

A Service of



Leibniz-Informationszentrum Wirtschaft Leibniz Information Centre

Görlitz, Katja; Tamm, Marcus

Working Paper
Parenthood and risk preferences

Diskussionsbeiträge, No. 2015/9

Provided in Cooperation with:

Free University Berlin, School of Business & Economics

Suggested Citation: Görlitz, Katja; Tamm, Marcus (2015): Parenthood and risk preferences, Diskussionsbeiträge, No. 2015/9, Freie Universität Berlin, Fachbereich Wirtschaftswissenschaft, Berlin

This Version is available at: https://hdl.handle.net/10419/108876

Standard-Nutzungsbedingungen:

Die Dokumente auf EconStor dürfen zu eigenen wissenschaftlichen Zwecken und zum Privatgebrauch gespeichert und kopiert werden.

Sie dürfen die Dokumente nicht für öffentliche oder kommerzielle Zwecke vervielfältigen, öffentlich ausstellen, öffentlich zugänglich machen, vertreiben oder anderweitig nutzen.

Sofern die Verfasser die Dokumente unter Open-Content-Lizenzen (insbesondere CC-Lizenzen) zur Verfügung gestellt haben sollten, gelten abweichend von diesen Nutzungsbedingungen die in der dort genannten Lizenz gewährten Nutzungsrechte.

Terms of use:

Documents in EconStor may be saved and copied for your personal and scholarly purposes.

You are not to copy documents for public or commercial purposes, to exhibit the documents publicly, to make them publicly available on the internet, or to distribute or otherwise use the documents in public.

If the documents have been made available under an Open Content Licence (especially Creative Commons Licences), you may exercise further usage rights as specified in the indicated licence.





Parenthood and risk preferences

Katja Görlitz Marcus Tamm

School of Business & Economics

Discussion Paper

Economics

2015/9

Parenthood and risk preferences

Katia Görlitz^a and Marcus Tamm^{b,*}

^a FU Berlin, RWI, and IZA

FU Berlin, Boltzmannstr. 20, 14195 Berlin, Germany,

katja.goerlitz@fu-berlin.de, Phone: +49-30-838-51242

^b RWI. and IZA

RWI, Invalidenstr. 112, 10115 Berlin, Germany,

tamm@rwi-essen.de, Phone: +49-30-2021598-18

Abstract. This study analyzes how risk attitudes change when individuals become parents

using longitudinal data for a large and representative sample of individuals. The results show

that men and women experience a considerable increase in risk aversion which already starts

as early as two years before becoming a parent, is largest shortly after giving birth and

disappears when the child becomes older. These findings show that parenthood leads to

considerable changes in individual risk attitudes over time. Thus, analyses using risk

preferences as the explanatory variable for economic outcomes should be careful in

interpreting the findings as causal effects.

JEL classification: D1, D81, J13, J16.

Keywords: Risk aversion, risk preferences, preference stability, parenthood, children, gender

differences

* We acknowledge helpful comments and suggestions by Frank Fossen and Stefanie Schurer.

Corresponding author: Marcus Tamm.

1. Introduction

Risk preferences are important determinants for almost every economically relevant decision of individuals such as financial investments, consumption, employment and occupational choice. While standard economic models assume risk preferences to be stable over time, the empirical literature has found that they do change as a consequence of negative macro shocks such as natural disasters (Hanaoka et al. 2014, Eckel et al. 2009), civil conflicts (Voors et al. 2012) or financial crises (Guiso et al. 2013). Yet, little is known about the impact of individual-specific life events on preferences. One exception is Sahm (2012) who sheds light on the importance of job displacements and serious health diagnoses. In this paper, we analyze the family-related event of becoming a parent. While some authors have assumed that parenthood increases risk aversion (e.g. DeLeire and Levy 2004), to our knowledge, we are the first to show this empirically.

The analysis is based on representative longitudinal household data containing information on individuals' general risk attitudes in seven panel waves in addition to demographic and family-related information. Applying individual fixed-effects regressions, we investigate whether individuals' risk attitudes change as a consequence of parenthood and (if so) whether it is a permanent change or a transitory one that fades away when children grow old.

This paper contributes to the empirical literature that is concerned with identifying the extent to which risk preferences predict individuals' behavior, for example, with respect to health-related behavior, financial investments (Barsky et al. 1997, Dohmen et al. 2011), occupational choice (Bonin et al. 2007, Caliendo et al. 2009), marriage and childbearing (Schmidt 2008). If risk preferences change as a result of life events that are themselves correlated with the behavioral outcomes of interest, a regression of outcomes on risk preferences without accounting for their endogeneity would produce biased estimates. Parenthood is likely to be correlated with a variety of individual choices which makes this aspect important for a wide range of applications.

Furthermore, due to the different roles of men and women in parenthood, e.g., in giving birth and raising children, we test whether the impact of parenthood on risk attitudes differs by gender. This contributes to the literature concerned with explaining gender differences in economic choices. It could also help to explain the stylized fact that women appear to be more risk averse than men (e.g. Eckel and Grossman 2008). If risk aversion of women reacted more sensitive after childbirth than that of men, this could explain some of the gender gap in risk aversion.

2. Data

The analysis relies on data from the German Socio-Economic Panel (SOEP), a representative longitudinal survey of private households in Germany. SOEP includes a question on general risk attitudes in 2004, 2006 and from 2008 to 2012, leaving us a total of seven panel waves of data. The general risk attitude is measured by a survey question on individuals' self-assessed willingness to take risks. The English translation of the risk question reads: "How do you see yourself: Are you generally a person who is willing to take risks or do you try to avoid risks? Please answer on a scale from 0 to 10, where 0 means 'not at all willing to take risks' and 10 means 'very willing to take risks'." Dohmen et al. (2011) show, also using the SOEP data, that this question is behavioral valid and a reliable predictor of actual risk-taking behavior in a paid lottery-type experiment.

The estmation sample is restricted to individuals aged 17 to 64. It covers 29,000 individuals and 107,000 person-year observations. There is hardly any item nonresponse in the question of risk attitudes (0.4 percent) and in the other variables used in the empirical analysis. On average, each individual reports information on his or her risk attitude in three to four panel waves. 6,700 individuals answer the question on risk attitudes in seven waves, 3,000 individuals in six waves and 1,800 in five waves. Taken together these individuals who answer the risk question at least five times account for almost 70 percent of all person-year observations. Within the time period under consideration 1,243 respondents experienced the birth of their first child. This provides ample opportunities to observe individual changes in risk attitudes and parental status over time.

Table 1 displays the summary statistics on the willingness to take risks by gender and parental status. The average risk attitude is 4.68, with women reporting a lower willingness to take risks (4.27) than men (5.14). For both genders alike, childless individuals have a higher risk attitude than parents. When splitting the sample further by years since birth of the first child, risk aversion varies a lot for both genders in a non-systematic way. However, since the descriptive statistics do not account for parents' age that has been shown to correlate with risk aversion (Schurer 2015), a final conclusion can only be drawn from regression results.

_

¹ The data has been extracted using the Stata add-on-package PanelWhiz v4.0 (Oct 2012). Haisken-DeNew and Hahn (2010) document PanelWhiz.

Table 1 – Summary statistics of individuals' risk attitudes

	Women		Men		All					
	Mean	SD	Mean	SD	Mean	SD				
All	4.27	2.21	5.14	2.21	4.68	2.26				
Without children	4.67	2.19	5.33	2.21	5.04	2.23				
With children	4.08	2.20	4.98	2.20	4.46	2.25				
By years since birth of first child										
< 1 year	4.12	2.11	5.19	2.16	4.61	2.20				
1-3 years	4.33	2.15	5.25	2.10	4.76	2.17				
4-6 years	4.28	2.14	5.19	2.13	4.70	2.18				
7-9 years	4.13	2.17	5.17	2.19	4.61	2.24				
10-18 years	4.16	2.17	5.10	2.18	4.60	2.22				
> 18 years	3.99	2.22	4.83	2.23	4.33	2.26				
Observations	56,023		51,061		107,084					

3. Estimation and results

The aim of the regression analysis is to show how parenthood changes individuals' willingness to take risks. The following regression model is estimated:

$$y_{it} = B_{it}'\beta + X_{it}'\delta + \alpha_i + \varepsilon_{it}$$

where y is the risk attitude of individual i at time t (with t=2004, 2006, 2008-2012). The vector B comprises dummy variables that indicate the time after the birth of the first child: the first year after childbirth (b_0) and for each following year up to when the child is aged 6 (b_1 to b_6). From 6 years after childbirth onwards, the following group dummies are defined: 7 to 9 years, 10 to 12 years, 13 to 15 years, 16 to 18 years and 19 years and above. To look at changes in risk attitudes up to three years before childbirth, we also include the three dummy variables b_{-3} to b_{-1} . Any changes before childbirth might indicate that individuals first settle down before becoming parents and that settling down already influences risk preferences. The vector of control variables X comprises age in a cubic specification, year dummies and a binary indicator for the birth of a second child. The latter enables us distinguishing the effect of becoming a parent (by having a first child) from the effect of a subsequent birth. The coefficient α_i comprises individual fixed-effects.

The regression is estimated by applying the linear fixed-effects model. Furthermore, we apply the following three fixed-effects models that take the ordinal nature of the dependent variable into account: the "blow-up and cluster" (BUC) estimator (Baetschmann et al. 2014), the two-step minimum estimator (Das and van Soest 1999), henceforth DvS, and the estimator suggested by Ferrer-i-Carbonell and Frijters (2004), henceforth FCF.

The regression results are presented in Table 2. Since the results from the linear fixed-effects model (Table 2, columns 1 and 5) are basically the same as those from the other models (Table 2, columns 2 to 4 and 6 to 8), the following discussion concentrates on the linear fixed-effects results. For illustrative purposes, Figure 1 depicts the point estimates (solid lines) together with the 90% confidence interval (dashed lines). It can be seen that both women's and men's risk attitude is significantly lower after the birth of the first child and that both become more risk averse as early as two years before the birth of the first child.²

For women, Table 2 shows that the size of the decrease in risk attitudes is 0.6 points in the year after the first child was born.³ This is equivalent to a drop by more than a fourth of a standard deviation. Between b₁ and b₇₋₉, the effects are still significantly negative ranging from -0.35 to -0.53 points. Afterwards, risk attitudes rise continuously and 12 years after the birth of the first child the effect becomes statistically insignificant. Men also become more risk averse around the birth of the first child, but the point estimate is somewhat smaller for them $(-0.36 \text{ points in } b_0)$. The level of risk attitudes is lowest in the year when the child is one year old (-0.42 points) which is equivalent to a decrease by a fifth of a standard deviation. Five years after childbirth, the effects for men cannot be rejected to differ from zero on a significance level of 10 percent.⁴ Figure 1 suggests that the overall pattern of how risk attitudes change as a consequence of parenthood is somewhat flatter for men than for women. However, this suggestion is not supported by statistical tests: There are neither significant gender differences when testing each birth-related coefficient against one another nor are they significant when running a joint test of equality of all gender-specific birth-related coefficients. For the interpretation of the results, note that the impact of having a second child is insignificant which indicates that becoming a parent for the first time changes risk attitudes and not the birth of a second child.

_

² The effects before childbirth do not appear to be driven by marriage. In specifications controlling for marriage or for marital status the results remain unchanged for the birth-related coefficients and the marriage dummy is always insignificant.

³ Also note that for women the effect in b_0 differs significantly from the effect in b_{-1} (p-value 0.0157).

⁴ When using the 5 percent significance level, this would be the case three years after childbirth. Also, fewer of the point estimates for men would be considered to differ from zero. In contrast, for women conclusions are the same irrespective of using the 5 or 10 percent significance level.

Figure 1 – Birth of the first child and the evolution of risk attitudes by gender

Note: The figures display the estimated regression coefficients (solid lines) together with the 90% confidence intervals (dashed lines) from the linear fixed-effects models presented in Table 2.

4. Conclusion

The results show that risk attitudes are not stable across time. Both men and women become more risk averse after the birth of their first child. This change of risk attitudes starts as early as two years before giving birth and is most pronounced shortly after the birth of the first child. The effect becomes smaller and statistically insignificant when the child grows older. Overall, we find clear evidence that family-related events can change individual risk attitudes. Therefore, in empirical applications that analyze how risk preferences predict individual decision making, risk attitudes cannot be treated as exogenous. Another finding is that the relationship between time since childbirth and risk attitudes does not differ significantly by gender. Thus, we conclude that the consequences of parenthood do not contribute to explaining the empirical finding of a gender gap in risk attitudes.

References

Barsky, Robert, Thomas Juster, Miles Kimball and Matthew Shapiro (1997), Preference Parameters and Behavioral Heterogeneity: An Experimental Approach in the Health and Retirement Study, The Quarterly Journal of Economics 112(2), 537-579.

Baetschmann, Gregori, Kevin Straub and Rainer Winkelmann (2014), Consistent estimation of the fixed effects logit model, Journal of the Royal Statistical Society: Series A (Statistics in Society), doi: 10.1111/rssa.12090.

Bonin, Holger, Thomas Dohmen, Armin Falk, David Huffman and Uwe Sunde (2007), Cross-sectional earnings risk and occupational sorting: The role of risk attitudes, Labour Economics 14(6), 926-937.

Caliendo, Marco, Frank Fossen and Alexander Kritikos (2009): Risk Attitudes of Nascent Entrepreneurs - New Evidence from an Experimentally-Validated Survey, Small Business Economics, 32(2), 153-167.

Das, Marcel and Arthur van Soest (1999), A panel data model for subjective information on household income growth, Journal of Economic Behavior & Organization 40(4), 409-426.

DeLeire, Thomas and Helen Levy (2004), Worker Sorting and the Risk of Death on the Job, Journal of Labor Economics 22(4), 925-953.

Dohmen, Thomas, Armin Falk, David Huffman, Uwe Sunde, Jürgen Schupp and Gert Wagner (2011), Individual Risk Attitudes: Measurement, Determinants and Behavioral Consequences, Journal of the European Economic Association 9(3), 522-550.

Eckel, Catherine, Mahmoud El Gamalb and Rick Wilson (2009), Risk Loving After the Storm: A Bayesian-Network Study of Hurricane Katrina Evacuees, Journal of Economic Behavior & Organization 69(2), 110-124.

Eckel, Catherine and Philip Grossman (2008), Men, Women and Risk Aversion: Experimental Evidence. In: Plott, Charles and Vernon Smith (eds.), Handbook of Experimental Economics, Vol. 1, 1061-1073. Amsterdam et al.: Elsevier.

Ferrer-i-Carbonell, Ada and Paul Frijters (2004), How important is methodology for the estimates of the determinants of happiness? The Economic Journal 114(497), 641-659.

Guiso, Luigi, Paola Sapienza and Luigi Zingales (2013), Time varying risk aversion, NBER working paper 19284.

Haisken-DeNew, John and Markus Hahn (2010), PanelWhiz: Efficient Data Extraction of Complex Panel Data Sets – An Example Using the German SOEP, Journal of Applied Social Science Studies 130(4), 643-654.

Hanaoka, Chie, Hitoshi Shigeoka and Yasutora Watanabe (2014), Do Risk Preferences Change? Evidence from Panel Data before and after the Great East Japan Earthquake. Available at SSRN: http://ssrn.com/abstract=2425396.

Sahm, Claudia (2012), How Much does Risk Tolerance Change? Quarterly Journal of Finance 2(4), 1250020.

Schmidt, Lucie (2008), Risk Preferences and the Timing of Marriage and Childbearing, Demography 45(2), 439-460.

Schurer, Stefanie (2015), Lifecycle Patterns in the Socioeconomic Gradient of Risk Preferences, IZA discussion paper 8821.

Voors, Maarten, Eleonora Nillesen, Philip Verwimp, Erwin Bulte, Robert Lensink and Daan van Soest (2012), Does Conflict Affect Preferences? Results from Field Experiments in Burundi, American Economic Review 102(2), 941-964.

Table 2 – The change of risk attitudes after becoming a parent

			_					
	Women				Men			
	Linear FE	BUC	DvS	FCF	Linear FE	BUC	DvS	FCF
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Birth of first child b ₋₃	-0.1086	-0.1309	-0.1072	-0.0990	-0.1501	-0.2391	-0.0732	-0.2702
	(0.1409)	(0.1815)	(0.1666)	(0.1703)	(0.1458)	(0.1960)	(0.1738)	(0.1908)
Birth of first child b ₋₂	-0.2542**	-0.3131**	-0.2080	-0.2763*	-0.3204**	-0.4842***	-0.5404***	-0.4588***
	(0.1232)	(0.1539)	(0.1457)	(0.1519)	(0.1321)	(0.1750)	(0.1589)	(0.1691)
Birth of first child b ₋₁	-0.3519***	-0.4371***	-0.4204***	-0.4773***	-0.2393*	-0.3461**	-0.2905*	-0.2365
	(0.1265)	(0.1610)	(0.1495)	(0.1574)	(0.1249)	(0.1722)	(0.1540)	(0.1709)
Birth of first child b ₀	-0.6042***	-0.7790***	-0.7081***	-0.6884***	-0.3623***	-0.5220***	-0.4263***	-0.4126**
·	(0.1214)	(0.1544)	(0.1462)	(0.1463)	(0.1246)	(0.1672)	(0.1473)	(0.1649)
Birth of first child b ₁	-0.4193***	-0.5389***	-0.4928***	-0.4251***	-0.4226***	-0.5646***	-0.4783***	-0.5865***
•	(0.1300)	(0.1644)	(0.1554)	(0.1605)	(0.1263)	(0.1709)	(0.1555)	(0.1703)
Birth of first child b ₂	-0.4324***	-0.5684***	-0.5049***	-0.4919***	-0.2487*	-0.3818**	-0.3007*	-0.4616***
	(0.1356)	(0.1724)	(0.1620)	(0.1675)	(0.1345)	(0.1795)	(0.1619)	(0.1742)
Birth of first child b ₃	-0.4546***	-0.5828***	-0.5671***	-0.4871***	-0.3678**	-0.5115***	-0.5166***	-0.4674**
	(0.1454)	(0.1834)	(0.1735)	(0.1735)	(0.1439)	(0.1891)	(0.1725)	(0.1844)
Birth of first child b ₄	-0.4414***	-0.5617***	-0.5018***	-0.4589**	-0.2521*	-0.3893*	-0.3522*	-0.4525**
·	(0.1522)	(0.1895)	(0.1791)	(0.1843)	(0.1527)	(0.1990)	(0.1797)	(0.1962)
Birth of first child b ₅	-0.5334***	-0.6631***	-0.6085***	-0.6098***	-0.3000*	-0.4356**	-0.4413**	-0.4771**
3	(0.1581)	(0.1972)	(0.1884)	(0.1935)	(0.1605)	(0.2089)	(0.1914)	(0.2009)
Birth of first child b ₆	-0.3512**	-0.4439**	-0.4093**	-0.3535*	-0.2480	-0.3563*	-0.3104	-0.5510***
v	(0.1632)	(0.2023)	(0.1927)	(0.1972)	(0.1643)	(0.2140)	(0.1970)	(0.2090)
Birth of first child b ₇₋₉	-0.4459***	-0.5704***	-0.5321***	-0.4833**	-0.2718	-0.3990*	-0.3680*	-0.4970**
	(0.1664)	(0.2071)	(0.1976)	(0.1992)	(0.1680)	(0.2181)	(0.2017)	(0.2130)
Birth of first child b ₁₀₋₁₂	-0.3755**	-0.4826**	-0.4035*	-0.3854*	-0.2903	-0.4214*	-0.3944*	-0.4955**
	(0.1804)	(0.2242)	(0.2147)	(0.2163)	(0.1832)	(0.2374)	(0.2211)	(0.2304)
Birth of first child b ₁₃₋₁₅	-0.2494	-0.3178	-0.2268	-0.3297	-0.1820	-0.2709	-0.2321	-0.4803*
	(0.1941)	(0.2408)	(0.2316)	(0.2318)	(0.1955)	(0.2530)	(0.2365)	(0.2464)
Birth of first child b ₁₆₋₁₈	-0.1056	-0.1348	-0.0573	-0.1480	-0.1801	-0.2630	-0.2225	-0.4185
	(0.2063)	(0.2553)	(0.2460)	(0.2464)	(0.2088)	(0.2684)	(0.2523)	(0.2617)
Birth of first child $b_{>18}$	-0.0762	-0.0958	-0.0177	-0.1521	-0.1636	-0.2423	-0.1844	-0.4341
	(0.2202)	(0.2721)	(0.2626)	(0.2627)	(0.2221)	(0.2844)	(0.2684)	(0.2779)
Birth of second child $b_{\geq 0}$	0.1106	0.1218	0.1416	0.0656	0.0036	0.0118	0.1026	0.1738
	(0.0870)	(0.1046)	(0.0969)	(0.1034)	(0.0934)	(0.1156)	(0.1105)	(0.1104)
Age and year effects	yes							
Observations	56,023	183,116	49,978	49,978	51,061	161,716	45,272	45,272

Note: The dependent variable is the self-reported risk attitude measured on a scale from 0 (extremely risk averse) to 10 (very willing to take risks). Estimates account for clustering at the person level.

Diskussionsbeiträge - Fachbereich Wirtschaftswissenschaft - Freie Universität Berlin Discussion Paper - School of Business and Economics - Freie Universität Berlin

2015 erschienen:

2015/1 GÖRLITZ, Katja und Christina GRAVERT

The effects of increasing the standards of the high school curriculum on school

dropout *Economics*

2015/2 BÖNKE, Timm und Clive WERDT

Charitable giving and its persistent and transitory reactions to changes in tax

incentives: evidence from the German Taxpayer Panel

Economics

2015/3 WERDT, Clive

What drives tax refund maximization from inter-temporal loss usage? Evidence from

the German Taxpayer Panel

Economics

2015/4 FOSSEN, Frank M. und Johannes KÖNIG

Public health insurance and entry into self-employment

Economics

2015/5 WERDT, Clive

The elasticity of taxable income for Germany and its sensitivity to the appropriate

model *Economics*

2015/6 NIKODINOSKA, Dragana und Carsten SCHRÖDER

On the Emissions-Inequality Trade-off in Energy Taxation: Evidence on the German

Car Fuel Tax *Economics*

2015/7 GROß, Marcus; Ulrich RENDTEL; Timo SCHMID; Sebastian SCHMON und Nikos

TZAVIDIS

Estimating the density of ethnic minorities and aged people in Berlin: Multivariate kernel density estimation applied to sensitive geo-referenced administrative data

protected via measurement error

Economics

2015/8 SCHMID, Timo; Nikos TZAVIDIS; Ralf MÜNNICH und Ray CHAMBERS

Outlier robust small area estimation under spatial correlation

Economics

2015/9 GÖRLITZ, Katja und Marcus TAMM

Parenthood and risk preferences

Economics