,1}$$
(2)

3.d. Calculation of cumulative abnormal CDS spread (CAS)

We also use standard event study techniques to calculate unexpected changes in CDS spreads in response to the announcement events. First, we calculate daily CDS abnormal spreads (AS) for each contract as:

$$AS_{i,t} = Spread_{i,t} - Rating Matched Index Spread_t$$
 (3)

where $Spread_{i,t}$ is the end-of-day spread quote of contract i on event day t, and is the rating-matched (investment-grade or speculative-grade) CDS index spread on event day t. We then calculate the three-day cumulative CDS abnormal spread (CAS) for each contract as:

$$CAS_{i,t} = AS_{i,1} - AS_{i,-1}$$
 (4)

where $AS_{i,1}$ is the CDS abnormal spread at the end of the next trading day after the equity issuance announcement by firm i, and $AS_{i,-1}$ is the CDS abnormal spread one day before the announcement.

¹¹ We also examined a specification where excess return is a function of the three Fama-French factors and an additional momentum factor. Results are similar to those reported in tables 3, 5, and 6. They are available from the authors on request.

Unexpected movements in the CDS spread measure the perceived impact on the cost of purchasing default insurance on the issuing institution. Unexpected changes in stock returns measure the valuation impact on the residual claimholders (shareholders) due to the firm's equity-raising action. Throughout the analysis, we use the three-day cumulative abnormal CDS spread (CAS) and cumulative abnormal stock return (CAR) as the key measures of the market reactions.

4. Results and discussions

4.a. Average unexpected reactions in CDS and equity markets

Table 2 summarizes CDS market reactions to 129 seasoned equity issuance announcements made by 56 financial firms and 356 announcements made by 172 nonfinancial firms between 2002 and 2013. We report the average CAS for the whole sample period and two sub periods, the "precrisis" period and "crisis" period. Panel A presents results based on the split using the collapse of Bear Stearns (3/24/2008) to delineate the beginning of the crisis period and panel B reports numbers based on the split using the TARP CPP announcement (10/14/2008) to delineate the beginning of the crisis period.

Over the entire sample period, 2002 through 2013, the market-adjusted CDS spread for financial firms (i.e., the default insurance premium) drops by an average 10.12 bps (see row "Total") over a three-day window in response to equity issuances announcements. This effect is statistically significant at the 1% level, suggesting that equity financing activities by financial firms are generally favorably received by CDS investors. That is, the price of default insurance, in the form of CDS, decreases at the announcement of stock issuances by financial firms.

Looking at results for the two sub-periods, we see that significantly stronger reactions are concentrated in the crisis period. Using the acquisition of Bear Stearns by J.P. Morgan to delineate

the start of the crisis, the three-day average CAS is -16.15 bps in the crisis period (panel A). Using the announcement of the TARP CPP program to delineate the start of the crisis, the two-day average CAS is -18.84 bps in the crisis period (panel B). In contrast, pre-crisis period estimated reactions appear to be insignificant both statistically and economically. Using the acquisition of Bear Stearns by J.P. Morgan to delineate the start of the crisis, the average CAS is -2.75 bps in the pre-crisis period (panel A). Using the announcement of the TARP CPP program to delineate the start of the crisis, the average CAS is -2.54 bps in the pre-crisis period (panel B). In both cases, the crisis period CAS estimates are highly significant and larger in magnitude than the pre-crisis period (the t-statistics for the differences in CAS between the crisis and pre-crisis period are 1.93 (significant at 5.7%) and 2.07 (significant at 4.2%) in panels A and B, respectively.

The estimated reactions appear economically significant when benchmarked with the normal market level in the 2002-2008 period. The raw CDS spread of the sample firms averages around 70 bps in this period. With this backdrop, the market-adjusted reaction in the crisis period accounts for more than 20% (e.g., -16.15 bps/70 bps) of the general market level in the pre-crisis period, regardless of the cutoff point we use for defining the crisis period. Thus, equity financing activities by financial firms are received most favorably by CDS investors in the crisis period. Financial firms that issued stock during the financial crisis experienced a steep drop in the price of default insurance.

Finally, comparing the results for financial firms to those of nonfinancial firms, we find differences. Using the acquisition of Bear Stearns by J.P. Morgan to delineate the start of the crisis period, the two-day average CAS for nonfinancial firms is -42.71 bps in the crisis period (panel A). Using the announcement of the TARP CPP program to delineate the start of the crisis period, the two-day average CAS is -45.70 bps in the crisis period (panel B). In contrast to the results for

financial firms, neither of these is statistically significant. However, estimated reactions in the precrisis period appear to be significant. The two-day average CAS is -3.26 bps (significant at 10%) in panel A and -4.80 bps (significant at 5%) in panel B. Finally and in contrast to the results for financial firms, for nonfinancial firms the crisis period CAS estimates are not significantly different from the pre-crisis period CAS estimates (the t-statistics for the differences in CAS between the crisis and pre-crisis period are 1.32 and 1.24 in panels A and B, respectively).

Table 3 presents evidence on the stock market reactions to stock issue announcements by financial and nonfinancial firms. Adjusted for systematic pricing factors, the stock market overall reacted negatively to seasoned stock offerings by financial firms over the period 2002 through 2013. Three-day cumulative abnormal stock returns (CARs) average -1.43%, significant at the 10% level. This magnitude is comparable with earlier evidence, e.g., Cornett et al. (1998) find an average two-day CAR of -1.62% for seasoned equity issuances by financial institutions.

Looking at the pre-crisis and the crisis periods separately, we find insignificantly larger (negative) reactions in the crisis period. Using the acquisition of Bear Stearns by J.P. Morgan to delineate the start of the crisis period, the mean CAR, -0.69%, is insignificant in the pre-crisis period. However, common stock issuances induce an average CAR of -2.02% (significant at 10%) in the crisis period. Using the announcement of the TARP CPP program to delineate the start of the crisis period, the pre-crisis mean CAR is -1.30%, while the crisis period mean CAR is -1.58% (both are insignificant). Further, the t-statistics for the differences in CAR between the crisis and pre-crisis period are 0.85 in panel A and 0.18 in panel B and are both are insignificant.

While the general conclusion from these results are that financial firm stock return reactions to seasoned equity issuances are not significantly different before versus during the

¹² We were only able to obtain stock market data for 123 of the 129 equity issue announcements by financial firms and 337 of the 356 announcements by nonfinancial firms. Thus, the number of events drops slightly.

financial crisis, when benchmarked with normal market conditions, the estimated market reactions appear rather significant. For example, during the crisis period, the average CAR (-2.02%, using the acquisition of Bear Stearns by J.P. Morgan as the start of the crisis period) is almost 40 times larger than the average raw return (0.051%) of the issuing institutions' stocks in the pre-crisis period. Thus, during the crisis, stock investors held markedly more negative views toward financial firms' equity financing activities.

As mentioned above, the TARP CPP was intended to inject equity into financial institutions that were suffering from temporary liquidity and other financial problems due to the financial crisis, but were otherwise in decent shape. These government capital injections need to be analyzed not only in terms of certification effects, but also in terms of how they impact investor perceptions of future regulatory action. While a capital injection certainly implies confidence in a financial firm's viability, it is different from a pure certification exercise like the Supervisory Capital Assessment Program (SCAP, more often known as bank stress tests) or the Depression-era bank holiday in that it alters the risk to taxpayers from a subsequent failure and may therefore influence regulators' behavior. Before the crisis, the U.S. government typically stood in a 'third-to-default' position (through the FDIC guarantee on deposits) behind shareholders, other capital providers, and unsecured creditors. After a TARP CPP injection, it moved to a "second-to-default" position behind shareholders.

It is worth pausing to consider what counts as "default" in this context. Under the terms of the FDICIA of 1991, banks are not subject to regular bankruptcy proceedings with normal creditor rights (the Wall Street Reform and Consumer Protection Act of 2010 has effectively extended this treatment to a range of financial holding companies). The event of default is the point at which the FDIC seizes the operating bank subsidiary; the common equity of the bank is typically wiped out

entirely and unsecured creditors are provided with a discounted payoff decided by the FDIC. Well ahead of seizure, however, the bank's supervisors might also force the bank to raise capital even if this issuance is not in the best interest of shareholders. The analytical objective for an analyst or investor is therefore to judge when regulatory intervention will occur and what losses they are subjected to post-intervention.

The immediate reaction to the announcement of capital injections through the TARP CPP focused on the certification effect; the program targeted "healthy" banks, the terms of the preferred stock seemed attractive, and the objective appeared to be to boost confidence in the financial system. The fact that government capital was now subordinate to unsecured creditors also served as a form of "seizure insurance" for anxious investors who had just experienced the shock of government-imposed losses on the collapse of Washington Mutual. But what happened next is instructive:

1. Bank capital has historically been measured mainly on the basis of Tier 1 capital ratios ((common equity + qualifying preferred stock) to risk-weighted assets). Differences in Tier 1 capital ratios across banks provided important information about banks' relative profitability, risk tolerance, and access to capital. Analysts and money managers used Tier 1 capital ratios almost exclusively to identify differences among banks. Preferred stock issued to the government through the TARP CPP qualified as Tier 1 capital. The injection of government capital, which amounted to approximately 3% of risk-weighted assets for virtually all publicly traded banks, made Tier 1 comparisons less meaningful. Not surprisingly, the focus of analysts and money managers shifted to the Tangible Common

¹³ One can imagine that investors, in an uncorrelated but otherwise identical set of banks, might decide to limit their losses on any one bank in order to preserve their upside gains on the portfolio as a whole.

¹⁴ Indeed, in a couple of cases that banks canceled planned equity raises once CPP was announced – a great real world example of Gresham's law at work.

Equity (TCE) ratio (common equity to tangible assets) since this measure did not include TARP CPP preferred stock: comparisons made on the basis of TCE remained undiluted. By early 2009 market participants were focusing almost exclusively on TCE and Tier 1 Common (common equity to risk-weighted assets) ratios. In the process, TARP CPP capital, as well as all other non-common equity Tier 1 capital, issued by banks had become essentially moot as measure used to evaluate bank health.

2. By design preferred stock issued through TARP CPP, was made senior to common equity in order to protect taxpayers from loss. So, while creditors obtained an additional buffer against losses in the event of default, shareholders did not – they remained in a first loss position. Shareholders might reasonably suspect that the government would be more prone to intervene if a TARP CPP bank got into trouble given that the government was now in a second loss position and had lost the buffer provided by other forms of regulatory capital and unsecured debt. Indeed, the higher loss outcomes could more than offset the certification effect. However, the continued decline in bank stocks in the period between the announcement of TARP CPP and the release of the first bank stress test (SCAP) results (see Figure 2) suggests equity investors did not in fact see TARP CPP capital injections as reducing their risk. Moreover, the unpopularity of the program, the addition of "strings" by Congress and the focus on getting the money back all made TARP CPP increasingly seem like expensive capital to the could not absorb losses or be leveraged like common equity.

¹⁵ As public outrage swelled over the rapidly growing cost of 'bailing out' financial institutions, the Obama administration and lawmakers attached more and more restrictions on banks that received TARP funds. For example, with the acceptance of TARP funds, banks were told to put off evictions and modify mortgages for distressed homeowners, let shareholders vote on executive pay packages, slash dividends, and withdraw job offers to foreign citizens. Some bankers stated that conditions of the TARP program had become so onerous that they wanted to return the bailout money as soon as regulators set up a process to accept the repayments. For example, just three months after receiving TARP funds, Signature Bank of New York announced that because of new executive pay restrictions assessed as a part of the acceptance of TARP funds, it notified the Treasury that it intended to return the \$120 million it had received.

Given these two considerations, it is not surprising that equity market reactions to new stock issuances by financial firm are larger in the crisis period relative to the pre-crisis period. However, while equity market reactions are larger in the crisis period, they are not significantly so. In part, this may be because the certification effect was more robust; differentiating amongst banks in terms of capital needs provided more information and the capital raised was fully loss absorbing. ¹⁶

Finally from table 3, comparing the results for financial firms to those of nonfinancial firms, we again find differences. Using the acquisition of Bear Stearns by J.P. Morgan to delineate the start of the crisis period, the two-day average CAR for nonfinancial firms is -1.10% in the precrisis period (panel A). Using the announcement of the TARP CPP program to delineate the start of the crisis period, the two-day average CAR is -1.77% in the pre-crisis period (panel B). In contrast to the results for financial firms, both of these are significant at 5% and 1%, respectively. However, neither is significantly different from the CARs for financial firms (-0.69% and -1.30%, respectively). Similarly, in the crisis period the two-day average CAR is -3.89% (significant at 1%) in panel A and -3.70% (significant at 1%) in panel B. In contrast to the results for financial firms, the crisis period CARs are significantly larger than the pre-crisis period CARs (the t-statistics for the differences in CARs between the crisis and pre-crisis period are 5.00 and 2.96 in panels A and B, respectively). Finally, the crisis period CARs for nonfinancial firms are larger than the crisis period CARs for financial firms (the t-statistics for the differences in CARs are 1.55 and 1.73 in panels A and B, respectively).

¹⁶ It is also instructive to contrast the reaction to TARP CPP with the subsequent release of stress test results and associated capital raising. Even though many banks were forced to undertake dilutive capital raises to satisfy the SCAP requirements and/or exit TARP, share prices generally rallied. This may also be because the certification effect was more robust. The government also played an entirely supervisory role and was not muddying their certification role by taking direct risk. Also, since the government was now at risk once a bank exhausted its TCE, it was logical for investors to focus on TCE levels as the best indicator of effective default risk - this adds to the arguments already listed in item 1.

Remembering that the TARP CPP program injected capital into financial institutions that were suffering from temporary problems, these results make sense. TARP CPP injections certified the long-term viability of the financial firms that received funds. However, TARP CPP funds were not provided to nonfinancial firms. As a result, these firms were left more susceptible to losses and even failure resulting from the financial crisis. Thus, it is not surprising that the crisis period CARs for financial firms (that were receiving capital injections from the federal government) are significantly larger (less negative) than the crisis period CARs for financial firms (that did not receive such capital injections).

Combining the results from tables 2 and 3, we find that price of default insurance (measured as abnormal changes in the CDS spreads) of financial firms that issue equity during the financial crisis is lowered significantly. As a result, as these financial firms issue equity after the crisis, their abnormal stock returns (CARs) are not significantly more negative than the pre-crisis period. In contrast, default risk for nonfinancial firms did not decrease significantly during the crisis. As a result, equity issuances by these firms resulted in significantly larger negative CARs. A reason for the differences between financial firms and nonfinancial firms may be due to the government sponsored TARP CPP program that injected equity into financial firms during the crisis. Equity injections through TARP CPP left the financial firms less exposed to failure during the crisis. The market reacted to this by assessing lower costs for default protection via credit default swaps and insignificant changes in returns on equity issue announcements. Nonfinancial firms had no such injections and were thus left more susceptible to the financial crisis. The market reacted to the crisis by assessing no abnormal decreases in costs for default protection via CDSs and larger negative returns on equity issue announcements.

4.b. Market reactions and firm/issuance characteristics

Having examined the CASs and CARs in isolation, we next relate the cross-sectional variation of market reactions to firm and issuance characteristics. Specifically, we regress the three-day CAS and CAR estimates on a risk measure-a dummy for being speculative-grade rated, 17 the issue size as a percent of the market capital of the firm, 18 and a set of firm characteristics 19 as controls (we include log of assets, leverage, stock volatility, and ROA). The high yield dummy is equal to 1 if the stock issuing financial firm is rated as speculative, and 0 otherwise. Ratings and issue size are obtained from S&P's Capital IQ. Market capital is measured as the market price of the firm's stock one month prior to the stock issue announcement (obtained from CRSP) times the number of shares of common stock outstanding. Assets is the book value of the firm's assets at quarter end before the stock issue announcement. Leverage is total debt divided by total assets at quarter end before the stock issue announcement. ROA is annualized net income divided by total assets at quarter end before the stock issue announcement. Volatility is the standard deviation in the firm's stock prices in the quarter prior to the stock issue announcement. All balance sheet data are obtained from Call Reports and/or SEC 10-Q filings. In the regressions, we cluster standard errors by bank.

Table 4 presents descriptive statistics on the regressions variables. The mean value of the sample firms is \$333.78 billion, ranging from \$2.42 billion to \$2.02 trillion. Thus, the sample includes both smaller and large financial firms. The mean issue size represents 19.76% of the pre-issue market capital, and ranges from 0.00% to 567.69%. The sample firms have a mean leverage of 86.38% before the equity issuance, ranging from 37.79% to 97.80%. The mean ROA for the sample firms is -0.55% and ranges from -76.03% to 48.17%. Thus, the sample includes firms that performed very well and very poorly over the sample period, a time frame that includes a period of

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¹⁷ There are 13 institutions, 11% of the sample firms, which are rated speculative grade.

¹⁸ We also examined the log of this variable. Results are generally of the same sign, but less significant.

¹⁹ The regression sample decreases due to the requirement of non-missing data for both LHS and RHS variables.

record profits for the financial institutions industry as well as the financial crisis which hurt financial firm profitability tremendously.

Table 5 presents the regressions results. Columns 1 through 3 present results using the three-day cumulative CDS abnormal spread (CAS) as the dependent variable, while columns 4 through 6 present results using the three-day cumulative abnormal stock returns (CAR) as the dependent variable. Regressions 1 and 4 include only the "high yield dummy" as an independent variable. Regressions 2 and 5 include the "high yield dummy" and the firm control variables. Finally, regressions 3 and 6 include the "high yield dummy," the relative issue size, and the firm control variables as independent variables.

Examining the regression results, it is evident that issuer risk is a significant determinant of the CDS reactions to stock issuances by financial firms. The "high yield dummy" is negative with persistent significance in all CAS regressions (models 1 through 3), e.g., the coefficient is -36.451, significant a 1%, in regression 1; the coefficient is -25.556, significant a 5%, in regression 2; and is -25.555, significant a 5%, in regression 3. The results suggest that equity issuances by riskier firms are associated with abnormal decreases in CDS spreads (i.e., the price of default insurance decreases), ranging from 25.555-36.451 bps. Similarly, regressions 4 through 6 suggest that stock market reactions to equity issuances by financial firms are more positive, albeit not consistently significant, for high risk firms, e.g., the coefficient is 2.853, significant at 10%, in regression 4 and is 3.396, insignificant, in both regressions 5 and 6. The average CAR of -1.43% (see table 3) is higher, by 2.853%-3.396%, if the firm is rated speculative grade.

Examining regressions 3 and 6, abnormal changes in CDS spreads in response to stock issuances by financial firms do not appear to be sensitive to the size of issuance; issue size as a percent of market capital is insignificant (coefficient = -0.007). Further, stock market reactions are

not significantly related to the issue size (coefficient = 0.001). These results are consistent with the notion that larger issue size (relative to firm value) is insignificant news for CDS (lower CAS) and stock investors. Despite the lack of significance of the issuance size estimate, signs are consistent with the notion that larger issuance size (relative to market capital) is favorable news for both CDS (lower CAS) and stock (higher CAR) investors. Finally, comparing regressions 2 and 3 vs regression 1 (and regressions 5 and 6 vs regression 4) we see that the inclusion of various controls for firm characteristics does not appear to alter the key findings. This is not surprising given that much of this firm specific financial information is incorporated in the credit rating measure. ²¹

Table 6 reports regression results split according to time period. Panel A of table 6 reports results using the acquisition of Bear Stearns by J.P. Morgan to delineate the start of the crisis period, while panel B reports results using the announcement of TARP CPP to delineate the start of the crisis period.²² The regression results appear to differ significantly across the pre-crisis and the crisis sub-periods. The results in the pre-crisis period for both panels A and B are similar to, yet more significant than, those in table 5. The "high yield dummy" is negative and significant at the 5% level in all CAS regressions (models 1 through 3), again suggesting that equity issuances by riskier firms are associated with decreases in CDS spreads (i.e., the price of default insurance decreases). Yet now we see that the increase in CDS spreads ranges from 20.951-34.680 bps. Further, regressions 4 through 6 suggest that stock market reactions to equity issuances by high risk financial firms are more positive, and now consistently significant, e.g., in panel A the coefficient is 7.246 in regression 4; the coefficient is 20.519 in regression 5; and is 20.245 in

²⁰ In unreported results, we also try to detect non-linear relationship using discrete dummy-based measures for different quartiles of the issuance size measure. We do not find any significant relationship, which we attribute to the limited size of our sample.

²¹ The signs of the controls are generally consistent with existing evidence, but there is no persistent significance in any control among a number of specifications that we tested.

²² In both panels, 3 of the high yield firms announce equity issuances in the pre-crisis period, while 10 announce equity issuances in the crisis period.

regression 6, all significant at 10% or better. Thus, issuer risk is a significant determinant of the CDS and stock market reactions to stock issuances by financial firms before the financial crisis.²³ Equity issuances by riskier firms are associated with significant abnormal decreases in CDS spreads and significant abnormal increases in stock market values.

Examining regressions 3 and 6, abnormal changes in CDS spreads in response to stock issuances by financial firms are now positively related to the size of issuance; issue size as a percent of market capital is significant (coefficient = 0.027 in Panel A and 0.020 in Panel B, both significant at 10%)). Further, stock market reactions are now negatively related to the issue size (coefficient = -0.009 in Panel A, insignificant, and -0.020 in Panel B, significant at 10%)). Thus, prior to the financial crisis the larger the issue size (relative to firm value), the larger the abnormal increase in CDS spreads and decrease in stock values. This is consistent with the notion that larger issuance size (relative to market capital) is negative news for both CDS and stock investors.

In both panels A and B of table 6 we see that results in the pre-crisis period virtually disappear and even reverse in the crisis period. For the crisis period regressions, the "high yield dummy" is insignificant in two of the three CAS regressions (only regression 1 reports a significant coefficient, -38.549, significant at 5%, in Panel A and -36.522, significant at 10%, in Panel B). Thus, the risk of the issuing firm has no or a much smaller effect on the price of default insurance during the financial crisis. Further, the coefficients on the "high yield dummy" are insignificant in the CAR regressions panel A models 4 through 6 and panel B model 4, are positive and significant (at 10%) in panel B model 4 (coefficient = 1.464), and are negative and significant (at 5%) in panel B model 6 (coefficient = -3.490). These results suggest that during the financial crisis higher risk firms see no consistent abnormal change in stock values at the

²³ We also examine regressions in which we interact the key variables with period dummies. Results are similar to those reported in tables 5 and 6. They are available from the authors on request.

announcement of an equity issue. As discussed above, preferred stock issued through the TARP CPP program was made senior common stock, but not to other capital providers. Thus, creditors obtained an additional buffer against losses in the event of default. As a result, the certification effect of an equity issue announcement prior to the financial crisis is not necessarily the dominant effect during the crisis. During the financial crisis, issuances of equity by higher risk financial firms did not provide the same signal of expected poor performance as they did before the crisis.

As a final test, we use S&P's Expected Default Frequency (EDF) measure to assess how equity issuances affect the perceived default likelihood during different time horizons. For 82 announcements, we are able to collect EDF data for horizons of 1, 2 and 3 years. ²⁴ Seven of these 82 announcements involve high yield firms. EDF is the percent change in the S&P Expected Default Frequency (EDF) for the firm for 1-year, 2-year, and 3-year horizons beginning from the quarter before the stock issue announcement. Descriptive statistics for percent change in the EDF are reported in Table 4. The mean change runs from 2.30% to 2.62% for the 1-year, 2-year, and 3-year EDF, respectively, and the minimum and maximum changes run from 0.01% to 24.63% for the three measures of change in default risk.

Regression results are reported in table 7. The regression specification in table 7 is identical to that in table 5, except that the dependent variable is the percentage change in the EDF for a 3-day period (t= -1 to +1) around the announcement date. Consistent with the CDS market reactions documented earlier in table 5, table 7 shows that the EDF measure decreases significantly more among riskier institutions, as reflected by the "high yield dummy" coefficient which is negative and significant (coefficient = -23.402 in regression 1, -22.581 in regression 2, and -21.296 in regression 3 (all are significant at 10%)). Further, changes in the EDF measures in

²⁴ S&P does not provide daily EDF data until June 2006. Given the focus on daily change in EDF around the equity issue announcement day, we are not able to include announcements before June 2006. This prohibits us from being able to isolate relations before versus during the financial crisis.

response to stock issuances by financial firms appear to be sensitive to the size of issuance; issue size as a percent of market capital is significant and negative in all three regressions. That is, larger issuances are associated with larger decreases in the EDF risk measure.

4.c. Sampling Alternatives

The results reported throughout the paper are based on a sample 129 equity issuances by financial firms. We expand the sample to include (1) equity issue announcements involving both common stocks and other security types such as preferred stock, options, warrants, etc. and (2) REITs-related financial firms. In both cases, all findings and interpretations remain unchanged and in many cases the estimates appear even stronger statistically.

5. Concluding remarks

In this paper, we study CDS and equity market reactions to seasoned equity issuances that were announced by financial companies between 2002 and 2013. We split the sample period, analyzing market reactions before and during the financial crisis. We measure the change in the market's perception of the probability of firm default by analyzing CDS data. We find that CDS prices respond quickly to new, default-relevant information. Over the full sample period, cumulative abnormal CDS spreads (CAS) drop significantly in response to equity issuance announcements. Further, the reactions are significantly stronger during the financial crisis. Cross-sectional analysis suggests equity issuances by riskier institutions (rated speculative grade) are received more favorably in the CDS market and by stock investors. However, the relations fail to hold during the financial crisis. During the financial crisis, the federal government injected equity into financial institutions. The equity was made senior to common stock, but not to other capital providers. Thus, the market reacted to this by assessing lower costs for default protection via

credit default swaps and insignificant changes in returns on equity issue announcements. Further, creditors obtained an additional buffer against losses in the event of default. As a result, the certification effect of an equity issue announcement prior to the financial crisis was not necessarily the dominant effect during the crisis. During the financial crisis, issuances of equity by higher risk financial firms did not provide the same degree of default-relevant information as they did before the crisis. Thus, the evidence indicates that single named CDSs based on financial firms' default probabilities are potentially useful for private investors and regulators.

Table 1

Number of Equity Issuances by Year

This table reports the year-by-year distribution of seasoned equity issue announcements by financial and nonfinancial firms. The initial sample is collected from S&P's Capital IQ database. We then limit the sample to those firms having CDS pricing data reported by Markit. The final sample includes data on 485 equity issuances by 56 financial institutions and 172 nonfinancial firms. Of the 485 announcements, 129 involve equity issuances by financial firms, with the rest including issuances by nonfinancial firms.

Year	All Issuances	Financial Firms	Nonfinancial Firms
2002	27	5	22
2003	38	10	28
2004	28	4	24
2005	31	12	19
2006	38	17	21
2007	27	6	21
2008	58	22	36
2009	73	21	52
2010	52	12	40
2011	38	8	30
2012	38	7	31
2013	37	5	32
Total	485	129	356

Table 2 Cumulative abnormal CDS spreads (CAS) in response to seasoned equity issuances

This table reports CDS market reactions to 485 seasoned equity issuance announcements made by 56 financial and 172 nonfinancial firms between 2002 and 2013. We calculate daily CDS Abnormal Spreads (AS) for each contract as:

$$AS_{i,t} = Spread_{i,t} - Rating Matched Index Spread_t$$

where *Spread*_{i,t} is the end-of-day spread quote of contract i on event day t, and is the rating-matched (investment-grade or speculative-grade) CDS index spread on event day t. We then calculate the three-day cumulative CDS Abnormal Spread (CAS) for each contract as:

$$CAS_{i,t} = AS_{i,1} - AS_{i,-1}$$

where $AS_{i,l}$ is the CDS abnormal spread at the end of the next trading day after the equity issuance announcement by institution i, and $AS_{i,-l}$ is the CDS abnormal spread one day before the announcement. Panel A presents results based on the split using the collapse of Bear Stearns (3/24/2008) to delineate the beginning of the crisis period, and panel B reports numbers based on the split using the TARP CPP announcement (10/14/2008) to delineate the beginning of the crisis period.

Panel A - Crisis period starts with Bear Stearns acquisition by J.P. Morgan announcement

Pre-crisis period: 1/1/2002 - 03/24/2008 Crisis period: 03/25/2008 - 01/20/2011

	Financial Firms			Nonfinancial firms					
Years	N	CAS (bps)	Mean Spread (bps)	Std Dev Spread (bps)	Years	N	CAS (bps)	Mean Spread (bps)	Std Dev Spread (bps)
Pre-crisis Crisis	58 71	-2.75 -16.15**	1.35 3.11	2.11 2.53	Pre-crisis Crisis	142 214	-3.26* -42.71	2.14 3.73	2.85 5.00
Total	129	-10.12***	2.32	2.50	Total	356	-26.97	3.10	4.34

Panel B – Crisis period starts with TARP CPP announcement

Pre-TARP period: 1/1/2002 - 10/14/2008 Post-TARP period: 10/14/2008 – 01/20/2011

	Financial Firms						Nonfinancial firms			
Years	N	CAS (bps)	Mean Spread (bps)	Std Dev Spread (bps)	Years	N	CAS (bps)	Mean Spread (bps)	Std Dev Spread (bps)	
Pre-crisis Crisis	69 60	-2.54 -18.84**	1.68 3.06	2.26 2.57	Pre-crisis Crisis	163 193	-4.80** -45.70	2.47 3.62	3.40 4.95	
Total	129	-10.12***	2.32	2.50	Total	356	-26.97	3.10	4.34	

^{***} p<0.01, ** p<0.05, * p<0.10

Table 3 Cumulative abnormal stock returns (CARs) in response to equity issuances

This table reports equity market reactions to 337 equity issuance announcements made by 56 financial and 172 nonfinancial firms between 2002 and 2013. We use standard event study techniques to calculate unexpected changes stock prices in response to the announcement events. Specifically, we calculate normal stock return performance using a 125 trading day window, ending five trading days before the announcement date. A Fama-French three factor model is used to predict the expected excess return:

$$Return_{excess i,t} = \alpha + \beta_1 (Market - R_f)_t + \beta_2 (SMB)_t + \beta_3 (HML)_t$$

where R_f is the risk-free return rate, SMB stands for "small (market capitalization) minus big," and HML stands for "high (book-to-market ratio) minus low." The abnormal return, $AR_{i,t}$, is then calculated as the actual, $Return_{excess}$, less the predicted $Return_{excess}$. Summation over the three-day window (the trading day before the announcement, the announcement day, and the following trading day) gives the three-day announcement period cumulative abnormal stock returns (CARs),

$$CAR (-1,1)_i = AR_{i,-1} + AR_{i,0} + AR_{i,1}$$

Panel A presents results based on the split using the collapse of Bear Stearns (3/24/2008) to delineate the beginning of the crisis period, and panel B reports numbers based on the split using the TARP announcement (10/14/2008) to delineate the beginning of the crisis period.

Panel A - Crisis period starts with Bear Stearns acquisition by J.P. Morgan announcement

Pre-crisis period: 1/1/2002 - 03/24/2008 Crisis period: 03/25/2008 - 01/20/2011

	Financial Firms				Nonfinancial firms				,	
Years	N	CAR (%)	Mean Return (%)	Std Dev Return (%)	_	Year	N	CAR (%)	Mean Return (%)	Std Dev Return (%)
Pre-crisis Crisis	54 69	-0.69 -2.02*	0.10 -3.66	3.40 8.66		Pre-crisis Crisis	128 209	-1.10** -3.89***	-0.73 -0.73	2.76 4.25
Total	123	-1.43*	-1.38	7.15		Total	337	-2.87***	-0.53	3.28

Panel B – Crisis period starts with TARP CPP announcement

Pre-TARP period: 1/1/2002 - 10/14/2008 Post-TARP period: 10/14/2008 – 01/20/2011

	Financial Firms				Nonfinancial firms					
Years	N	CAR (%)	Mean Return (%)	Std Dev Return (%)	_	Year	N	CAR (%)	Mean Return (%)	Std Dev Return (%)
Pre-crisis Crisis	64 59	-1.30 -1.58	-0.83 -3.31	7.02 6.86		Pre-crisis Crisis	149 188	-1.77*** -3.70***	-0.97 -0.37	3.27 3.96
Total	123	-1.43	-1.38	7.15		Total	337	-2.87***	-0.53	3.28

^{***} p<0.01, ** p<0.05, * p<0.10

Table 4

Descriptive statistics for regression variables

This table reports descriptive statistics for variables used in the regression analysis. The high yield dummy is equal to 1 if the stock issuing financial firm is rated as speculative. Ratings and issue size are obtained from S&P's Capital IQ. Market capital is measured as the market price of the firm's stock one month prior to the stock issue announcement (obtained from CRSP) times the number of shares of common stock outstanding. Assets is the book value of the firm's assets at quarter end before the stock issue announcement. Leverage is total debt divided by total assets at quarter end before the stock issue announcement. ROA is annualized net income divided by total assets at quarter end before the stock issue announcement. Volatility is the standard deviation in the firm's stock prices in the quarter prior to the stock issue announcement. All balance sheet data are obtained from Call Reports and/or SEC 10-Q filings. EDF is change in the S&P Expected Default Frequency (EDF) for the firm from the quarter before to the quarter after the stock issue announcement.

Variable	N	Mean	Median	Std Dev	Minimum	Maximum
Issue size as a % of market cap (%)	120	19.76	9.11	54.09	0.00	567.69
Total assets (millions of \$)	121	333,780	120,931	449,557	2,425	2,020,966
Leverage (%)	121	86.38	89.40	10.70	37.79	97.80
ROA (%)	121	-0.55	0.86	11.36	-76.03	48.17
Stock volatility (%)	120	3.84	2.30	3.51	14.57	19.54
EDF (1-yr) (%)	82	2.30	0.65	3.72	0.01	16.11
EDF (2-yr) (%)	82	2.53	0.79	4.05	0.01	24.19
EDF (3-yr) (%)	82	2.62	0.91	4.01	0.01	24.63
High yield dummy	1: 13					
	0: 110					

Table 5

Cross-sectional regression analysis of CAS and CAR market reactions to firm and issuance characteristics

This table presents cross-sectional regressions of CASs and CARs on firm and issuance characteristics. Specifically, we regress the three-day CAS and CAR estimates on a risk measure (a dummy for being speculative-grade rated), a measure of issuance size to market capital of the firm, and a set of firm characteristics as controls: log of assets, leverage, stock volatility and ROA). The high yield dummy is equal to 1 if the stock issuing financial firm is rated as speculative. Ratings and issue size are obtained from S&P's Capital IQ. Market capital is measured as the market price of the firm's stock one month prior to the stock issue announcement (obtained from CRSP) times the number of shares of common stock outstanding. Assets is the book value of the firm's assets at quarter end before the stock issue announcement. Leverage is total debt divided by total assets at quarter end before the stock issue announcement. ROA is annualized net income divided by total assets at quarter end before the stock issue announcement. Volatility is the standard deviation in the firm's stock prices in the quarter prior to the stock issue announcement. All balance sheet data are obtained from Call Reports and/or SEC 10-Q filings. Robust t-statistics are in parentheses.

Dependent Variable	Cumulativ	e abnormal C	DS spread	Cumulativ	Cumulative abnormal stock return		
	(1)	(2)	(3)	(4)	(5)	(6)	
Constant	-6.315*	-33,232	-32.185	-1.692	6.408	6.270	
	(0.524)	(10.379)	(14.025)	(0.425)	(2.857)	(3.633)	
High yield dummy	-36.451***	-25.556**	-25.555**	2.853*	3.396	3.396	
	(0.524)	(0.469)	(0.423)	(0.415)	(1.766)	(1.767)	
Issue size as % of market cap			-0.007			0.001	
			(0.024)			(0.005)	
Control for asset size, leverage, ROA and volatility	NO	YES	YES	NO	YES	YES	
Observations	123	120	120	123	120	120	
Adjusted R-squared	5.8%	7.7%	6.9%	0.2%	-1.5%	-2.4%	

^{***} p<0.01, ** p<0.05, * p<0.1

Table 6
Cross-sectional regression analysis of CAS and CAR market reactions to firm and issuance characteristics in the pre-crisis and crisis periods

This table presents cross-sectional regressions of CASs and CARs on firm and issuance characteristics. Specifically, we regress the three-day CAS and CAR estimates on a risk measure (a dummy for being speculative-grade rated), a measure of issuance size to market capital of the firm, and a set of firm characteristics as controls: log of assets, leverage, stock volatility and ROA). The high yield dummy is equal to 1 if the stock issuing financial firm is rated as speculative. Ratings and issue size are obtained from S&P's Capital IQ. Market capital is measured as the market price of the firm's stock one month prior to the stock issue announcement (obtained from CRSP) times the number of shares of common stock outstanding. Assets is the book value of the firm's assets at quarter end before the stock issue announcement. Leverage is total debt divided by total assets at quarter end before the stock issue announcement. ROA is annualized net income divided by total assets at quarter end before the stock issue announcement. Volatility is the standard deviation in the firm's stock prices in the quarter prior to the stock issue announcement. All balance sheet data are obtained from Call Reports and/or SEC 10-Q filings. Panel A presents results based on the split using the collapse of Bear Stearns (3/24/2008) to delineate the beginning of the crisis period, and panel B reports numbers based on the split using the TARP announcement (10/14/2008) to delineate the beginning of the crisis period. Robust standard errors are in parentheses.

Pre-crisis period: 01/01/2002 – 03/	25/2008						
Dependent Variable	Cumulativ	ve abnormal C	DS spread	Cumulative abnormal stock return			
	(1)	(2)	(3)	(4)	(5)	(6)	
Constant	-1.794	-14.628*	-19.985*	-1.007	17.106***	18.880**	
	(1.107)	(2.181)	(2.838)	(1.040)	(0.036)	(0.464)	
High yield dummy	-20.951**	-23.438**	-22.610**	7.246*	20.519**	20.245**	
	(1.107)	(0.783)	(0.976)	(1.040)	(0.980)	(1.096)	
Issue size as % of market cap			0.027*			-0.009	
			(0.003)			(0.002)	
Control for asset size, leverage, ROA and volatility	NO	YES	YES	NO	YES	YES	
Observations	54	53	53	54	53	53	
R-squared	4.4%	9.9%	9.1%	3.0%	30.3%	30.1%	

Panel A - Crisis period starts with Bear Stearns acquisition by J.P. Morgan

Crisis period: 03/25/2008 – 01/20/2011									
Dependent Variable	Cumulati	ve abnormal C	DS spread	Cumulative abnormal stock return					
	(1)	(2)	(3)	(4)	(5)	(6)			
Constant	-10.223	-185.276*	-156.639	-2.296	4.136	-0.009			
	(2.485)	(14.742)	(58.429)	(0.516)	(23.321)	(32.895)			
High yield dummy	-38.549**	17.358	20.519	1.933	-1.478	-1.935			
	(2.485)	(8.662)	(8.734)	(0.516)	(0.711)	(1.057)			
Issue size as % of market cap			-0.407			0.059			
			(0.572)			(0.127)			
Control for asset size, leverage, ROA and volatility	NO	YES	YES	NO	YES	YES			
Observations	69	67	67	69	67	67			
R-squared	4.7%	12.0%	12.5%	-1.4%	-5.1%	-3.6%			

Panel B – Crisis period starts wit	h TARP CPP	announcement					
Pre-crisis period: 01/01/2002 – 10/	14/2008						
Dependent Variable	Cumulati	ve abnormal Cl	DS spread	Cumulative abnormal stock return			
	(1)	(2)	(3)	(4)	(5)	(6)	
Constant	-1.547	-3.312*	-6.916*	-1.586	13.423*	15.244*	
	(1.134)	(0.442)	(0.761)	(0.816)	(1.196)	(1.382)	
High yield dummy	-21.197**	-34.680**	-34.223**	7.826*	15.512***	15.281***	
	(1.134)	(0.766)	(0.724)	(0.816)	(0.171)	(0.195)	
Issue size as % of market cap			0.020*			-0.010*	
-			(0.002)			(0.001)	
Control for asset size, leverage, ROA and volatility	NO	YES	YES	NO	YES	YES	
Observations	64	62	62	64	62	62	
R-squared	3.5%	10.4%	9.3%	2.2%	16.2%	15.8%	
Crisis period: 10/14/2008 – 01/20/2	2011						
Dependent Variable	Cumulati	ve abnormal Cl	DS spread	Cumulative abnormal stock return			
	(1)	(2)	(3)	(4)	(5)	(6)	
Constant	-12.250	-187.774**	-158.956	-1.826*	-6.895	-7.996	
	(5.760)	(4.063)	(47.961)	(0.161)	(17.856)	(25.014)	
High yield dummy	-36.522*	20.019	22.532	1.464*	-3.394	-3.490**	
	(5.760)	(8.993)	(11.340)	(0.161)	(0.683)	(0.140)	
Issue size as % of market cap			-0.600			0.023	
			(0.946)			(0.147)	
Control for asset size, leverage, ROA and volatility	NO	YES	YES	NO	YES	YES	
Observations	59	58	58	59	58	58	

R-squared*** p<0.01, ** p<0.05, * p<0.1

3.9%

8.7%

9.7%

-1.8%

6.1%

4.6%

Cross-sectional regression analysis of changes in the S&P expected default frequency around a stock issue announcement to firm and issuance characteristics

Table 7

This table presents cross-sectional regressions of the change in the S&P expected default frequency (EDF) around a stock issue announcement by a financial firm on firm and issuance characteristics. Specifically, we regress the change in the EDF for the issuing firm for 1-year, 2year, and 3-year horizons beginning from the quarter before the stock issue announcement on a risk measure-a dummy for being speculative-grade rated-a normalized issuance size measure (using market cap), and a set of firm characteristics as controls: log of assets, leverage, stock volatility and ROA). The high yield dummy is equal to 1 if the stock issuing financial firm is rated as speculative. Ratings and issue size are obtained from S&P's Capital IQ. Market capital is measured as the market price of the firm's stock one month prior to the stock issue announcement (obtained from CRSP) times the number of shares of common stock outstanding. Assets is the book value of the firm's assets at quarter end before the stock issue announcement. Leverage is total debt divided by total assets at quarter end before the stock issue announcement. ROA is annualized net income divided by total assets at quarter end before the stock issue announcement. Volatility is the standard deviation in the firm's stock prices in the quarter prior to the stock issue announcement. All balance sheet data are obtained from Call Reports and/or SEC 10-Q filings. Standard errors are clustered by bank. Robust standard errors are in parentheses.

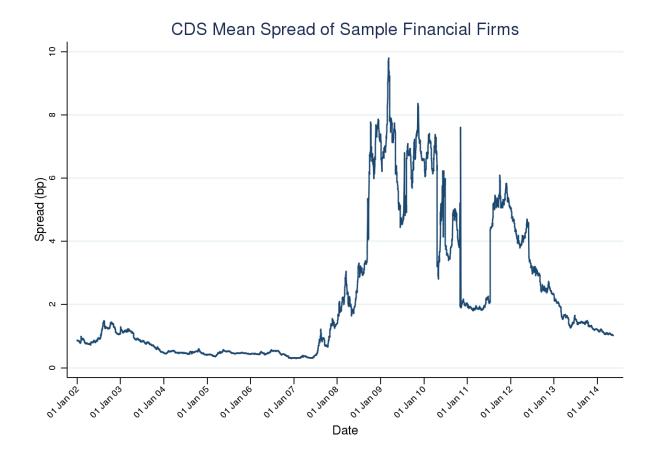
	1-yr horizon	2-yr horizon	3-yr horizon
Constant	69.342	69.297	61.769
	(44.866)	(41.732)	(37.949)
High Yield Dummy	-23.402*	-22.581*	-21.296*
·	(12.440)	(11.775)	(10.947)
Issuance Size as % of Market Cap	-0.079**	-0.066**	-0.063**
•	(0.030)	(0.027)	(0.024)
Controls for Asset Size, Leverage, ROA	1		
and Volatility	YES	YES	YES
Observations	82	82	82
R-squared	4.2%	5.8%	6.7%

^{***} p<0.01, ** p<0.05, * p<0.10

Figure 1

CDS Mean Spreads for the Sample Banks in the 2002-2013 Period

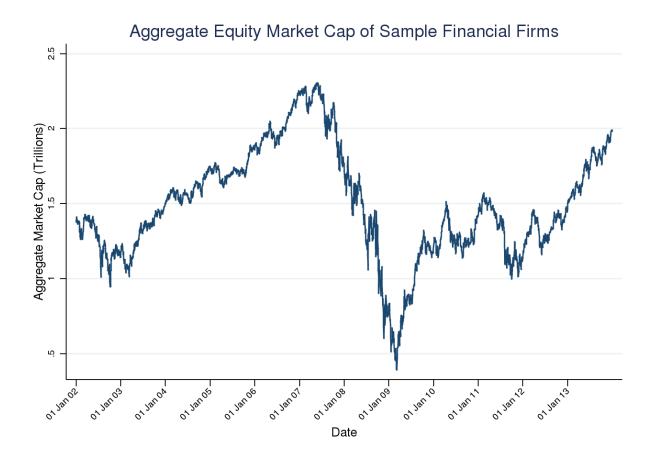
This figure plots the daily series of the mean CDS spreads of the 56 financial institutions that announced equity issuances between 2002 and 2013. This sample is collected from S&P's Capital IQ database. We then limit the sample to those financial firms having CDS pricing data reported by Markit. The firm provides CDS end-of-day quotes for the CDS contracts



Aggregate Market Equity Capitalization for the Sample Banks in the period 2002-2013

Figure 2

This figure plots the daily evolution of the combined equity market capitalization of the 56 financial institutions that announced equity issuances between 2002 and 2013. Equity market capitalization is calculated as the daily stock price times the number of shares of common stock outstanding (from CRSP data tapes).



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