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Is there a "Double Disadvantage" vis-à-vis Stayers?**

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ABSTRACT

East-West Migration and Gender: Is there a “Double Disadvantage” vis-à-vis Stayers?*

This paper documents whether female East-West migrants in Germany after the reunification experience a gain or a disadvantage after they moved compared to both stayers and males. It employs panel data techniques to take account of unobserved heterogeneity. I find that migrant women after migration neither experience a drop in relative employment, nor lower relative hourly wages. They do, however, work less hours and have a lower annual income. The results also suggest that for them, the income effect dominates the substitution effect and they substitute market work with home production, specifically with childcare.

JEL Classification: J16, J61, R23

Keywords: migration, gender, panel data

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1 Introduction

Gender differences among migrants are often more substantial than among the local population in general. Boyd (1984) postulates that "...in addition to the status of being a migrant, immigrant women experience additional difficulties in the labor force as women..." (p. 1092). Empirical studies have investigated whether immigrant women face a so-called "double disadvantage" of being both a female and a foreign-born with respect to labour force participation and employment (De Jong and Madamba, 2001, Raijman and Semyonov, 1997, Boyd, 1984, Kats, 1982), wages (De Jong and Madamba, 2001, Haberfeld, 1993, Kossoudji and Ranney, 1984), occupational status and job mobility (Raijman and Semyonov, 1997, Boyd, 1984) and job mismatch (De Jong and Madamba, 2001). All of them compare female immigrants to male immigrants and to native-born females in the receiving country.¹

On the other hand, understanding how female migrants perform relative to *stayers* is crucial in order to complete the picture. A parallel question of interest is: Do female migrants experience a gain in their relative labour market outcomes after migration, or do they experience a (double) disadvantage with respect to their male counterparts as well as the sending country's population? This paper attempts to answer this question.

Neoclassical theory of migration (Todaro, 1969, Harris and Todaro, 1970) postulates that migration occurs if the present discounted value of the lifetime income stream in the destination region, net of migration costs, is higher than the one in the source region.

¹A large and related literature analyzes the effects of family migration and finds that it has a negative impact on married women's labour force participation, employment, weeks worked, hours worked, income and attitudes towards work (see, for example, Boyle et al, 2002 for a review and the references therein). Moreover, two growing strands of literature focus separately on female immigrants' assimilation (see, for example, Amuedo-Dorantes and de la Rica, 2007, Blau and Kahn, 2005, Blau et al. 2002, Antecol et al., 2003, Schoeni, 1998, and Funkhouser and Trejo, 1998), and on the gender gap among immigrants (see, for example, Antecol, 2000 and the references therein). Finally, Adsera and Chiswick (2007, forthcoming) analyze labour market performance of immigrants by gender in the fifteen EU countries. They find a significant negative effect of immigrant status on earnings upon arrival and that gender differences are more important among non-EU born migrants.

Migrants are often viewed as being positively self-selected with respect to the sending country population, and thus, on average, being more likely to engage in labor market activities and to earn higher wages in the receiving country. However, family migration models (starting with Mincer, 1978) emphasize that the decision to move is a joint decision by the family and women are typically viewed as "tied" movers. Being tied to their husbands they do not necessarily experience a gain from migration. Family investment models, on the other hand, postulate that, willing to maximize the joint returns to migration, married female migrants in the early stage of emigration engage in labor market activities and are the main income earners, while their husbands invest in the human capital of the host country and are expected to contribute to the family budget later on (Duleep and Dowhan, 2002, Baker and Benjamin, 1997, Duleep and Sanders, 1993).

A second contribution of this paper is that it analyzes gender and migration in the context of the transition economy of East Germany, which is becoming an increasingly relevant issue in light of the recent EU enlargements and European East-West migration.

In East Germany, labor force participation of women was high under communism (more than 80 percent), and remained relatively high after the reunification (72 percent in May 2000) (Bonin and Euwals, 2005). However, their employment fell more than that of men, and both occupational segregation and gender gap in wages existed (Hunt, 2002). At the same time, fertility declined (Lechner, 2001), as did the availability of childcare facilities, which is, however, still better in eastern than in western Germany (Wrohlich, 2004). Finally, the majority of migrants from East to West Germany were women (see Figure 1).

This paper describes the labor market performance of female migrants from East to West Germany over 1990 - 2001, comparing them both to stayers and males. It documents the relation between being both female and a migrant and four outcomes of interest: individual annual income, employment, hours worked and hourly earnings. It uses the

German Socio-Economic Panel (GSOEP), which is a longitudinal dataset that contains information on both pre- and post- migration histories of migrants as well as information on stayers. Given that there are both pre- and post- data available for the same individuals, I use a sort of "difference-in-difference-in-difference" approach. Having panel data allows differencing away time-invariant unobservable confounders. Thus, to the extent that self-selection into migration is influenced by time-invariant unobservables, this approach identifies a causal effect of being a female migrant on labor market performance.²

The main results of this study are the following. I find that migrant women after migration neither experience a drop in relative employment, nor earn lower hourly wages, compared to the change in relative outcomes of stayers. They do, however, work fewer hours and have a lower relative annual income. The results also suggest that engaging in childcare activities, having a husband in the West or a partner with a higher income contribute to the explanation of this effect, indicating that female migrants in the West seem to substitute some market work with home production, in particular childcare. This negative effect is most likely attributable to the combination of both demand factors on the one hand, such as availability of part-time vacancies in the West, and supply factors on the other hand, such as reduced labor supply due to the childcare activities and availability of childcare institutions.

The rest of the paper is structured as follows. Section 2 describes the data used and presents descriptive evidence. Section 3 outlines econometric methodology and discusses the results. Some explanations are suggested in section 4, section 5 explores migration within western German states, and section 6 provides a robustness analysis. Finally, section 7 presents the conclusions.

²Note, however, that if unobservables are not time-invariant it still leaves the selection bias associated with the non-random selection of movers.

2 Data and descriptive evidence

I use 1990-2001 waves of the eastern sample of the German Socio-Economic Panel survey (GSOEP)³. In the GSOEP easterners are traced if they move to western Germany. Thus, the main advantage of this dataset is that it has both pre- and post-migration information for the same individuals, while the main disadvantage, however, is a small number of observations for movers.

The outcomes of interest are constructed as follows. The total annual individual income is a sum of individual earnings from the main job, second job and self-employment, and the social security benefits (such as unemployment benefits, maternity benefits etc.). The mean income is set to missing only if information on all the components is missing.⁴ Employment is a dummy that equals one if an individual is working, and zero otherwise. Hours per week are reported hours worked per week. Finally, hourly earnings are calculated as monthly earnings (wages and salaries from main job, second job and self-employment) divided by the number of hours worked per week, further divided by 4.3⁵. All financial variables are inflated to 2001 by regional CPIs and are expressed in DM.

The migrants' group includes all persons that experienced migration during 1990-2001, and the stayers' group comprises those who stayed in the East during 1990-2001. I concentrate on working age individuals (18-54 years old) for whom the data on the key variables are not missing⁶. Final sample sizes vary with the dependent variables used and range from 18,126 to 8,984 observations, 425-1,169 of whom belong to the migration group (actual and potential migrants).⁷

³See SOEP Group (2001) for a description of the dataset.

⁴I also exclude the obvious outliers from the sample, i.e. individuals whose average annual income is less than 1,000 DM or greater than 130,000 DM.

⁵I also exclude the outliers, i.e. those earning less than 100 DM or more than 20,000 DM per month.

⁶Individuals in full-time education and military service as well as return migrants are excluded. I also exclude commuters from stayers, since they constitute a specific group, but keep them in the robustness checks. Finally, I keep only those for whom the data is observed for both "before" and "after" periods.

⁷In the sample, around 6 percent are migrants. This number is consistent with the aggregate figures.

Table 1 provides socioeconomic characteristics for migrants and stayers by gender for the periods "before" and "after".⁸ As can be seen from this table, migrants on average are younger and are less likely to be married than stayers both "before" and "after". The proportion of both male and female migrants with university degree is almost the same as that of stayers "before", although this proportion is much higher for migrants after migration, with male migrants on average being more educated than females. Since not all human capital acquired in the East is transferrable to the West, it seems that migrants do invest in their human capital in the West (probably part-time).

Table 2 presents labor force behavior for males and females. The pooled data reveals that there are fewer unemployed among male migrants than among stayers after migration, however this trend does not hold for migrant women, with 95 per cent of males and 75 per cent of females being employed after migration. Large differences between genders exist in part-time work before and after migration: while 7 per cent of female migrants work part-time before move (18 per cent of female stayers and 3 per cent of male migrants do so), the proportion increases to 40 per cent after they move (compared to 18 per cent of female stayers and 0.004 per cent of male migrants). Table 2 shows also occupational distribution of males and females before and after migration. Females, both stayers and migrants before and after the move, tend to be concentrated in the technician and associate professional jobs, the second largest group being service and sales workers. Males are concentrated in craft, construction and related trades occupations, the second largest group being plant and machinery operators. Finally, there seems to exist no descriptive evidence of the downward occupational mobility after migration.

Tables 1 and 2 display also the outcomes of interest. A number of features are worth noting. First, table 1 shows that the annual income of male migrants after migration is

⁸Note that for the analysis below I have to define pre- and post- periods also for stayers. I define 1990-1995 as a "pre-" period, and 1996-2001 as a "post-" period. While this definition is somewhat arbitrary, the main results hold also with different year thresholds.

much higher than their initial income before migration, and is also higher than the income of stayers. The income of females, however, does not follow this trend: in the period after migration, the annual income of female migrants seems to be even lower than the annual income of female stayers. Second, while the differences in employment and hourly wages in all groups are not large before migration (with the exception of employment for males), they become obvious after migration with male migrants working more and female migrants working less than stayers (see table 2). Third, there is a striking difference in hours worked between migrant men and women (table 2): while before migration migrants and stayers have an almost equal amount of hours worked per week (with males on average working roughly 5 hours more than females), after the move, the number of hours increases for migrant men from 46.49 to 46.70, but drops for women from 41.41 to 32.80, and the proportion of migrant women working part-time increases from 7 per cent to 40 per cent. Fourth, both male and female migrants earn higher hourly wages than stayers after migration, and the gender pay gap exists for migrants both before and after migration (table 1). Finally, it is worth noting that overall there exist some systematic differences in outcomes of migrants and stayers even before migration occurs. It is possible that the endogeneity of migration decision generates these differences, or that they are due to differences in observable characteristics between migrants and stayers discussed above. These issues are addressed in the following section.

Before controlling for observed heterogeneity, however, it is also useful to undertake another descriptive exercise and to compare the differences in means between migrants and stayers by gender (a sort of "difference-in-difference-in-difference" approach). Table 3 illustrates these unadjusted estimates. Each panel compares the change in the respective outcome along three dimensions of variation. The first is the comparison of the periods before and after migration, the second difference is between migrants and stayers, and the third one is between men and women. Each cell contains the mean average outcome for

the group labeled on the axes, along with the standard errors. Does the differentiation of the labor force by sex and mobility status operate to the "double disadvantage" of migrant women? The answer appears to be "yes" in terms of annual income, employment and hours worked per week. However, this does not seem to be the case for hourly wages - there is a fall in relative hourly earnings of female migrants compared to the change in relative hourly earnings of stayers, however it is not significant. This descriptive exercise provides some evidence that female migrants face a decrease in certain labour market outcomes relative to male migrants and stayers after they move. However, the causal interpretation of this effect is problematic, since there may be important observable and unobservable characteristics that confound it. The econometric analysis below addresses these issues.

3 Regression framework and estimation results

Table 3 does not control for the observed heterogeneity between the groups, such as human capital and demographics. The regression equation that controls for these observable characteristics has the following form⁹:

$$Y_{i,t} = \beta_1(F_i \times M_i \times A_t) + \beta_2(F_i \times M_i) + \beta_3(F_i \times A_t) + \beta_4(M_i \times A_t) + \beta_5 M_i + \beta_6 F_i + \beta_7 A_t + \delta X_{i,t} + \eta_t + \varepsilon_{i,t} \quad (1)$$

where $Y_{i,t}$ is the outcome variable of individual i in year t , F_i indicates if an individual i is a female, M_i indicates if she is in the migrants' group, A_t is a dummy that equals 1 for the period "after" and is 0 otherwise, $X_{i,t}$ is a vector of control variables including a

⁹This specification is similar to the one in Ichino et al. (2006), where the authors use a "difference-in-difference-in-difference" approach to estimate the effects of age and being displaced on the subsequent labor market outcomes.

constant, η_t are year fixed effects, and $\varepsilon_{i,t}$ is an error term assumed to be uncorrelated with other variables.

The coefficient β_1 on the third-level interaction is the parameter of interest. It captures all variation in labor market outcomes specific to migrants (relative to stayers) to females (relative to males) in the years after migration (relative to before). The second-level interactions control for time-invariant characteristics of the migrant females (β_2), changes over time for all females (β_3), and changes over time for the migrants' group (β_4). Finally, the migrant dummy controls for any time-invariant characteristics of the migrants (β_5), female dummy - for time-invariant characteristics of females (β_6), and "after" dummy - for the time-series changes in outcome.

Tables 4-7 present the estimation results for different outcomes.¹⁰ In all tables column (1) reports the estimates from an OLS regression without controls, column (2) adds standard controls such as age and its square as a proxy for experience, marital status, number of children less than 14 years old, education, blue-collar and public sector employment dummies, year and region fixed effects, and column (3) adds a lagged hourly wage as a proxy for skills. Further columns include additional controls and pre-determined covariates. Finally, in the last columns I also control for individual fixed effects.¹¹

Do female migrants after migration experience a significant income loss relative to males and stayers? The first row of Table 4 presents the estimate of the effect on the annual income. As can be seen from this table, the answer appears to be yes. The effect holds with the addition of controls, and even the fixed effects estimation indicates that female migrants face a significant 24-32 percent drop in relative annual income after they

¹⁰Note that sample size changes when lagged hourly earnings and pre-determined controls are included, however, the main results hold in spite of the changes in composition. To compare the effect, columns with the same number of observations have to be considered.

¹¹Note that if unobservables are not time-invariant and are positively correlated with the probability to move, the estimation results are biased upwards and constitute the "upper bound" of the true effect.

move. Other coefficients are also worth noting. The effect for all migrants after they move relative to stayers is positive, although is significant only when pre-determined controls are included. The same holds for the effect for all females in the period "after". Neither the time-invariant effect for female migrants nor the migration dummy is statistically significant. Females receive lower annual income than males. The coefficients on the other covariates are as expected¹²: experience has a concave profile, university graduates earn higher annual income, those with a general schooling degree receive lower income relative to apprentices, public sector employees have higher income, and the coefficient on the lagged wage is positive and significant. Thus, even after having controlled for skills and individual fixed effects, the relative effect for female migrants on annual income remains negative and highly significant.

Table 5 shows analogous estimates for employment probabilities. In contrast to annual income, female migrants do not face significantly lower relative employment probability after they move. The effect changes sign, however remains statistically insignificant across all specifications, with the exception of OLS(1) and OLS(3). The positive and significant effect of migrants after migration disappears in the fixed effects estimation, but all females in the period "after" seem to have higher relative probability of being employed. Employment prospects seem to worsen over time, and both coefficients on time-invariant migrant and female dummies are negative and significant in the majority of specifications. Coefficients on age, marital status, number of children, schooling and lagged wage have the expected signs.

Table 6 presents the results for hours worked per week.¹³ If being a female migrant does not influence the relative employment outcome after migration, it does appear to influence the relative weekly hours worked. The coefficient on the third-level interaction is

¹²available upon request.

¹³These are hours worked conditional on being employed. Note that due to the high labor force participation of females in East Germany, the selection into the labor force problem can be ignored here. Labor supply and wages can be modelled jointly, however, this is beyond the scope of this paper.

negative and statistically significant in all model specifications. Moreover, the magnitude of the effect diminishes only slightly with the inclusion of additional controls: female migrants experience a 22-27 percent decrease in relative weekly hours worked after they move. Moreover, the coefficient on the second-level interaction $female \times after$ is also negative and statistically significant, indicating that in the period "after" all females face a drop in their weekly working hours. Finally, a negative and significant coefficient on female dummy indicates that, in general, females work less hours relative to males. The coefficients on all other controls have the expected signs. Overall, even after having controlled for skills and individual fixed effects, the relative effect for female migrants on hours worked remains negative and highly significant.¹⁴

Finally, Table 7 presents estimation results for hourly earnings. Estimates are largely statistically insignificant (with the only exception of OLS(2)). Coefficients on the second-level interaction $female \times after$ is positive and significant in all specifications but one, indicating a 3-6 per cent increase in the relative hourly wages of females in the period "after". The gender wage gap, however, still exists with females earning 7-15 per cent on average less than males. The remaining coefficients are as expected: experience has a concave profile, university graduates earn more and those with general schooling degree earn less relative to apprentices; blue-collar workers earn less than white-collar employees and public sector employees have higher hourly earnings.

Overall, the regression analysis indicates that female migrants after migration experience a decrease in their relative annual incomes and hours worked, but not in relative employment and hourly earnings.

¹⁴I also estimated a Tobit model for the weekly hours worked (not reported). The results were the same - the coefficient on the third-level interaction was negative and highly significant.

4 Searching for explanations

So far, we have established that compared to stayers and male migrants, female migrants in the West face a drop in their relative weekly working hours and annual incomes. Migrant women seem to switch to part-time work after migration, and thus receive a lower relative annual income. But is this effect equally distributed across all female migrants? Is it a voluntary choice or a disadvantage, preferences or demand? It is difficult to disentangle true preferences, and the analysis below attempts to at least suggest the possible answers.

Table 8 shows the effect for different groups indicated in each row. For example, when the effect is estimated for a subpopulation of married individuals (see the first panel), it is still negative and significant. The same holds for the subpopulation of married with children (second panel), married before migration (third panel) and for those with children before migration (fourth panel). On the other hand, when the estimations are run for a subpopulation of singles, the effect becomes insignificant. Thus, fertility and marriage constitute potential explanation of the negative effect, and the potential endogeneity of fertility and marital status does not seem to be a problem. In addition, I also reestimated all the models conditioning on having worked before migration, and the negative effect for hours and income was still present.

These findings indicate several interesting facts. First, the results in Table 8 suggest that the effect is heterogenous across different demographic groups. Second, the group that reduces relative supply of working hours (and thus receives lower annual income) is the one consisting of married female migrants after migration (with or without children). For singles, being a female and a migrant influences their relative labor market outcomes insignificantly. Thus, family background matters, and merits a more detailed exploration.

Could family characteristics explain the negative effect of being a female migrant on relative annual income and weekly hours? Table 9 provides an answer. In this table, I have reestimated the baseline model in equation (1) including the additional interactions

of the effect ($F_i \times M_i \times A_t$) with other variables.¹⁵ If it is true that these variables reinforce the effect for female migrants, we should see a significant coefficient on these additional interactions and β_1 should decrease. Indeed, the first panel of Table 9 indicates that the marital status contributes to the explanation. The interaction with the spouse dummy is negative and statistically significant for both annual income and hours worked, while the ($F_i \times M_i \times A_t$) dummy becomes insignificant. Having children younger than 14 years old and being a female migrant reduces the relative annual income, but not hours worked after migration. On the other hand, there seems to be no additional effect from already having a spouse before migration.¹⁶ While having children *per se* does not seem to contribute to the explanation of the effect on hours supplied, spending time for childcare contributes significantly, and the ($F_i \times M_i \times A_t$) dummy becomes again either insignificant (in income equation) or smaller in magnitude (in hours equation). Finally, the income of the other household members also contributes significantly to the explanation of this effect.¹⁷ A higher partners' (or other household members') income significantly reduces the relative annual income and hours worked for female migrants, while the ($F_i \times M_i \times A_t$) effect becomes insignificant or smaller.

This exercise suggests that family background is indeed a potential reason behind a negative relative effect for migrant women: having a husband, having a higher income of other members of the household, and spending time for childcare is associated with working less hours and receiving a lower annual individual income. In contrast, being married already before migration does not have any additional effect on the outcomes of interest. Migrant women who have a spouse in the West or who live in "rich" households

¹⁵similar to Ichino et al. (2006).

¹⁶"Married before" dummy is equal to 1 if an individual was married one year before and in the year of migration.

¹⁷This variable is constructed as the difference between monthly household income and monthly income of an individual i in year t . I also experimented with individual partner's income, although the sample size dropped significantly. While the results were qualitatively the same, my preferred variable is the income of other household members.

reduce their supply of labor market hours. Since the wage rate increases after migration (see Table 1), it appears that the income effect dominates the substitution effect for these females. After migration they may switch either to more leisure, or to household production including childcare. Indeed, since spending time for childcare has a separate negative effect for these females, it implies that childcare is another potential explanation. Since the availability of childcare facilities is better in East Germany than in the West¹⁸, and since relatives, who could potentially contribute to the childcare activities, are usually left in the East, at the margin, female migrants substitute their market work in the West with childcare. The opportunity costs of home production seem to be higher for female migrants. On the other hand, there may be also not enough full-time jobs available for females in the West, or the part-time jobs are now more available. Overall, the negative relative effect on working hours and incomes is most likely to be due to the combination of both supply and demand factors.

5 The effect in western Germany

Is the double negative effect found above specific for transition economies or does it also hold in other contexts? This section aims at answering this questions and presents the analysis only for western German migration.

In order to be consistent with the analysis above, I have tried to follow closely definitions and sample selection rules. In particular, the four labor market outcomes of interest are defined as in Section 2, and the time period is the same: 1990-2001. The definition of migrants, however, changes, since here migrants include those individuals who change their residence from one western German state ("*Bundesland*") to another over 1990-2001.

¹⁸For example, in 1990 there were 54.2 child care places available per hundred children under the age of three in East Germany, while there were only 1.8 such places in the West. The number has dropped for the East and increased slightly for the West, but there are still significant differences with 36.3 places and 2.8 places in 1998, respectively (see Wrohlich, 2004).

I again concentrate on working age individuals (18-54 years old) for whom the data on the key variables are not missing and drop return and multiple movers¹⁹ The final sample size ranges from 26,378 to 32,377 observations.

Table 10 presents estimation results. As can be seen from this table, female inter-state migrants in western Germany experience either insignificant or a positive relative effect. A fixed-effects estimation suggests that, compared to migrant males and stayers, females are more likely to work and have both higher working hours and higher annual incomes after they move, relative to before. They do not face a significant relative effect on their hourly wages, though.

Thus, it seems that the negative relative effect on the supply of working hours and annual incomes is specific for women who move between regions with different institutions, such as childcare facilities and availability of part-time jobs.

6 Sensitivity analysis

In addition to changes in the specification of the baseline model (see Tables 4-7), I also performed several robustness checks. Table 11 shows this sensitivity analysis.

First, I have controlled for additional household-level characteristics, such as total household size and household income (panel A). The fixed effects estimation results show that migrant females after migration experience a 24 percent drop both in their relative annual income and in relative weekly hours worked. Second, in panel B, I have included detailed controls for occupation (nine occupational groups according to ISCO88 definition). Again, the results were not affected: while the negative effect on income is 25

¹⁹Again, individuals in full-time education and military service as well as return migrants are excluded. However, inter-state commuters are included in the "stayers" group, since it is not possible to identify them in the data for West Germany (nevertheless, this is comparable with the robustness checks for East-West analysis below). Finally, I keep only those for whom the data is observed for both "before" and "after" periods.

percent in fixed effects estimation, the effect on hours equals to 26 percent. Finally, to check how robust the results are to different definitions of the control group, in panel C, I have retained commuters in stayers' group. These are individuals who reside in the East and work in the West, and for whom the impact of family background is likely to be more similar to stayers than to migrants. The negative relative effect for annual income is now 25 percent, and the effect for weekly hours is equal to 23 percent. In all panels the relative effects for employment and hourly earnings are insignificant. Overall, the main results remain robust to changes in the definition of the control group and to the inclusion of additional controls: female East-West migrants experience a drop in their relative hours worked and annual incomes after migration.

7 Conclusions

This paper documents the relative labor market performance of female migrants from East to West Germany over 1990-2001, compared to migrant males and stayers. I use a sort of "difference-in-difference-in-difference" methodology and panel data techniques to purge away time invariant unobservable confounders.

The main results indicate that female migrants experience a decrease in their relative annual incomes and hours worked after migration, but not in the relative employment probabilities or hourly wages, compared to the changes in the relative outcomes of stayers. This is consistent with standard labor supply model, and for these females the income effect dominates the substitution effect. Moreover, the negative effect is heterogeneous across different demographic groups and is not present for single female migrants. The family background thus serves as a potential explanation.

The results also suggest that having a husband in the West, having a higher partner's income, having children and spending time for childcare indeed contribute to the expla-

nation of this negative effect. Thus, female migrants in the West seem to substitute some market work with home production, in particular with childcare.

Overall, the negative relative effect on working hours and incomes is specific to the transition economy structure of East Germany, and is most likely to be attributable to the combination of both supply and demand factors. This effect remains robust to changes in specification and in the sample.

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8 Appendix

Table 1: Socioeconomic characteristics by gender and mobility status "before" and "after"

| | Males | | Females | |
|--|------------|------------|------------|------------|
| | Stayers | Migrants | Stayers | Migrants |
| | before | | | |
| Age | 36.12 | 31.79 | 35.91 | 31.01 |
| | (8.84) | (9.39) | (8.61) | (9.33) |
| Married | 0.71 | 0.51 | 0.77 | 0.49 |
| Number of kids<14 y.o. | 1.07 | 0.83 | 1.13 | 1.03 |
| | (1.00) | (0.98) | (0.99) | (0.94) |
| General school | 0.07 | 0.16 | 0.08 | 0.20 |
| University | 0.12 | 0.15 | 0.08 | 0.09 |
| Other technical or vocational training | 0.16 | 0.08 | 0.19 | 0.13 |
| Observations | [3670] | [196] | [4562] | [304] |
| Annual income | 39209.18 | 37921.29 | 29799.78 | 29832.13 |
| | (17643.66) | (21807.53) | (17010.30) | (22217.62) |
| Observations | [3419] | [184] | [4206] | [274] |
| Hourly earnings | 19.92 | 19.01 | 19.02 | 18.12 |
| | (11.95) | (11.42) | (14.97) | (12.23) |
| Observations | [2894] | [142] | [3008] | [184] |
| | after | | | |
| Age | 39.61 | 37.55 | 39.63 | 35.35 |
| | (8.72) | (8.60) | (8.50) | (9.22) |
| Married | 0.67 | 0.64 | 0.73 | 0.69 |
| Number of kids<14 y.o. | 0.82 | 0.84 | 0.88 | 0.91 |
| | (0.94) | (0.96) | (0.93) | (0.90) |
| General school | 0.04 | 0.04 | 0.05 | 0.08 |
| University | 0.12 | 0.25 | 0.08 | 0.14 |
| Other technical or vocational training | 0.25 | 0.18 | 0.28 | 0.24 |
| Observations | [3866] | [259] | [4859] | [410] |
| Annual income | 40096.95 | 57016.90 | 31943.42 | 30998.65 |
| | (19357.46) | (25349.15) | (19566.65) | (22053.43) |
| Observations | [3339] | [228] | [4162] | [329] |
| Hourly earnings | 20.88 | 27.87 | 20.96 | 25.37 |
| | (16.30) | (14.34) | (19.14) | (24.90) |
| Observations | [2765] | [215] | [3080] | [270] |

Note: standard deviations in parentheses. "Before" stands for a period before moving West for migrants and before 1996 for stayers, "after" stands for a period after individual move for migrants and after 1996 for stayers. See text for definitions. Annual income is a sum of labour income and social security benefits. Hourly wage includes wages and salaries from main job, second job and self-employment. All financial variables are inflated to 2001 by regional CPIs and expressed in DM. Reference categories: single, apprenticeship.

Table 2: Labour force behavior by gender and mobility status "before" and "after"

| | Males | | Females | |
|-----------------------------------|---------|----------|---------|----------|
| | Stayers | Migrants | Stayers | Migrants |
| | before | | | |
| Employed | 0.90 | 0.84 | 0.76 | 0.73 |
| Blue collar | 0.50 | 0.46 | 0.17 | 0.11 |
| Observations | [3670] | [196] | [4562] | [304] |
| Part-time work | 0.04 | 0.03 | 0.18 | 0.07 |
| Hours per week | 45.19 | 46.49 | 40.01 | 41.41 |
| | (9.14) | (9.21) | (9.02) | (7.19) |
| Observations | [3096] | [151] | [3247] | [200] |
| Occupation (in %): | | | | |
| managers | 7.08 | 7.05 | 3.66 | 2.79 |
| professionals | 12.10 | 5.13 | 13.80 | 10.70 |
| technicians, assoc. professionals | 9.63 | 7.69 | 30.37 | 40.47 |
| clerks | 2.99 | 5.13 | 17.69 | 17.67 |
| service, sales workers | 4.77 | 10.26 | 18.61 | 17.21 |
| agricultural, fishery workers | 1.79 | 1.92 | 1.59 | 1.40 |
| craft, construction workers | 39.55 | 23.72 | 5.25 | 3.72 |
| machinery operators | 14.71 | 23.72 | 3.30 | 0 |
| elementary occupations | 7.17 | 7.69 | 5.72 | 6.05 |
| armed forces | 0.22 | 7.69 | 0 | 0 |
| | after | | | |
| Employed | 0.84 | 0.95 | 0.76 | 0.75 |
| Blue collar | 0.46 | 0.55 | 0.17 | 0.13 |
| Observations | [3866] | [259] | [4562] | [410] |
| Part-time work | 0.03 | 0.004 | 0.18 | 0.40 |
| Hours per week | 45.89 | 46.70 | 40.00 | 32.80 |
| | (9.47) | (9.35) | (9.02) | (11.90) |
| Observations | [3153] | [239] | [3247] | [301] |
| Occupation (in %) | | | | |
| managers | 7.29 | 5.42 | 3.66 | 2.07 |
| professionals | 12.78 | 14.17 | 13.80 | 11.72 |
| technicians, assoc. professionals | 10.48 | 11.67 | 30.37 | 37.24 |
| clerks | 3.66 | 3.75 | 17.69 | 18.97 |
| service, sales workers | 4.51 | 1.67 | 18.61 | 20.00 |
| agricultural, fishery workers | 2.56 | 0.42 | 1.59 | 0 |
| craft, construction workers | 38.37 | 29.58 | 5.25 | 1.03 |
| machinery operators | 12.31 | 22.92 | 3.30 | 4.14 |
| elementary occupations | 8.02 | 9.17 | 5.72 | 4.83 |
| armed forces | 0.03 | 1.25 | 0 | 0 |

Note: standard deviations in parentheses. See footnote of Table 1.

Table 3: Differences in labour market outcomes by gender and mobility status

| | Before | | After | | After - before | |
|----------|-------------------------|----------------------|---------------------|----------------------|----------------------|----------------------|
| | Males (1) | Females (2) | Males (3) | Females (4) | Males (5) | Females (6) |
| | log total annual income | | | | | |
| Stayers | 10.450 (0.010) | 10.104 (0.011) | 10.456 (0.010) | 10.150 (0.011) | 0.006 (0.015) | 0.047*** (0.016) |
| Migrants | 10.333 (0.054) | 10.006 (0.050) | 10.840 (0.034) | 9.994 (0.052) | 0.508*** (0.062) | -0.011 (0.073) |
| M-S | -0.117*** (0.043) | -0.098** (0.044) | 0.385*** (0.041) | -0.156*** (0.043) | 0.502*** (0.064) | -0.058 (0.075) |
| DDD | -0.560*** (0.099) | | | | | |
| | employment | | | | | |
| Stayers | 0.896 (0.005) | 0.759 (0.006) | 0.845 (0.006) | 0.743 (0.006) | -0.051*** (0.008) | -0.016* (0.009) |
| Migrants | 0.842 (0.026) | 0.727 (0.026) | 0.946 (0.014) | 0.751 (0.021) | 0.104*** (0.028) | 0.024 (0.033) |
| M-S | -0.054** (0.023) | -0.032 (0.025) | 0.101*** (0.023) | 0.008 (0.022) | 0.155*** (0.029) | 0.040 (0.034) |
| DDD | -0.115** (0.045) | | | | | |
| | log weekly hours | | | | | |
| Stayers | 3.787 (0.004) | 3.656 (0.005) | 3.800 (0.005) | 3.625 (0.006) | 0.013** (0.006) | -0.030*** (0.008) |
| Migrants | 3.818 (0.017) | 3.701 (0.017) | 3.826 (0.012) | 3.394 (0.028) | 0.008 (0.020) | -0.307*** (0.037) |
| M-S | 0.032 (0.020) | 0.046** (0.021) | 0.026 (0.017) | -0.231*** (0.022) | -0.005 (0.021) | -0.277*** (0.038) |
| DDD | -0.272*** (0.043) | | | | | |
| | log hourly earnings | | | | | |
| Stayers | 2.879 (0.009) | 2.806 (0.009) | 2.907 (0.009) | 2.882 (0.010) | 0.029** (0.013) | 0.076*** (0.014) |
| Migrants | 2.808 (0.045) | 2.666 (0.054) | 3.217 (0.032) | 2.993 (0.039) | 0.409*** (0.054) | 0.327*** (0.066) |
| M-S | -0.071 (0.041) | -0.140*** (0.040) | 0.310*** (0.034) | 0.111*** (0.035) | 0.380*** (0.055) | 0.251*** (0.067) |
| DDD | -0.129 (0.087) | | | | | |

Note: standard errors in parenthesis. Cells contain means of the respective labour market outcome. Before / after and migrants and stayers are defined in the text. DDD is the "difference-in-difference" for females minus that for males. *** significant at 1% level, ** significant at 5% level, * significant at 10% level.

Table 4: Estimation results: annual income

| | OLS (1) | OLS (2) | OLS (3) | OLS (4) | FE (5) | FE (6) |
|-------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| female*migr*after | -0.443*** (0.095) | -0.474*** (0.082) | -0.456*** (0.097) | -0.513*** (0.111) | -0.239*** (0.074) | -0.316*** (0.099) |
| migr*after | 0.057 (0.544) | 0.297 (0.409) | 0.681*** (0.042) | 0.983*** (0.208) | 0.342*** (0.119) | 0.853*** (0.325) |
| female*after | 0.019 (0.019) | 0.029 (0.016) | -0.010 (0.016) | 0.061*** (0.019) | 0.017 (0.012) | 0.050*** (0.014) |
| female*migr | -0.089 (0.071) | -0.024 (0.059) | 0.004 (0.057) | 0.038 (0.078) | | |
| after | 0.127** (0.057) | 0.006 (0.050) | -0.062 (0.053) | 0.005 (0.063) | -0.057* (0.035) | -0.140*** (0.050) |
| migr | -0.021 (0.047) | 0.057 (0.039) | 0.033 (0.036) | -0.069 (0.057) | | |
| female | -0.225*** (0.013) | -0.310*** (0.012) | -0.240*** (0.011) | -0.363*** (0.013) | | |
| Controls | No | Yes | Yes | No | Yes | No |
| Controls at t=0 | No | No | No | Yes | No | Yes |
| Earnings in 1991 | No | No | Yes | No | No | No |
| Year dummies | Yes | Yes | Yes | Yes | Yes | Yes |
| Region dummies | Yes | Yes | Yes | Yes | Yes | Yes |
| R ² | 0.08 | 0.31 | 0.37 | 0.25 | - | - |
| Observations | 13119 | 13119 | 9244 | 12854 | 13119 | 12854 |

Note: robust standard errors are given in parenthesis. *** significant at 1% level, ** significant at 5% level, * significant at 10% level. The dependent variable is the log of individual total annual income (labour income plus social security benefits). Additional controls include age and its square, marital status, number of children less than 14 years old, education, blue-collar and public sector employment dummies, year and region fixed effects. In fixed effects (FE) estimation time-invariant covariates are dropped.

Table 5: Estimation results: employment

| | OLS(1) | OLS(2) | OLS(3) | OLS(4) | OLS(5) | FE(6) | FE(7) | FE(8) |
|-------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| female*migr*after | -0.106** (0.046) | -0.069 (0.045) | -0.164*** (0.053) | 0.032 (0.045) | -0.082* (0.046) | 0.034 (0.049) | 0.039 (0.049) | -0.014 (0.051) |
| migr*after | 0.431*** (0.040) | 0.526*** (0.039) | 0.094*** (0.030) | 0.083 (0.075) | 0.520*** (0.041) | 0.109 (0.085) | -0.087 (0.116) | 0.093 (0.059) |
| female*after | 0.035*** (0.012) | 0.021** (0.011) | -0.003 (0.011) | 0.030*** (0.011) | 0.031*** (0.011) | 0.008 (0.008) | 0.019** (0.009) | 0.030*** (0.010) |
| female*migr | 0.012 (0.037) | 0.033 (0.036) | 0.105*** (0.037) | 0.032 (0.036) | 0.025 (0.037) | | | |
| after | -0.107*** (0.028) | -0.123*** (0.027) | -0.060* (0.034) | -0.100*** (0.027) | -0.112*** (0.027) | -0.088*** (0.024) | -0.077*** (0.026) | -0.100*** (0.026) |
| migr | -0.075*** (0.027) | -0.065*** (0.026) | -0.072*** (0.026) | -0.047* (0.026) | -0.084*** (0.027) | | | |
| female | -0.137*** (0.008) | -0.030*** (0.008) | 0.008 (0.008) | -0.006 (0.008) | -0.100*** (0.008) | | | |
| Controls | No | Yes | Yes | Yes | No | Yes | Yes | No |
| Controls at t=0 | No | No | No | No | Yes | No | No | Yes |
| Earnings in 1991 | No | No | Yes | No | No | No | No | No |
| Others' income | No | No | No | Yes | No | No | Yes | No |
| Year dummies | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Region dummies | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| R ² | 0.03 | 0.16 | 0.15 | 0.17 | 0.07 | - | - | - |
| Observations | 18126 | 18126 | 11464 | 15990 | 18126 | 18126 | 15990 | 18126 |

Note: robust standard errors are given in parenthesis. *** significant at 1% level, ** significant at 5% level, * significant at 10% level. The dependent variable is a binary employment status of an individual. Additional controls include age and its square, marital status, number of children less than 14 years old, education, blue-collar dummies, year and region fixed effects. In fixed effects (FE) estimation time-invariant covariates are dropped.

Table 6: Estimation results: weekly hours

| | OLS(1) | OLS(2) | OLS(3) | OLS(4) | OLS(5) | FE(6) | FE(7) | FE(8) |
|-------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| female*migr*after | -0.261*** (0.041) | -0.261*** (0.041) | -0.266*** (0.053) | -0.220*** (0.041) | -0.233*** (0.047) | -0.245*** (0.053) | -0.246*** (0.056) | -0.264*** (0.055) |
| migr*after | 0.329*** (0.038) | 0.310*** (0.039) | -0.071 (0.053) | 0.219*** (0.041) | 0.266*** (0.044) | 0.152** (0.077) | -0.143 (0.105) | 0.144** (0.074) |
| female*after | -0.040*** (0.010) | -0.040*** (0.010) | -0.020** (0.010) | -0.043*** (0.010) | -0.025*** (0.009) | -0.031*** (0.008) | -0.032*** (0.009) | -0.020** (0.008) |
| female*migr | 0.003 (0.025) | 0.001 (0.025) | -0.015 (0.034) | -0.008 (0.026) | 0.006 (0.027) | | | |
| after | 0.006 (0.024) | 0.006 (0.024) | 0.028 (0.029) | 0.007 (0.024) | 0.015 (0.022) | 0.004 (0.021) | 0.002 (0.024) | 0.013 (0.024) |
| migr | 0.036* (0.020) | 0.044** (0.020) | 0.045* (0.025) | 0.046** (0.021) | 0.014 (0.021) | | | |
| female | -0.130*** (0.007) | -0.145*** (0.007) | -0.145*** (0.008) | -0.129*** (0.007) | -0.140*** (0.007) | | | |
| Controls | No | Yes | Yes | Yes | No | Yes | Yes | No |
| Controls at t=0 | No | No | No | No | Yes | No | No | Yes |
| Earnings in 1991 | No | No | Yes | No | No | No | No | No |
| Others' income | No | No | No | Yes | No | No | Yes | No |
| Year dummies | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Region dummies | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| R ² | 0.10 | 0.12 | 0.12 | 0.14 | 0.11 | - | - | - |
| Observations | 13729 | 13729 | 9729 | 12475 | 12130 | 13729 | 12475 | 12130 |

Note: robust standard errors are given in parenthesis. *** significant at 1% level, ** significant at 5% level, * significant at 10% level. The dependent variable is the log of hours worked per week. Additional controls include age and its square, marital status, number of children less than 14 years old, education, blue-collar and public sector employment dummies, year and region fixed effects. In fixed effects (FE) estimation time-invariant covariates are dropped.

Table 7: Estimation results: hourly earnings

| | OLS(1) | OLS(2) | OLS(3) | OLS(4) | OLS(5) | FE(6) | FE(7) | FE(8) |
|-------------------|----------------------|----------------------|----------------------|----------------------|----------------------|---------------------|---------------------|----------------------|
| female*migr*after | -0.131 (0.087) | -0.179** (0.073) | -0.064 (0.084) | -0.094 (0.072) | -0.062 (0.086) | 0.036 (0.069) | 0.051 (0.071) | 0.082 (0.076) |
| migr*after | 0.050 (0.390) | 0.314 (0.270) | 0.759*** (0.052) | -0.302*** (0.065) | -0.242*** (0.076) | 0.434*** (0.147) | 0.345** (0.151) | 0.661*** (0.215) |
| female*after | 0.044** (0.018) | 0.053*** (0.016) | 0.014 (0.016) | 0.061*** (0.016) | 0.058*** (0.017) | 0.034*** (0.013) | 0.044*** (0.013) | 0.048*** (0.013) |
| female*migr | -0.075 (0.069) | 0.019 (0.056) | 0.048 (0.052) | 0.017 (0.054) | 0.002 (0.066) | | | |
| after | 0.153*** (0.051) | 0.042 (0.043) | -0.067 (0.049) | 0.040 (0.043) | 0.053 (0.048) | -0.023 (0.040) | -0.017 (0.039) | -0.114*** (0.040) |
| migr | -0.042 (0.045) | 0.015 (0.040) | -0.001 (0.035) | 0.027 (0.039) | -0.009 (0.048) | | | |
| female | -0.073*** (0.013) | -0.151*** (0.012) | -0.090*** (0.011) | -0.116*** (0.012) | -0.157*** (0.013) | | | |
| Controls | No | Yes | Yes | Yes | No | Yes | Yes | No |
| Controls at t=0 | No | No | No | No | Yes | No | No | Yes |
| Earnings in 1991 | No | No | Yes | No | No | No | No | No |
| Others' income | No | No | No | Yes | No | No | Yes | No |
| Year dummies | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Region dummies | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| R ² | 0.04 | 0.27 | 0.37 | 0.31 | 0.25 | - | - | - |
| Observations | 12461 | 12461 | 8984 | 12279 | 11088 | 12461 | 12279 | 11088 |

Note: robust standard errors are given in parenthesis. *** significant at 1% level, ** significant at 5% level, * significant at 10% level. The dependent variable is the log of hourly earnings. Additional controls include age and its square, marital status, number of children less than 14 years old, education, blue-collar and public sector employment dummies, year and region fixed effects. In fixed effects (FE) estimation time-invariant covariates are dropped.

Table 8: Heterogeneity of the effect

| | Annual income | | Weekly hours | |
|--------------------------|----------------------|----------------------|----------------------|----------------------|
| | OLS | FE | OLS | FE |
| Married | -0.669*** (0.096) | -0.150** (0.076) | -0.227*** (0.054) | -0.265*** (0.070) |
| | [9605] | | [9137] | |
| Married with children | -0.910*** (0.119) | -0.192** (0.101) | -0.318*** (0.060) | -0.142** (0.073) |
| | [6679] | | [6369] | |
| Married before | -0.453*** (0.106) | -0.157** (0.082) | -0.160*** (0.064) | -0.265*** (0.087) |
| | [9459] | | [8856] | |
| With children before | -0.615*** (0.116) | -0.227*** (0.090) | -0.243*** (0.066) | -0.235*** (0.064) |
| | [8246] | | [7710] | |
| Single | -0.173 (0.143) | -0.097 (0.132) | -0.097 (0.074) | -0.042 (0.065) |
| | [3514] | | [3338] | |
| Worked before | -0.392*** (0.090) | -0.302*** (0.079) | -0.227*** (0.048) | -0.266*** (0.056) |
| | [11752] | | [12195] | |

Note: robust standard errors are given in parenthesis, sample size - in brackets. *** significant at 1% level, ** significant at 5% level, * significant at 10% level. Only the coefficient on the third-level interaction in equation (1) is reported. Rows define the subpopulations for which the model is estimated. Additional controls in OLS include age and its square, university degree, general schooling degree, vocational training (reference-apprenticeship), blue-collar worker, public sector employee, year and region fixed effects (as well as others' income in the equation for weekly hours). In fixed effects estimation (FE) time invariant covariates are dropped.

Table 9: The effect of additional interactions

| | Annual income | | Weekly hours | |
|---------------------------|-------------------------|-------------------------|-------------------------|-------------------------|
| | OLS | FE | OLS | FE |
| married*f*m*a | -0.425*** (0.101) | -0.319*** (0.119) | -0.171*** (0.063) | -0.194** (0.084) |
| female*migr*after | -0.181* (0.106) | -0.048 (0.105) | -0.102 (0.064) | -0.129 (0.080) |
| kids*f*m*a | -0.526*** (0.099) | -0.260** (0.111) | -0.184*** (0.060) | -0.078 (0.074) |
| female*migrant*after | -0.227*** (0.088) | -0.131 (0.083) | -0.133*** (0.050) | -0.213*** (0.067) |
| married before*f*m*a | -0.012 (0.096) | 0.108 (0.116) | 0.115*** (0.045) | 0.071 (0.083) |
| female*migrant*after | -0.473*** (0.097) | -0.280*** (0.099) | -0.248*** (0.046) | -0.266*** (0.065) |
| Observations | 13119 | | 12475 | |
| hours for childcare*f*m*a | -0.081*** (0.019) | -0.072*** (0.022) | -0.039*** (0.008) | -0.037*** (0.010) |
| female*migrant*after | -0.314*** (0.094) | -0.099 (0.096) | -0.121*** (0.046) | -0.168*** (0.068) |
| Observations | 10994 | | 10566 | |
| others' inc*f*m*a | -0.0002*** (0.00003) | -0.0002*** (0.00004) | -0.0001*** (0.00001) | -0.00004** (0.00002) |
| female*migrant*after | 0.001 (0.086) | 0.024 (0.089) | -0.056 (0.046) | -0.178*** (0.062) |
| Observations | 12937 | | 12475 | |

Note: robust standard errors are given in parenthesis. *** significant at 1%, ** significant at 5%, *significant at 10% level. "f*m*a" stands for the third-level interaction "female*migrant*after". "Hours for childcare" are reported hours spent per weekday on childcare; "others' income" stands for monthly income of other members of the household. Controls include age and its square, married, number of kids less than 14 years old, university degree, general schooling degree, vocational training (reference-apprenticeship), blue-collar worker, public sector employee, year and region fixed effects (as well as others' income in the equation for weekly hours). In fixed effects estimation (FE) time invariant covariates are dropped.

Table 10: The effect in western Germany

| | Annual income | | Employment | | Weekly hours | | Hourly earnings | |
|-------------------|----------------------|---------------------|----------------------|--------------------|----------------------|---------------------|----------------------|-------------------|
| | OLS | FE | OLS | FE | OLS | FE | OLS | FE |
| female*migr*after | -0.074 (0.067) | 0.247** (0.119) | -0.032 (0.030) | -0.074 (0.064) | 0.069* (0.038) | 0.212*** (0.072) | -0.061 (0.053) | -0.136 (0.085) |
| migr*after | -0.010 (0.048) | 0.052 (0.101) | 0.023 (0.021) | 0.006 (0.052) | -0.002 (0.025) | -0.013 (0.065) | -0.114*** (0.041) | 0.007 (0.075) |
| female*after | -0.063*** (0.016) | -0.019* (0.011) | 0.016** (0.008) | 0.014** (0.006) | -0.036*** (0.010) | -0.010 (0.008) | 0.008 (0.014) | 0.004 (0.012) |
| female*migr | 0.263*** (0.059) | | 0.015 (0.026) | | 0.033 (0.033) | | 0.058 (0.045) | |
| after | -0.007 (0.028) | -0.095** (0.048) | -0.052*** (0.013) | 0.005 (0.026) | -0.003 (0.016) | -0.039 (0.026) | 0.003 (0.023) | -0.011 (0.047) |
| migr | -0.070* (0.040) | | -0.030* (0.017) | | -0.004 (0.019) | | 0.040 (0.032) | |
| female | -0.627*** (0.011) | | -0.043*** (0.005) | | -0.276*** (0.008) | | -0.151*** (0.010) | |
| R ² | 0.30 | - | 0.14 | - | 0.27 | - | 0.22 | - |
| Observations | 28111 | | 32377 | | 26766 | | 26378 | |

Note: robust standard errors are given in parenthesis. *** significant at 1%, ** significant at 5%, *significant at 10% level. Controls include age and its square, married, number of kids less than 14 years old, university degree, general schooling degree, vocational training (reference - apprenticeship), blue-collar worker, public sector employee (not for employment equation), year and region fixed effects (as well as others' income in equations for employment, hours and hourly wages). In fixed effects estimation (FE) time invariant covariates are dropped.

Table 11: Additional robustness checks

| | Annual income | | Employment | | Weekly hours | | Hourly earnings | |
|----------------------|--|-----------|------------|---------|--------------|-----------|-----------------|---------|
| | OLS | FE | OLS | FE | OLS | FE | OLS | FE |
| | A: controlling for household characteristics | | | | | | | |
| female*migrant*after | -0.456*** | -0.239*** | -0.062 | 0.030 | -0.255*** | -0.240*** | -0.164** | 0.028 |
| | (0.080) | (0.073) | (0.044) | (0.048) | (0.041) | (0.052) | (0.071) | (0.069) |
| Observations | 12937 | | 17683 | | 13451 | | 12302 | |
| | B: controlling for occupations | | | | | | | |
| female*migrant*after | -0.441*** | -0.236*** | | | -0.238*** | -0.259*** | -0.059 | 0.074 |
| | (0.080) | (0.072) | | | (0.042) | (0.057) | (0.070) | (0.072) |
| Observations | 12842 | | | | 12215 | | 12027 | |
| | C: retaining commuters in stayers' group | | | | | | | |
| female*migrant*after | -0.474*** | -0.239*** | 0.032 | 0.039 | -0.220*** | -0.246*** | -0.094 | 0.051 |
| | (0.082) | (0.074) | (0.045) | (0.049) | (0.041) | (0.056) | (0.072) | (0.071) |
| Observations | 13119 | | 15990 | | 12475 | | 12279 | |

Note: robust standard errors are given in parenthesis. *** significant at 1%, ** significant at 5%, *significant at 10% level. Only the coefficient on the third-level interaction is reported. Controls include age and its square, married, number of kids less than 14 years old, university degree, general schooling degree, vocational training (reference - apprenticeship), blue-collar worker, public sector employee (not for employment equation), year and region fixed effects (as well as others' income in equations for employment, hours and hourly wages in panels B and C). In fixed effects estimation (FE) time invariant covariates are dropped. In panel A household characteristics include household size and log of monthly household income, and the number of children is dropped. In panel B nine major occupational groups (ISCO88) are included (reference - elementary occupations), and blue collar and public sector dummies are dropped. In panel C commuters are included in stayers' group.

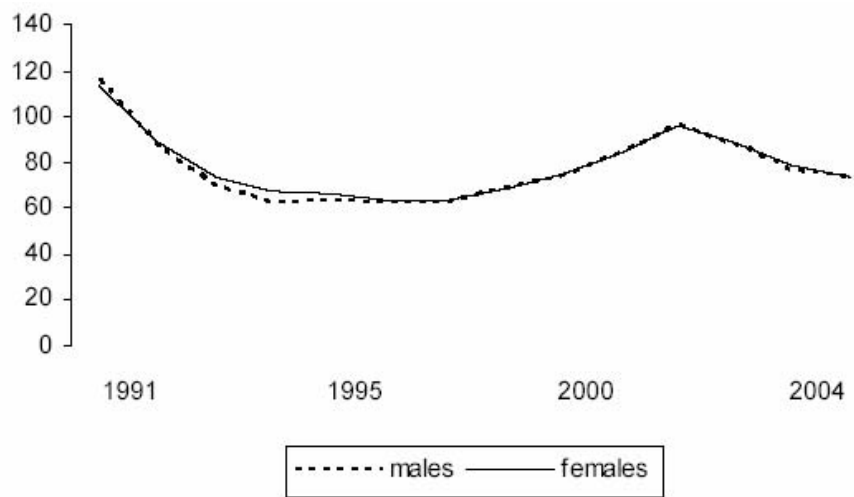


Figure 1: Flow (in thousands) of East-West migrants in Germany by gender, 1991-2004.
 Source: Statistisches Bundesamt, 2005.