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## The effect of educational mismatch on wages for 25 countries

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KÖZGAZDASÁGTUDOMÁNYI INTÉZET

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**BUDAPEST WORKING PAPERS ON THE LABOUR MARKET**  
**BWP – 2008/8**

**The effect of educational mismatch on  
wages for 25 countries**

PÉTER GALASI

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The effect of educational mismatch on wages for 25 countries

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# **The effect of educational mismatch on wages for 25 countries<sup>1</sup>**

Péter Galasi

## **Abstract**

By making use of the Duncan&Hoffman model, the paper estimates returns to educational mismatch using comparable microdata for 25 European countries. Our aim is to investigate the extent to which the main empirical regularities produced by other papers on the subject are confirmed by our data base. On the basis of tests proposed by Hartog&Oosterbeek, we also consider whether the observed empirical patterns are in line with the Mincerian basic human capital model and Thurow's job competition model. Using Heckman's sample-selection estimator, we find that results are rather consistent with those found in the literature, and that both the job-competition model, and the Mincerian human capital model could be rejected for most of the countries.

JEL: J21, J23, J24, J31

**Keywords:** job-education matching, overeducation, undereducation, returns to over- and undereducation, international comparison

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# **A túl- és az alulképzés bérhozama 25 országban**

Galasi Péter

## Összefoglaló

A tanulmányban 25 európai ország, kétezres évek közepi állapotot tükröző, reprezentatív keresztmetszeti mintáin egyrészt a Duncan – Hoffman modellre támaszkodva megvizsgáljuk, hogy adatbázisunk milyen mértékben tükrözi az illeszkedés bérhozamával foglalkozó irodalom legfontosabb empirikus következtetéseit, másrészt - a Hartog - Oosterbeek által javasolt statisztikai próbák segítségével – azt elemezzük, hogy empirikus modellünk eredményei alapján mit mondhatunk az emberi tőke minceri alapmodelljének és a thurowi állásverseny modelljének érvényességéről. A becsléseket Heckman szelekciós torzítást kiküszöbölő becslőfüggvényével végeztük, s azt találtuk, hogy eredményeink többnyire egybeesnek az irodalomban található empirikus szabályosságokkal. A statisztikai próbák az országok többségében cáfolják mind a minceri, mind a thurowi modell empirikus érvényességét.

**Tárgyszavak:** a munka és az iskolai végzettség illeszkedése, túlképzés, alulképzés, bérhozam, nemzetközi összehasonlítás

## 1. INTRODUCTION

Recently, a growing number of studies<sup>2</sup> have been concerned with the labour market consequences of inadequate education. The paper estimates returns to educational mismatch using comparable microdata for 25 European countries. It complements the literature<sup>3</sup> on the effect of underschooling and overschooling on wages. First, using Duncan&Hoffman (1981) model we try to investigate to what extent the main empirical findings of the literature are valid for our samples. Bauer (2002), drawing on Hartog (2000), summarises the main empirical results as follows: the returns to attained years of schooling are lower than the returns to required years of schooling (Result 1 –  $R1$ ); the returns to overeducation are positive but smaller than the wage premium associated with a year of required education ( $R2$ ); the returns to years of undereducation are negative ( $R3$ ), but their absolute value is smaller than the returns to required education ( $R4$ ), and smaller than the returns to overeducation ( $R5$ ); the estimated returns to overeducation are always significant ( $R6$ ), whereas those to undereducation are not so ( $R7$ ). Second, applying the statistical tests proposed by Hartog&Oosterbeek (1988), we also consider whether the observed empirical patterns are consistent with the standard Mincerian human capital model (Mincer 1974) or Thurow's job-competition model (Thurow 1975).

The paper is divided into five sections. Section 2 presents the Duncan&Hoffman model and the Hartog&Oosterbeek tests, Section 3 provides information on data, definitions, variables, and estimation techniques, Section 4 reports the main empirical findings, Section 5 concludes.

## 2. THE MODEL

The Duncan&Hoffman model decomposes attained years of education ( $S$ ) into years of education required on a job ( $R$ ), years of over- ( $O$ ) and underschooling ( $U$ ):  $S = R + O - U$ . This implies  $S = R$  for an adequately educated individual,  $S = R + O$  ( $O > 0$ ) if the worker is overeducated, and  $S = R - U$  ( $U > 0$ ) in the case of undereducation. For a sample of individuals and using linearised specification, the earnings equation estimating the (constant) average returns to a year of required, over- and undereducation can be written as

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<sup>2</sup> Theoretical and measurement problems are summarised in Hartog (2000), Green&McIntosh&Vignoles (1999), van der Velden&van Smoorenburg (1997), Borghans&de Grip (1999). Groot&Maassen van den Brink (2000). Hartog (2000) and Rubb (2003b) report many empirical results.

<sup>3</sup> For example Chevalier (2003), Cohn&Khan (1995), Cohn&Ng (2000), Daly&Büchel&Duncan (2000), Dolton&Vignoles (2000), Groot (1996), Mendes de Oliveira&Santos&Kiker (2000), Rubb (2003a), Vahey (2000)).

$$w = \alpha_0 + \alpha_R R + \alpha_O O + \alpha_U U ,$$

where  $w$  denotes log earnings,  $\alpha_R = \partial w / \partial R$ ,  $\alpha_O = \partial w / \partial O$ , and  $\alpha_U = \partial w / \partial U$  stand for the returns to educational (mis)match to be estimated. If the expected estimation results hold then  $\alpha_O > 0$  and  $\alpha_R > \alpha_O$  (R2);  $\alpha_U < 0$  (R3);  $\alpha_R > |\alpha_U|$  (R4);  $\alpha_O > |\alpha_U|$  (R5); and  $\alpha_R > \alpha_O > |\alpha_U|$  (R2, R4, R5).

According to Hartog&Oosterbeek, since in a Mincerian setting the educational requirements of the job do not affect wages, then the equality of the following parameters should hold:  $\alpha_R = \alpha_O = |\alpha_U|$ ; as for Thurow's model, wages are determined by the marginal-productivity requirements of the job, individuals' assets are irrelevant, thus the estimation should produce  $\alpha_O = |\alpha_U|$ . Note that if the Mincerian model is accepted, R2 and/or R4 and/or R5 have to be rejected, if Thurow's model seems to be valid, then R5 cannot hold.

### 3. DATA, VARIABLES, ESTIMATORS

We use European Social Survey data<sup>4</sup> (Jowel et al. 2005).<sup>5</sup> The data were collected in 2004 to 2006<sup>6</sup> for a large number of countries using a common questionnaire. The overall sample used in the paper consists of 25 European countries out of which 20 and 2 are members of the EU,<sup>7</sup> and the European Economic Area,<sup>8</sup> respectively. The remaining countries are Switzerland, Turkey, and Ukraine. The size of the pooled sample, for which earnings equations can be estimated - that is, that the number of persons with non-zero earnings - is about 13500. Individual countries' sample size varies between 200 and 900 that might adversely affect the precision of estimations, especially in the case of smaller samples.<sup>9</sup>

The first key variable we use here is the usual before-tax monthly wage (in Euro). The original variable refers to different time periods for different countries. As regards the number of persons with non-zero wage in the pooled sample, about 74 per cent have monthly earnings data. For 13, 8, and 5 per cent of the respondents the periods are a year, four weeks, and one or two weeks, respectively. Thus in the case of 26 per cent of the observations the

<sup>4</sup> ESS round 2; the data file used in the paper was released on 12-12-2006.

<sup>5</sup> The data archive and distributor of ESS data is the *Norwegian Social Science Data Services*; for information see <http://ess.nsd.uib.no>.

<sup>6</sup> 62 per cent of interviews were done in 2004, 34 per cent in 2005, four per cent in 2006.

<sup>7</sup> Austria, Belgium, Czech Republic, Denmark, Estonia, Finland, France, Germany, Great Britain, Greece, Hungary, Ireland, Luxembourg, Netherlands, Poland, Portugal, Slovakia, Slovenia, Spain, Sweden.

<sup>8</sup> Iceland and Norway.

<sup>9</sup> For sample sizes see Table A1.

original variable was recoded.<sup>10</sup> The average worker of the pooled sample earns 1218 euros (standard deviation: 3.53) and the country averages range from 47 (Ukraine) to 3300 euros (Denmark).<sup>11</sup>

Schooling is measured as years of full-time education completed (highest level).<sup>12</sup> Since the survey provides no information about part-time education, it is likely that education attainment is underestimated and as a consequence returns to education will be overestimated.

Educational mismatch depends on how education required on a job is measured. Out of the three standard methods (job-analyst method, subjective method, and the method based on realised matches),<sup>13</sup> here we can use the second one that relies on self-reporting. Two variables are available to produce the required education variable. The exact wording of these questions is as follows: 1. if someone was applying nowadays for the job you do now, would they need any education or vocational schooling beyond compulsory education?, and 2. about how many years of education or vocational schooling beyond compulsory education would they need? The second variable is not fully continuous and it is truncated from above. The truncation does not seem serious since only four per cent of workers report the highest number of years (at least 10 years of required education beyond compulsory school). In case of interval coding we use interval midpoints.<sup>14</sup> Required years of education are reported in relation to compulsory education, so we need to know the years of compulsory education. They are available from the UNESCO statistical data base and the information refers to 2004.<sup>15</sup> The years of required education variable is produced with the help of the two original required education variables and the UNESCO data. Then, years of under and overeducation are computed using the required and attained education variables.

Table 1 presents the incidence of educational mismatch. The proportion of properly educated workers amounts to 8 per cent for the pooled sample varying between one (Turkey) and 19 per cent (Austria); when the samples are pooled 33 per cent of workers are overeducated, and this ranges from 15 (Netherlands) to 79 per cent (Estonia). On average, the number of the undereducated is 59 per cent with important cross-country differences; Estonia represents the lowest value (13 per cent), and Netherlands do the highest one (82 per cent).

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<sup>10</sup> Weekly and bi-weekly earnings were multiplied by 4.4, and 2.2, respectively. Wages referring to four weeks were multiplied by 1.075, annual earnings were divided by 12.

<sup>11</sup> For means and standard deviations see Table A1.

<sup>12</sup> Means and standard deviations are reported in Table 2.

<sup>13</sup> see Hartog (2000), Green&McIntosh&Vignoles (1999), van der Velden&van Smoorenburg (1997), Borghans&de Grip (1999).

<sup>14</sup> The variable is coded as: less than 1 year (beyond compulsory school); about 1 year; about 2 years; about 3 years; about 4-5 years; about 6-7 years; about 8-9 years; 10 years or more (beyond compulsory school).

<sup>15</sup> See Table A1 in the Appendix



Table 1

**The number of properly, over- and undereducated workers (per cent)**

| Country           | Properly<br>educated | Over- | Under- | Together |
|-------------------|----------------------|-------|--------|----------|
| Austria           | 18.8                 | 46.7  | 34.5   | 100      |
| Belgium           | 12.6                 | 25.1  | 62.3   | 100      |
| Czech<br>Republic | 6.2                  | 49.5  | 44.3   | 100      |
| Denmark           | 8.4                  | 52.6  | 39.1   | 100      |
| Estonia           | 8.2                  | 78.9  | 12.9   | 100      |
| Finland           | 8.4                  | 52.6  | 39.0   | 100      |
| France            | 8.0                  | 26.6  | 65.3   | 100      |
| Germany           | 9.1                  | 19.5  | 71.3   | 100      |
| Great Britain     | 9.2                  | 28.2  | 62.6   | 100      |
| Greece            | 4.8                  | 77.1  | 18.2   | 100      |
| Hungary           | 5.2                  | 31.1  | 63.7   | 100      |
| Iceland           | 4.7                  | 47.7  | 47.7   | 100      |
| Ireland           | 10.7                 | 67.4  | 21.9   | 100      |
| Luxembourg        | 5.6                  | 45.1  | 49.3   | 100      |
| Netherlands       | 3.3                  | 14.7  | 82.0   | 100      |
| Norway            | 7.6                  | 41.3  | 51.2   | 100      |
| Poland            | 4.1                  | 59.1  | 36.8   | 100      |
| Portugal          | 17.5                 | 33.3  | 49.2   | 100      |
| Slovakia          | 15.1                 | 46.7  | 38.2   | 100      |
| Slovenia          | 4.8                  | 17.5  | 77.7   | 100      |
| Spain             | 5.6                  | 50.2  | 44.3   | 100      |
| Sweden            | 8.9                  | 40.1  | 51.0   | 100      |
| Switzerland       | 10.0                 | 22.4  | 67.6   | 100      |
| Turkey            | 1.4                  | 27.9  | 70.8   | 100      |
| Ukraine           | 11.5                 | 24.1  | 64.4   | 100      |
| Pooled sample     | 8.0                  | 32.9  | 59.1   | 100      |
| N                 | 13488                |       |        |          |

The number of properly educated is low, and that of undereducated is high when compared to the results of other surveys. As regards studies based on the self-reporting method, Hartog (2000), reviewing the empirical results of the literature, finds that the lowest value for the properly educated is 47 per cent, and the highest one for the over- and undereducated are 33 and 32 per cent, respectively. In Groot&Maassen van den Brink (2000)'s paper focussing on meta-analysis the upper limit for overeducation incidence is 42 and that for undereducation amounts to 20 per cent in the case of studies applying the subjective method.

Years of attained, required, over-, and undereducation, and their standard deviations are shown in Table 2.

Table 2

**Years of attained, required, over-, and undereducation**

| Country        | Attained |      | Required<br>education |      | Over- |      | Under- |      |
|----------------|----------|------|-----------------------|------|-------|------|--------|------|
|                | mean     | s.d. | mean                  | s.d. | mean  | s.d. | mean   | s.d. |
| Austria        | 12.6     | 2.9  | 12.2                  | 2.8  | 2.4   | 2.0  | 1.5    | 1.2  |
| Belgium        | 13.5     | 3.5  | 14.9                  | 2.0  | 2.2   | 1.7  | 3.1    | 2.7  |
| Czech Republic | 12.8     | 2.3  | 12.6                  | 2.5  | 2.1   | 1.2  | 1.6    | 0.9  |
| Denmark        | 14.5     | 3.2  | 13.9                  | 3.1  | 2.7   | 2.1  | 2.0    | 1.9  |
| Estonia        | 13.1     | 2.9  | 10.9                  | 2.2  | 3.1   | 2.1  | 1.6    | 1.0  |
| Finland        | 14.1     | 3.6  | 13.6                  | 2.7  | 2.7   | 1.9  | 2.2    | 1.5  |
| France         | 12.9     | 3.8  | 14.5                  | 2.8  | 2.1   | 1.7  | 3.3    | 2.4  |
| Germany        | 13.7     | 3.1  | 15.7                  | 2.5  | 2.3   | 1.7  | 3.1    | 1.7  |
| Great Britain  | 12.9     | 3.0  | 13.7                  | 2.4  | 2.6   | 1.7  | 2.4    | 1.8  |
| Greece         | 12.8     | 3.7  | 10.5                  | 2.1  | 3.7   | 2.1  | 2.9    | 1.4  |
| Hungary        | 12.8     | 2.6  | 14.2                  | 3.1  | 1.6   | 1.0  | 2.7    | 1.6  |
| Iceland        | 14.3     | 4.0  | 14.9                  | 3.6  | 2.8   | 1.9  | 3.5    | 3.3  |
| Ireland        | 13.4     | 3.2  | 12.3                  | 2.5  | 3.0   | 1.8  | 2.4    | 2.1  |
| Luxembourg     | 12.2     | 4.5  | 12.7                  | 3.2  | 2.9   | 2.1  | 3.4    | 2.7  |
| Netherlands    | 13.6     | 3.4  | 16.6                  | 2.5  | 1.9   | 1.7  | 3.9    | 2.4  |
| Norway         | 14.1     | 3.4  | 14.4                  | 2.7  | 2.4   | 1.8  | 2.5    | 1.8  |
| Poland         | 12.9     | 2.9  | 12.0                  | 2.9  | 2.8   | 1.9  | 1.7    | 1.3  |
| Portugal       | 9.2      | 4.4  | 10.5                  | 2.7  | 3.0   | 1.7  | 3.6    | 1.7  |
| Slovakia       | 12.9     | 2.7  | 12.4                  | 2.9  | 2.9   | 2.0  | 1.9    | 2.0  |
| Slovenia       | 12.5     | 3.3  | 14.3                  | 2.7  | 1.8   | 1.4  | 2.7    | 1.9  |
| Spain          | 13.3     | 5.1  | 12.9                  | 2.8  | 4.3   | 3.4  | 3.4    | 2.6  |
| Sweden         | 13.2     | 3.1  | 13.5                  | 2.7  | 2.3   | 1.5  | 2.2    | 1.6  |
| Switzerland    | 10.9     | 3.3  | 12.9                  | 2.8  | 2.0   | 1.6  | 3.6    | 2.1  |
| Turkey         | 9.6      | 3.9  | 11.5                  | 3.4  | 2.5   | 1.4  | 3.7    | 2.5  |
| Ukraine        | 12.4     | 2.4  | 13.5                  | 2.0  | 2.1   | 1.3  | 2.2    | 1.4  |
| Pooled sample  | 13.0     | 3.5  | 13.8                  | 3.0  | 2.7   | 2.1  | 2.9    | 2.1  |
| N              | 13488    |      | 13488                 |      | 4443  |      | 7971   |      |

Years of over- and undereducation were computed for the over- and undereducated

The average worker in the pooled sample has completed 13 years of schooling. The data exhibit a great deal of cross-country heterogeneity, however. Education attained is less than 10 years in Portugal and Turkey, and it is higher than 14 years in four Scandinavian countries (Denmark, Finland, Iceland, and Norway). On average, required education exceeds attained education by 0,8 years; for ten countries, however, the opposite holds, the former is lower than the latter.<sup>16</sup> For the typical Dutch worker required education amounts to 16.6 years, whereas the value of the same indicator seems much lower for the Greek worker (10.5 years). The average overeducated worker of the pooled sample has 2.7 years of surplus schooling. The values range from 1.6 (Hungary) to 4.9 (Spain) years for individual countries. As regards years of underschooling, the data exhibit 2.9 years of undereducation and country means fall within the interval of 1.5 (Austria) to 3.9 (Netherlands) years.

<sup>16</sup> Austria, Czech Republic, Denmark, Estonia, Finland, Greece, Ireland, Poland, Spain, Slovakia

Testing *R1* requires estimating earnings functions so as to obtain parameter estimates for the attained education variable. In addition, we need earnings equations to be estimated with educational mismatch variables formulated according the Duncan&Hofmann model. The equations are of the standard Mincerian type. The dependent variable is the log monthly wage, the key explanatory variables are years of education attained - for the earnings functions, and the three matching variables (years of required, over- and undereducation) as regards the earnings equations with the educational mismatch variables. Sex (female =1), and potential labour market experience and its square are also inserted in all the equations. When the samples are pooled a series of country dummies are present in the equations in order to control for cross-country differences (reference category: Austria). The equations are estimated by ols<sup>17</sup> and Heckman's (1979) selectivity-bias-corrected estimator.<sup>18</sup> As regards participation equations, estimated by probit, and needed for the Heckman model, control variables include a female dummy, age, age-squared, the number and the sex of dependent children in the family. Participation equations are estimated for ILO's able-bodied persons (aged 15-74).

#### 4. RESULTS

As for the earnings equations with the educational-mismatch variables, the sign of the estimated coefficients for the selection variable is negative (except for one equation), implying that wages would be overestimated by ols. In addition, for 17 countries and the pooled sample the estimation yields parameter estimates for the selection variable that are significant (at the  $p = 0.05$  level). Thus, the selectivity-bias assumption can be accepted.<sup>19</sup> Out of 78 estimated coefficients, the selectivity-bias-corrected estimates of the three educational-mismatch variables<sup>20</sup> produce 51 parameters that are significantly different from zero (at the  $p = 0.05$  level), and for 28 coefficient estimates, Heckman's estimator yields at least a 0.5 percentage-point high difference in returns to educational mismatch compared to ols estimates.<sup>21</sup> Therefore, selectivity bias is detected for 68 per cent of the countries, and using

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<sup>17</sup> Although the main empirical tool relating to the effect of educational mismatch on wages is still ols, some authors choose other estimators even for cross-sectional data. For example, Budria&Moro-Egido (2008), and McGuinness&Bennett (2007) use quantile regression; Guironnet&Peypoch (2007) and Jensen&Gartner&Rässler (2006) apply stochastic production frontier models. Note that a growing number of estimates are based on panel data (see Battu&Belfield&Sloane 1999, Bauer 2002, Daly&Büchel&Duncan 2000, Dolton&Silles 2008, Dolton&Vignoles 2000, Rubb 2006).

<sup>18</sup> The estimator has rarely been used in this literature. Exceptions are Sloane&Battu&Seaman (1999), Dolton&Vignoles (2000), and di Pietro&Urwin (2006).

<sup>19</sup> For parameter estimates of the selection correction variable and the independence of earnings and participation equations see Table A2.

<sup>20</sup> Estimation results are presented in Table 4.

<sup>21</sup> Ols estimates for the educational mismatch variables are reported in Table A3.

Heckman's estimator, as opposed to ols estimates, results in sensible differences in returns to educational mismatch for about one third of the estimated parameters.

The results of selectivity-bias-corrected estimations are mixed, and reporting practice is not unanimous in the literature. Sloane&Battu&Seaman (1999) arrive at non-significant parameter estimates for the selection-correction variable, but the authors report only estimates produced by the Heckman's estimator. Dolton&Vignoles (2000), in their study on higher-education graduates, estimate selectivity-bias-corrected regression for women only, and they conclude that while there is some sign of the presence of sample selection, the effect of self-selection on returns to educational mismatch is negligible; therefore they do not report results from selectivity-bias-corrected estimations. Di Pietro&Urwin (2006) find that the parameter estimates for the selection variable is significantly different from zero, but the coefficients estimated by ols and Heckman's estimator are not significantly different; thus parameters estimated by ols are not reported.

As regards our standard earnings equations, estimates for the selection variable are significant for the pooled sample and 15 countries, and they are mostly negative.<sup>22</sup> Using Heckman's<sup>23</sup> or ols<sup>24</sup> estimator result in only slight differences in returns to education for the majority of the countries; only one fifth of the estimated coefficients show at least 0.05 percentage point differences. Note that as for the returns to education variable, all parameter estimates produced by both estimators are significant (at the  $p = 0.05$  level).

Our ols estimates for returns to education are in general higher than those estimated by ols and reported in Trostel&Walker&Woolley (2002), and in Flabbi&Paternostro&Tiongson (2007). Putting aside difference in time periods for which the data are collected, this may be due partly to the specification of the independent variable. Both papers work with after-tax wage rates, whereas our measure of remuneration is before-taxes and monthly. Using before-tax earnings yields higher returns when there is a progressive income tax system (as might be the case for most of the countries in our database). Making use of monthly earnings would produce higher wage premium for the better educated than wage rates do, if higher levels of education are associated with longer working hours.

Most of the empirical regularities found in the literature seem to hold for our samples. As regards  $R1$ ,<sup>25</sup> the expected results are observed for 22 countries; the exceptions are the Czech Republic, Slovenia and Ukraine. Estimates for the pooled sample also show higher returns to required than to attained education; an extra year of required schooling yields 2.5 percentage point higher wages than an additional year of observed schooling (see Table 3).

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<sup>22</sup> For details, see Table A4.

<sup>23</sup> See Table A4.

<sup>24</sup> See Table A5.

<sup>25</sup> returns to attained education are lower than returns to required education

A brief inspection of coefficients in Table 4 also provides some support for the empirical results found elsewhere. As for  $R_2$ ,<sup>26</sup> returns to overeducation are positive but smaller than the wage premium associated with a year of required education for each country. Our estimations based on ESS data also confirm  $R_3$ <sup>27</sup> in the case of 23 countries – only the Irish and Slovene samples fail to exhibit the general empirical pattern, since the sign of their estimated parameters are positive.  $R_4$ <sup>28</sup> holds for all countries, meaning that returns to required education are always higher than the absolute value of those to undereducation.

Table 3

**Returns to required and attained education, per cent**

| Country        | Required education | Attained |
|----------------|--------------------|----------|
| Austria        | 7.5                | 6.2      |
| Belgium        | 9.4                | 3.5      |
| Czech Republic | 7.1                | 7.1      |
| Denmark        | 6.7                | 4.6      |
| Estonia        | 13.8               | 8.8      |
| Finland        | 8.7                | 5.4      |
| France         | 15.0               | 7.6      |
| Germany        | 11.2               | 7.6      |
| Great Britain  | 14.3               | 9.7      |
| Greece         | 9.3                | 3.4      |
| Hungary        | 13.5               | 12.5     |
| Iceland        | 7.6                | 4.0      |
| Ireland        | 8.7                | 6.7      |
| Luxembourg     | 10.0               | 5.7      |
| Netherlands    | 12.9               | 7.3      |
| Norway         | 6.9                | 5.3      |
| Poland         | 9.7                | 8.2      |
| Portugal       | 16.4               | 8.0      |
| Slovakia       | 9.1                | 6.4      |
| Slovenia       | 8.5                | 8.7      |
| Spain          | 8.6                | 5.1      |
| Sweden         | 7.4                | 5.8      |
| Switzerland    | 8.5                | 5.5      |
| Turkey         | 10.9               | 8.6      |
| Ukraine        | 6.2                | 8.6      |
| Pooled sample  | 9.7                | 7.2      |

Parameters estimated by Heckman's selection-correction estimator

All estimates are significant at least at the p=0.05 level

For returns to observed education see Table A4

For returns to required education see Table 4

<sup>26</sup> returns to overeducation are positive but smaller than returns to required education,  $\alpha_R > \alpha_O$

<sup>27</sup> returns to undereducation are negative,  $\alpha_U < 0$

<sup>28</sup> the absolute value of the returns to undereducation is smaller than returns to required education,  $\alpha_R > |\alpha_U|$

As regards  $R5$ ,<sup>29</sup> it holds for 16 countries. In order to obtain  $R6$ <sup>30</sup> we would need significant parameter estimates for the overeducation variable in each of the countries, but only more than a half of them exhibit this pattern. The estimated coefficient for the years of undereducation variable is not significant for 14 countries, thus  $R7$ <sup>31</sup> is also supported by the data. By taking a look at the parameters for the pooled sample we can conclude that  $R2$  to  $R6$  hold true, and that the parameter estimated for returns to undereducation significantly differs from zero (see Table 4).

We also have run Wald tests in order to check whether the expected results involving two coefficients hold or not; this concerns  $R2$ ,  $R4$ , and  $R5$ . The null hypothesis is that the two coefficients are equal (see Table 5).

As regards  $R2$ , the equality of the coefficients could not be rejected for four countries, but we can accept observed differences in the parameter estimates in the case of 16 countries and the pooled sample at the  $p = 0.05$  level, and five countries at the  $p = 0.1$  level.

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<sup>29</sup> the absolute value of the returns to undereducation is smaller than returns to overeducation,  $\alpha_O > |\alpha_U|$

<sup>30</sup> returns to overeducation are always significant

<sup>31</sup> returns to undereducation are not always significant

Table 4

**Returns to educational mismatch, per cent**

| Country        | Required education |      |       | Overeducation |     |      | Undereducation |      |       |
|----------------|--------------------|------|-------|---------------|-----|------|----------------|------|-------|
|                | Coefficient        | %    | z     | Coefficient   | %   | z    | Coefficient    | %    | z     |
| Austria        | 0.072              | 7.5  | 6.60  | 0.030         | 3.1 | 1.79 | -0.025         | -2.5 | -0.97 |
| Belgium        | 0.089              | 9.4  | 6.29  | 0.027         | 2.8 | 1.58 | -0.028         | -2.8 | -1.67 |
| Czech Republic | 0.069              | 7.1  | 5.78  | 0.065         | 6.7 | 3.38 | -0.002         | -0.2 | -0.07 |
| Denmark        | 0.065              | 6.7  | 6.41  | 0.015         | 1.5 | 1.17 | -0.034         | -3.4 | -1.96 |
| Estonia        | 0.129              | 13.8 | 13.25 | 0.044         | 4.5 | 3.72 | -0.022         | -2.2 | -0.74 |
| Finland        | 0.083              | 8.7  | 14.40 | 0.008         | 0.8 | 0.80 | -0.025         | -2.5 | -2.23 |
| France         | 0.139              | 15.0 | 12.77 | 0.035         | 3.5 | 1.41 | -0.078         | -7.5 | -6.93 |
| Germany        | 0.106              | 11.2 | 9.78  | 0.049         | 5.1 | 2.31 | -0.033         | -3.2 | -2.75 |
| Great Britain  | 0.134              | 14.3 | 5.20  | 0.013         | 1.3 | 0.54 | -0.025         | -2.4 | -1.16 |
| Greece         | 0.088              | 9.3  | 5.03  | 0.072         | 7.4 | 4.67 | -0.046         | -4.5 | -1.72 |
| Hungary        | 0.127              | 13.5 | 7.50  | 0.051         | 5.2 | 1.71 | -0.069         | -6.7 | -3.25 |
| Iceland        | 0.074              | 7.6  | 7.41  | 0.027         | 2.8 | 1.54 | -0.001         | -0.1 | -0.11 |
| Ireland        | 0.083              | 8.7  | 3.66  | 0.035         | 3.6 | 1.52 | 0.015          | 1.5  | 0.49  |
| Luxembourg     | 0.096              | 10.0 | 11.92 | 0.081         | 8.5 | 6.74 | -0.005         | -0.5 | -0.67 |
| Netherlands    | 0.121              | 12.9 | 11.58 | 0.008         | 0.8 | 0.34 | -0.034         | -3.3 | -3.36 |
| Norway         | 0.067              | 6.9  | 9.58  | 0.030         | 3.1 | 2.67 | -0.014         | -1.4 | -1.11 |
| Poland         | 0.092              | 9.7  | 8.91  | 0.061         | 6.3 | 4.75 | -0.006         | -0.6 | -0.28 |
| Portugal       | 0.152              | 16.4 | 15.72 | 0.066         | 6.8 | 4.43 | -0.028         | -2.8 | -2.00 |
| Slovakia       | 0.087              | 9.1  | 10.18 | 0.050         | 5.2 | 2.21 | -0.010         | -1.0 | -0.64 |
| Slovenia       | 0.082              | 8.5  | 4.24  | 0.014         | 1.4 | 0.32 | 0.002          | 0.2  | 0.06  |
| Spain          | 0.083              | 8.6  | 7.85  | 0.039         | 4.0 | 3.77 | -0.035         | -3.5 | -2.72 |
| Sweden         | 0.071              | 7.4  | 16.14 | 0.020         | 2.0 | 2.03 | -0.023         | -2.3 | -2.71 |
| Switzerland    | 0.081              | 8.5  | 8.95  | 0.030         | 3.1 | 1.70 | -0.026         | -2.5 | -2.32 |
| Turkey         | 0.103              | 10.9 | 4.15  | 0.042         | 4.3 | 1.23 | -0.058         | -5.6 | -2.26 |
| Ukraine        | 0.060              | 6.2  | 2.32  | 0.002         | 0.2 | 0.05 | -0.051         | -5.0 | -1.26 |
| Pooled sample  | 0.093              | 9.7  | 12.93 | 0.028         | 2.8 | 4.12 | -0.021         | -2.1 | -3.06 |

Parameters estimated by Heckman's selection-correction estimator.

The equations estimated

earnings equation

dependent variable: log usual monthly before-tax earnings

control variables: sex, potential labour market experience and its square; country dummies in the case of pooled sample (reference: Austria)

participation equation

dependent variable: worker

control variables: sex, age, age squared, number of children in the family, sex of children; country dummies in the case of pooled sample (reference: Austria)

We arrive at  $R_4$  for 21 countries and the pooled sample at the  $p=0.05$  level, and two countries at the  $p=0.1$  level. However, in the case of the remaining two countries, we cannot reject the hypothesis that the two estimated parameters are equal.

An important finding from the present estimations is that  $R_5$  does not seem to hold as a general rule. As we can see from Table 4, the absolute value of returns to underschooling is not smaller than returns to overschooling for 9 countries. In addition, the statistical tests displayed in Table 5 show that the equality of the two coefficients cannot be rejected for 21

countries and the pooled sample, implying that for an overwhelming majority of the cases  $R5$  does not hold. That is, the wage penalty associated with a year of undereducation is not lower than the wage premium due to a year of overeducation.

Let us turn to the Hartog&Oosterbeek (1988) tests elaborated for the Duncan&Hoffman (1981) model. The job competition model can be rejected, since the hypothesis that both coefficients are zero does not seem to hold for 17 countries and the pooled sample, although for the remaining 8 countries it can be accepted. Similar results are arrived at with regard to the standard Mincerian model. The equality of the three coefficients is rejected for 23 countries and the pooled sample, thus the Mincerian model can be rejected.

Table 5

### Wald tests for coefficients' equality

| Country        | $R2$  |       | $R4$   |       | $R5$  |       | Mincer |       | Thurow |       |
|----------------|-------|-------|--------|-------|-------|-------|--------|-------|--------|-------|
|                | chi2  | p     | chi2   | p     | chi2  | p     | chi2   | p     | ch2    | p     |
| Austria        | 5.38  | 0.020 | 3.68   | 0.055 | 0.02  | 0.882 | 11.41  | 0.003 | 5.04   | 0.081 |
| Belgium        | 7.20  | 0.073 | 16.35  | 0.000 | 0.00  | 0.985 | 22.10  | 0.000 | 6.88   | 0.032 |
| Czech Republic | 0.05  | 0.826 | 8.58   | 0.003 | 2.73  | 0.098 | 14.65  | 0.001 | 13.63  | 0.001 |
| Denmark        | 11.82 | 0.001 | 3.26   | 0.071 | 0.72  | 0.396 | 15.53  | 0.000 | 5.66   | 0.059 |
| Estonia        | 45.78 | 0.000 | 12.58  | 0.000 | 0.37  | 0.542 | 73.15  | 0.000 | 35.33  | 0.000 |
| Finland        | 54.00 | 0.000 | 28.70  | 0.000 | 1.21  | 0.271 | 102.49 | 0.000 | 6.69   | 0.035 |
| France         | 18.89 | 0.000 | 31.86  | 0.000 | 2.26  | 0.133 | 65.68  | 0.000 | 53.69  | 0.000 |
| Germany        | 6.20  | 0.013 | 38.95  | 0.000 | 0.39  | 0.530 | 47.55  | 0.000 | 15.63  | 0.000 |
| Great Britain  | 13.23 | 0.000 | 20.59  | 0.000 | 0.11  | 0.745 | 26.55  | 0.000 | 1.95   | 0.377 |
| Greece         | 0.68  | 0.411 | 2.45   | 0.118 | 0.56  | 0.454 | 3.19   | 0.203 | 32.94  | 0.000 |
| Hungary        | 6.23  | 0.013 | 12.26  | 0.001 | 0.22  | 0.637 | 26.99  | 0.000 | 15.25  | 0.001 |
| Iceland        | 6.56  | 0.010 | 27.20  | 0.000 | 1.18  | 0.278 | 37.05  | 0.000 | 2.57   | 0.276 |
| Ireland        | 3.31  | 0.069 | 24.51  | 0.000 | 1.96  | 0.161 | 32.82  | 0.000 | 2.82   | 0.244 |
| Luxembourg     | 1.97  | 0.160 | 80.01  | 0.000 | 25.84 | 0.000 | 84.32  | 0.000 | 47.69  | 0.000 |
| Netherlands    | 19.87 | 0.000 | 47.57  | 0.000 | 0.88  | 0.347 | 61.96  | 0.000 | 12.66  | 0.002 |
| Norway         | 8.15  | 0.004 | 22.65  | 0.000 | 0.77  | 0.381 | 40.43  | 0.000 | 10.66  | 0.005 |
| Poland         | 4.65  | 0.031 | 19.41  | 0.000 | 3.94  | 0.047 | 32.91  | 0.000 | 24.63  | 0.000 |
| Portugal       | 38.87 | 0.000 | 63.29  | 0.000 | 3.56  | 0.059 | 94.40  | 0.000 | 22.72  | 0.000 |
| Slovakia       | 3.68  | 0.055 | 23.82  | 0.000 | 1.91  | 0.167 | 42.02  | 0.000 | 6.06   | 0.048 |
| Slovenia       | 2.74  | 0.098 | 14.60  | 0.000 | 0.09  | 0.765 | 22.53  | 0.000 | 0.10   | 0.949 |
| Spain          | 11.36 | 0.001 | 6.25   | 0.012 | 0.04  | 0.846 | 13.92  | 0.001 | 0.37   | 0.542 |
| Sweden         | 24.94 | 0.000 | 30.48  | 0.000 | 0.05  | 0.824 | 72.95  | 0.000 | 17.22  | 0.000 |
| Switzerland    | 7.01  | 0.008 | 20.49  | 0.000 | 0.03  | 0.858 | 30.09  | 0.000 | 12.27  | 0.002 |
| Turkey         | 3.51  | 0.061 | 4.15   | 0.042 | 0.13  | 0.714 | 9.53   | 0.009 | 6.94   | 0.031 |
| Ukraine        | 1.90  | 0.169 | 0.06   | 0.799 | 0.67  | 0.413 | 2.27   | 0.322 | 1.64   | 0.441 |
| Pooled         | 81.34 | 0.000 | 106.00 | 0.000 | 0.43  | 0.514 | 178.63 | 0.000 | 26.14  | 0.000 |

$R2$  null:  $\alpha_R = \alpha_O$ ,

$R4$  null:  $\alpha_R = |\alpha_U|$ ,

$R5$  null:  $\alpha_O = |\alpha_U|$

Thurow null:  $\alpha_O = |\alpha_U| = 0$

Mincer null:  $\alpha_R = \alpha_O = |\alpha_U|$



Other studies draw mixed conclusions. Groot (1996), using ols estimates, rejects both models, and this is consistent with the findings of this paper. Bauer (2002), using panel models, examines the parameter estimates for men and women separately. He concludes that the Mincerian model can be accepted for females only, and the job-competition model can be rejected for both sexes.

## **5. CONCLUSION**

The paper investigates the wage effect of over- and undereducation using comparable microdata for 25 European countries. The results confirm most of the conclusions of the existing empirical literature. However, some earlier findings seem not to hold as a general empirical rule. Namely, the present estimations produce  $R1$ ,  $R2$ ,  $R3$ ,  $R4$ , and  $R7$ , whereas  $R5$  and  $R6$  are not confirmed in our database. We have also found that both the job-competition model, and the Mincerian human capital model can be rejected for most of the countries.

## APPENDIX

Table A1

### Sample characteristics

| Country        | No of workers | Monthly earnings |      |      |      | YCE |
|----------------|---------------|------------------|------|------|------|-----|
|                |               | log mean         | sd   | mean | sd   |     |
| Austria        | 415           | 7.5              | 0.51 | 1781 | 1.67 | 9   |
| Belgium        | 519           | 7.7              | 0.69 | 2208 | 1.99 | 13  |
| Czech Republic | 544           | 6.1              | 0.42 | 429  | 1.52 | 10  |
| Denmark        | 672           | 8.1              | 0.82 | 3316 | 2.26 | 10  |
| Estonia        | 722           | 5.8              | 0.64 | 322  | 1.90 | 9   |
| Finland        | 787           | 7.7              | 0.46 | 2148 | 1.59 | 10  |
| France         | 641           | 7.5              | 0.79 | 1834 | 2.20 | 11  |
| Germany        | 688           | 7.6              | 0.68 | 1962 | 1.98 | 13  |
| Great Britain  | 601           | 7.4              | 1.05 | 1683 | 2.85 | 12  |
| Greece         | 355           | 7.1              | 0.56 | 1252 | 1.75 | 9   |
| Hungary        | 440           | 6.1              | 0.55 | 447  | 1.73 | 10  |
| Iceland        | 239           | 8.1              | 0.58 | 3160 | 1.79 | 11  |
| Ireland        | 609           | 7.5              | 1.01 | 1773 | 2.75 | 10  |
| Luxembourg     | 438           | 7.8              | 0.67 | 2423 | 1.96 | 10  |
| Netherlands    | 607           | 7.7              | 0.68 | 2175 | 1.96 | 13  |
| Norway         | 910           | 8.0              | 0.58 | 3084 | 1.79 | 11  |
| Poland         | 481           | 5.9              | 0.58 | 359  | 1.78 | 9   |
| Portugal       | 265           | 6.5              | 0.55 | 685  | 1.74 | 9   |
| Slovakia       | 332           | 5.8              | 0.46 | 327  | 1.59 | 9   |
| Slovenia       | 433           | 5.3              | 0.51 | 202  | 1.66 | 10  |
| Spain          | 418           | 7.1              | 0.70 | 1229 | 2.02 | 11  |
| Sweden         | 923           | 7.8              | 0.39 | 2348 | 1.48 | 10  |
| Switzerland    | 689           | 8.0              | 0.67 | 3049 | 1.96 | 9   |
| Turkey         | 252           | 6.6              | 0.69 | 738  | 2.00 | 9   |
| Ukraine        | 508           | 3.9              | 0.85 | 47   | 2.34 | 12  |
| Pooled sample  | 13488         | 7.1              | 1.26 | 1218 | 3.53 |     |

YCE: Years of compulsory education

Source: UNESCO Institute for Statistics

<http://stats.uis.unesco.org/TableViewer/tableView.aspx?ReportId=210>

Table A2

**Parameter estimates for the selection-correction variable (lambda) and tests for independence of equations**

| Country        | Lambda | Wald test |        |
|----------------|--------|-----------|--------|
|                |        | chi2      | p      |
| Austria        | -0.428 | 16.98     | 0.0000 |
| Belgium        | 0.839  | 114.58    | 0.0000 |
| Czech Republic | -0.342 | 17.20     | 0.0000 |
| Denmark        | 0.046  | 0.80      | 0.3717 |
| Estonia        | -0.284 | 8.30      | 0.0040 |
| Finland        | -0.153 | 2.35      | 0.1249 |
| France         | 0.721  | 165.27    | 0.0000 |
| Germany        | -0.685 | 72.20     | 0.0000 |
| Great Britain  | -1.173 | 66.78     | 0.0000 |
| Greece         | 0.711  | 27.15     | 0.0000 |
| Hungary        | -0.073 | 0.15      | 0.6991 |
| Iceland        | -0.071 | 0.39      | 0.5348 |
| Ireland        | -1.275 | 54.82     | 0.0000 |
| Luxembourg     | -0.623 | 38.88     | 0.0000 |
| Netherlands    | -0.103 | 1.23      | 0.2668 |
| Norway         | -0.370 | 7.15      | 0.0075 |
| Poland         | -0.157 | 4.25      | 0.0393 |
| Portugal       | 0.659  | 138.73    | 0.0000 |
| Slovakia       | -0.015 | 0.02      | 0.8897 |
| Slovenia       | -1.260 | 370.83    | 0.0000 |
| Spain          | -0.064 | 1.47      | 0.2249 |
| Sweden         | -0.214 | 9.50      | 0.0021 |
| Switzerland    | -0.440 | 6.96      | 0.0083 |
| Turkey         | -0.058 | 0.17      | 0.6793 |
| Ukraine        | -1.196 | 98.56     | 0.0000 |
| Pooled         | -0.563 | 26.95     | 0.0000 |

Table A3

**Returns to educational mismatch, ols**

| Country        | Required education |      |       | Overeducation |      |       | Undereducation |       |       |
|----------------|--------------------|------|-------|---------------|------|-------|----------------|-------|-------|
|                | coeff              | %    | t     | coeff         | %    | t     | coeff          | %     | t     |
| Austria        | 0.077              | 8.1  | 7.27  | 0.031         | 3.2  | 2.05  | -0.029         | -2.8  | -1.02 |
| Belgium        | 0.078              | 8.1  | 5.00  | 0.019         | 1.9  | 1.14  | -0.008         | -0.8  | -0.36 |
| Czech Republic | 0.076              | 7.9  | 7.32  | 0.072         | 7.5  | 3.85  | -0.007         | -0.7  | -0.24 |
| Denmark        | 0.063              | 6.5  | 6.28  | 0.013         | 1.4  | 1.07  | -0.032         | -3.2  | -1.88 |
| Estonia        | 0.136              | 14.5 | 13.92 | 0.049         | 5.0  | 4.30  | -0.030         | -2.9  | -1.00 |
| Finland        | 0.087              | 9.1  | 16.73 | 0.011         | 1.1  | 1.24  | -0.031         | -3.1  | -2.85 |
| France         | 0.129              | 13.8 | 10.67 | 0.017         | 1.7  | 0.68  | -0.056         | -5.4  | -4.93 |
| Germany        | 0.116              | 12.3 | 11.22 | 0.049         | 5.0  | 2.04  | -0.039         | -3.8  | -3.09 |
| Great Britain  | 0.147              | 15.8 | 3.96  | 0.054         | 5.5  | 1.93  | -0.058         | -5.6  | -1.89 |
| Greece         | 0.057              | 5.9  | 3.93  | 0.026         | 2.6  | 2.21  | -0.011         | -1.0  | -0.30 |
| Hungary        | 0.129              | 13.7 | 8.22  | 0.053         | 5.5  | 1.79  | -0.071         | -6.8  | -3.21 |
| Iceland        | 0.076              | 7.9  | 8.47  | 0.029         | 2.9  | 1.64  | -0.003         | -0.3  | -0.25 |
| Ireland        | 0.100              | 10.5 | 3.82  | 0.065         | 6.7  | 3.37  | -0.023         | -2.3  | -0.74 |
| Luxembourg     | 0.095              | 10.0 | 9.15  | 0.086         | 9.0  | 5.36  | -0.004         | -0.4  | -0.34 |
| Netherlands    | 0.124              | 13.2 | 11.84 | 0.009         | 0.9  | 0.35  | -0.037         | -3.6  | -3.48 |
| Norway         | 0.077              | 8.0  | 12.84 | 0.040         | 4.1  | 3.72  | -0.021         | -2.1  | -1.98 |
| Poland         | 0.099              | 10.5 | 9.81  | 0.068         | 7.0  | 5.31  | -0.017         | -1.7  | -0.83 |
| Portugal       | 0.133              | 14.2 | 14.85 | 0.048         | 5.0  | 2.94  | -0.008         | -0.8  | -0.41 |
| Slovakia       | 0.088              | 9.2  | 11.09 | 0.051         | 5.2  | 2.39  | -0.010         | -1.0  | -0.65 |
| Slovenia       | 0.095              | 10.0 | 4.53  | -0.019        | -1.9 | -0.34 | -0.097         | -9.3  | -2.65 |
| Spain          | 0.084              | 8.7  | 7.88  | 0.040         | 4.1  | 3.86  | -0.036         | -3.5  | -2.74 |
| Sweden         | 0.076              | 7.9  | 18.16 | 0.023         | 2.4  | 2.50  | -0.028         | -2.8  | -3.27 |
| Switzerland    | 0.086              | 8.9  | 10.63 | 0.029         | 2.9  | 1.71  | -0.029         | -2.8  | -2.82 |
| Turkey         | 0.109              | 11.6 | 6.59  | 0.047         | 4.8  | 1.62  | -0.062         | -6.0  | -3.38 |
| Ukraine        | 0.069              | 7.1  | 2.98  | -0.015        | -1.5 | -0.33 | -0.110         | -10.4 | -3.25 |
| Pooled         | 0.108              | 11.4 | 20.45 | 0.044         | 4.5  | 7.57  | -0.041         | -4.1  | -7.05 |

Table A4

**Parameters of returns to attained education from earnings functions estimated  
by Heckman selection-correction estimator, and tests of independence of  
equations**

| Country        | Coefficient | Per cent | z    | Lambda | Wald test of independent equations |
|----------------|-------------|----------|------|--------|------------------------------------|
| Austria        | 0.060       | 6.2      | 6.2  | -0.494 | 29.57                              |
| Belgium        | 0.034       | 3.5      | 3.7  | 0.854  | 122.19                             |
| Czech Republic | 0.069       | 7.1      | 8.2  | -0.322 | 10.35                              |
| Denmark        | 0.045       | 4.6      | 5.4  | 0.076  | 1.99                               |
| Estonia        | 0.085       | 8.8      | 10.2 | -0.292 | 8.69                               |
| Finland        | 0.053       | 5.4      | 10.7 | -0.024 | 0.05                               |
| France         | 0.073       | 7.6      | 7.3  | 0.743  | 129.15                             |
| Germany        | 0.073       | 7.6      | 9.9  | -0.714 | 69.82                              |
| Great Britain  | 0.093       | 9.7      | 8.3  | -1.189 | 62.14                              |
| Greece         | 0.033       | 3.4      | 3.9  | 0.684  | 26.25                              |
| Hungary        | 0.118       | 12.5     | 7.7  | -0.086 | 0.39                               |
| Iceland        | 0.039       | 4.0      | 4.4  | -0.156 | 1.83                               |
| Ireland        | 0.065       | 6.7      | 6.2  | -1.273 | 53.86                              |
| Luxembourg     | 0.056       | 5.7      | 7.8  | -0.618 | 24.73                              |
| Netherlands    | 0.071       | 7.3      | 8.5  | -0.116 | 1.15                               |
| Norway         | 0.051       | 5.3      | 10.6 | -0.374 | 6.91                               |
| Poland         | 0.079       | 8.2      | 9.3  | -0.100 | 1.28                               |
| Portugal       | 0.077       | 8.0      | 8.0  | -0.033 | 0.07                               |
| Slovakia       | 0.062       | 6.4      | 6.2  | -0.381 | 8.00                               |
| Slovenia       | 0.083       | 8.7      | 6.6  | -1.268 | 255.83                             |
| Spain          | 0.050       | 5.1      | 9.3  | -0.059 | 0.92                               |
| Sweden         | 0.056       | 5.8      | 13.9 | -0.194 | 3.27                               |
| Switzerland    | 0.053       | 5.5      | 8.9  | -0.497 | 8.29                               |
| Turkey         | 0.083       | 8.6      | 6.6  | 0.000  | 0.00                               |
| Ukraine        | 0.082       | 8.6      | 4.7  | -1.183 | 105.94                             |
| Pooled         | 0.070       | 7.2      | 21.4 | -0.523 | 16.63                              |

The equations estimated

earnings equation

dependent variable: log usual monthly before-tax earnings

control variables: sex, potential labour market experience and its square; country dummies in the case of pooled sample (reference: Austria)

participation equation

dependent variable: worker

control variables: sex, age, age squared, number of children in the family, sex of children; country dummies in the case of pooled sample (reference: Austria)

Table A5

**Parameters of returns to attained education from earnings functions estimated by ols**

| Country        | Coefficient | Per cent | t    | R2    |
|----------------|-------------|----------|------|-------|
| Austria        | 0.059       | 6.1      | 6.2  | 0.330 |
| Belgium        | 0.029       | 2.9      | 2.2  | 0.103 |
| Czech Republic | 0.068       | 7.1      | 7.7  | 0.275 |
| Denmark        | 0.044       | 4.5      | 5.4  | 0.063 |
| Estonia        | 0.085       | 8.9      | 10.1 | 0.254 |
| Finland        | 0.053       | 5.4      | 10.8 | 0.334 |
| France         | 0.083       | 8.7      | 8.4  | 0.243 |
| Germany        | 0.074       | 7.7      | 9.4  | 0.324 |
| Great Britain  | 0.095       | 9.9      | 6.0  | 0.167 |
| Greece         | 0.031       | 3.2      | 3.2  | 0.191 |
| Hungary        | 0.117       | 12.5     | 7.4  | 0.334 |
| Iceland        | 0.039       | 3.9      | 4.4  | 0.317 |
| Ireland        | 0.059       | 6.1      | 3.5  | 0.044 |
| Luxembourg     | 0.059       | 6.0      | 7.8  | 0.224 |
| Netherlands    | 0.070       | 7.2      | 8.5  | 0.320 |
| Norway         | 0.052       | 5.3      | 10.4 | 0.247 |
| Poland         | 0.080       | 8.3      | 9.4  | 0.218 |
| Portugal       | 0.077       | 8.0      | 9.1  | 0.394 |
| Slovakia       | 0.060       | 6.2      | 5.5  | 0.240 |
| Slovenia       | 0.079       | 8.2      | 4.2  | 0.063 |
| Spain          | 0.050       | 5.2      | 9.2  | 0.237 |
| Sweden         | 0.055       | 5.7      | 15.1 | 0.347 |
| Switzerland    | 0.055       | 5.7      | 9.3  | 0.358 |
| Turkey         | 0.083       | 8.7      | 7.1  | 0.279 |
| Ukraine        | 0.069       | 7.1      | 3.5  | 0.035 |
| Pooled         | 0.070       | 7.3      | 22.3 | 0.757 |

The equation estimated

dependent variable: log usual monthly before-tax earnings

control variables: sex, potential labour market experience and its square; country dummies in the case of pooled sample (reference: Austria)

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