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## Investigating the exponential age distribution of firms

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## Investigating the Exponential Age Distribution of Firms

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**Abstract** While several plots of the aggregate age distribution suggest that firm age is exponentially distributed, we find some departures from the exponential benchmark. At the lower tail, we find that very young establishments are more numerous than expected, but they face high exit hazards. At the upper tail, the oldest firms are older than the exponential would have predicted. Furthermore, the age distribution of international airline companies displays multimodality. Although we focused on departures from the exponential, we found that the exponential was a useful reference point and endorse it as an appropriate benchmark for future work on industrial structure.

**JEL** L20, L25, L11

**Keywords** Age distribution; exponential distribution; firm size distribution, survival

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# 1 Introduction

A very large literature has focused on the firm size distribution (for a recent survey see de Wit (2005)). In fact, there is even a JEL classification code that specifically seeks to accommodate research on the firm size distribution.<sup>1</sup> Among this vast body of literature, some authors seek to determine the best functional fit to the empirical firm size distribution, generally focusing on the lognormal, the Pareto, and the generalized beta as the most suitable candidates. Other theoretical models seem to focus on generating the empirically-observed size distribution as one of the main ‘reality checks’ of their model’s predictions. Why has the empirical firm size distribution received so much attention? We suggest a few reasons here. First, the size distribution gives a useful summary representation of the structure of industries and economies, that allows for comparisons across samples and over time. Second, the size distribution displays a smooth, regular shape at the aggregate level that matches closely to theoretical densities. Third, the familiar right-skewed shape of the aggregate size distribution is remarkably robust across datasets and is referred to as a stylized fact of industry structure. Fourth, data on firm size is relatively easy to obtain. Fifth, the regular shape of the size distribution provides an explanandum that has inspired theoretical models (e.g. Gibrat’s (1931) celebrated model of firm growth).

In contrast to the size distribution, the age distribution has barely been investigated. In this paper, we suggest that the age distribution is a useful summary representation of the structure of industries, that it displays a regular shape that is robust across datasets and is a close match to the exponential distribution. The age distribution is also observed to make a useful contribution to theoretical modelling of the firm growth process.

One reason why the aggregate age distribution has not been investigated in previous empirical work is presumably because of the difficulty in obtaining data on the ages of large samples of firms. Headd and Kirchhoff (2009, p. 548) comment on “the dearth of information by business age” and write that “Simply stated, industrial organization and small business researchers are deprived of firm-age data.” The following reasons can help explain why data on firm age is harder to obtain than data on variables such as firm size. First, while the size distribution is constructed from current data (i.e. a firm’s current size), the age distribution is constructed from historical data concerning a firm’s initial conditions (i.e. a firm’s year of founding). Furthermore, in the construction of administrative datasets firms are required to give information on variables such as sales and employment (e.g. for tax reasons), while this requirement does not exist for age data. While we acknowledge that data on firm age has not always been easy to obtain, we also observe that the situation has been improving, which leads us to consider that the firm age distribution will receive more attention in future research.

Although the prior literature has not focused on the age distribution *per se*, many studies have focused on the related issue of firms’ survival rates. An early

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<sup>1</sup>This JEL code is L11: Production, Pricing, and Market Structure; Size Distribution of Firms. This information is taken from the February 2009 update to the JEL classification system (see [http://www.aeaweb.org/journal/jel\\_class\\_system.php](http://www.aeaweb.org/journal/jel_class_system.php)).

contribution coined the term ‘liability of newness’ to describe how young organizations face higher risks of failure (Stinchcombe, 1965). More recently, however, authors have referred to the ‘liability of adolescence’ (Bruderl and Schussler, 1990, Fichman and Levinthal, 1991) to explain why firms face an initial ‘honeymoon’ period in which they are buffered from sudden exit by their initial stock of resources. Still others have identified liabilities of senescence and obsolescence (Barron et al., 1994) according to which older firms are expected to face higher exit hazards. As such, the literature on firm age and survival has given conflicting predictions, and scholars who are not familiar with the subtleties of these conflicting concepts may not have a clear idea about the age structure of firms in an industry.

Furthermore, the existing literature on firm survival has often focused on tracking small samples of firms in specific industries (for example, Delacroix and Carroll (1983) on Argentinian and Irish newspapers, Barron et al. (1994) on credit unions in New York City, Klepper (2002) on the automobile, tyre, television, and penicillin industries in the US, and Thompson (2005) on the iron and steel shipbuilding industry in the US). In this vein, some studies have provided evidence that there are distinct periods of high entry and high exit at specific stages in the life cycle of some industries and submarkets (see for example Klepper and Thompson (2006) on the US laser industry, Guenther (2009) on the German machine tools industry, and Buenstorf and Klepper (2009) on the US tyre industry). While we acknowledge that detailed analysis of specific industries has been a fruitful field of research, in this paper we complement the existing literature by focusing on the age distribution at the aggregate level. In the absence of detailed information on the survival histories of specific age cohorts, it may be preferable to focus on the cross-sectional age distribution at a point in time. Instead of focusing on mortality rates over time for small samples of firms, the age distribution corresponds to a snapshot of accumulated mortality rates for all firms from all age cohorts combined. In addition, the age distribution might shed light on the structure of the age of technology used in an industry, and also the degree of adoption of general purpose technologies throughout the economy, if firms are assumed to be characterized by the capital vintage of the period in which they enter (as in the theoretical model in Salter (1960)). Furthermore, to the extent that organizations remain fundamentally inert once they are founded (Hannan and Freeman, 1984), the age distribution can elucidate the variety of different types of organization operating in an industry.

The aim of this paper is to draw attention to the exponential age distribution, and also to discuss cases in which the empirical age distribution drifts away from the exponential. Even in this situations, however, we argue that the exponential is a helpful reference point. Section 2 presents the theoretical interest in the exponential age distribution, and shows how empirically-observed age distributions from a number of different aggregate datasets seem to match well to the exponential case. Section 3 investigates the age distribution of young establishments in the US. Section 4 investigates the age distribution of the world’s oldest firms. Section 5 presents a disaggregated analysis of the international airlines sector. Section 6 concludes.

## 2 Theoretical modelling

We now demonstrate how the age distribution is of interest in theoretical models. In particular, an exponential age distribution is assumed in the following model of the firm size distribution. In this mathematical model, a Gibrat growth process is shown to give a lognormal firm size distribution within cohorts, which is then combined with an exponential distribution of firm age to give a Pareto firm size distribution at the aggregate level. The basic mathematical model (i.e. integrating a lognormal distribution over an exponential distribution to obtain a Pareto) was previously used by Huberman and Adamic (1999) to explain the number of web pages on internet sites, before being brought into economics by Reed (2001), who focused mainly on explaining the distributions of earnings and city sizes. Coad (2008) applies this model to modelling the firm size distribution, and presents some preliminary analysis on the firm age distribution.

Let  $x_t$  be the size of a firm at time  $t$ , and let  $\varepsilon_t$  be random variable representing an iid idiosyncratic, multiplicative growth shock over the period  $t - 1$  to  $t$ , with mean  $\bar{\varepsilon}$ . We have

$$x_t - x_{t-1} = \varepsilon_t x_{t-1} \quad (1)$$

which can be developed to obtain

$$x_t = (1 + \varepsilon_t)x_{t-1} = x_0(1 + \varepsilon_1)(1 + \varepsilon_2) \dots (1 + \varepsilon_t) \quad (2)$$

It is then possible to take logarithms in order to approximate  $\log(1 + \varepsilon_t)$  by  $\varepsilon_t$  to obtain<sup>2</sup>

$$\log(x_t) \approx \log(x_0) + \varepsilon_1 + \varepsilon_2 + \dots + \varepsilon_t = \log(x_0) + \sum_{s=1}^t \varepsilon_s \quad (3)$$

In the limit, as  $t$  becomes large, the  $\log(x_0)$  term will become insignificant, and we obtain:

$$\log(x_t) \approx \sum_{s=1}^t \varepsilon_s \quad (4)$$

Central Limit Theorem implies that  $\log(x_t)$  is normally distributed, which means that firm size (i.e.  $x_t$ ) is lognormally distributed:

$$P(x_t) = \frac{1}{x_t \sqrt{2\pi\sigma^2 t}} e^{-\left(\frac{(\ln x_t - \bar{\varepsilon}t)^2}{2\sigma^2 t}\right)} \quad (5)$$

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<sup>2</sup>This logarithmic approximation is only justified if  $\varepsilon_t$  is ‘small’ enough (i.e. close to zero), which can be reasonably assumed by taking a short time period (Sutton (1997)).

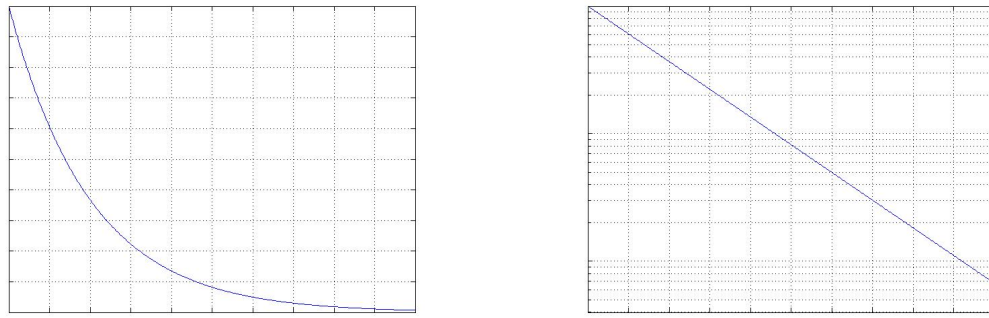


Figure 1: Exponential distribution plotted on linear axes (left) and plotted again with a logarithmic  $y$ -axis (right).

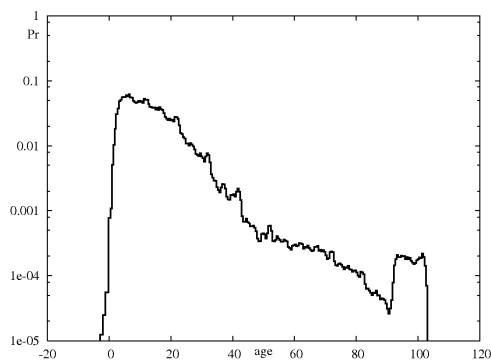


Figure 2: Kernel density of the age distribution of Indian small scale industries in 2003. Source: Coad and Tamvada (2008).

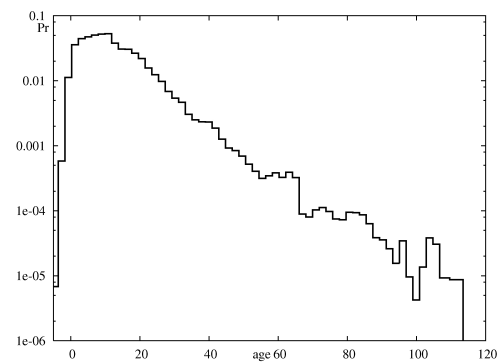


Figure 3: Kernel density of the age distribution of Spanish firms in 2005. Source: Coad (2008), based on the data in Segarra et al. (2008), page 92.

This lognormal firm size distribution corresponds to firms of the same age, within the same cohort.<sup>3</sup> In an extension of the model, however, we need no longer assume that  $t$  has the same value for all firms. Instead, we suggest that  $t$  is itself a random variable. It seems reasonable to assume the distribution of firm age to be exponentially distributed. If  $t$  is exponentially distributed, we have:

$$P(t) = \lambda e^{-\lambda t} \quad (6)$$

Figure 2 shows what an exponential distribution looks like on linear axes (left) and also with a logarithmic  $y$ -axis (right).

<sup>3</sup>An interesting and recent strand of literature has investigated how the firm size distribution for young cohorts of firms evolves over time as these cohorts grow older (Cabral and Mata, 2003, Angelini and Generale, 2008, Cirillo, 2010). These studies generally observe that the firm size distribution for very young cohorts is particularly skewed to the right, but that  $\log(\text{size})$  becomes less skewed and more symmetric (i.e. size approaches the log-normal) as cohorts grow older. The model presented here assumes that the size distribution within cohorts is lognormal, which therefore might not be entirely appropriate for cohorts of very young firms.

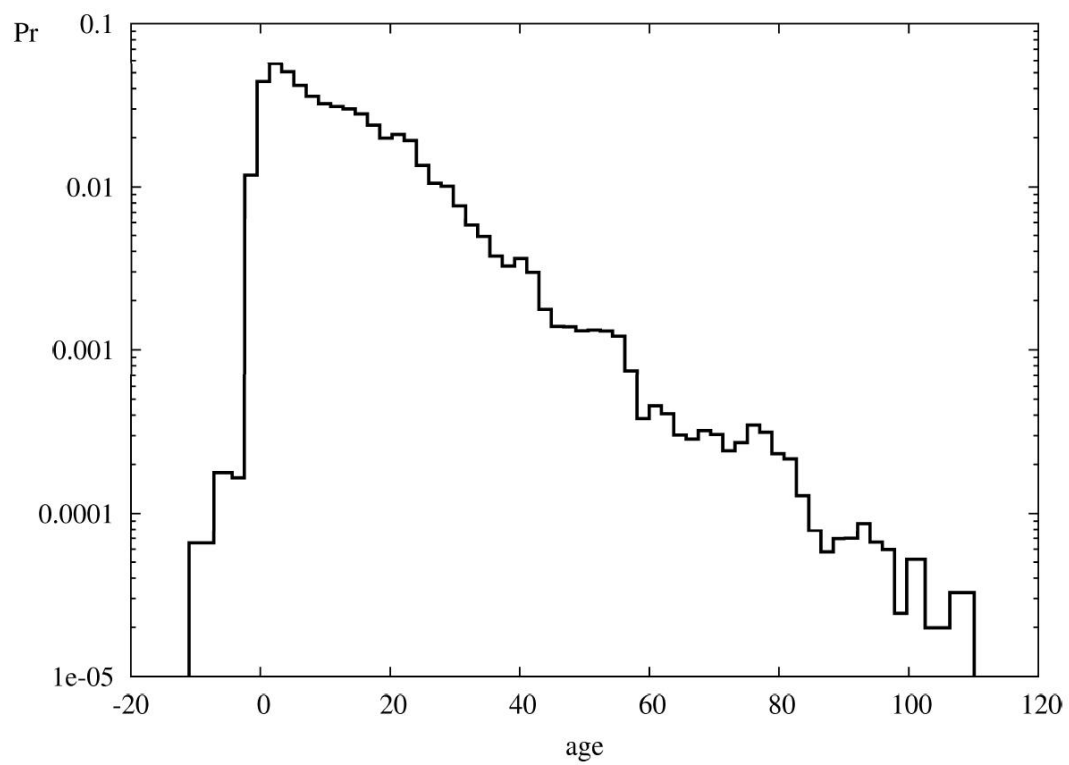


Figure 4: Kernel density of the age distribution of Italian firms in 2000, based on the Unicredit dataset in Bottazzi et al. (2008).

In order to obtain the mixture of these two distributions, we apply the following rule: if the distribution of a variable  $a$ ,  $p(a, b)$ , depends on a parameter  $b$  which in turn is distributed according to its own distribution  $r(b)$ , then the distribution of  $a$  is given by  $p(a) = \int r(b) \cdot p(a, b) db$  (Adamic and Huberman (1999), Huberman and Adamic (1999)).

This gives us the following:

$$P(x_t) = \int \lambda e^{\lambda t} \cdot \frac{1}{x_t \sqrt{2\pi\sigma^2 t}} e^{-\left(\frac{(\ln x_t - \bar{\varepsilon}t)^2}{2\sigma^2 t}\right)} dt \quad (7)$$

and, as in Adamic and Huberman (1999), this can be developed to yield:

$$P(x_t) = C \cdot x_t^{-\beta} \quad (8)$$

where  $C$  is a constant and is given by  $C = \lambda/\sigma(\sqrt{(\bar{\varepsilon}/\sigma)^2 + 2\lambda})$ . The exponent  $\beta$  is in the range  $[1, \infty]$  and is determined by  $\beta = 1 - \frac{\bar{\varepsilon}}{\sigma^2} + \frac{\sqrt{(\bar{\varepsilon}^2 + 2\lambda\sigma^2)}}{\sigma^2}$ . When the mean growth rate is close to 0%,  $\bar{\varepsilon}$  will be close to 1. As a result, if  $\lambda$  is small (implying that the exponential decay is relatively weak, i.e. that it is not uncommon to find firms with an age much greater than one)<sup>4</sup>, and if  $\sigma$  is small (which is not implausible either), then the exponent  $\beta$  will be close to Zipf's value of 1, which has been observed in empirical work on US firms (Axtell (2001)).

The scant empirical evidence on the age distribution suggests that the exponential distribution is a valid heuristic.<sup>5</sup> Figures 2 and 3 shows the age distribution for Indian small scale businesses and also for Spanish firms. Figure 2 shows that, even in a sample of small businesses, these firms have very different ages. Most firms are relatively young, but some are extremely old. Figure 4 shows the age distribution of a census of Italian firms of all sizes in the year 2000, based on the data in Bottazzi et al. (2008). Furthermore, analysis of the age distribution of Irish firms can be found in Kinsella (2009). These distributions appear to be well approximated by a straight line of negative slope over most of the support, covering several orders of magnitude, which on these semi-log axes would signify that the empirical distribution is well approximated by an exponential distribution. Although data on firm age may contain a certain amount of measurement error, the fact that these three diagrams constructed from independent datasets are in conformity with each other is encouraging.

Figures 2, 3 and 4 show the exponential distribution appears to be more appropriate for describing the central part of the distribution, but little attention was given to the age distribution for the youngest firms. Furthermore, the age structure for the oldest firms was not well explored. In the rest of the paper, we investigate the age distribution for these two extremes, focusing on the age distribution for very

<sup>4</sup>This condition is trivial since the duration  $t$  of a Gibrat-type 'shock' can be made arbitrarily short.

<sup>5</sup>Cook and Ormerod (2003) suggest a power law distribution for age, but this is not based on any direct observation of data, but observation of aggregated data and speculation about the disaggregated structure generating the aggregated data.



young establishments and also very old firms, taking the exponential distribution as a reference point. We also show that, while the exponential is a good representation of the age distribution at the aggregate level, it is not always valid for individual sectors (such as the international airline sector).

### 3 The age distribution of young establishments

#### 3.1 Previous literature

In an economic system characterized by constant entry rates, an exponential age distribution in a cross-section of firms implies a constant survival rate for firms. Consider again the exponential age distribution:

$$P(t) = \lambda e^{-\lambda t} \quad (9)$$

The probability of a firm being of age  $t$  is  $\lambda e^{-\lambda t}$ , while the probability of a firm being of age  $t + 1$  is equal to  $\lambda e^{-\lambda(t+1)} = \lambda e^{-(\lambda t + \lambda)} = \lambda e^{-\lambda t} \cdot e^{-\lambda}$ , where  $e^{-\lambda} < 1$  is the survival rate. In other words, the probability of a firm surviving to age  $t + 1$  is equal to the probability of it surviving to age  $t$ , multiplied by the survival rate  $e^{-\lambda}$ .

While the assumption of constant entry rates is seen to be approximately valid at the aggregate level (we explore this later), the hypothesis of constant survival rates for firms of different ages is rejected for small firms. Figure 5 summarizes results from previous research that shows how the exit hazard decreases over time for new plants and firms. Figure 5 is plotted with a logarithmic  $y$ -axis, because a constant survival probability for different years would be represented by a straight line on these axes. The lines do appear to be approximately straight, but a closer inspection suggests that they are slightly ‘droopy’ or convex with respect to the origin. As such, it is worth investigating whether or not survival probabilities are constant or increasing over time in the case of new plants and firms. Our analysis on the BDS dataset on new US establishments in the following section complements these studies by providing stronger evidence that annual survival rates increase in the years immediately following entry.

#### 3.2 Database

The database we analyze is publicly available and can be found online at the following URL: [http://www.ces.census.gov/index.php/bds/bds\\_database\\_list](http://www.ces.census.gov/index.php/bds/bds_database_list).<sup>6</sup> The Business Dynamics Statistics (BDS) database is a comprehensive government database on the population of young US establishments (also referred to hereafter as plants), which contains relatively detailed information on the number of young establishments and their ages.<sup>7</sup>

<sup>6</sup>The online database was accessed and downloaded by the author on the 21st November 2009.

<sup>7</sup>Since the data series on Business Dynamics Statistics are based on administrative rather than sample data, there are no issues related to sampling error. Nonsampling error, however, still exists. Nonsampling errors can occur for many reasons, such as the employer submitting corrected

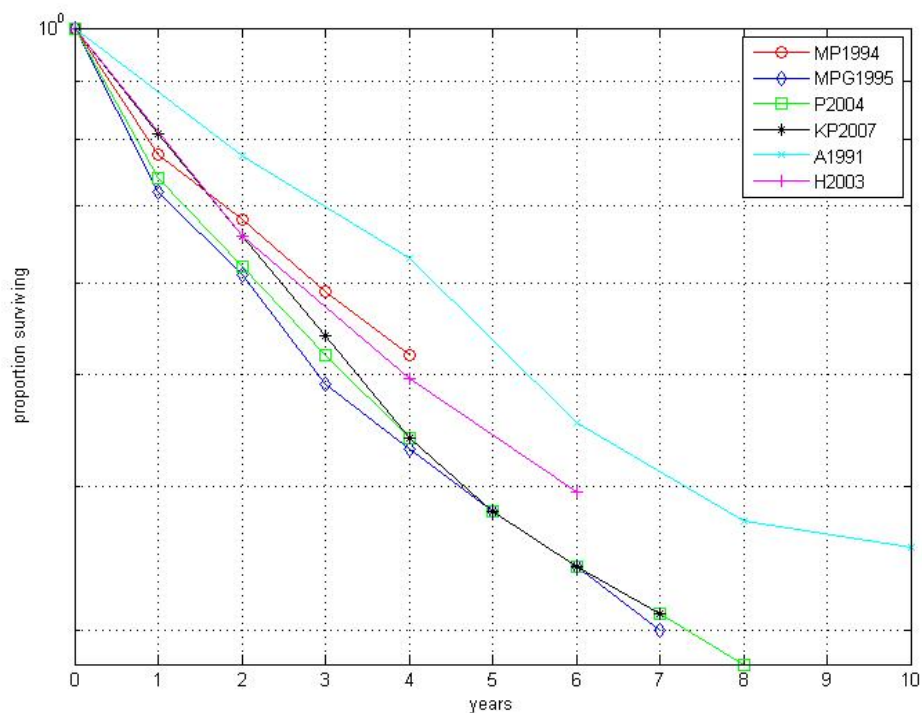


Figure 5: Percentage of surviving plants and firms reported in previous analyses. Note the log scale on the  $y$ -axis – if the survival probability is constant across years, the data should be represented as a straight line on these semi-log axes. The legend refers to the following data respectively: Mata and Portugal (1994, p. 235) on Portuguese firms (1983 cohort), Mata et al. (1995, p. 468) on Portuguese plants (1983 cohort), Persson (2004, p. 428) on Swedish plants (1987 cohort), Audretsch (1991, p. 443) on US firms (1976 cohort), Headd (2003, p. 59) on US firms (1989-1998) and Knaup and Piazza (2007, p. 8) for US plants (1998 cohort).

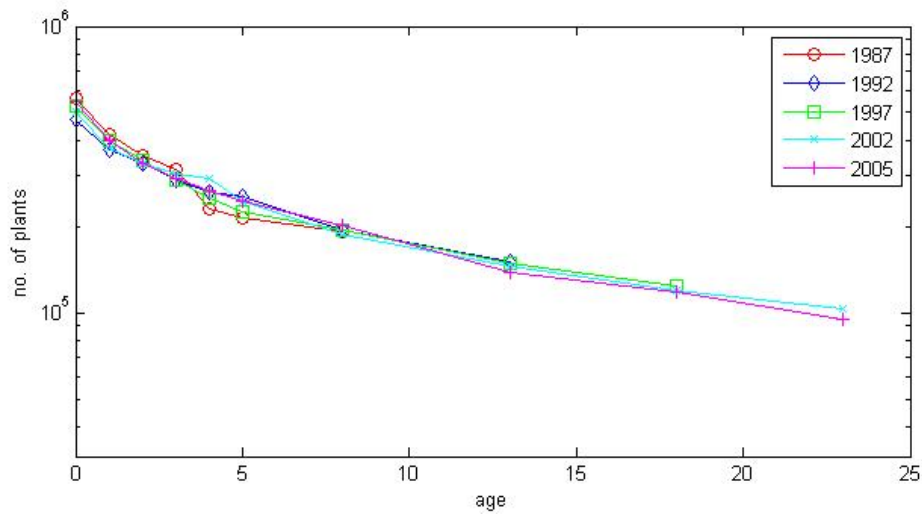


Figure 6: Aggregate age distribution for different years

Birth year is defined as the year an establishment first reports positive employment in the US LBD database. Establishment age is computed by taking the difference between the current year of operation and the birth year. Given that the LBD series starts in 1976 observed age is by construction left censored at 1975.

In the case of multi-plant firms, establishments are assigned a firm age based upon the age of the parent firm. This reflects the idea that new establishments that are set up by incumbent parent firms can already benefit from their parents market experience, and so are not considered to be entirely new establishments. The age of the parent firm, in turn, is based on the age of the oldest establishment in the firm. The vast majority of new firms are single-unit firms, however.

The vast majority of establishment openings are true greenfield entrants. Similarly, the vast majority of establishment closings are true establishment exits. Note, however, that mergers and acquisitions and divestitures could lead to abrupt changes in firm age purely from establishment composition issues if we defined firm age in each year using age of the oldest establishment owned in that year. Unfortunately there is no way to control for this effect in the database.

We begin by taking the number of plants of age 0, 1, 2, 3, 4, 5, 6-10, 11-15, 16-20, 21-25, and 26+. We take the midpoint of those classes that span more than one year (e.g. firms in the 6-10 year class are represented by the age 8), and divide the total number of plants in the class by the number of years spanned by the class, to obtain a representative frequency for the midpoint. We ignore the last category (26+) because it is unbounded.

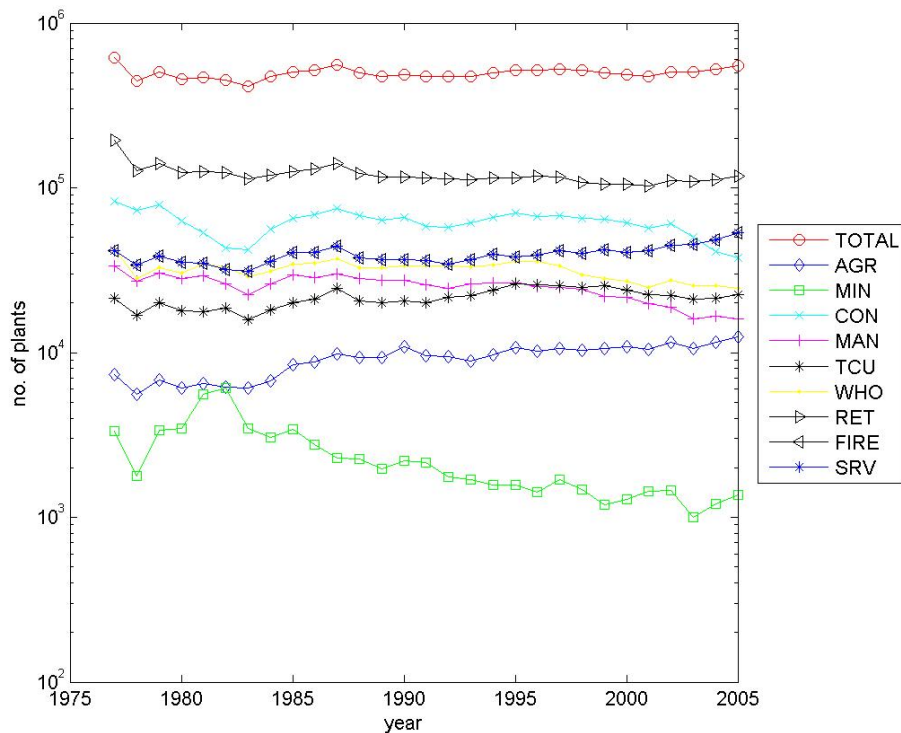


Figure 7: Number of establishments aged zero in each year. By order of appearance in the legend, the sectors are: Agricultural services, forestry and fishing; Mining; Construction; Manufacturing; Transportation and public utilities; Wholesales trade; Retail trade; Finance, insurance and real estate; Services.

### 3.3 Analysis

Figure 6 plots the aggregate age distribution for different years. Instead of pooling the years together, we focus on the age distribution for individual years.<sup>8</sup> Nonetheless, we observe that the age distribution changes little over time. We also observe that the age distribution is visibly convex with respect to the origin, whereas an exponential age distribution for this sample of young plants would suggest a straight line.

An aggregate age distribution such as the one observed here can be decomposed into two distinct factors. First, it could arise because the number of entrants in each year is steadily increasing (i.e. approximately exponentially increasing). Second, it could be the artifact of the survival rates within cohorts, such that a roughly

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employment data after the end of the year as well as late filers. Other sources of error include typographical errors made by businesses when providing information. Such errors, however, are likely to be distributed randomly throughout the dataset.

<sup>8</sup>It doesn't make good sense to pool years together, because the age distribution taken from different years is not independent. For instance, if in one year an unusually large number of establishments are observed to enter, this will probably be visible in cross-sections of the age distribution taken in subsequent years.

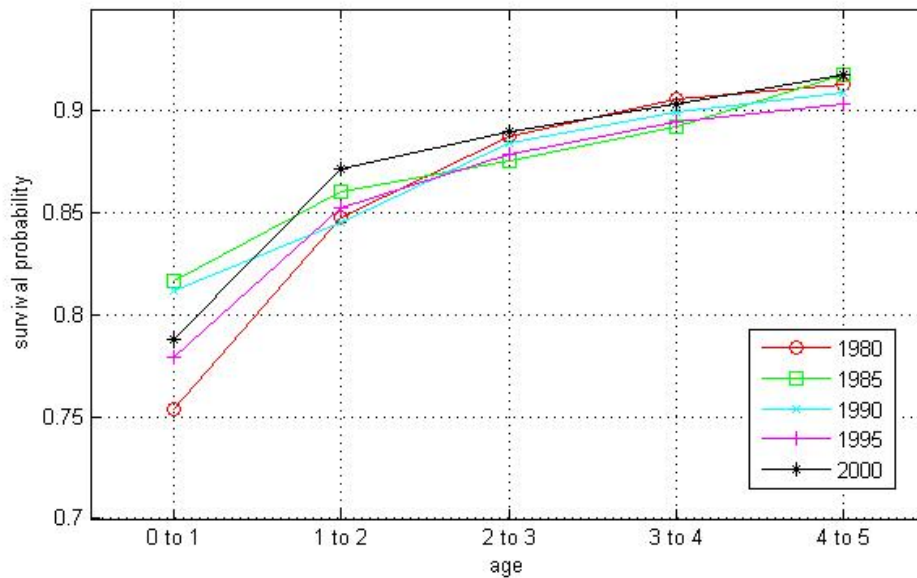


Figure 8: Survival rates of cohorts at different ages

constant proportion of plants within any cohort exit each year. The first scenario is investigated in Figure 7. The number of entering plants appears to be roughly constant over the time period at the aggregate level, although at a sectorally disaggregated level the pattern is much messier. For instance, the number of entrants per year appears to be steadily increasing in the Agricultural services, forestry and fishing sector, for example, while it tends to decrease in the Mining sector (which is a relatively mature industry). The second scenario is investigated in Figure 8, which plots the survival rates for different cohorts. In each year a certain proportion of establishments are observed to exit. This proportion is not constant over time, however. The observation that the survival rate is lower for very young plants is consistent with the unexpectedly high number of very young plants in the aggregate age distribution that was observed in Figure 6.

Figure 8 shows that survival rates are lowest over the period 0-1 year, and that they rise steadily over time. Survival of the first year is hardest, but survival gradually becomes easier over time. While only 79% survive their first year, 91% of establishments survive from their 4th year to their 5th year, on average. Needless to say, these differences in survival probabilities are highly statistically significant (for details see Table 2 in the Appendix).

To summarize, the exponential age distribution does not hold in the case of very young plants because, although the number of entrants is roughly constant across years, the youngest plants are observed to have a higher exit hazard. This stands in contrast to a constant exit hazard over time predicted by the exponential age distribution benchmark.

Table 1: The World's oldest family companies. Dates of founding are approximate in some cases. Source: The Economist (2004), based on data from familybusiness-magazine.com

	Date of founding	Country
Kongo Gumi	578	Japan
Hoshi Ryokan	718	Japan
Château de Goulaine	1000	France
Fonderia Pontificia Marinelli	1000	Italy
Barone Ricasoli	1141	Italy
Barovier & Toso	1295	Italy
Hotel Pilgrim Haus	1304	Germany
Richard de Bas	1326	France
Torrini Firenze	1369	Italy
Antinori	1385	Italy
Camuffo	1438	Italy
Baronnie de Coussergues	1495	France
Grazia Deruta	1500	Italy
Fabbrica D'Armi Pietro Beretta	1526	Italy
William Prym	1530	Germany

## 4 The age distribution of the oldest firms

In this section we investigate the upper tail of the firm age distribution — the case of ‘Methuselah’ firms. Our dataset on the oldest firms includes a high proportion of family firms, because joint-stock corporations do not face the same survival imperatives as family firms, where firm failure might be interpreted in terms of letting the family down. To give an idea of the kind of ages these firms reach, Table 1 shows the ages of the world's oldest family firms.

Some old firms, such as family firms, go to great lengths to continue their operations Bertrand and Schoar (2006). Consider for example the gun-maker Beretta (founded in 1526): the current president is Ugo Gussalli Beretta, who was adopted by his childless uncle in order to inherit the Beretta name and keep the succession within a direct family line.<sup>9</sup> In cases such as this, firms may continue for reasons that are not purely commercial, and as a result we may expect departures from the exponential age distribution benchmark due to the extreme longevity of a small number of firms at the upper tail of the age distribution.

In this section we the age distribution of the world's oldest firms, basing our analysis on databases available on the internet.<sup>10</sup> Figure 9 shows a Zipf plot of

<sup>9</sup>See <http://www.familybusinessmagazine.com/worldsoldest.html> (accessed 18th January 2010).

<sup>10</sup>The main dataset comes from [http://en.wikipedia.org/wiki/List\\_of\\_oldest\\_companies](http://en.wikipedia.org/wiki/List_of_oldest_companies) (accessed on the 14th December 2009), and is based on data from Tokyo Shoko Research, Japan. This dataset of the oldest companies includes brands and companies, excluding associations, educational, government or religious organizations. To be listed, a brand or company name must remain, either in whole or in part, since inception. If the original name has subsequently changed due to acquisition

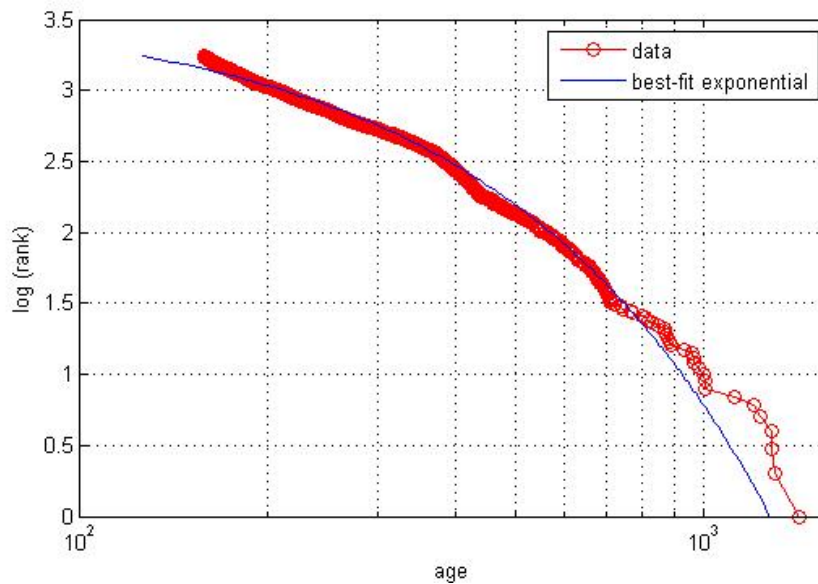


Figure 9: Zipf plot of the age distribution for the world's oldest companies. A best fit exponential is also plotted as a reference.

the world's oldest companies (on double-log axes). Alongside the plotted data is the best-fit exponential distribution. The empirical distribution is noticeably more right-skewed than the best-fit exponential distribution, indicating that the world's oldest firms are much older than the exponential distribution would predict. On these double-log axes, a straight line corresponds to a Pareto or power law distribution. The concavity of the empirical distribution on these axes therefore indicates that the empirical age distribution is less skewed than the Pareto case.

We consider these extremely old companies as meaningful observations with a plausible economic explanation. We can only remark, however, that these extremely old firms are small in number, and that even though the exponential benchmark for the empirical age distribution is not verified exactly, it remains a useful approximate benchmark in practical terms. Indeed, one might even suppose that the popular em-

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or renaming, this must be verifiable. Age is calculated as 2009 minus year of founding. Another useful data source is [http://www.servinghistory.com/topics/List\\_of\\_oldest\\_companies](http://www.servinghistory.com/topics/List_of_oldest_companies) (accessed on the 25th January 2010), which contains a list of the world's oldest companies as well as describing their line of business and providing external links to the companies' websites (interestingly enough, all of these firms seem to have their own websites). Some minor discrepancies can be found between the Wikipedia (W hereafter) database, the Serving History (SH) database, and the Economist (E) database on old family firms presented in Table 1. W and SH report the age of Hoshi Ryokan as 717, whereas it is 718 in E. E describes an Italian firm called Camuffo, founded in 1438; W and SH report that a German firm called Andechs was founded in 1438, but there is no Italian firm called Camuffo. W reports the age of Chivas as 1802 whereas it is recorded as 1801 in SH. Furthermore, SH has 14 'missing' observations when compared to W, corresponding to firms founded in the years 1295, 1398, 1399, 1498, 1499, 1599, 1649, 1699, 1749, 1774, 1799, 1824 (2 firms), and 1851. Thus, SH has only  $n=1726$  observations for the years up to 1851, whereas W has  $n=1740$ . The analysis in Figure 9 is based on W, but it comes as no surprise to discover that when we repeated the analysis on SH data the results were virtually identical.

pirical methodology of excluding extreme observations as ‘outliers’ may well overlook this upper-tail phenomenon completely.

## 5 Sector-level analysis

Concerning the size distribution, a number of researchers have suggested that the smooth shape observed at the aggregate level is merely a statistical artifact generated through aggregation, and that the disaggregated size distribution observed at the level of individual sectors is much less regular (Dosi et al., 1995). For example, Bottazzi and Secchi (2005) observe significant bimodality in their analysis of the firm size distribution of the worldwide pharmaceutical industry, and relate this to a cleavage between the industry leaders and fringe competitors. Bottazzi et al. (2010) also observe significant bimodality in the French clothing sector.

In this section, we investigate the possibility of multimodality in the age distribution of the international airline industry. We focus on this particular industry because we consider it to be a special case that is particularly likely to show a multimodal age distribution. In the early 20th Century, when airline technology began to take off, many countries invested heavily in national airlines. International air regulations (in particular, bilaterally-negotiated air traffic rights which were allocated by foreign government departments to specific airline companies by name) provided incentives to governments to subsidize their airlines even if they made losses – and these national airlines did frequently make losses. (For instance, the Belgian carrier Sabena only reported a positive financial result twice in its 78 year history.)<sup>11</sup> In other words, failure of airlines was often artificially avoided through government intervention. Furthermore, new firms were often denied air traffic rights on key international routes. As a result, we anticipate that the airline industry contains an unnaturally high frequency of old airline companies — a historical characteristic that will presumably be discernable in the empirical age distribution.

Airlines are included in the dataset based on whether they are IATA members. This list of members corresponds to the population of major international airlines. We identified 231 airlines as IATA members on the basis of the member list downloaded from the IATA website.<sup>12</sup> After scanning the internet for the relevant websites, age data was obtained for all 231 of these companies, and age is measured relative to the year 2010.<sup>13</sup>

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<sup>11</sup>These two profitable years were hardly impressive – Sabena made a profit in 1958 only because of the EXPO held in Brussels, and in 1998 due to some financial window-dressing by means of a sales-and-lease-back deal with Flightlease (Swissair) (Knorr and Arndt, 2004)

<sup>12</sup>Data was downloaded from [http://www.iata.org/membership/airline\\_members\\_list?All=true](http://www.iata.org/membership/airline_members_list?All=true) on the 18th January 2010. The main advantages of IATA membership concern international transport (as opposed to transport within domestic airspace), being useful for such issues as interline transport agreements, global distribution systems, foreign currency management, and baggage handling. National airlines and low-cost airlines are therefore not likely to be interested in becoming IATA members.

<sup>13</sup>For example, since Adria Airways was founded in 1961, its age will be calculated as  $2010 - 1961 = 49$ . In the few cases where year of founding is not recorded as being the same as the year of commencement of operations, we calculate a company’s age on the basis of its year of founding. In these cases where year of founding and year of commencement of operations are different, year of



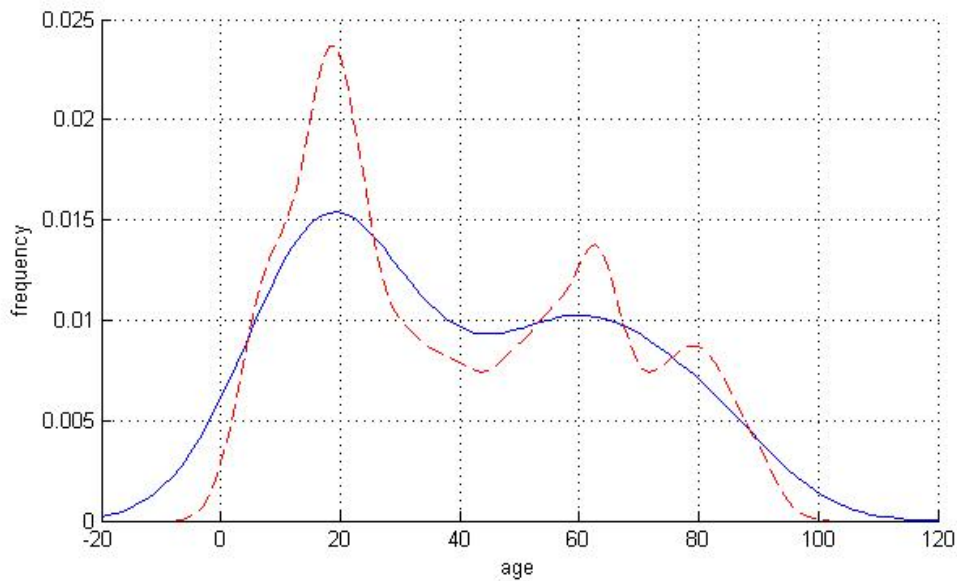


Figure 10: Kernel density of the age distribution for international airlines (IATA members in 2010). Kernel densities obtained using the normal kernel function. The smoother line is the Matlab 7.9.0 default for estimating normal densities ( $u=10.5758$ ). The dotted line is obtained using a kernel bandwidth that is three times smaller than this default value.

The age distribution is presented in Figure 10. The empirical age distribution displays clear multimodality, reflecting the fact that the international airline industry contains many old companies. The exponential distribution is therefore not a useful approximation to the empirical age distribution in this industry, although it provides a useful benchmark which allows us to comment on the unexpectedly large number of old international airline companies.

## 6 Conclusion

We began the paper by showing some age distribution plots at the aggregate level, observing that the exponential distribution appeared to be a useful approximation for the empirical distribution. In later sections of the paper, however, we focused on situations in which the exponential gave only an imperfect representation. In contrast to the exponential benchmark, we observed that young establishments seem to be especially numerous, the oldest firms seem to be exceptionally long-lived, and at the disaggregated level of certain specific sectors we can observe a particularly irregular age structure.

In spite of these departures from the exponential distribution, we argue that the exponential is still a useful benchmark for understanding the age structure of industries. Concerning the large number of young establishments, it could be that

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founding precedes year of commencement of operations by only one year in the majority of cases.

this excess weight in the age distribution corresponds to over-entry by inefficient establishments who exit shortly afterwards (the case of ‘hopeful monsters’); excess entrepreneurship undertaken by overoptimistic entrepreneurs. (Santarelli and Vivarelli (2007) provide an interesting survey of this phenomenon of over-entry.) To the extent that departures from the exponential benchmark among young establishments represent over-entry, then the exponential age distribution could be used to gauge the magnitude of this phenomenon.

Departures from the exponential benchmark in the case of the oldest firms also have a ready economic explanation, in that certain long-lived firms, and especially family firms, do not pursue economic rationality in the sense of maximization of expected profits, but instead they may seek to maximize their chances of survival (e.g. by pursuing risk-averse strategies).

We also presented evidence that the exponential distribution may not always be a valid heuristic at the disaggregated level of individual sectors. We focused on a particular sector that we suspected of having an irregular age distribution — the international airline industry. In contrast to the smooth shape observed at the aggregate level, the age distribution of this particular sector is much messier and displayed conspicuous multimodality.

To summarize, therefore, there are a number of situations in which the empirical age distribution strays from the exponential benchmark. Nonetheless, we consider the exponential to be a useful approximation. In the words of Herbert Simon, “statistically significant deviations of data from a generalization should not always, or usually, lead us to abandon the generalization” (Simon, 1968, p. 454). We argue that even in those situations where the exponential can be rejected on statistical grounds, it still serves as a useful benchmark against which these distortions can be gauged. For example, we suggest that theoretical models of firm entry, exit, and industry evolution would do well to generate an exponential age distribution as part of their output, even though the empirical data is not exactly exponentially distributed.

## APPENDIX

Table 2: Survival rates for the years following entry for different cohorts of young establishments. Survival rates for individual cohorts are followed by average survival rates, and pairwise two-sample  $t$ -tests that reject the hypotheses that the survival rates are constant over time within ageing cohorts.

	0-1	1-2	2-3	3-4	4-5
1977	0.7702	0.8610	0.8780	0.8617	0.9221
1978	0.8974	0.8238	0.8552	0.8905	0.9221
1979	0.7349	0.8589	0.8737	0.9046	0.9153
1980	0.7539	0.8479	0.8876	0.9057	0.9132
1981	0.7969	0.8590	0.8855	0.9047	0.9100
1982	0.7717	0.8538	0.8855	0.8952	0.9062
1983	0.8282	0.8493	0.8751	0.9074	0.9071
1984	0.8727	0.8605	0.8931	0.8993	0.9055
1985	0.8168	0.8602	0.8757	0.8922	0.9180
1986	0.8059	0.8408	0.8708	0.9035	0.9032
1987	0.7571	0.8398	0.8854	0.8910	0.9047
1988	0.7826	0.8627	0.8816	0.8890	0.9114
1989	0.8121	0.8643	0.8719	0.9019	0.9101
1990	0.8121	0.8455	0.8843	0.8991	0.9093
1991	0.7824	0.8596	0.8815	0.9020	0.9142
1992	0.7749	0.8526	0.8831	0.9007	0.9075
1993	0.7765	0.8535	0.8828	0.8997	0.9069
1994	0.7765	0.8563	0.8781	0.8926	0.9054
1995	0.7794	0.8521	0.8787	0.8951	0.9034
1996	0.7758	0.8488	0.8740	0.8927	0.9028
1997	0.7689	0.8485	0.8781	0.8941	0.9116
1998	0.7943	0.8702	0.8976	0.9117	0.9230
1999	0.7874	0.8628	0.8942	0.9064	0.9120
2000	0.7881	0.8718	0.8899	0.9038	0.9175
Average	0.7924	0.8543	0.8809	0.8977	0.9109
Std Dev	0.0357	0.0105	0.0090	0.0098	0.0062
$t$ -stat	0-1 & 1-2		-8.1541		
$p$ -value			1.75E-10		
DoF			46		
$t$ -stat	1-2 & 2-3		-9.4322		
$p$ -value			2.53E-12		
DoF			46		
$t$ -stat	2-3 & 3-4			-6.1847	
$p$ -value				1.53E-07	
DoF				46	
$t$ -stat	3-4 & 4-5				-5.5873
$p$ -value					1.20E-06
DoF					46

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