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Long working hours and alcohol use: systematic review and meta-analysis of published studies and unpublished individual participant data

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Appendix 2

Although risky alcohol use is common in the workforce, little attention has been paid to any association with work. A recent large scale meta-analysis of data from individual participants provided no clear support for an association between perceived “job strain” (an index of work stress) and risky alcohol use. Another work related factor—long working hours—has been associated with several adverse health outcomes that are linked to alcohol use, including cardiovascular diseases, depression and anxiety, sleep deprivation, and occupational injuries. Direct evidence on the association between long working hours and alcohol use, however, is based on relatively small studies with insufficient power to detect weak or moderate associations. Three narrative reviews, published in 1999-2006, summarised findings from up to six studies that reported inconclusive results, some studies showing increased alcohol use among people with long working hours, and other studies observing no relation. To date, there has been no systematic quantification of the link between long working hours and alcohol use.

Meta-analysis is a powerful tool to examine associations between exposures and health outcomes as it is based on a greater range of participants, a larger sample size, and more events than any individual study. A traditional meta-analysis of published studies is a quantitative synthesis of published aggregate data with a limited ability to control for the effects of publication bias (that is, selective reporting of positive findings) and heterogeneity in the methods used (for example, cut-off levels, method of measurement, analytic strategies, etc). Meta-analysis of individual participant data provides an opportunity to deal with these biases by inclusion of unpublished individual level data to the analysis and by standardisation of definitions of the exposures, outcomes, and statistical methods within these data.

We carried out a meta-analysis of 36 published studies with aggregate data (34 cross sectional and two prospective), identified by a systematic review and supplemented with unpublished individual participant data from 27 studies, of which 18 provided data for prospective analysis. We examined whether long working hours are associated with alcohol use cross sectionally, longitudinally, and in relation to sustained risky alcohol use. We also examined whether any association is apparent across sociodemographic groups, geographical regions, and characteristics of the study cohort. Finally, we compared the findings from published studies and studies with unpublished individual participant data to evaluate publication bias.
(working and hours) or (overtime and work) or overtime). Outcome search terms were “alcohol or alcohols or ethanol or wine or beer or liquor or spirit or alcoholism or (alcohol and dependence) or (alcohol and abuse) or (substance) or (drinking and behaviour) or (alcohol and drinking) or (alcoholic and beverages) or (alcohol and consumption) or (alcohol and misuse) or (heavy and drinking) or lifestyle or (life and style) or (health and behaviours) or (health and behaviour)”. We also scrutinised the reference list of all eligible publications and performed a cited reference search of these using the Institute of Scientific Information Web of Science (to May 2014) to identify all studies citing the included studies.

Study selection criteria

We included published studies providing aggregate data if they met the following criteria: empirical and peer reviewed study; at least an abstract with estimates and/or full results published in English; individual level exposure and outcome data; examined the association between working hours and alcohol use (reported either by a defined high level of alcohol consumption or as a continuous variable); reported either estimates of relative risk, odds ratios, or hazard ratios with 95% confidence intervals or data to calculate these. Cross sectional, case-control, prospective studies, and trials were eligible for this meta-analysis.

Data extraction from studies with aggregate data

We extracted the following information from each retrieved article: name of first author, start of follow-up (year) or publication year if start year not reported, study location (country), population, number of participants, mean follow-up time, mean age or age range, proportion of women, method of assessment and definition of long working hours, proportion of employees with long hours, method of assessment of alcohol use, definition of risky alcohol use (when applicable), proportion of employees with risky use at baseline and proportion of new cases at follow-up, and covariates included in the adjusted models.

Individual participant data

We also identified 28 datasets with relevant individual participant data. Of these, we excluded the Intervention Project on Absence and Well-being (IPAW) study22 from the analysis because there were no cases of risky alcohol use among employees with long working hours. Of the remaining 27 datasets, we obtained 15 from the Individual-Participant-Data Meta-analysis in Working Populations (IPD-Work) Consortium, which we lead, and 12 from open access data collections. The IPD-Work consortium is a network of investigators established at a meeting in London in 2008 to pool cohort studies with data on work life and health. Individual participant data from 15 cohort studies with measurements of working hours and alcohol consumption were available from IPD-Work. Another source of unpublished data was from open access collections. These were retrieved from collections curated by the Inter-University Consortium for Political and Social Research (ICPSR; www.icpsr.umich.edu/icpsrweb/ICPSR/) and the UK Economic and Social Data Service (http://ukdataservice.ac.uk/). These repositories seek research data and pertinent documents from researchers, research agencies, and government entities and process, preserve, and disseminate the data and documents for researchers. We identified 12 additional studies with data on working hours and alcohol use from these two sources (see appendix 1 for a more detailed description). Each constituent study was approved by the relevant local or national ethics committee, and all participants gave informed consent to participate.

In line with previous evidence suggesting ≥55 hours of work a week to be harmful for health,23 and the European Union Directive of a recommended limit of 48 hours a week,24 we used the following categories of hours worked: <35, 35-40 (reference), 41-48, 49-54, and ≥55 hours. In the Helsinki Health Study,25,26 weekly working hours were categorised as <30, 30-40, 41-50, and ≥50 in the response format. In our analyses, we used the ≥50 working hours group in the Helsinki Health Study to represent the ≥55 hours group and the 30-40 hours group as the reference group.

In all individual participant data studies, information on alcohol use was based on survey responses. We defined risky alcohol use by standard categorisations.27 One drink was considered equivalent to one unit or one glass of alcoholic drink.

Participants were classified as follows: non-use or moderate use (up to 14 drinks/week for women and up to 21 drinks/week for men), and risky alcohol use (>14 drinks/week and ≥21 drinks/week, respectively). According to public health guidelines, these limits define risky alcohol use among men and women in many countries.28,29 In the German Socioeconomic Panel Survey cohort,28 participants were asked how often they drank beer, wine, spirits, and mixed drinks (0=never, 1=seldom, 2=once in a while, 3=regularly); and risky alcohol use was defined as score of 7 or more on the summed score of these four questions. In the US National Longitudinal Survey of Youth 1979 cohort,29 risky alcohol use was defined as “binge drinking”—that is, having consumed six or more alcohol units in one evening at least twice during the past month. In the prospective cohort analyses of new onset risky alcohol use, we excluded cases of risky alcohol use at baseline.

Statistical analysis

We used a two stage meta-analysis27 with study specific estimates obtained from published studies with aggregate data and studies with unpublished individual participant data at the first stage. The second stage was an analysis that pooled all the study specific estimates.

First stage analysis in studies with aggregate data

We converted means, correlation coefficients, and regression coefficients reported for continuous variables to odds ratios according to published formulas.27 The studies with aggregate data had used different covariate adjustments. In two published studies, the exposure measure was a continuous variable estimating one hour increase in working hours with a dichotomous alcohol use outcome.28 29 In these cases, we raised the estimate to the 18th power to get an estimate comparable with that in the individual participant data (≥55 hours v 35-40 hours—that is, about an 18 hour difference).

First stage analysis in studies with individual participant data

We analysed data from the individual participant data studies to generate study level odds ratios and 95% confidence intervals within each study. These analyses were adjusted for age, sex, race/ethnicity (white/ethnic majority v other when available), and socioeconomic status, which was categorised as high, intermediate, and low. In the analyses of individual participant data, this three level socioeconomic status variable was based
on register or survey based occupational position or educational qualification.

Second stage analysis

In both types of data, we used random effects meta-analysis to combine the results from the individual participant data analyses and the (converted) estimates of the aggregate data. We used odds ratios as an indicator of effect size and their 95% confidence intervals as an indicator of precision. We examined heterogeneity of the estimates by computing an I² statistic, presented the summary estimates of the random effects analyses, and assessed the possibility of publication bias in the analysis of published studies by using a funnel plot of the estimates against their standard errors and Egger’s test for small study effects.

We used cross sectional aggregate and individual participant data for prespecified meta-regression and stratified meta-analyses to examine sex (men only, women only, both), region (Europe, North America, Australia/New Zealand, Asia), number of adjustment variables (none, sociodemographics, multivariable), and population (random sample of the population, occupational cohort with several occupations, occupational cohort with a single occupation (except physicians), physicians, patients as sources of heterogeneity between studies.

In the prospective analysis the corresponding variables were sex, age (<50 v ≥50), socioeconomic status, region (North America, United Kingdom, other Europe, Australia), sample type (population based/community sample v occupational cohort), the baseline level of alcohol use in the cohort (<9% v ≥9%), and participation rate in the prospective cohort (<75% v ≥75%). We used SAS 9.4 (Cary, NC) and Stata 13.1 (StataCorp, College Station, TX) to analyse study specific data and Stata 13.1 for meta-analyses.

Results

Aggregate data from published studies

The literature search identified 3409 citations from PubMed and 2314 from Embase, of which 651 were selected for further review (fig 1i). Of these, 26 fulfilled the eligibility criteria. We identified 10 additional eligible studies from a manual search of the references of the retrieved relevant publications (including reviews and books) and the cited reference search of the articles that were selected for meta-analysis. In total, there were 36 eligible published studies with aggregate data, of which 34 were cross sectional and two were cohort studies.

The total number of participants included in the meta-analysis of published studies with aggregate data was 139 112 in the 34 cross sectional analyses and 6873 in the two prospective analyses (see table A in appendix 2). There were 80 345 men and 61 983 women; in one study (n=3657) sex distribution was not reported. The study entry year/publication year ranged from 1970 to 2013. Only two published studies were prospective. 28 30 Of all published studies, 13 were from Japan, 31-41 seven from the United States, 31-40 four from Canada, 29 30 51-53 three from Australia, 34 35 54 two from the UK, 33 35 57 and one each from New Zealand, 36 Denmark, 38 Taiwan, 40 France, 38 Spain, 39 and Sweden. 36 One study included participants from both Australia and New Zealand. 38 Altogether 11 studies were general population based; 30 31 33 34 41 43 52 53-56, 60 physicians (six studies), 31 32 33 34 41 42 industrial workers (four), 31 32 41 42 office workers (three), 34 37 42 public sector employees (two), 38 39 nurses (two), 40 41 transport employees (two), 40 41 patients (two), 39 40 information technology engineers (one), 39 police officers (one), 39 managers (one), 39 and university graduates (one). 39 In 15 studies 28 32 37 40-42 43 50 55 59 62 the association between long working hours and alcohol use was not the main topic of the study—that is, either long working hours or alcohol use were covariates, and unadjusted estimates were usually displayed in a descriptive table.

There was large variation in the assessment of working hours in published studies with aggregate data: 20 studies used a threshold of 45 hours a week or higher (the highest threshold being >80 hours/week) while three studies used >40 hours/31 32 52 and one study used “frequent overtime” (see table A in appendix 2). Of these, 10 studies 28 31 32 41 46 48 54 59 60 62 used standard hours (such as 35-40 hours/week) as a reference group—most often all hours below the exposure group were included in the reference group. In the 12 remaining studies, working hours were treated as a continuous variable.26 30 31 32 42 44 50 51 53 56 57 61

With the exception of two studies, 28 29 the studies that treated working hours as a continuous variable also treated alcohol use measure as a continuous variable and reported a correlation or linear regression coefficient between the exposure and outcome (see table A in appendix 2). Two studies calculated the difference in mean alcohol use between groups of working long and shorter hours. 30 62 Probably in part because of large variation in the definition, the prevalence of working long hours also varied substantially, ranging from 4.1% to 69.8%. Alcohol use was based on self reported frequency from questionnaires or amount of alcohol use, except in one study that used the Composite International Diagnostic Interview (CIDI) to determine DSM-IV diagnoses of alcohol use disorders. 36 There was also large variation in the definition of risky alcohol use; and its prevalence between studies ranged from 3.4% to 64.0%.

Individual participant data

The 27 studies with individual participant data (fig 1j) and table B in appendix 2 drew on populations from the US (eight studies), 26 64-69 Germany (five), 25 70-71 Denmark (four), 27-29 the UK (three), 72-74 Finland (three), 75 76 Sweden (two), 77-78 Australia (one), and Belgium (one). 79-80 Eighteen studies were population based or community studies, 26 27 66-68 70-74 76 77 79 80 five were occupational cohorts, 24 75 78 80-82 84 and two were samples of college graduates and their siblings. 24 Of the studies, 26 provided hourly data on daily or weekly working hours, and in one study the number of hours was pre-categorised. 24

The cross sectional data from the 27 individual participant data studies came from 194 581 participants (89 799 men, 104 782 women). The proportion of employees working long hours (≥55/week) at baseline ranged from 0.8% to 20.2% across the studies (overall 7.5%). The prevalence of risky alcohol use at baseline (n=22 857 cases) varied from 2.3% to 38.8%, and the overall prevalence was 11.7%. Eighteen prospective cohorts provided data from a second wave, which allowed us to calculate new onset risky alcohol use. The prospective data included 93 729 participants (36 342 men, 57 387 women), the follow-up time ranged from 3.4 to 21.9 years (mean 6.0 person years), and the proportion of participants with new onset risky alcohol use (n=5878 new cases) ranged from 1.2% to 17.4%; overall new onset was 6.3%.

Cross sectional association between working hours and alcohol use

The overall pooled odds ratio of the association between long working hours and alcohol use in the meta-analysis of 61 cross sectional studies was 1.11 (95% confidence interval 1.05 to 1.18). Similar associations were found in the 27 individual participant data studies (1.10, 1.04 to 1.18; fig 2j) and published
studies with aggregate data (1.12, 1.02 to 1.22; fig 3). The $I^2$ statistics suggested significant heterogeneity among published studies with aggregate data ($I^2=81\%$, $P<0.001$) but not individual participant data ($I^2=18\%$, $P=0.21$). The total pooled $I$ was 72%, with $P<0.001$. In the individual participant data, the overall prevalence of risky alcohol use was 11.8%.

We examined the potential sources for heterogeneity in the aggregate data: sex (men only, women only, both), region (Europe, North America, Australia/New Zealand, Asia), covariate adjustments (none, sociodemographics, multivariate), and population (random sample of the population, occupational cohort with several occupations, occupational cohort with a single occupation (excluding physicians), physicians, patients). Meta-regression models provided no evidence that any of these factors explained the heterogeneity (data not shown; all $P>0.11$).

To examine whether a stricter definition of the exposure and outcome would provide different results in the cross sectional studies with aggregate data, we carried out a meta-analysis in a refined subgroup of studies that fulfilled the following criteria: the exposure and outcome were categorical; the cut point defining long hours was high enough (that is, it did not include employees with working hours close to standard hours, such as 41 hours); the reference group did not include employees with long working hours (such as 50 hours); the outcome was based on a specific definition of risky alcohol use. Six studies were eligible for this analysis. The odds ratio for risky alcohol use was 1.14 (95% confidence interval 0.89 to 1.47) and the $I^2$ was 68.6% ($P=0.007$; data not shown), again suggesting a similar estimate but with considerably high heterogeneity among published studies.

The funnel plot for published cross sectional studies appears symmetrical (fig 4) and is not dependent on whether long working hours and alcohol consumption was the main research question in the study or not. No evidence of an association between study size and the estimates was found in Egger’s test: 0.12 (95% confidence interval −0.38 to 0.63), $P=0.62$ for the total data; 0.10 (−0.71 to 0.90), $P=0.81$ for studies with long hours and alcohol use as the main research question; 0.22 (−0.58 to 1.02), $P=0.57$ for studies with another main research question.

**Prospsective association between long working hours and new onset risky alcohol use**

In the meta-analysis of the 18 prospective individual participant data studies and two studies with published aggregate data, participants who used alcohol above the recommended limits at baseline were excluded. The adjusted pooled odds ratio for new onset risky alcohol use was 1.12 (95% confidence interval 1.04 to 1.20) for those working long hours (≥55 in the individual participant data studies), compared with those working standard hours (35-40) (fig 5). The associations were similar in the individual participant data (1.12, 1.01 to 1.25) and two studies with published aggregate data (1.10, 0.66 to 1.84). The $I^2$ statistic for heterogeneity was 0% ($P=0.96$) overall; 0% ($P=0.99$) in the individual participant data studies and 56% ($P=0.10$) in the aggregate data.

We further examined whether the prospective associations were similar in men and women, in older (≥50) and younger (<50) employees, among socioeconomic groups, in different geographical areas, in different study samples, and different levels of risky drinking (<9.0% $v$ ≥9.0%) in the cohort (fig 6). The analyses of sex, age, and socioeconomic group were carried out in the individual participant data studies; for geographical region and type of study sample, all prospective studies were included and for the level of drinking, one study with aggregate data$^2$ provided data. Although the odds ratio among women seemed greater (1.36, 95% confidence interval 1.10 to 1.68) than among men (1.11, 0.98 to 1.25), the difference was not significant ($P=0.09$). The associations were not dependent on age or socioeconomic group, region, sample type, or the level of risky drinking in the cohort (all $P>0.09$ for heterogeneity).

To examine whether the association was dependent on sample attrition, we performed the analyses in subgroups with follow-up participation rate of <75% $v$ ≥75%. The corresponding odds ratio in the former group was 1.18 (1.02 to 1.36) and 1.10 (1.01 to 1.20) in the latter, with $P=0.44$ for heterogeneity (data not shown).

In the individual participant data studies, we examined the association between each working hours category (<35, 41-48, 49-54, and ≥55) in comparison with 35-40 hours a week associated with risky alcohol use. In the prospective analysis that excluded those who were risky users at baseline (fig 7), the overall incidence of risky alcohol use was 6.3% and the incidence among those working 35-40 hours was 6.2%. A part time job (<35 hours) and working 41-48 hours were not associated with new onset risky alcohol use whereas working 49-54 hours was associated with an odds ratio of 1.13 (95% confidence interval 1.02 to 1.26; adjusted difference in incidence 0.8 percentage points) and working 55 hours or more was associated with an odds ratio of 1.12 (1.01 to 1.25; adjusted difference in incidence 0.7 percentage points).

In the corresponding cross sectional analysis the adjusted odds ratio of risky alcohol use was 0.94 (95% confidence interval 0.88 to 1.02) for <35 hours, 1.03 (0.97 to 1.09) for 41-48 hours, 1.13 (1.05 to 1.22) for 49-54 hours, and 1.10 (1.04 to 1.18) for ≥55 hours when compared with standard 35-40 weekly working hours (data not shown). The adjusted difference in prevalence was 1.5 percentage points among those with 49-54 weekly working hours and 1.1 percentage points among those with 55 hours or more, when compared with the prevalence among employees working 35-40 hours (11.3%).

Finally, we used the individual participant data studies to examine whether long working hours were associated with sustained risky alcohol use among those who were risky users at baseline (data not shown). The odds ratio for sustained risky alcohol use among risky users who worked ≥55 hours at baseline was 1.16 (95% confidence interval 0.83 to 1.61) compared with those working standard 35-40 hours a week, suggesting a similar though non-significant association to that found for new onset risky use among non-risky users at baseline.

**Discussion**

This systematic review and meta-analysis of 63 studies provides comprehensive evidence of an association between long working hours and alcohol use. Long working hours were associated with 1.11-fold likelihood of higher levels of alcohol use in the cross sectional analysis of 333 693 participants from 14 countries. A similar association (odds ratio 1.12) for new onset risky alcohol use was observed in a prospective analysis of 20 studies including 100 602 participants from nine countries. The associations were not dependent on participants’ sex, age, socioeconomic status, geographical region, type of study sample, prevalence of risky alcohol use in the cohort, or participation rate at follow-up. Eighteen prospective studies with individual participant data allowed us to perform a more refined comparison between work hour categories. Compared with the standard (35-40) weekly working hours, working 49-54 hours was associated with an odds ratio of 1.13 and working ≥55
hours a week was associated with an odds ratio of 1.12 for new onset risky alcohol use.

**Strengths and limitations**

Our systematic review and meta-analysis provides the first quantitative summary estimate of the association between long working hours and alcohol use. This has not been provided in previous narrative reviews.\(^{11-13}\) \(^{18}\) We found only two published studies with a prospective design. Our meta-analysis, however, included 18 prospective cohort studies with unpublished individual participant data, which allowed us to examine the prospective association between long working hours and the onset of risky alcohol use in a large dataset with a harmonised exposure and outcome measurements.

Comparison of published studies with aggregate data and studies with unpublished individual participant data allowed us to account for publication bias, which can bias results from meta-analyses that use only published studies.\(^{17}\) We observed no evidence for publication bias. Firstly, we found similar associations among both types of study. Secondly, statistical assessment of publication bias among published studies, stratified by whether long working hours and alcohol use was the main or secondary research question, and use of a funnel plot and Egger’s test of a small study effect suggested no indication of publication bias.

The present meta-analysis has limitations. With the exception of three published studies that used roster or payroll data,\(^{29-31}\)\(^{58-59}\) exposure to long working hours was based on self reports and measured only once. Prolonged exposure to long working hours, as identified in repeat data, might carry a greater risk of risky alcohol use than long working hours measured only once. Self reported working hours might involve recall bias arising from the inability of the employees to accurately recall hours worked. Some investigators, however, have shown that self reported work hours is a reliable measure (for example, two week test-retest intraclass correlation coefficient r=0.91).\(^{45-61}\) and a major bias in longitudinal analysis is likely only if the error in estimating the number of hours worked varies according to alcohol use.

Alcohol use was based on self reported frequency or amount of alcohol used, except in one cross sectional study that used the Composite International Diagnostic Interview (CIDI) to assess DSM-IV diagnoses of alcohol use disorders.\(^{58}\) Under-reporting of alcohol use has been shown to occur in survey questionnaires\(^{47}\) and might be particularly common in heavy drinkers. A major bias would occur if under-reporting was also associated with reporting of working hours. Misclassification of alcohol use could have introduced heterogeneity into the meta-analyses.

Furthermore, this meta-analysis was based on observational studies, which are generalisable only within the study context. Despite cultural variations in drinking habits, even between developed countries,\(^{1}\) our findings regarding long working hours and alcohol use were similar across the different geographical regions examined. Our results are also in accordance with global reports of alcohol use, which have shown large differences in alcohol consumption between countries.\(^{1}\) We found, for example, that the prevalence of risky alcohol use was lower in the US than in Europe. Although the prevalence of risky alcohol use varied substantially among the cohorts, we found no evidence to suggest that the association between long working hours and alcohol use would be dependent on the prevalence of risky alcohol use in the cohort.

Even though our IPD-Work collaboration includes several major cohort studies in Europe and the two open access databases are large, including cohorts from the US, UK, and Australia, the meta-analysis of individual participant data is based on a convenience sample and does not include all potential data. Thus, availability bias might have affected the findings.\(^{19}\) Finally, it is not possible to draw causal inference based on observational studies, which means that our data cannot provide direct evidence for policy recommendations.\(^{86}\) Prospective analyses, however, allowed us to get some indication of temporal order of associations, which is consistent with a causal effect, and partially overcomes the problem of reverse causality (risky alcohol use leading to longer working hours), which can bias cross sectional analyses.

**Comparison with previous research and sources of heterogeneity**

Previous narrative reviews of long working hours and alcohol use including up to six studies have reported mixed findings.\(^{11-13}\)\(^{18}\) Similarly, we observed substantial heterogeneity among published studies with aggregate data (I\(^2\)=81%, P<0.001). Importantly, this heterogeneity was not present in the meta-analysis of individual participant data studies with harmonised exposure and outcome measures (I\(^2\)=18%, P=0.21 in the cross-sectional analysis; 0%, P=0.99 in the prospective analysis). None of the potential sources of heterogeneity examined in this study explained the heterogeneity found among the cross sectional studies with aggregate data. We explored whether use of a definition of the exposure and outcome close to that used in the individual participant data studies would remove the heterogeneity observed in the six published studies but found no such reduction (I\(^2\)=68.6%, P=0.007). It seems that the sources of heterogeneity must relate to other differences between published studies with aggregate data.

**Interpretation of findings**

In this meta-analysis, working very long hours (≥55 a week) was associated with a 1.12-fold increase in the odds ratio of new onset risky alcohol use among employees whose alcohol use was within the recommended limits at baseline. A similar (1.13-fold) association was found among those who worked 49-54 hours a week. To protect employee health and safety, the European Union Working Time Directive requires EU countries to guarantee all workers the right to work a maximum of 48 hours a week, on average, including overtime.\(^{27}\) This threshold is supported by our analyses. At risk alcohol use has previously been associated with increased risk of liver cirrhosis, cancers, seizure disorders, and stroke.\(^{2}\) A recent meta-analysis of 56 epidemiological studies using Mendelian randomisation analysis suggested that reduction of alcohol consumption, even for light to moderate drinkers, might be beneficial for cardiovascular health.\(^{65}\)

Possible explanations for the association between long working hours and risky alcohol use might involve the work environment as well as individual characteristics.\(^{40}\) One view is that alcohol use alleviates stress that is caused by work pressure and working conditions. Working overtime and characteristics such as high demands and lack of control might contribute to stress at work.\(^{39-40}\) Although our recent individual participant data meta-analysis found little evidence to support a combination of high demands and low job control as a predictor of risky alcohol use,\(^{7}\) it remains to be investigated whether a combination of long working hours and low control over work increases the risk. It has also been hypothesised that employees who are not
well integrated and supervised by the work organisation are at a higher risk of developing alcohol use problems. Depression and sleep problems might contribute to the link between long working hours and alcohol use. Other factors, such as personality traits, could confound the association because they might make people work long hours and also be related to a tendency to risky alcohol use. For example, individuals with "type A" behaviour pattern, which is characterised by aggressiveness and irritability and a chronic struggle to achieve more and more in less time, can end up in highly competitive jobs and work communities characterised by a "work hard play hard" culture with positive attitudes towards heavy alcohol use. Personality traits of impulsivity and novelty seeking have also been found to be associated with risky alcohol use.

Although the legislation in European countries provides employees with the right to restrict maximum working hours to 48 hours a week, many people still work long hours either as their own choice or because of external pressure to do so. Choosing to work long hours is common among the higher socioeconomic groups, with highly educated managers and professionals expecting dedication to work from themselves and others. This strategy can be highly rewarding in terms of rapid promotions, salary increases, and control over work. On the other hand, involuntary long hours could reflect, for example, financial difficulties and striving for a living wage or coercion. We did not find differences in the association between long working hours and risky alcohol use between socioeconomic groups. It is possible that voluntary and involuntary overtime is more evenly distributed across socioeconomic groups than previously thought.

Conclusions and unanswered questions

The workplace is an important setting for the prevention of alcohol misuse because more than half of the adult population are employed. Risky alcohol use is an issue in the workplace because it can have adverse and serious effects on employees, such as absenteeism, inefficiency, poor performance, impaired decision making, damaged customer relations, and injuries at work. Brief interventions by healthcare personnel that involve assessment of the individual’s drinking habits in combination with feedback about health risks might be useful in efforts to reduce problems with alcohol use in the working population. Our findings suggest that alcohol consumption is more likely to rise to risky levels among employees who work more than 48 hours a week compared with those with standard working hours. In absolute terms, however, the difference between these groups was relatively small because the adjusted incidence in new onset risky alcohol use was only 0.8 and 0.7 percentage points higher among individuals who worked 49-54 and 55 hours or more compared with those who worked standard hours. Further research is needed to assess whether preventive interventions against risky alcohol use could benefit from information on working hours.

Contributors: MV and MK developed the hypothesis and study design and supervised this study. MJ, MV, SN, JM, and TL performed statistical analyses. All authors contributed to study concept and design, analysis and interpretation of data, and drafting or critical revision of the manuscript for important intellectual content, or in addition, data acquisition. MV and KA performed the literature searches and/or screened papers. MK, TT, RR, and ND obtained funding for the IPD-Work Consortium. MJ, SN, and MK had full access to individual participant data in the study and take responsibility for the integrity of the unpublished data and the accuracy of the data analysis. MV is guarantor.

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Competing interests: All authors have completed the ICMJE uniform disclosure form at www.icmje.org/coi_disclosure.pdf and declare: no support from any organisation for the submitted work other than those mentioned above; no financial relationships with any organisations that might have an interest in the submitted work in the previous three years; no other relationships or activities that could appear to have influenced the submitted work.

Ethical approval: Each constituent study with individual participant data was approved by the relevant local or national ethics committee, and all participants gave informed consent to participate.

Data sharing: No additional data from our meta-analysis are available. Data for ACL, ALAMEDA, BCS, HELDA, MIDUS, NCDS, NHANES-I, NLSS, NSFH, SOEP, WLSG, and WLSS studies are publicly available to researchers. In addition, most constituent studies of the IPD-Work consortium have an established data sharing policy.

Transparency: MV affirms that this manuscript is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted; and that any discrepancies from the study as planned (and, if relevant, registered) have been explained.

18 Caruso CC. Possible broad impacts of long working hours. Ind Health 2006;44:531-6.
What is already known on this topic

The European Union Work Time Directive threshold for long hours is 48 hours a week.

Long working hours are thought to increase the risk of cardiovascular diseases, sleep deprivation, occupational injuries, and mental health problems.

The association between long working hours and risky alcohol use has not been systematically evaluated.

What this study adds

This meta-analysis of 36 published studies, identified by systematic literature searches, and 27 studies with unpublished individual participant data synthesised evidence on the association between long working hours and alcohol use.

Long working hours were linked to increased alcohol use in cross-sectional and prospective analyses. People working more than 48 hours a week at baseline had 1.13 times higher likelihood of new cases of risky alcohol use compared to those working standard hours.

No heterogeneity was observed between men and women or by age group, socioeconomic status, geographical region, or characteristics of the study cohort.


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Figures

![Diagram of search strategy and selection of studies for meta-analysis of effect of long working hours on alcohol use](image)

**Fig 1** Search strategy and selection of studies for meta-analysis of effect of long working hours on alcohol use
Fig 2 Cross sectional associations between long working hours and alcohol use from individual participant data adjusted for sex, age, socioeconomic status, and ethnicity
### Fig 3 Cross sectional associations between long working hours and alcohol use from published data

<table>
<thead>
<tr>
<th>Study</th>
<th>Odds ratio (95% CI)</th>
<th>Odds ratio (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Au (women)</td>
<td>1.29 (1.16 to 2.72)</td>
<td>1.46 (1.23 to 2.16)</td>
</tr>
<tr>
<td>Baldwin</td>
<td>0.39 (0.21 to 0.73)</td>
<td>0.20 (0.10 to 0.41)</td>
</tr>
<tr>
<td>Razanan</td>
<td>0.40 (0.20 to 1.11)</td>
<td>1.15 (0.92 to 1.42)</td>
</tr>
<tr>
<td>Etoe (men)</td>
<td>0.39 (0.21 to 0.73)</td>
<td>0.20 (0.10 to 0.41)</td>
</tr>
<tr>
<td>Etoe (women)</td>
<td>1.15 (0.92 to 1.42)</td>
<td>1.16 (0.99 to 1.35)</td>
</tr>
<tr>
<td>Gibb5</td>
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<td>1.73 (1.05 to 2.89)</td>
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<td>Gu2</td>
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<tr>
<td>Hasegawa (men)44</td>
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<td>1.21 (0.93 to 1.57)</td>
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<td>0.62 (0.43 to 0.90)</td>
<td>0.72 (0.58 to 0.90)</td>
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<td>Holtermann (men)59</td>
<td>1.63 (1.31 to 2.03)</td>
<td>0.93 (0.60 to 1.46)</td>
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<td>Jones (men)57</td>
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<td>0.84 (0.60 to 1.19)</td>
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<td>Jones (women)57</td>
<td>0.91 (0.27 to 3.00)</td>
<td>0.80 (0.49 to 1.34)</td>
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<td>Kawakami (men)31</td>
<td>1.04 (1.00 to 1.07)</td>
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<td>Kawakami (women)31</td>
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<td>1.05 (0.78 to 1.42)</td>
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<tr>
<td>Kobayashi (men)41</td>
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<td>0.67 (0.45 to 0.99)</td>
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<td>Lallukka (men)38</td>
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<td>1.08 (0.44 to 2.69)</td>
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<tr>
<td>Lallukka (women)38</td>
<td>3.25 (2.51 to 4.22)</td>
<td>1.51 (1.05 to 2.14)</td>
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<tr>
<td>Liu (men)36</td>
<td>1.15 (0.93 to 1.43)</td>
<td>1.00 (0.85 to 1.17)</td>
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<td>Marchand19</td>
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<td>2.89 (1.34 to 6.23)</td>
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<td>Marchand11</td>
<td>2.89 (1.34 to 6.23)</td>
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<tr>
<td>Nakata42</td>
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<td>0.64 (0.42 to 0.98)</td>
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<td>Nash24</td>
<td>1.90 (1.36 to 2.64)</td>
<td>1.35 (1.08 to 1.68)</td>
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<td>1.09 (0.82 to 1.48)</td>
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<td>1.12 (1.02 to 1.22)</td>
<td>1.00 (0.85 to 1.17)</td>
</tr>
</tbody>
</table>

Pooled: P<0.001, I²=81%
**Fig 4** Funnel plot with pseudo 95% confidence limits stratified by whether long working hours and alcohol use was main or secondary research question

**Fig 5** Associations between long working hours and new onset risky alcohol use (individual participant data adjusted for sex, age, socioeconomic status, and ethnicity at baseline)
**Fig 6** Pooled associations between long working hours and new onset risky alcohol use by sex, age group, socioeconomic status, geographic region, sample type, and prevalence of risky alcohol use in cohort.

<table>
<thead>
<tr>
<th>Pooled estimates</th>
<th>Odds ratio (95% CI)</th>
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<tr>
<td><strong>Sex</strong></td>
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<tr>
<td>Men</td>
<td>1.11 (0.98 to 1.25)</td>
</tr>
<tr>
<td>Women</td>
<td>1.36 (1.10 to 1.68)</td>
</tr>
<tr>
<td><strong>Age</strong></td>
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<tr>
<td>&lt;50 years</td>
<td>1.12 (0.99 to 1.27)</td>
</tr>
<tr>
<td>&gt;50 years</td>
<td>1.18 (0.95 to 1.47)</td>
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<tr>
<td><strong>Socioeconomic status (SES)</strong></td>
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<tr>
<td>Low SES</td>
<td>1.09 (0.88 to 1.35)</td>
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<tr>
<td>Intermediate SES</td>
<td>1.22 (0.97 to 1.53)</td>
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<tr>
<td>High SES</td>
<td>1.16 (0.99 to 1.36)</td>
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<tr>
<td><strong>Region</strong></td>
<td></td>
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<tr>
<td>North America</td>
<td>1.11 (1.01 to 1.23)</td>
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<tr>
<td>United Kingdom</td>
<td>1.18 (0.99 to 1.40)</td>
</tr>
<tr>
<td>Other Europe</td>
<td>1.06 (0.90 to 1.26)</td>
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<tr>
<td>Australia</td>
<td>1.19 (0.78 to 1.82)</td>
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<tr>
<td><strong>Sample</strong></td>
<td></td>
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<tr>
<td>Community sample</td>
<td>1.11 (1.02 to 1.21)</td>
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<tr>
<td>Occupational cohort</td>
<td>1.14 (0.97 to 1.34)</td>
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<tr>
<td><strong>Prevalence of alcohol use</strong></td>
<td></td>
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<tr>
<td>Low (&lt;9%)</td>
<td>1.11 (1.00 to 1.22)</td>
</tr>
<tr>
<td>High (&gt;9%)</td>
<td>1.13 (1.00 to 1.28)</td>
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</tbody>
</table>

**Fig 7** Pooled association between weekly working hours and new onset risky alcohol use, adjusted for sex, age, socioeconomic status, and ethnicity at baseline.

<table>
<thead>
<tr>
<th>Weekly working hours</th>
<th>Odds ratio (95% CI)</th>
<th>Odds ratio (95% CI)</th>
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<tr>
<td>&lt;35</td>
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<td>Reference</td>
</tr>
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<td>35-40</td>
<td>1.04 (0.95 to 1.14)</td>
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<td>41-48</td>
<td>1.13 (1.02 to 1.26)</td>
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<tr>
<td>49-54</td>
<td>1.12 (1.01 to 1.25)</td>
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</table>