

Villaverde, José

Conference Paper

Regional convergence, polarization and mobility in the European Union, 1980-1996

42nd Congress of the European Regional Science Association: "From Industry to Advanced Services - Perspectives of European Metropolitan Regions", August 27th - 31st, 2002, Dortmund, Germany

Provided in Cooperation with:

European Regional Science Association (ERSA)

Suggested Citation: Villaverde, José (2002) : Regional convergence, polarization and mobility in the European Union, 1980-1996, 42nd Congress of the European Regional Science Association: "From Industry to Advanced Services - Perspectives of European Metropolitan Regions", August 27th - 31st, 2002, Dortmund, Germany, European Regional Science Association (ERSA), Louvain-la-Neuve

This Version is available at:

<https://hdl.handle.net/10419/115623>

Standard-Nutzungsbedingungen:

Die Dokumente auf EconStor dürfen zu eigenen wissenschaftlichen Zwecken und zum Privatgebrauch gespeichert und kopiert werden.

Sie dürfen die Dokumente nicht für öffentliche oder kommerzielle Zwecke vervielfältigen, öffentlich ausstellen, öffentlich zugänglich machen, vertreiben oder anderweitig nutzen.

Sofern die Verfasser die Dokumente unter Open-Content-Lizenzen (insbesondere CC-Lizenzen) zur Verfügung gestellt haben sollten, gelten abweichend von diesen Nutzungsbedingungen die in der dort genannten Lizenz gewährten Nutzungsrechte.

Terms of use:

Documents in EconStor may be saved and copied for your personal and scholarly purposes.

You are not to copy documents for public or commercial purposes, to exhibit the documents publicly, to make them publicly available on the internet, or to distribute or otherwise use the documents in public.

If the documents have been made available under an Open Content Licence (especially Creative Commons Licences), you may exercise further usage rights as specified in the indicated licence.

REGIONAL CONVERGENCE, POLARISATION AND MOBILITY IN THE EUROPEAN UNION, 1980-1996

José Villaverde Castro

University of Cantabria and FUNCAS

Abstract

This paper deals with three relevant questions for the European Union: the degrees of regional inequality, polarisation and mobility in the per capita regional income distributions between 1980 and 1996. Using different techniques, the paper shows that inter-regional inequality has slightly decreased, there is no apparent polarisation and the degree of mobility is relatively high. Despite this, the speed of all these three aspects has declined over time, which is interesting for polarisation but raises some worries about convergence and mobility for the future.

Key words

Inequality, convergence, polarisation, mobility, European regions.

JEL code

R11, R23

1. Introduction

Regional disparities in the European Union (EU) have been a subject of great interest and debate since, at least, the early seventies. In fact, the extent of these disparities was the main reason to establish the ERDF in 1975 and the European regional policy. Although most of the studies agree that regional disparities in the EU decreased during the sixties and seventies, it is also widely admitted that there has been little progress since then and that its actual degree is too high. Furthermore, there are some concerns that regional disparities in the EU may increase in the near future as a result of both the widening and deepening of the integration process.

Along with the new developments of the theory of economic growth questioning the relevance of automatic spatial convergence, the European integration process itself has encouraged the research on regional convergence. This paper, inserted in this fruitful tradition, tries to identify some of the most relevant features of the economic evolution of the European regions between 1980 and 1996. To accomplish this aim, the statistical information used comes from the REGIO and CRENoS databanks and refers to 141 European regions, both NUTS 2 and NUTS 1 (See appendix for the whole list of regions); the key variable of the analysis is per capita GDP expressed in purchasing power standards.

The remainder of the paper is organised as follows. Section 2 analyses the level and recent evolution of regional inequalities in the

EU. Section 3 examines the external shape of the European regional income distribution, while section 4 deals with its dynamics. A final section summarises the main conclusions.

2. Regional disparities in the EU

According to the latest Eurostat release (February 2002), averaging over the years 1997-1999 and taking into consideration purchasing power standards, per capita GDP of the richest region in the EU, Inner London, was 4.7 times higher than that in the poorest regions, Ipeiros, Reunion and Guayana. This is a fairly large number that roughly illustrates the actual extent of regional inequalities in the EU. Other official publications by the European Commission show that, during the last ten years, the extent of these disparities has not changed very much and that the mobility in the income distribution has been quite low; in fact the ranking of the ten (twenty five) most and least prosperous regions is, nowadays, roughly the same as a decade ago.

Although important, these conclusions need to be somehow qualified, because they are critically dependent on three factors: the period under study, the number of regions considered in the analysis and the statistic indicators used to measure inequality. It is then necessary to deal more in-depth with regional inequalities in the EU and this is done through the computation of some inequality indicators (Gini, Theil and Atkinson, with different degrees of poverty aversion). The results (Table 1) permit to highlight two main conclusions. The first, that the degree of inequality in the regional per

capita income (GDP) distribution in the EU shows a clear-cut reduction. Nevertheless, the strength of this reduction is too varied, the lowest rate realised by the Gini index (with a fall equal to 9.3%) and the highest rate corresponding to the Atkinson index A(2), with a fall equal to 26,5%. The second conclusion is that although there is an apparent similarity in the inequality reduction processes, as shown by all the indices (Figure 1), the evolution has not been uniform at all over time: in particular the highest volatility is associated to some Atkinson indices, mainly those displaying a higher poverty aversion, A(50) and A (100).

Spatial disparities in the EU are higher inside countries than among them. As the Theil index is additively decomposable, we have used it to show which part of the European regional inequality is due to the inequality among countries (external inequality) and which one comes from the inequalities inside countries (internal inequality). The results (Table 2) show that about three fourths of total inequality originates within each country and that just one fourth comes from the external inequality. These results support the idea of a European regional policy. At the same time, Table 2 shows that the decrease in the global inequality index has been achieved thanks to both the reduction of the internal and external inequalities. Nevertheless, this has been attained to a major extent due to the reduction of external inequality (the fall was of 41.3%) rather than that of internal inequality (the fall was 10.5%), which means that the convergence process was stronger among countries than among regions in each country. As a result of this evolution, internal inequalities have gained weight in the global inequality; to say in a different way, this implies that the richest regions of the poorest countries have been the main contributors to the decline of the European regional inequality.

Even though the general trend of European regional inequality is a decreasing one, Figure 1 also shows the existence of some cyclical movements around it. Are these ups and downs somehow linked to the evolution of GDP in the EU? A well known hypothesis maintains that spatial inequalities go hand in hand with the rate of growth of economic activity: inequalities increase during expansions and decrease during contractions or periods of stagnating activity. Has this been happening in the EU? To give a simple answer to this question we have regressed the degree of regional inequality (as measured by the Theil index) on the growth of GDP in the EU. The results obtained

$$\text{Theil}_t = 0,03276 + 0,00042 \Delta \text{PIB}_{t, t-1} \quad R^2 = 0,35$$

(30,5) (2,8) (t statistics in brackets)

show that the hypothesis of a positive nexus between both elements cannot be rejected. Thus, as much as the economic cycle can be smoothed in the EU, a reduction of cyclical movements in the general tendency of reduction of the European regional inequality can also be expected.

Taking into account that in the EU coexist very different regional situations, we have also studied how the convergence process has taken place. For this, we have considered seven groups of regions according to their per capita GDP in 1980, being the EU average equal to 100: 1<50%; 50≥2>75%; 75≥3>90%; 90≥4>110%; 110≥5>130%; 130≥6>150%; 7≥150%¹. The results obtained for the Theil index (Figure 2) show that, except for group 4, regional inequality has increased. As a consequence, the fall in the overall

regional inequality has its roots in the fall of inter-group inequality. Furthermore, according to Garcia Greciano (1997) we have calculated the inter-groups inequality as the difference between the overall and intra-group inequalities, the latter computed as the simple average of the inequality in the seven groups previously mentioned. The results (see Figure 2 again) show that the bulk of overall inequality (between 81 and 94%) is due to inter-groups inequality.

Finally, another interesting point that emerges when dealing with regional inequalities in per capita GDP is whether there is any association between this variable and the degree of regional concentration of GDP. This is a relevant point since it is not the same for the reduction in the overall inequality index to come, for instance, from losing population the poorest regions than from a relative lower rate of economic growth in the richest regions.

In the EU, leaving aside some particular situations (i.e. Darmstadt and Luxemburg on the positive side and Asturias on the negative one), convergence or reduction of overall regional inequality has not implied relevant shifts in the regional distribution of output and population. The Gini-Hirschman² concentration index shows (Table 3) that changes have been very modest (0.20% for output and 0.29% for population) and that the indices have also been very stable over time: the coefficient of variation for both variables is very low. This means that, contrary to what happened in previous years, European regional convergence in the eighties and the first half of the nineties has not taken place at the expense of depopulation of the least developed regions.

3. Per capita income distribution in the EU. Is there any polarisation?

Conventional inequality indicators (Gini, Theil, Atkinson) mainly capture the spread of an income distribution, underlying only the deviations from the global average and ignoring whether, or not, there is clustering around different local poles. On the contrary, polarisation places more emphasis on clustering, so as to compare the homogeneity of a group with the overall heterogeneity of a given population.

A simple way to deal with the potential existence of regional polarisation in the EU is through the estimation of density functions for the regional distribution of per capita GDP. Figure 3 plots simple and weighted density functions for four selected years estimated using a gaussian kernel with bandwidth selected as in Silverman (1986). A straightforward visual inspection of these functions shows that our main and previous conclusion –decreasing inequality or convergence in regional per capita income across the EU- is also supported in this analysis. First of all, the ratio between the extreme values of the distribution has declined over time; that is, the spread of relative incomes has decreased. Secondly, the external shape of the distribution for the relative incomes shows a clear unimodality situation when dealing with simple density functions. On the contrary, when dealing with weighted density functions, an incipient bimodality, or polarisation in two groups, can be observed: the first group refers to regions with a per capita GDP similar to the mean (which, once again, implies convergence) while the second group is made up of regions with very high relative income levels. Thirdly, density functions are lightly skewed to the right, although, in general terms, the degree of skewness has declined over time. And finally,

given that the probabilistic mass has been shifting little by little to the right, not only the convergence conclusion is supported once again but also the fact that this convergence is moving towards the European average. Thus, contrary to the conclusion by Boldrin and Canova (2001), we obtain that the distributions collapse towards their central value.

4. Mobility within the European regions

Density functions, as those shown in Figure 3, offer an appealing description of some interesting characteristics of the relative income distribution of the European regions at different times. Unfortunately, they do not offer any insight about the intra-distribution dynamics or, using Quah words, “*churning-like behavior when individual economies transit from one part of the distribution to another*” (Quah, 1996). Taking into consideration that, from a policy-oriented point of view, mobility in the per capita income ranking of the European regions is important, we are interested in finding out how many of these European regions have changed their position in the ranking over time. In the case of many regions changing their position, it is said that there exists mobility; otherwise, it is said there is persistence in the distribution.

The Sixth Periodic Report on the Social and Economic Situation and Development of Regions in the European Union (European Commission, 1999) analyses this question and, considering that the Spearman’s rank correlation coefficient between the rankings in 1986 and 1996 was 0.91, concludes that “*there has been little change in the*

ranking of particular regions and the order in terms of GDP per head was much the same in 1996 as ten years earlier". Is this a right conclusion? No doubt it is when dealing with the context of the Sixth Report; however, our context, with 141 regions and a larger period of analysis (1980-1996), presents a much more attractive profile.

A standard methodology currently used to trace the movements within a distribution is based on the estimation of transition probability matrices. This is so because Markov chains -in which these transition matrices are based- provide useful representations of dynamic processes like the one we are dealing with. Thus, if we denote by F_0 and F_t the initial and final distributions, the link between them can be written as $F_t = M^t * F_0$, where M^t represents the transition probability matrix. What this expression describes is simply the time evolution of F_0 , by mapping F_0 into F_t . Operator M^t is approximated by discretizing our set of values (of regional per capita income) into intervals. Due to the lack of sound theoretical methods proposed to obtain a (more or less) appropriate partition of the distribution (i.e. the minimum variance criterion of Cochran (1966)) we have somewhat arbitrarily partitioned our distribution into the same 7 classes or states used when computing groups inequalities³.

According to this grouping, the transition probability matrix obtained -computing it in one-step- is displayed in Table 4. In particular, for the whole sample period (80-96) the main features are the following:

1. Persistence is important however not too high. Although, on average, 58.9% of the regions have remained in their initial state, the

figure goes up to 80% for the regions in state 2 and down to 48,2% for the regions in state 5.

2. The mobility degree has been quite high: 41.1% of the regions have moved from one state to another. This mobility has been a little bit higher in descending than in ascending terms. This suggests a decreasing convergence, as shown in section 2.

3. Provinces changing the class they were in simply move to a contiguous one. This means, according to Jone's terminology, that there are neither growth miracles nor growth disasters: this is quite standard result even when the matrix is computed in one-step. To be precise, there are just three "miracles" (Ireland, jumping from class 2 to class 4, Utrecht from 3 to 5 and Luxembourg from 5 to 7) and two "disasters" (Drenthe and Picardie, both moving from class 5 to class 3).

4. The visual inspection of the density functions for the initial and final years of the sample confirms convergence towards the average: the number of regions with an income per capita similar to the mean (those in class 4) increases from 29.8% in 1980 to 34% in 1996.

5. And finally, another important result that can be seen in the table is that the ergodic distribution is unimodal, with a maximum in the fourth class: this implies convergence in the long run (and convergence to the average) instead of polarisation. Nevertheless, the long run regional income distribution is also skewed to the right: 38.5% of the European regions will be in the third class and another 43% will be in the fourth.

Taking into consideration that the transition matrix has been estimated in only one step, we have no clear idea as to how mobility has evolved over time. In order to see this evolution, it is necessary to

estimate transition matrices step by step and then calculate a mobility index, i.e. Shorrocks' index which is given by the expression

$$M(A) = [m - \text{tr}(A)] / (m - 1)$$

where A is the transition matrix, $\text{tr}(A)$ is its trace and m the number of elements in the distribution. When applying this expression to our distribution ($m=141$ regions) the results, shown in Figure 4, are:

- 1.- Between 1980-81 and 1995-96 the degree of mobility fell by 25%;
- 2.- The degree of mobility is very low every two consecutive years;
- 3.- The reduction in the mobility degree has not been uniform over time but very volatile, except between 1989-90 and 1993-94.

These results, although interesting, should only be considered as an approximation to the actual mobility since, in order to obtain them, only the elements of the principal diagonal of the A (step by step) matrices have been considered. Nevertheless, these results suggest that, although not negligible, the interregional mobility degree in the EU has been declining over time. Furthermore, although there have been some advances in the process of regional convergence in the EU this has coexisted with a decline on the probability of a region to move to another income class.

5.- Conclusions

Per capita regional income disparities in the EU are very large (about twice the level of USA). Despite of that, this paper has obtained some results that mitigate their seriousness. First of all, it has been shown

that regional disparities (when computed by different inequality indices) have been declining, although to a decreasing rate. Secondly, there are no clear signs of polarisation in two or more local poles. Thirdly, the mobility degree in the regional income distributions is quite high (more than 41% of the European regions have experienced a change in their income classes between 1980 and 1996). Overall, the degree of mobility year to year is not only much lower but has also declined, although following a very volatile pace. In other words, regional income disparities in the EU in 1996 are less serious than in 1980. Nevertheless, the reduction in both the speed of convergence and the mobility within regional per capita GDP distribution raise some concerns for the foreseeable future.

REFERENCES

- Amstrong, H.W. (1995) Convergence among regions of the European Union, 1950-1990. *Papers in Regional Science*, 74, pp.143-152.
- Azzoni, C. (2001) Economic growth and regional income inequality in Brazil. *The Annals of Regional Science*, nº 1, pp. 133-152.
- Barro, R. y Sala-i-Martin, X. (1991) Convergence across states and regions. *Brookings Papers on Economic Activity*, nº 1, pp. 107-182.
- Boldrin, M. y Canova, f. (2001) Inequality and convergence in Europe's regions: reconsidering European regional policies. *Economic Policy*, nº 32, pp. 207-253.
- Chatterji, M. (1993) Convergence clubs and endogenous growth. *Oxford Review of Economic Policy*, 8, pp. 57-69.
- Cuadrado, J.R. (2001) Regional convergence in the European Union: From hypothesis to actual trends. *The Annals of Regional Science*, vol 35, pp. 333-356.
- Dalgaard C-J y Vastrup, J. (2001) On the measurement of σ -convergence. *Economic Letters*, nº 70, pp.283-287.
- Drennan. M.P. y Lobo, J. (1999) A simple test for convergence of metropolitan income in the United States. *Journal of Urban Economics*, nº 3, pp. 350-359.
- Esteban, J.M. (1996) Desigualdad y polarización. Una aplicación a la distribución interprovincial de la renta en España. *Revista de Economía Aplicada*, nº 11, pp. 5-26.

- Esteban, J.M. y Ray, D. (1994) On the measurement of polarisation. *Econometrica*, 62, pp.819-81.
- Friedman, M. (1992) Do old fallacies ever die?. *Journal of Economic Literature*, 30, pp.2129-2132.
- García Greciano, B. (1997) Distribución de renta, crecimiento y convergencia regional en España. Mimeo (Doctoral dissertation, Universidad Complutense de Madrid).
- Goerlich, F. y Mas. M (2001) Tres cuestiones (¿marginales?) sobre convergencia. Mimeo (Universidad de Valencia e IVIE).
- Lopez Bazo, E., Vayá, E. , Mora, J. y Suriñach, J. (1998). Regional economic dynamics and convergence in the European Union. *The Annals of Regional Science*, nº 3, pp. 343-370.
- Magrini, S. (1999) The evolution of income disparities among the regions of the European Union. *Regional Science and Urban Economics*, vol. 59, pp. 257-281.
- Quah, D. (1993) Empirical cross-section dynamics in economic growth. *European Economic Review*, nº 2-3, pp.426-434.
- Quah, D. (1997) Empirics for growth and distribution: Stratification, polarisation and convergence clubs. *Journal of Economic Growth*, 2, 27-59.
- Overman, H. y Puga, D. (2000) Unemployment clusters across European regions and countries. CEP, Discussion Paper nº 434.
- Neven, D. y Gouyette, C (1995) Regional convergence in the European Community. *Journal of Common Market Studies*, 33, pp.47-65.
- Rapún, M., Gil, C., Escurra, R. y Pascual. P. (2001) Movilidad y desigualdad regional en la Unión Europea. Mimeo.
- Rodriguez Pose, A. (1998) Convergencia y modelos de crecimiento regional en Europa. En Cuadrado et al. "Convergencia regional en España. Hechos, tendencias y perspectivas", F. Argenteria y Visor, Madrid, pp. 71-109.
- Shorrocks, A.F. (1978) The measurement of mobility. *Econometrica*, vol. 46, pp. 1013-1024.
- Villaverde, J. (2001) La distribución espacial de la renta en España: 1980-1995. *Papeles de Economía Española*, nº 88, pp. 166-181.
- Villaverde, J. y Sánchez-Robles, B. (2001) Polarización, convergencia y movilidad entre las provincias españolas:1955-1997. *Revista Asturiana de Economía*, nº 20, pp. 7-26.
- Villaverde, J. y Sánchez-Robles, B. (2002) Convergence or Twin Peaks? The Spanish case. En Meeusen y Villaverde (eds.) "Convergence issues in the European Union", Edward Elgar, Cheltenham, (en prensa)

1 This partition is more illustrative of the European situation than the conventional one which considers only five intervals, <75, 75-90, 90-110, 110-125 y >125.

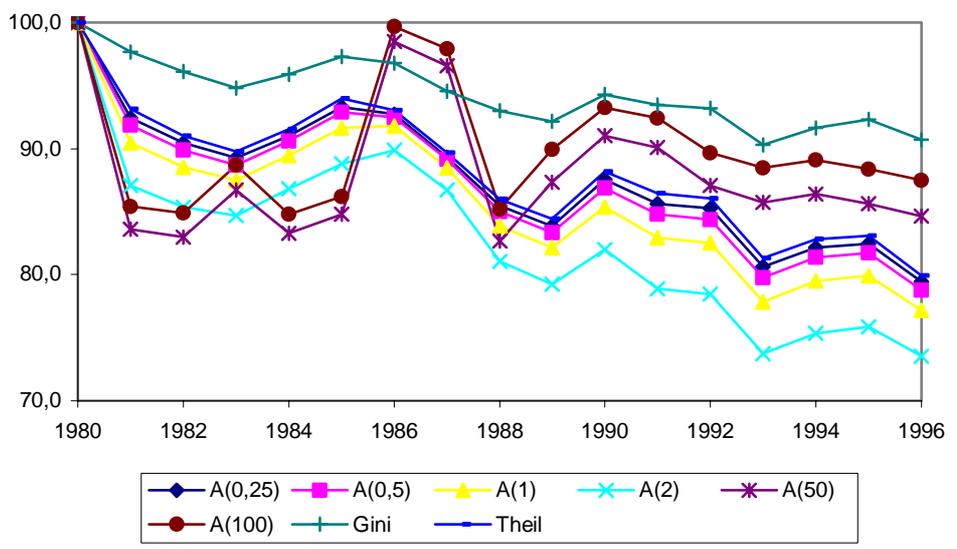
2 The Gini-Hirschman index is given by the expression

$$C_i = \left[\sum_i (x_i / X)^2 \right]^{1/2}$$

where X represents the variable under consideration (GDP and population in our case) and the sub-index i denotes the European regions.

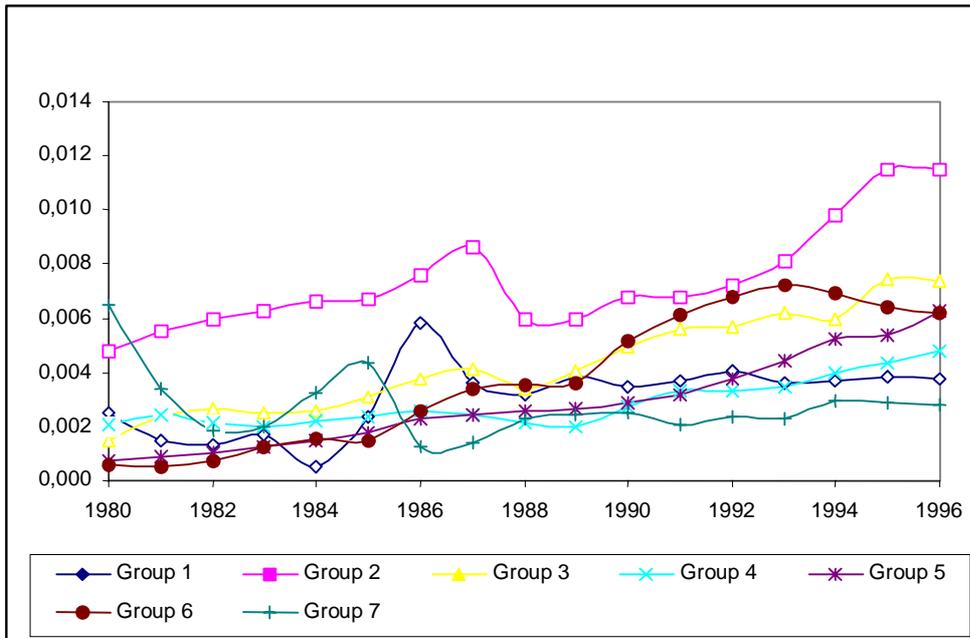
- 3 The computing of stochastic kernels solves these problems by using a continuous approach instead of a discrete one. The main drawback of this approach is that its interpretation is not as direct and clear as in the discrete approach, because it does not offer quantitative information about the degree of mobility.

**Figure 1.- REGIONAL INEQUALITY IN THE E.U.
(1980=100)**



**Figure 2.- DECOMPOSITION OF REGIONAL INEQUALITY IN THE E.U.
(Theil Index)**

a) Groups, inequality



b) Intra-group, inter-group and total inequality

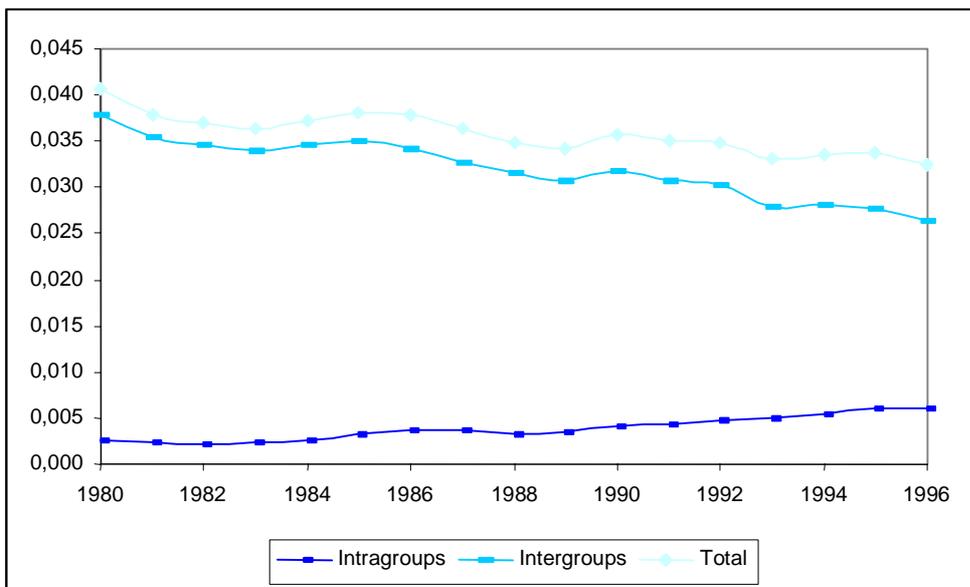
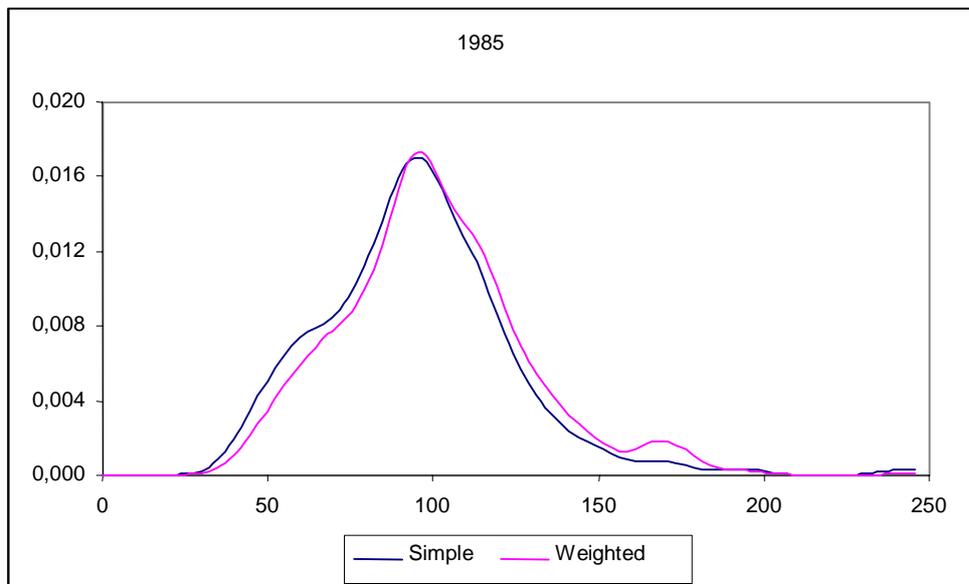
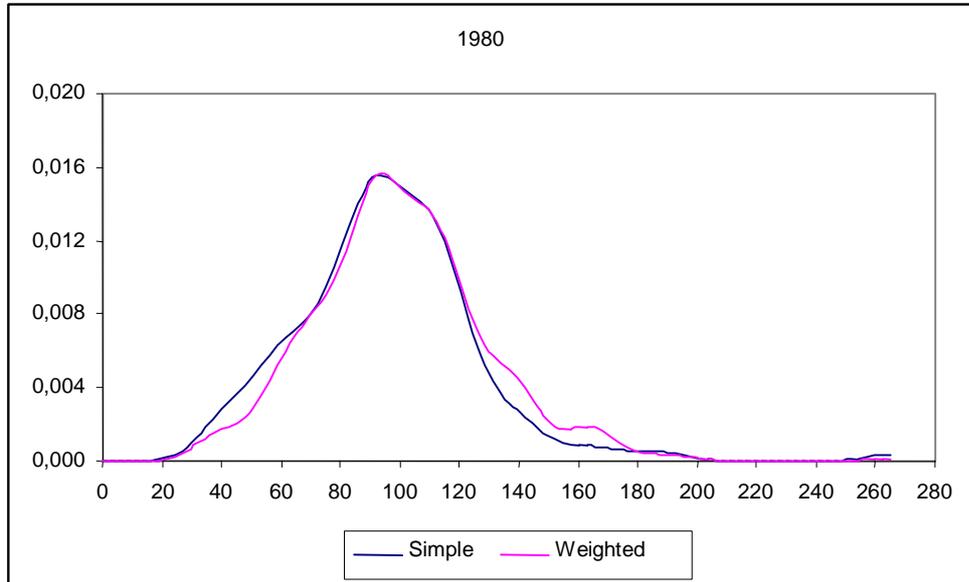
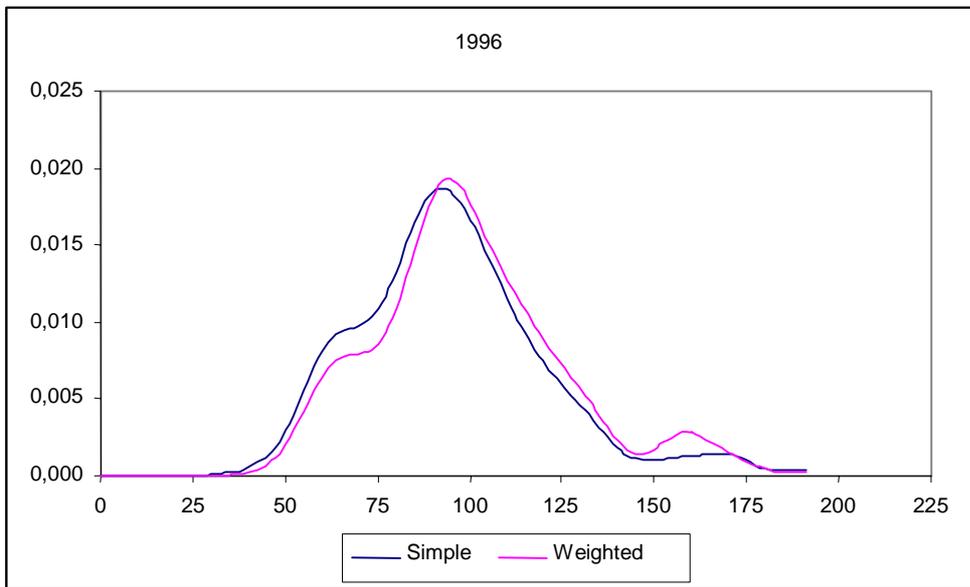
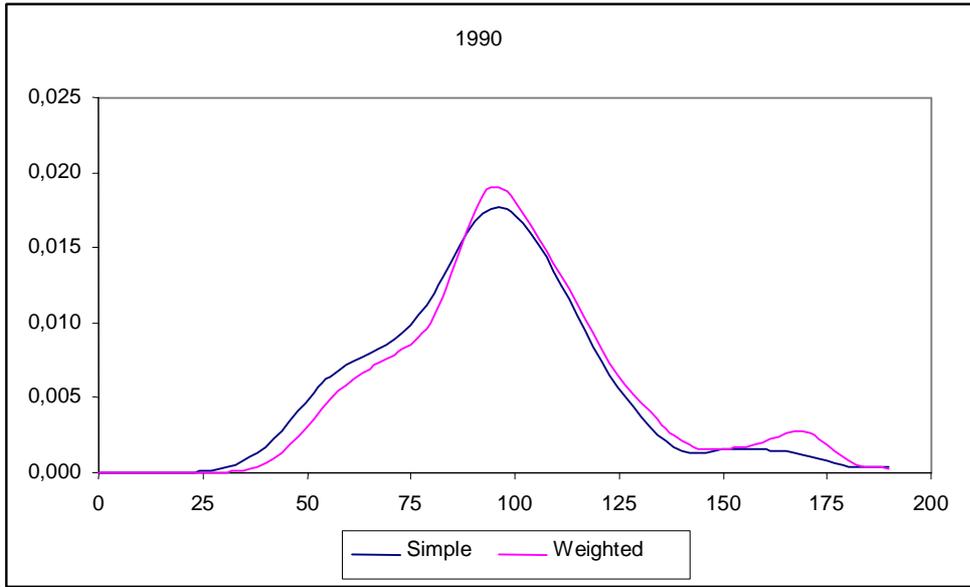


Figure 3.- DENSITY FUNCTIONS



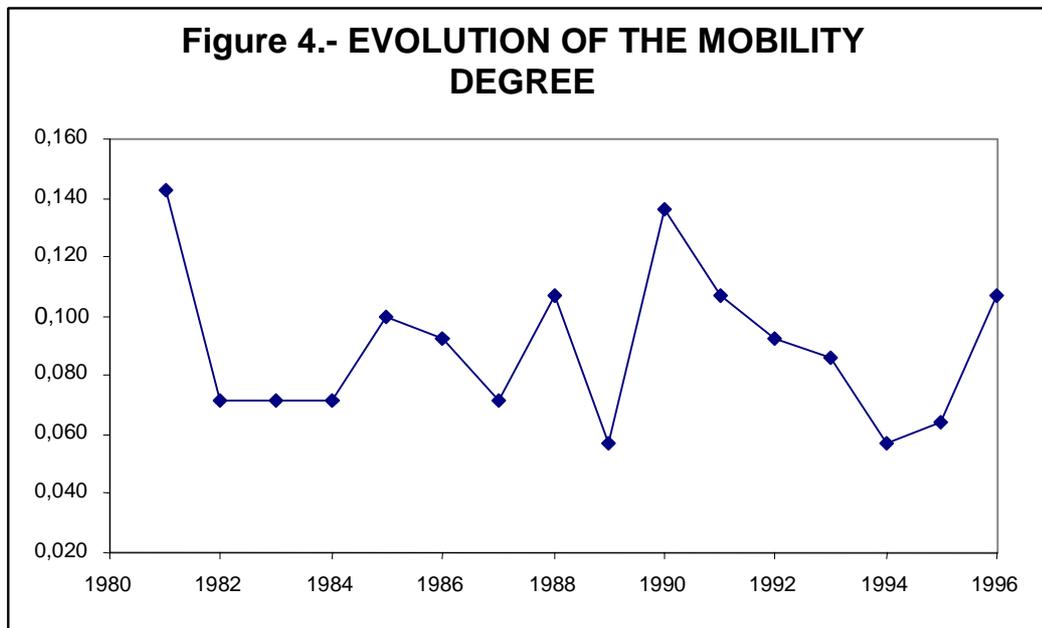


Table 1.- REGIONAL INEQUALITY IN THE E.U.

	A(0,25)	A(0,5)	A(1)	A(2)	A(50)	A(100)	Gini	Theil
1980	0,0100	0,0202	0,0410	0,0851	0,5923	0,6117	0,1549	0,0406
1981	0,0093	0,0185	0,0371	0,0741	0,4953	0,5224	0,1513	0,0378
1982	0,0091	0,0181	0,0363	0,0726	0,4914	0,5194	0,1489	0,0370
1983	0,0090	0,0179	0,0358	0,0720	0,5134	0,5424	0,1468	0,0364
1984	0,0091	0,0183	0,0367	0,0739	0,4933	0,5186	0,1486	0,0372
1985	0,0094	0,0187	0,0375	0,0755	0,5022	0,5274	0,1507	0,0382
1986	0,0093	0,0187	0,0376	0,0765	0,5835	0,6097	0,1499	0,0378
1987	0,0090	0,0180	0,0362	0,0738	0,5719	0,5989	0,1465	0,0364
1988	0,0086	0,0172	0,0344	0,0690	0,4896	0,5212	0,1441	0,0349
1989	0,0084	0,0168	0,0337	0,0674	0,5173	0,5501	0,1427	0,0343
1990	0,0088	0,0175	0,0350	0,0697	0,5392	0,5703	0,1461	0,0358
1991	0,0086	0,0171	0,0340	0,0671	0,5337	0,5653	0,1447	0,0351
1992	0,0086	0,0170	0,0338	0,0667	0,5156	0,5485	0,1444	0,0349
1993	0,0081	0,0161	0,0319	0,0627	0,5077	0,5411	0,1398	0,0330
1994	0,0082	0,0164	0,0326	0,0641	0,5118	0,5449	0,1420	0,0336
1995	0,0083	0,0165	0,0327	0,0645	0,5072	0,5406	0,1430	0,0337
1996	0,0080	0,0159	0,0316	0,0625	0,5014	0,5352	0,1405	0,0325
% change	20,5	21,2	22,8	26,5	15,3	12,5	9,3	20,1

Table 2.- THEIL INDEX DECOMPOSITION

	IndEX	Inequality			
		External	External(%)	Internal	Internal(%)
1980	0,0406	0,0126	31,1	0,0280	68,9
1981	0,0378	0,0121	31,9	0,0258	68,1
1982	0,0370	0,0117	31,6	0,0253	68,5
1983	0,0364	0,0109	30,0	0,0255	70,0
1984	0,0372	0,0113	30,4	0,0259	69,6
1985	0,0382	0,0112	29,4	0,0269	70,6
1986	0,0378	0,0108	28,6	0,0270	71,4
1987	0,0364	0,0098	26,9	0,0266	73,1
1988	0,0349	0,0098	28,0	0,0251	72,0
1989	0,0343	0,0096	27,9	0,0247	72,1
1990	0,0358	0,0106	29,5	0,0252	70,5
1991	0,0351	0,0104	29,7	0,0247	70,3
1992	0,0349	0,0100	28,6	0,0250	71,4
1993	0,0330	0,0084	25,3	0,0247	74,7
1994	0,0336	0,0089	26,3	0,0248	73,7
1995	0,0337	0,0089	26,5	0,0248	73,5
1996	0,0325	0,0074	22,9	0,0250	77,1

Table 3
G.D.P. AND POPULATION REGIONAL CONCENTRATION
(Gini-Hirschman Index)

	PIB	Población
1980	11,95	10,99
1981	11,96	11,00
1982	12,04	11,01
1983	12,05	11,01
1984	12,01	11,01
1985	12,06	11,02
1986	12,10	11,02
1987	12,15	11,02
1988	12,12	11,03
1989	12,13	11,04
1990	12,11	11,03
1991	12,08	11,03
1992	12,05	11,02
1993	12,03	11,02
1994	12,01	11,02
1995	11,97	11,02
1996	11,98	11,02
Coefficient of variation	0,0051	0,0011
% change	0,20	0,29

Table 4.- TRANSITION MATRIX
(1980-1996)

Classes	1	2	3	4	5	6	7
1	0,1429	0,8571	0,0000	0,0000	0,0000	0,0000	0,0000
2	0,0000	0,8000	0,1600	0,0400	0,0000	0,0000	0,0000
3	0,0000	0,1111	0,5185	0,3333	0,0370	0,0000	0,0000
4	0,0000	0,0000	0,2857	0,6429	0,0714	0,0000	0,0000
5	0,0000	0,0000	0,0741	0,4074	0,4815	0,0000	0,0370
6	0,0000	0,0000	0,0000	0,0000	0,1250	0,6250	0,2500
7	0,0000	0,0000	0,0000	0,0000	0,0000	0,4000	0,6000
Initial distrib.	0,0496	0,1773	0,1915	0,2979	0,1915	0,0567	0,0355
Final distrib.	0,0071	0,2057	0,2270	0,3404	0,1277	0,0496	0,0426
Ergodic distrib.	0,0000	0,0484	0,3835	0,4300	0,0932	0,0276	0,017

ANNEX LIST OF REGIONS

Belgium

Bruxelles-Brussels
Antwerpen
Limburg (B)
Oost-Vlaanderen
Vlaams Brabant
West-Vlaanderen
Brabant Wallon
Hainaut
Liège
Luxembourg (B)
Namur

Denmark**Germany**

Stuttgart
Karlsruhe
Freiburg
Tübingen
Oberbayern
Niederbayern
Oberpfalz
Oberfranken
Mittelfranken
Unterfranken
Schwaben
Bremen
Hamburg
Darmstadt
Gießen
Kassel
Braunschweig
Hannover
Lüneburg
Weser-Ems
Düsseldorf
Köln
Münster
Detmold
Arnsberg
Koblenz
Trier

Merseyside)

Rheinessen-Pfalz
Saarland
Schleswig-Holstein

Greece

Anatoliki Makedonia, Thraki
Kentriki Makedonia
Dytiki Makedonia
Thessalia
Ipeiros

Ionia Nisia

Dytiki Ellada

Stereia Ellada

Peloponnisos

Attiki

Voreio Aigaio

Notio Aigaio

Kriti

Spain

Galicia

Principado de Asturias

Cantabria

Pais Vasco

Comunidad Foral de Navarra

La Rioja

Aragón

Comunidad de Madrid

Castilla y León

Castilla-la Mancha

Extremadura

Cataluña

Comunidad Valenciana

Baleares

Andalucía

Murcia

Canarias (ES)

France

Île de France

Champagne-Ardenne

Picardie

Haute-Normandie

Centre

Basse-Normandie

Bourgogne

Nord - Pas-de-Calais

Lorraine

Alsace

Franche-Comté

Pays de la Loire

Bretagne

Poitou-Charentes

Aquitaine

Midi-Pyrénées

Limousin

Rhône-Alpes

Auvergne

Languedoc-Roussillon

Provence-Alpes-Côte d'Azur

Corse

Ireland**Italy**

Piemonte

Valle d'Aosta

Liguria

Lombardia

Trentino-Alto Adige

Veneto

Friuli-Venezia Giulia

Emilia-Romagna

Toscana

Umbria

Marche

Lazio

Abruzzo

Molise

Campania

Puglia

Basilicata

Calabria

Sicilia

Sardegna

Luxembourg**Netherlands**

Groningen

Friesland

Drenthe

Utrecht

Noord-Holland

Zuid-Holland

Zeeland

Noord-Brabant

Limburg (NL)

Portugal

Norte

Centro (P)

Lisboa e Vale do Tejo

Alentejo

Algarve

United Kingdom

North East

North West (including

Yorkshire and The Humber

East Midlands

West Midlands

Eastern

South East

South West

Wales

Scotland

Northern Ireland