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International migration: a panel data analysis of economic and non-economic determinants

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and Non-Economic Determinants**

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ABSTRACT

International Migration: A Panel Data Analysis of Economic and Non-Economic Determinants*

In this paper I empirically investigate economic and non-economic determinants of migration inflows into fourteen OECD countries by country of origin, between 1980 and 1995. The annual panel data set used makes it possible to exploit both the time-series and cross-country variation in immigrant inflows. I focus on both supply and demand determinants of migration patterns and find results broadly consistent with the theoretical predictions of a standard international-migration model. Both first and second moments of the income distribution in the destination and origin countries shape international migration movements. In particular, I find evidence of robust and significant pull effects, that is the positive impact on immigrant inflows of improvements in the mean income opportunities in the host country. Inequality in the origin and destination economies affects the size of migration rates as predicted by Borjas (1987) selection model. Finally, among the non-economic determinants, I investigate the impact on emigration rates of geographical, cultural, and demographic factors as well as the role played by changes in destination countries' migration policies.

JEL Classification: F22

Keywords: international migration, push and pull factors, network effects

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1 Introduction

Do flows of international migrants respond to economic incentives, in spite of destination countries' restrictive immigration policies? Which economic determinants affect international migration flows? Do non-economic variables - such as geographical, cultural and demographic factors - shape cross-country immigration patterns? What role do host countries' restrictive immigration policies play? Are network effects at work? In this paper, I address these questions using a yearly panel data set that allows me to exploit both the time-series and cross-country variation in international immigrant flows.

Cross-country migration patterns vary considerably over time, and across destination and origin countries (see Appendix 2). Some OECD countries have experienced a decrease in the size of the annual immigrant inflow between 1980 and 1995.¹ Over the same years, the number of immigrants per year has increased in several other OECD countries.² The percentage change of the yearly immigrant inflow in this period ranges between negative 42% (in Japan) and positive 48% (in Canada) (OECD 1997). For all destinations, such changes are anything but monotonic (see Figure 3). The variation in terms of origin countries is remarkable as well.

Both economic and non-economic factors are likely to influence the size, origin, and destination of labor movements at each point in time and contribute to the variation observed in the data. While clearly it is important to understand the driving forces behind recent international migration patterns, a limited amount of empirical research has been devoted to this topic, perhaps due to past unavailability of cross-country data on migration.

In this paper, I empirically investigate economic and non-economic determinants of bilateral immigrant flows, across destination and origin countries.³ I first derive testable predictions about the main factors affecting international migration, using a simple theoretical framework. My model focuses on both supply and demand determinants of immigration patterns. Next, I empirically relate bilateral migrant flows (normalized by origin country's population) to the economic, geographical, cultural and demographic determinants suggested by the theory. The explanatory variables I identify are mean income opportunities in the source and destination country, their relative income inequality, the distance and common borders between the two countries, their colonial links and common languages, and the immigration-policy legislation in the host country. The age structure of the origin country's population and network effects are additional determinants suggested by the theoretical model (Clark, Hatton and Williamson 2002).

To analyze migration patterns across countries, I use yearly data on immigrant inflows

¹For example, Australia, France, Japan, Netherlands, and the United Kingdom.

²For example, in Belgium, Canada, Denmark, Germany, Luxembourg, Norway, Sweden, Switzerland, and the United States.

³While I distinguish between "economic" and "non-economic" determinants, there is an economic rationale behind the impact of both sets of variables on migration flows.

into fourteen OECD countries by country of origin, between 1980 and 1995. The source of this data is the International Migration Statistics for OECD countries (OECD 1997), based on the OECD's Continuous Reporting System on Migration (SOPEMI).

My empirical analysis both delivers estimates broadly consistent with the theoretical predictions of a standard international-migration model and generates empirical puzzles.

Among the economic determinants, I investigate the effect on international immigrant flows of both first and second moments of the income distributions in the destination and origin countries.⁴ I find that pull factors - that is, improvements in the mean income opportunities in the destination country - significantly increase the size of emigration rates. This result is very robust to changes in the specification of the empirical model. In addition, both absolute and relative pull factors matter. The emigration rate to a given destination is estimated to be an increasing function of that country's per worker GDP and a decreasing function of the average per worker GDP of all the other host countries in the sample⁵ (each weighted by the inverse of distance from the origin country).

Positive and significant pull effects may appear, at first sight, inconsistent with restrictive immigration policies of several destination countries in the sample.⁶ From a theoretical point of view, the impact of pull (and push) factors is a function of whether immigration is quantity-constrained. If immigration quotas are binding in the host country, pull (and push) factors should have no effect. However, my results show that pull effects matter, in spite of immigration restrictions of most destination countries in the sample. One interpretation of this finding is that the estimated coefficient simply captures an average effect, across country pairs characterized by different immigration-policy arrangements: as long as immigration constraints are not binding for each country pair, this average effect should - according to the theory - be positive. I find evidence consistent with this interpretation when I differentiate the effect of pull and push factors according to each host country's restrictiveness towards immigrants. In particular, I interact mean income opportunities in, respectively, the host and source country with an indicator variable of changes in destination countries' migration policies. I find that, when immigration laws become less restrictive in a host country, the effect of pull (push) factors in that country turns more positive (negative).

Another explanation of my result - that is, positive and significant pull effects - is that even countries with binding official immigration quotas often accept unwanted (legal) immigration. Restrictive immigration policies are often characterized by loopholes, that leave room for potential migrants to take advantage of economic incentives. For example, immigration to Western European countries still took place after the late Seventies, despite the official closed-door policy (Joppke 1998). Family-reunification policies are thought to be one

⁴The income distributions of interest are the ones which characterize the origin country's population in each location.

⁵Since the host countries in the sample include the great majority of destinations of immigrants in the world, it is not overly restrictive to focus on them.

⁶The data set I use only covers legal migration.

reason behind continuing migration inflows.⁷

Controlling for destination and origin countries' fixed effects, the sign of the impact of push factors - that is, declining levels of per worker GDP in the origin country - is seldom consistent with the theoretical predictions of the basic model. In those cases, the size of the effect is smaller than for pull factors and is almost always insignificant. This is surprising given that, in the basic model, push and pull factors have similar-sized effects (with opposite signs). An explanation of my result is that the impact of income opportunities in the origin country is likely to be affected by poverty constraints, due to fixed costs of migration and credit-market imperfections (Schiff 1998 and Yang 2003). Lower levels of per worker GDP in the source country both strengthen incentives to leave and make it more difficult to overcome poverty constraints. In the empirical analysis I find (weak) evidence that my result on push factors is driven by poverty constraints in the origin country.

Yet another interpretation of my findings is that positive and significant pull effects and insignificant push effects are explained by the demand side of international migration (i.e., migration policies), and not by the supply side as is often assumed in the previous literature. The theoretical model makes the assumption that migration quotas are exogenous. In reality changes in mean income opportunities in the destination country affect both migrants' incentive to move and the political process behind the formation of migration policies. If migration quotas are binding, only the latter channel will be at work and the determinants on the supply side won't have any effect.⁸

Inequality in the source and host economies is related to the size of emigration rates as predicted by Borjas (1987) selection model. An increase in the origin country's relative inequality has a non-monotonic effect on the size of the emigration rate: the impact is estimated to be positive if there is positive selection, negative if there is negative selection.

Next, I investigate the role played by geographical, cultural and demographic factors. Among the variables affecting the costs of migration, distance between destination and origin country appears to be of considerable importance. Its effect is negative, significant and steady across specifications. On the other hand, controlling for destination and origin countries' fixed effects, I don't find evidence that cultural variables play a significant role. Demographics - in particular, the share of the origin country's population who is young - shapes bilateral flows as predicted by the theory.

Finally, I empirically investigate the importance of network effects. Since immigrants are likely to receive support from compatriots already established in the host country, they will have an incentive to choose destinations with larger communities of fellow citizens. Network

⁷Another channel of entry of immigrants in countries with official closed-door policy is asylum-seekers laws. Joppke (1998) writes about Germany's experience (p.285): "Since the recruitment stop of 1973, the chain migration of families of guest workers was (next to aylum) one of the two major avenues of continuing migration flows to Germany, in patent contradiction to the official no-immigration policy."

⁸Notice that a deeper analysis of the demand side of international migration is outside the scope of this paper and it's an interesting direction of future research.

effects imply that bilateral migration flows are highly correlated over time, which is what I find in my analysis. However, the interpretation of this result is ambiguous as it could be driven by either supply factors (that is, network effects) or by demand factors (family reunification policies, for example).

The empirical literature on the determinants of migration includes a number of works, some of which date back to the nineteenth century (Ravenstein 1885). More recently, Clark, Hatton and Williamson (2002) and Karemera, Oguledo and Davis (2000) both focus on the fundamentals explaining immigrant inflows into the United States by country of origin in the last decades. Helliwell (1997 and 1998) sheds light on factors affecting labor movements in his investigation of the magnitude of immigration border effects, using data on Canadian interprovincial, US interstate and US-Canada cross-border immigration.

The contribution of this paper to the literature is threefold. First, my work is the first one I am aware of to use the OECD (1997) data on international migration to systematically investigate the economic and non-economic determinants of international flows of migrants. Previous works have either used country cross-sections (Borjas 1987 and Yang 1995, for example), or have focused on a single destination country over time (Borjas and Bratsberg 1996, Clark, Hatton and Williamson 2002, and Karemera, Oguledo and Davis 2000, for example) or a single origin country over time (for example, Yang 2003).⁹ By extending the focus of the analysis to a multitude of origin and destination countries and taking advantage of both the time-series and cross-country variation in the data, I can test the robustness and broader validity of the results found in the previous literature.¹⁰

Second, this paper carefully reviews and proposes solutions to various econometric issues that arise in the empirical analysis, such as endogeneity and reverse causality. Some of these econometric complications have not been addressed in the previous literature. Finally, this paper puts greater emphasis than previous works on the demand side of international migration, i.e. destination countries' migration policies. This change of perspective is important, given restrictive immigrant policies in the great majority of host countries.

The framework used in this work to study migration flows is related to the gravity model of trade, which is used to analyze bilateral *trade* flows across countries.¹¹ As a matter of fact, I use several variables that appear frequently in the trade gravity literature (*distance*, *border*, *common language*, and *colony*). There also exists a gravity model of immigration, developed in the geography literature and sometimes used in economics papers. However, the empirical

⁹From a methodological point of view, the paper most related to mine is Clark, Hatton, and Williamson (2002).

¹⁰Since I began my work on international migration, I have become aware of another related, but independent paper analyzing cross-country migration patterns: Pedersen, Pytlikova, and Smith (2004). This paper uses another data set to explore the determinants of international migration.

¹¹A number of works empirically analyzes trade flows within this setting (Helpman 1987 and Hummels and Levinsohn 1995). The same type of framework is used to explain bilateral cross-border equity flows across countries (Portes and Rey 2002) as well as FDI flows (Brenton *et al.* 1999, Frankel and Wei 1996, and Mody, Razin and Sadka 2002).

specification I use, suggested by economic theory, differs in part from the standard equation estimated by geographers.¹²

The rest of the paper is organized as follows. Section 2 presents a simple model of international migration. In Section 3 I describe the data sets used in the empirical part, while in Section 4 I discuss the estimating equations and some econometric issues that complicate the analysis. Finally, Section 5 presents the main empirical results and Section 6 concludes.

2 Theoretical framework

Both supply and demand factors determine international migration flows. Migrants' decision to move, according to economic and non-economic incentives, shape the supply side of labour movements. The host country's immigration policy represents the demand side, i.e. the demand for immigrants in the destination country. The latter one, in turn, can be thought of as the outcome of a political-economy model in which individual attitudes toward immigrants, interest-groups pressure, policy-makers preferences and the institutional structure of government interact with each other and give rise to a final immigration-policy outcome (Rodrik 1995, Facchini and Willmann 2004, Mayda 2003).

The theoretical framework used in this paper is closely related to the previous literature. The main difference is the greater emphasis in my model on demand factors, i.e. destination countries' immigration policy. I will consider a world with two economies: country 0, which is the country of origin of immigrant flows and country 1, which is the country of destination. I will first focus on the supply side of immigration, that is migrants' decision to move. In particular, I will look at the probability that an individual chosen randomly from the population of country 0 will migrate to country 1.

In each country, wages are a function of the individual skill level (s_i). In the origin country, the wage of individual i equals:

$$w_{0i} = \alpha_0 + \theta_0 \cdot s_i + disturbance_{0i}, \quad (1)$$

while in the country of destination:

$$w_{1i} = \alpha_1 + \theta_1 \cdot s_i + disturbance_{1i}, \quad (2)$$

where the two disturbances have both zero means over the origin country's population. Since in the empirical analysis I will use aggregate data, I need to rewrite individual i 's wages

¹²The standard equation estimated by geographers looks as follows (Gallup 1997): $flow_{ij} \propto \frac{P_i P_j}{dist_{ij}^2}$. As Helliwell (1997) points out, there is still a contrast between economic and gravity explanations of immigrant flows.

in the two locations as a function of the mean wages of the origin country's population at home and abroad respectively:¹³

$$w_{0i} = \mu_0 + v_{0i}, \text{ where } v_{0i} \sim N(0, \sigma_0^2), \quad (3)$$

$$w_{1i} = \mu_1^0 + v_{1i}, \text{ where } v_{1i} \sim N(0, \sigma_1^2), \quad (4)$$

where the correlation coefficient between v_{0i} and v_{1i} equals ρ_{01} .

Notice that μ_0 equals $\alpha_0 + \theta_0 \cdot \bar{s}_0$ and μ_1^0 equals $\alpha_1 + \theta_1 \cdot \bar{s}_0$, where \bar{s}_0 is the mean skill level of the origin country's population. The parameter μ_1^0 is likely to be different from μ_1 , the mean wage in country 1 of the destination country's population: $\mu_1 = \alpha_1 + \theta_1 \cdot \bar{s}_1$, where \bar{s}_1 represents the mean skill level of the destination country's population (Borjas 1999a, and Clark, Hatton and Williamson 2002). This point will become relevant below, in the empirical analysis.

I assume that each individual has Cobb Douglas preferences for the two goods produced in the world (x_1 and x_2):

$$U(x_1, x_2) = Ax_1^{1-\delta}x_2^\delta, \quad 0 < \delta < 1, \quad A > 0, \quad (5)$$

which implies an indirect utility (function) from having an income y given by:¹⁴

$$v(p_1, p_2; y) = \bar{A}(p_1, p_2) \cdot y. \quad (6)$$

I assume that each country is a small open economy characterized by free trade with the rest of the world:¹⁵ therefore goods' prices p_1 and p_2 are given and equal - and $\bar{A}(p_1, p_2)$ also does not vary - across countries.¹⁶ An individual in country 0 will migrate to country 1 if the utility of moving is greater than the utility of staying at home i.e., given the assumptions above, if the expected income in the destination country net of migration costs is greater than the expected income in the origin country. Following the literature (for example, Borjas 1999a, and Clark, Hatton and Williamson 2002), I can define an index I_i that measures the net benefit of moving relative to staying at home for a risk-neutral individual i :

$$I_i = \eta_{01} \cdot (w_{1i} - w_{0i} - C_i) + (1 - \eta_{01}) \cdot (-w_{0i} - C_i), \quad (7)$$

$$\Rightarrow I_i = \eta_{01} \cdot w_{1i} - w_{0i} - C_i, \quad (8)$$

¹³More in general, in (3) and (4), w_{0i} and w_{1i} are rewritten as a function of first and second moments of the income distribution of the origin country's population at home and abroad respectively.

¹⁴In the following expression: $\bar{A}(p_1, p_2) = A\left(\frac{1-\delta}{p_1}\right)^{1-\delta}\left(\frac{\delta}{p_2}\right)^\delta$.

¹⁵Given free trade, what explains differences in rates of return to labor across countries? The answer is that the conditions for factor price equalization are not satisfied: for example, if international productivity differences exist, as in Treffer (1993), then only *adjusted* factor-price-equalization holds.

¹⁶In the empirical analysis I adjust for international differences in goods' prices, using PPP income levels.

where η_{01} is the probability that the migrant from country 0 will be allowed to stay in country 1, w_{0i} and w_{1i} are respectively the wage in the origin and destination country, and $C_i = \mu_C + v_i^C$, with $v_i^C \sim N(0, \sigma_C^2)$, represents the level of individual migration costs.¹⁷ The correlation coefficients between v_i^C and (v_{0i}, v_{1i}) are equal to (ρ_{0C}, ρ_{1C}) .

This model focuses on *labor* mobility. Migration allows an individual to take advantage of differences in rates of return to *labor* across countries. Migrants may own capital, either at home or in the destination country, but their capital income opportunities are assumed to be independent of their residence. That is why the index I_i does not include a capital-income term.¹⁸

In addition, the implicit assumption in (8) is that, if the migrant is not allowed to stay in the destination country, he still incurs the migration costs C_i and gives up the home wage w_{0i} . In other words, the individual moves to the host country before knowing whether he will be able to stay (for a longer period of time) and gain the income w_{1i} .¹⁹ The immigrant from country 0 may not be allowed to stay in country 1 because of quotas due to a restrictive immigration policy, as is explained below.

The probability that an individual chosen randomly from the population of the origin country will migrate from country 0 to country 1 equals:

$$P = \Pr[I_i > 0] = \Pr[\eta_{01} \cdot (\mu_1^0 + v_{1i}) - (\mu_0 + v_{0i}) - (\mu_C + v_i^C) > 0], \quad (9)$$

which can be rewritten as:

$$P = 1 - \Phi(z), \quad (10)$$

where $z = -\frac{(\eta_{01} \cdot \mu_1^0 - \mu_0 - \mu_C)}{\sigma_v}$, σ_v is the standard deviation of $(\eta_{01} \cdot v_{1i} - v_{0i} - v_i^C)$, and $\Phi(\cdot)$ is the cumulative distribution function of a standard normal.²⁰ The probability of emigration from country 0 to country 1 in (10) is approximated by the *supply* emigration rate $\frac{I_{01}^S}{P_0}$, i.e. the emigration rate on the supply side of the migration model.

Consider now a situation in which the destination country's immigration policy implies either explicit or implicit quantity constraints for immigrants coming from each origin country.²¹ Let I_{01}^D represent the maximum number of migrants from country 0 allowed each period

¹⁷I assume that each individual knows the wage levels w_{1i} and w_{0i} he would get in each location, the migration costs C_i and the probability η_{01} . Notice that, while each individual takes the probability η_{01} as given, this probability is endogenously determined in the model.

¹⁸In other words, capital is internationally mobile. The migrant can own capital in the origin and destination country and receive income from it, no matter where he resides.

¹⁹This assumption is consistent with the evidence that immigrants often arrive to a destination country with tourist or student visas with the hope of being able to stay.

²⁰In particular, $\sigma_v^2 = (\eta_{01}^2 \sigma_1^2 + \sigma_0^2 + \sigma_C^2 - 2\eta_{01}\rho_{01}\sigma_0\sigma_1 - 2\eta_{01}\rho_{1C}\sigma_1\sigma_C + 2\rho_{0C}\sigma_0\sigma_C)$.

²¹My model differs from previous ones in the literature in the way it analyzes the impact of quantity restrictions induced by immigration policy. Clark, Hatton, and Williamson (2002) and Hatton and Williamson (2002) model immigration policy as affecting the level C_i of migration costs.

into country 1. These immigration quotas may or may not be binding.

What we can observe in the data is the actual emigration rate $\frac{I_{01}}{P_0}$, i.e. the number of immigrants from country 0 to country 1, divided by the population of country 0. Only in the absence of binding immigration quotas does the actual emigration rate $\frac{I_{01}}{P_0}$ equal the *supply* emigration rate $\frac{I_{01}^S}{P_0}$. In the presence of binding quantity-constraints, $\frac{I_{01}}{P_0}$ will be less than $\frac{I_{01}^S}{P_0}$. That is, the actual emigration rate is equal to the minimum between $\frac{I_{01}^S}{P_0}$ and $\frac{I_{01}^D}{P_0}$:

$$\frac{I_{01}}{P_0} = \min\left(\frac{I_{01}^S}{P_0}, \frac{I_{01}^D}{P_0}\right), \quad (11)$$

where the immigration quota I_{01}^D represents country 1's demand for immigrants from country 0, which is a function of country 1's immigration policy. The heavy lines in Figure 1 give the actual emigration rate as a function of μ_1^0 and μ_h , $h = 0, C$. In this paper I assume that I_{01}^D is exogenous, thus it is not affected by μ_1^0 neither by μ_h , $h = 0, C$.²² This is a strong assumption that will be questioned in the interpretation of the empirical results.

I assume that the probability η_{01} - that the migrant from country 0 will be allowed to stay in country 1 - is equal to $\min\{1, \frac{I_{01}^D}{P_0 \cdot P}\}$ (the number of people, from country 0 to country 1, who are allowed in, divided by the number of those who try to get in).

Given (10) and (11), it is possible to derive testable predictions for the impact of μ_1^0 , μ_0 , and μ_C on the actual emigration rate from country 0 to country 1:²³

$$\frac{d(\frac{I_{01}}{P_0})}{d\mu_1^0} = \begin{cases} \frac{\phi(z)}{\sigma_v} > 0, & \text{if } \frac{I_{01}^S}{P_0} < \frac{I_{01}^D}{P_0}; \\ 0, & \text{if } \frac{I_{01}^S}{P_0} \geq \frac{I_{01}^D}{P_0} \end{cases} \quad (12)$$

$$\frac{d(\frac{I_{01}}{P_0})}{d\mu_h} = \begin{cases} -\frac{\phi(z)}{\sigma_v} < 0, & \text{if } \frac{I_{01}^S}{P_0} \leq \frac{I_{01}^D}{P_0}; \\ 0, & \text{if } \frac{I_{01}^S}{P_0} > \frac{I_{01}^D}{P_0} \end{cases} \quad (13)$$

where $\phi(\cdot)$ is the density function of a standard normal and $h = 0, C$. The comparative-static results in (12) show the effect of pull factors - that is, improvements of the mean income opportunities in the destination country - according to whether the immigration quotas are binding or not. Pull effects are positive and strongest when restrictions are not binding neither ex-ante nor ex-post, they are positive but smaller in size when the quota is binding ex-post but not ex-ante and, finally, they are equal to zero in a quantity-constrained world.

²²Alternatively, I_{01}^D can be explicitly modeled within a political-economy framework. In that case, the immigration quotas are likely to depend on the capital-labor ratio of the median voter (see Benhabib 1996), on the size of past immigration flows from the same origin country, both because of family-reunification policies and because of pro-immigration votes of naturalized immigrants (Ortega 2004), and on the extent of political organization of various interest groups (Grossman and Helpman 1994 and Facchini and Willmann 2004).

²³The total differential of P equals:

$$dP = \frac{\phi(z)}{\sigma_v} d(\eta_{01}\mu_1^0 - \mu_0 - \mu_C) + \phi(z) (\eta_{01}\mu_1^0 - \mu_0 - \mu_C) \left(-\frac{1}{\sigma_v^2}\right) d\sigma_v.$$

A parallel interpretation explains the comparative-static results in (13), which describe push effects (changes of μ_0 , that is mean income opportunities in the origin country) and the impact of mean migration costs (changes of μ_C), according to the immigration-policy regime.

Thus, according to this simple model, pull and push factors have either similar-sized effects (with opposite signs), when quotas are not binding, or they both have no effect on emigration rates, when quotas are binding. In the empirical analysis I will not be able to control for whether a country-pair/year observation is characterized by binding migration quotas or not (since I do not have data on I_{01}^D). Therefore I will estimate an average effect across pairs of countries characterized by either one of them. However, I will be able to use information on changes of I_{01}^D . I should find that pull (push) effects are more positive (negative) than average, for a given destination country, if that host country's migration policy becomes less restrictive.²⁴

Focusing for simplicity on the region where immigration quotas are not binding, it is straightforward to derive predictions for the impact on $\frac{I_{01}}{P_0}$ of the second moments of the income distributions of the origin country's population, at home and abroad respectively. In particular, assuming that $\sigma_C = 0$, we obtain the following expressions (where $k < 0$):²⁵

$$\frac{d(\frac{I_{01}}{P_0})}{d\sigma_1} = k \cdot (\eta_{01}\mu_1^0 - \mu_0 - \mu_C) \cdot (\sigma_1 - \rho_{01}\sigma_0), \quad (14)$$

$$\frac{d(\frac{I_{01}}{P_0})}{d\sigma_0} = k \cdot (\eta_{01}\mu_1^0 - \mu_0 - \mu_C) \cdot (\sigma_0 - \rho_{01}\sigma_1). \quad (15)$$

In my discussion I will assume that $(\eta_{01}\mu_1^0 - \mu_0 - \mu_C) > 0$ so that, based on first-moments considerations, on average immigrants have an incentive to migrate. The results in (14) and (15) imply that, if $\frac{\sigma_0}{\sigma_1} < 1$ and ρ_{01} is sufficiently high ($\rho_{01} > \frac{\sigma_0}{\sigma_1}$), then $d\sigma_0 > 0$ or $d\sigma_1 < 0$ (i.e., an increase in the relative inequality $\frac{\sigma_0}{\sigma_1}$) will increase the emigration rate. The intuition about this result is straightforward. Given that income inequality in the origin country is lower than in the destination country ($\frac{\sigma_0}{\sigma_1} < 1$) and that ρ_{01} is sufficiently high, this is a situation with positive selection of immigrants from country 0 to country 1 (Borjas 1987): migrants are selected from the upper tail of the income distribution at home and end up in the upper tail of the income distribution abroad (in both cases, the relevant distribution is the origin country's population one). For example, consider potential migrants from Portugal to the United States (Figure 2). Given that income inequality is lower in Portugal than in the U.S. and assuming that ρ_{01} is sufficiently high, among Portuguese workers it's the better-off

²⁴The reason is that, with higher I_{01}^D , the range of μ_1^0 (μ_0) for which the effect is strictly positive (negative) is wider.

²⁵Formulas (14) and (15) are based on the expression for dP in footnote 22 (given that immigration quotas are not binding, then $\frac{I_{01}}{P_0} = P$). If quotas are not binding ($\eta_{01} = 1$), assuming that $\sigma_C = 0$, then: $d\sigma_v = [\sigma_1^2 + \sigma_0^2 - 2\rho_{01}\sigma_0\sigma_1]^{-\frac{1}{2}}[(\sigma_1 - \rho_{01}\sigma_0)d\sigma_1 + (\sigma_0 - \rho_{01}\sigma_1)d\sigma_0 - \sigma_0\sigma_1d\rho_{01}]$. Notice that, in formula (14) and (15), $k = \phi(z)(\sigma_1^2 + \sigma_0^2 - 2\rho_{01}\sigma_0\sigma_1)^{-\frac{1}{2}}(-\frac{1}{\sigma_v^2}) < 0$.

who have an incentive to migrate while those at the lower tail of the income curve have an incentive to stay. The reason is that the probability of both very high and very low incomes is higher in the U.S. than in Portugal. An increase in income inequality in Portugal will make individuals at the bottom of the income distribution relatively worse-off at home and will increase their incentive to leave.

Similarly, if $\frac{\sigma_0}{\sigma_1} > 1$ and ρ_{01} is sufficiently high ($\rho_{01} > \frac{\sigma_1}{\sigma_0}$), then $d\sigma_0 > 0$ or $d\sigma_1 < 0$ (i.e., an increase in the relative inequality $\frac{\sigma_0}{\sigma_1}$) will decrease the emigration rate. Given that income is more dispersed at home than abroad, this is the case of negative selection of immigrants from country 0 to country 1 (Borjas 1987): migrants are selected from the lower tail of the income distribution at home and end up in the lower tail of the income distribution abroad. An example of this situation is migration from Brazil to the U.S., given that income inequality in the latter is lower than in the former (Figure 2). An increase in income inequality in Brazil will lower the emigration rate because those who were not migrating beforehand, the better-off, will have even less incentive to do so afterwards.

Notice that, in a way similar to the impact of first moments, the effect of second moments will be affected by whether immigration quotas are binding or not.

Finally, a few extensions of this simple theoretical framework deliver interesting predictions that I will discuss and investigate in the empirical section.

3 Data

In this paper, I combine an international panel data set on bilateral immigration flows with additional macroeconomic and non-economic data on the origin and destination country of each immigrant flow. Data on immigration comes from the International Migration Statistics (IMS) data set for OECD countries (OECD 1997), which contains information on immigrant flows by country of origin, based on the OECD's Continuous Reporting System on Migration (SOPEMI).²⁶ Population registers and residence and work permits are the main sources of these statistics. In particular, I use data on yearly immigrant inflows into fourteen OECD countries by country of origin, in the period 1980-1995. Appendix 2, at the end of the paper, presents summary statistics on immigrant inflows by host and source country, averaged over the years 1980-1995 (see also Figure 3).

Appendix 3 shows that the disaggregate (i.e., by country of origin) IMS statistics on migrant flows don't add up to 100% of the total flow into each destination country. The percentage of the total immigrant inflow covered by the disaggregate data ranges between 45% (Belgium) and 84% (United States). Put differently, the data set includes zero flows in correspondence of some country pairs (immigrant inflows from Italy to the United States, for example). Some of these observations correspond to truly zero flows, while others are likely to correspond to very small flows. If the latter observations are recorded as zeros in

²⁶Note that the IMS data does not cover illegal immigration.

the disaggregate data set, there will be a discrepancy between total flows and the sum of flows by origin country. In the empirical analysis I will keep zero-flows observations in the data set.²⁷ I will also investigate the robustness of my results using a Tobit model.²⁸

Summary statistics and data sources of each variable used in the empirical model are documented in Appendix 1. Data on macroeconomic variables comes from various sources: the 2001 World Development Indicators (World Bank 2001), the Penn World Tables (versions 5.6 and 6.1), and the World Bank's Global Development Network Growth Database, Macro Time Series (Easterly and Sewadeh 2002). Geographical and cultural information, such as on great-circle distance, land border, common language, and colonial ties, comes from Glick and Rose's (2001) data set on gravity-model variables. I also use statistics on the average number of schooling years in the total population (over age 15) from Barro and Lee's (2000) data set.²⁹ Data on Gini coefficients of destination and origin countries, used to construct the origin country's relative inequality variable, comes from Deininger and Squire (1996) data set (only high-quality observations were used). Finally, information on origin countries' share of young population comes from the United Nations.

A separate Appendix to the paper documents the main characteristics of the migration policies of the destination countries in the sample and the timing (after 1980) of changes in their legislations (Mayda and Patel, 2004).³⁰ A data set of destination countries' migration-policy changes, between 1980 and 1995, was constructed on the basis of the information in this Appendix and used in the empirical analysis.

4 Empirical model

The theoretical framework in Section 2 suggests an empirical specification characterized by the emigration rate as the dependent variable and, among the explanatory variables, the mean wage of the origin country's population in, respectively, the origin and destination countries. As approximations for the latter two variables, I use the (log) level of per worker GDP, PPP-adjusted (constant 1996 international dollars) in the two countries.³¹ Based on the basic model, I expect pull effects to be positive on average and push effects to be negative on average.

Another determinant of bilateral immigration flows implied by the model of Section 2 is the distance between the two locations, which affects migration costs C_i . The further

²⁷The results don't change when I drop zero observations (see previous version of this paper).

²⁸See regression (4), Table 4.

²⁹Since this panel only contains data at five-year intervals (in the period I consider, the years covered are 1980, 1985, 1990, 1995), I linearly extrapolate figures for the in-between years (by assigning one fifth of the five-year change in the variable to each year).

³⁰The Appendix can be found in my website.

³¹Per worker GDP is not a direct measure of the mean wage of the origin country's population at home and abroad. See below for robustness checks related to this point.

away the two countries are, the higher the monetary travel costs for the initial move, as well as for visits back home. Remote destinations may also discourage migration because they require longer travel time and thus higher foregone earnings. Another explanation as to why distance may negatively affect migration is that it is more costly to acquire information ex-ante about far-away countries (Greenwood 1997 and Lucas 2000). I introduce additional variables that affect the level of migration costs C_i . A common land border is likely to encourage migration flows, since land travel is usually less expensive than air travel. Linguistic and cultural similarity are also likely to reduce the magnitude of migration costs, for example by improving the transferability of individual skill from one place to the other. Past colonial relationships should increase emigration rates, to the extent that they translate into similar institutions and stronger political ties between the two countries, thus decreasing the level of migration costs C_i .

In a cross-country analysis, such as in this paper, unobserved country-specific effects could result in biased estimates. For example, I may estimate a positive coefficient on the destination country's per worker GDP. Based on this result, it is not clear whether immigrants go to countries with higher wages or, alternatively, whether countries with higher wages have other features that attract immigrants. Along the same lines, a negative coefficient on income at home leaves open the question of whether immigrants leave countries with lower wages or, alternatively, whether countries with lower wages have certain characteristics that push immigrants to leave. To (partly) get around this problem, I exploit the panel structure of the data set and I introduce dummy variables for both destination and origin countries. This allows me to control for unobserved country-specific effects which are additive and time-invariant. All specifications in Tables 1, 4, 5, and 6 include destination and origin countries' fixed effects and have robust standard errors clustered by country pair, to address heteroscedasticity and allow for correlation over time of country-pair observations.³² Notice that destination country fixed effects allow me to control for features of destination countries' immigration policy which don't change over time and are common across origin countries. In order to capture the effect of changes in destination countries' migration policies, I introduce two interaction terms of an indicator variable of such changes with pull and push factors, respectively. According to the theory, if the migration policy of a destination country becomes less restrictive, the effect of pull (push) factors should turn more positive (negative).

Finally, I introduce the share of the origin country's population who is young (between 15 and 29 years old) as a demographic determinant of migration flows. Consider an extension of the basic model in Section 2 to a multi-period setting. In this set-up, the individual cares not only about current wage differentials net of moving costs, but about future ones too. As Clark, Hatton, and Williamson (2002) point out, this implies that a potential migrant from country 0 will have a bigger incentive to migrate the younger he is, as the present discounted

³²Most specifications in the tables include separate dummy variables for destination and origin countries, while the remaining regressions control for country-pair fixed effects.

value of net benefits will be higher the more extended the length of the remaining work life time (for positive I_i in each year). We would then expect the share of young population in the origin country to increase the emigration rate out of that country.

The basic empirical specification thus looks as follows:

$$\begin{aligned} \frac{flow_{ijt}}{P_{it}} = & \beta + \beta_0 p w g d p_{it-1} + \beta_1 p w g d p_{jt-1} + \beta_2 d i s t_{ij} + \beta_3 b o r d e r_{ij} + \beta_4 c o m l a n g_{ij} + \beta_5 c o l o n y_{ij} + \\ & + \beta_6 y o u n g p o p_{it-1} + \beta_7 p w g d p_{it-1} \cdot i m m i g p o l_{jt} + \beta_8 p w g d p_{jt-1} \cdot i m m i g p o l_{jt} + \varepsilon_{ijt} \end{aligned} \quad (16)$$

where i is the origin country, j the destination country and t time. $\frac{flow_{ijt}}{P_{it}}$ is the emigration rate from i to j at time t ($flow_{ijt}$ is the inflow into country j from country i at time t , P_{it} is the population of the origin country at time t). $p w g d p$ is the (log) per worker GDP, PPP-adjusted (constant 1996 international dollars) and $d i s t$ measures the (log) great-circle distance between the two countries. The variable $b o r d e r$ equals one if the two countries in the pair share a land border. $c o m l a n g$ and $c o l o n y$ are two dummy variables equal to one, respectively, if a common language is spoken in both locations, and for pairs of countries which were, at some point in the past, in a colonial relationship. Finally, $y o u n g p o p$ is the share of the population in the origin country aged 15-29 years old and the variable $i m m i g p o l$ equals one (minus one) if in that year the destination country's immigration policy became less (more) restrictive (zero otherwise). According to the theory, I expect that $\beta_0 < 0$, $\beta_1 > 0$, $\beta_2 < 0$, $\beta_3 > 0$, $\beta_4 > 0$, $\beta_5 > 0$, $\beta_6 > 0$, $\beta_7 < 0$, and $\beta_8 > 0$.

I will next discuss a few econometric issues that affect the interpretation of pull and push coefficients. First, assuming that per worker GDP proxies for the mean income opportunity of the origin country's population in each location (see below for a discussion of this point), an empirical complication is the possibility of reverse causality and, more in general, of endogeneity in the time-series dimension of the analysis. The theoretical model in Section 2 predicts that, *ceteris paribus*, higher (lower) income opportunities in the destination (origin) country increase emigration rates. However, a positive β_1 (negative β_0) may just reflect causation in the opposite direction, i.e. the impact of immigrant flows on wages (or levels of per worker GDP) in the host and source country. After all, this channel is the focus of analysis in most labour-economics papers (see Friedberg and Hunt 1995 for a survey of this literature). More broadly, other time-variant third factors may drive contemporaneous wages and immigrant flows.

As for reverse causality, notice that it is likely to bias the estimates toward zero. The reason is that immigrant inflows are likely to decrease wages in the destination country and outflows are likely to increase wages in the origin country. While the opposite signs are a theoretical possibility (for example, in the economic-geography literature, because of economies of scale), the empirical evidence in the labor-economics literature is that immigrant

inflows have a negative impact on the destination country's wages (Borjas 2003) and that immigrant outflows have a positive impact on the origin country's wages (Mishra 2003).

I address reverse-causality and endogeneity issues in two ways. First of all, in the basic specification, I relate *current* emigration rates to *lagged* values of (log) per worker GDP, at home and abroad. Indeed, while it is hard to claim that wages at home and abroad are strictly exogenous, it is plausible to assume that they are predetermined, in the sense that immigrant inflows - and third factors in the error term - only affect contemporaneous and future wages.³³

I next use instrumental-variables estimation with countries' terms of trade as an instrument for PPP-adjusted income levels in the destination and origin country. Papers in the literature where shocks to terms of trade are used as instruments for growth rates of income are, for example, Pritchett and Summers (1996) and Easterly, Kremer, Pritchett and Summers (1993).

To capture the effect of income opportunities for a potential migrant at home and abroad, I use data on per worker GDP (PPP-adjusted) in the origin and destination country. However, per worker GDP is not a direct measure of such income opportunities, since it depends on rates of return to both capital and labor and on per worker endowments of each factor. In other words, a higher per worker GDP in the destination country does not necessarily mean better income opportunities on average for a worker from the origin country, since it could be due to a higher capital-labor ratio or to a more skilled labor force in the destination country's population. In order to focus on rates of return to labor, I will run a robustness check and control for the mean skill level and per worker endowment of capital in each country.³⁴

The last set of predictions of the model in Section 2 (formulas (14) and (15)) is related to the second moment of the income distributions in the origin and destination country. According to the theory, given low values of $\frac{\sigma_0}{\sigma_1}$, if $\frac{\sigma_0}{\sigma_1}$ increases, the emigration rate will increase, while given high values of $\frac{\sigma_0}{\sigma_1}$, if $\frac{\sigma_0}{\sigma_1}$ increases, the emigration rate will decrease. In order to test this prediction, I introduce in the model the origin country's relative inequality ($\frac{\sigma_0}{\sigma_1}$) in linear and quadratic form. I expect the coefficient on the linear term to be positive and on the quadratic term to be negative (Clark, Hatton, and Williamson 2002).

A few extensions of the simple theoretical framework in Section 2 deliver interesting predictions that I investigate empirically. First, it is possible to incorporate poverty constraints into the model, due to fixed costs of migration and credit market imperfections in the origin country. As Young (2003) shows, these assumptions imply that the effect on emigration rates of income opportunities at home is non-monotonic, positive at very low levels of income and

³³Strict exogeneity of an explanatory variable implies $E[X_{it}\varepsilon_{is}] = 0$, for $\forall s, t$, while predeterminedness implies $E[X_{it}\varepsilon_{is}] = 0$, for $\forall s > t$. In one of the following specifications, I also control for lagged values of the emigration rate, since if the emigration rate is autocorrelated, predeterminedness of the regressors does not guarantee consistency of the estimates.

³⁴International differences in rates of return to capital also matter but, as a first approximation, I will assume that capital is internationally mobile.

negative for higher levels. Accordingly, I extend the empirical model previously specified by introducing both a linear and a quadratic term in per worker GDP of the origin country.

Second, the theoretical model can be modified by considering in (8) the *expected* wage, both in the origin and destination country, given uncertainty of finding a job in each place. The probability of finding a job can be approximated by one minus the unemployment rate, which I introduce as a regressor in the empirical specification.

Finally, an important extension is to consider the possibility for workers to choose among multiple destination countries. In this framework, potential migrants compare mean income opportunities in their origin country to those in the destination country considered and in any other host country. For each pair of source and host economies, I construct and control for a *multilateral pull* term which is an average of per worker GDPs of all the other destination countries in the sample (each weighted by the inverse of distance from the origin country).

To conclude, I investigate the role of past migration flows to the destination country from the same origin country. The latter ones affect the current emigration rate through both the supply and the demand channel. On the supply side, lagged emigration rates or, alternatively, the size of the immigrant stock from the same source country, proxy for network effects, which are likely to reduce the cost C_i of migration. On the demand side, past migration flows influence the emigration rate in two ways: through family-reunification immigration policies and through political-economy factors (see, for example, Goldin (1994) and Ortega (2004), where the votes of naturalized immigrants affect immigration policy outcomes). The introduction of the lagged emigration rate among the explanatory variables makes the model a dynamic one. I use Arellano and Bond’s GMM estimator to deal with the incidental parameter problem that arises with fixed-effects estimation of such a dynamic equation.³⁵

5 Empirical results

Table 1, at the end of the paper, presents the results from estimation of equation (16) exploiting both the cross-country and time-series variation in the data.

The estimates of Table 1 show a systematic pattern, broadly consistent with the theoretical predictions of the basic international-migration model. The emigration rate is positively related to the destination country’s (log) per worker GDP, as predicted in Section 2. According to the estimate in regression (1), a ten percent increase in the level of per worker GDP in the destination country increases emigration by 2.6 emigrants per 100,000 individuals of the origin country’s population. In other words, a 10% increase in the host country’s per worker GDP implies a 19% increase in the emigration rate (as the mean of the dependent variable is,

³⁵In a dynamic equation, the fixed effects (or within) estimator of the coefficient of the lagged dependent variable is consistent as $T \rightarrow \infty$, for given N , but it is not consistent for given T , as $N \rightarrow \infty$ (incidental parameter problem). The intuition behind this result is that, in the latter case, the number of parameters to be estimated tends to infinity, while the information used to estimate each parameter does not increase.

in regression (1), 14 emigrants per 100,000 individuals). The impact on the emigration rate of a change in the income opportunities at home is not consistent with the predictions of the basic model. Push effects are estimated to be insignificantly different from zero. As pointed out above, one reading of this result is that it is driven by the effect of poverty constraints in the origin country. I will investigate this possibility in Table 5.

In Table 1 I also explore the role played by geographical (*log distance* and *land border*), cultural (*common language* and *colony*), and demographic (*share of young population (origin)*) determinants, respectively. Controlling for destination and origin countries' fixed effects, the picture that emerges from my results is one in which geography and demographics are the most important non-economic determinants of migration flows. According to the estimate in column (1), doubling the great-circle distance between the source and host country decreases the number of emigrants by 40 per 100,000 individuals in the origin country (significant at the 1% level). On the other hand, a common land border does not appear to play a significant role. The impact of a common language, though of the right sign, is not always statistically significant and, surprisingly, past colonial relationships do not appear to significantly affect migration rates. Finally, the share of the origin country's population who is young has a positive and significant impact on emigration rates. A ten percentage point increase in the origin country's 15-29 years old population raises the emigration rate by 24 emigrants per 100,000 individuals (regression (3)).

Next, I investigate the interaction between changes in destination countries' migration policies and, respectively, pull and push factors (column (4), Table 1). Consistent with the theoretical predictions, positive pull factors are bigger (smaller) than average for a destination country whose migration policy becomes less (more) restrictive. Push factors turn negative and significant once migration restrictions are relaxed.³⁶

Finally, in regression (6) I only exploit the variation over time within country pairs, by introducing dummy variables for each combination of origin and destination countries.³⁷ These country-pairs fixed effects allow me to control for time-invariant features of the destination country's immigration policy which are specific for each origin country.

The specifications in Tables 2 and 3 are the most closely related to trade gravity-model regressions (which are usually estimated in a cross section). In Table 2 I focus on a subset of determinants and estimate their coefficients by exploiting only the cross-country variation in the data. I divide the period between 1981 and 1995 into three segments and I focus on each at a time. I relate average emigration rates in each subperiod to the average income opportunities at home and abroad in the previous five-year interval (plus time-invariant variables). In Table 3 I perform a similar exercise by estimating the model year by year. Due to the low number of observations in each regression, in Table 2 and Table 3 I don't

³⁶Interestingly, the coefficients of the two interaction variables are equal in absolute value.

³⁷Regression (6) does not include the regressors *log distance*, *land border*, *common language* and *colony* since they are constant within country pairs and, therefore, they would be perfectly collinear with the dummy variables.

control for country-specific fixed effects, which explains the difference in the magnitude of the effects relative to the regressions in Table 1. The coefficients are still qualitatively consistent with the panel-data results, though less precisely estimated.

In Table 4 I run a few robustness checks of the panel-data results. In the first regression, I use (within-country variation in) the terms of trade to instrument for (within-country variation in) the level of per worker GDP of both destination and origin countries. Terms of trade are correlated with the average real income of a country since they affect countries' purchasing power vis a vis goods produced by the rest of the world. In the first stage, the impact of the terms of trade on per worker GDP is positive and significant at the 1% level, for both destination and origin countries. In addition, if the assumption of small open economies holds, terms of trade are unlikely to affect emigration rates directly or to be correlated with other country-level characteristics that have an impact on migration patterns (exclusion restriction). Regression (1) shows that pull and push coefficients are robust to endogeneity issues as the instrumental variable estimates are consistent with my earlier results.

In columns (2) and (3), I ask whether per worker GDP in the two locations is a good measure of the mean income opportunities for migrant workers at home and abroad. I first control for the average schooling level in both countries in specification (2). Pull effects are still estimated to be positive and significant (at the 1% level). In line with the theoretical predictions, the average skill level in the population of the destination (origin) country has a negative (positive) impact on the emigration rate. In column (3) I control for per worker endowments of both skill and capital and again find that my prior findings are robust.

In Table 5 I examine the predictions based on the extensions of the theoretical framework of Section 2. I find only weak evidence of poverty constraints in regression (1) and confirm, in the following specification, my results on push and pull factors in terms of unemployment rates. Third-country effects shape bilateral migration flows as expected, given that the coefficient on the *multilateral pull* term is negative and significant.³⁸ Finally, the size of emigration rates is affected by the second moments of the income distributions in the two location in the way predicted by Borjas (1987) selection model.

To conclude, I analyze network effects by introducing the lagged emigration rate(s) among the explanatory variables (Table 6). Emigration rates clearly show a lot of inertia. However, it's unclear whether supply or demand factors explain the autocorrelation in bilateral migration flows.³⁹

³⁸The multilateral pull term places migrants' decision to move in a multi-country framework. It is inspired by the multilateral trade resistance term in Anderson and van Wincoop (2003) (even though mine is an atheoretical measure).

³⁹Fixed-effects estimation of a dynamic model with a short panel (small T) produces biased estimates. This might explain the estimates in the first two columns of push and pull factors (which are different from previous ones). Once I use Arellano and Bond's estimator, the coefficients become consistent with what I previously found.

6 Conclusions

In this paper, I investigate economic and non-economic determinants of international migration flows. This analysis both delivers estimates consistent with the predictions of a standard international-migration model and generates empirical puzzles.

In particular, I find that pull factors, that is improvements in the income opportunities in the destination country, significantly increase the size of emigration rates. This result, which appears to be very robust to changes in the specification of the empirical model, is surprising, given very restrictive immigration policies of the destination countries included in the sample. The sign of the impact of push factors - that is, declining levels of per worker GDP in the origin country - is seldom consistent with the theoretical predictions of the model and, when it is, the size of the effect is smaller than for pull factors and is almost always insignificant. Among the variables affecting the costs of migration, distance appears to be one of the most important ones. Its effect is negative, significant and quite steady across specifications.

The investigation of the determinants of international migration leads to other interesting research questions. This analysis provides a framework to address policy-related issues, as has been done in the trade gravity-model literature. The analysis of the impact of labour movements on source and host economies - on their standards of living, for example - faces the inherent problems of endogeneity of migration flows and reverse causality. Since this work helps isolate the exogenous determinants of immigrant flows, it is possible to use it to construct a *first stage* for this type of analyses (as in Frankel and Romer 1999, for example).

By taking advantage of both the time-series and cross-country variation in an annual panel data set, this paper makes progress in explaining the economic and non-economic determinants of international migration flows.

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Table 1. Panel data regressions

Equation	1	2	3	4	5	6
Dependent variable	Emigration rate					
log per worker gdp (destination)	0.00026	0.00026	0.00029	0.00028	0.00028	0.0002
	0.00011*	0.00011*	0.00011*	0.00011*	0.00011*	0.00004**
log per worker gdp (origin)	0.00005	0.00005	0.00003	0.00004	0.00004	-0.00007
	0.00009	0.00008	0.00008	0.00008	0.00008	0.00003*
log distance	-0.0004	-0.00039	-0.00041	-0.00041	-0.00037	
	0.00009**	0.00009**	0.00009**	0.00009**	0.00008**	
land border	-0.00026	-0.00036	-0.00037	-0.00037		
	0.00019	0.00023	0.00023	0.00023		
common language		0.00027	0.00022	0.00022	0.00019	
		0.00016+	0.00016	0.00016	0.00012	
colony		-0.00001	0.00003	0.00003		
		0.00017	0.00017	0.00017		
share of young population (origin)			0.00242	0.00222	0.00223	0.00129
			0.00110*	0.00108*	0.00109*	0.00040**
per worker gdp (destination)*immig policy change				0.00006	0.00006	0
				0.00002**	0.00002**	0.00001
per worker gdp (origin)*immig policy change				-0.00006	-0.00006	0
				0.00002**	0.00002**	0.00001
constant	0.00053	0.00044	-0.00027	-0.00008	-0.00045	-0.00168
	0.00112	0.00107	0.00117	0.00115	0.0011	0.00050**
number of observations	8294	8294	8010	8010	8010	9767
R-squared	0.24	0.25	0.25	0.25	0.25	0.85

OLS estimates. Destination and origin countries' dummy variables are included in each specification, except the last one (in regression 6, country-pairs dummy variables are included). Standard errors, clustered by country pairs, are presented under each estimated coefficient. + significant at 10%; * significant at 5%; ** significant at 1%

per worker gdp is the level of per worker GDP, PPP-adjusted (constant 1996 international dollars), lagged by one year.

distance is the great-circle distance. *Land border* equals one if the destination and origin countries share a land border.

common language equals one if a common language is spoken in both destination and origin countries.

colony is a dummy variable for pairs of countries ever in a colonial relationship.

share of young population (origin) is the share of the population in the origin country aged 15-29.

immigration policy change equals one if in that year the destination country's immigration policy became less restrictive, minus one if it became more restrictive, zero if there was no change in immigration policy.

See Appendix 1 for data sources.

Table 2. Cross-country regressions

Equation	1	2	3
	1981-1985	1986-1990	1991-1995
Dependent variable	Emigration rate		
log per worker gdp (destination)	0.00028 0.00015+	0.00068 0.00022**	0.00088 0.00020**
log per worker gdp (origin)	-0.00002 0.00003	-0.00004 0.00004	-0.00004 0.00003
log distance	-0.00007 0.00003*	-0.0001 0.00004*	-0.0001 0.00003**
common language	0.00037 0.00007**	0.0004 0.00008**	0.00025 0.00007**
constant	-0.00208 0.00166	-0.00592 0.00240*	-0.00809 0.00222**
number of obs	571	628	709
R-squared	0.06	0.06	0.06

OLS estimates. The standard errors are presented under each estimated coefficient.

+ significant at 10%; * significant at 5%; ** significant at 1%

per worker gdp is the level of per worker GDP, PPP-adjusted (constant 1996 international dollars), averaged over the five years preceding the relevant period (1976-1980 for the regression in column (1), for example).

distance is the great-circle distance.

common language is a dummy variable equal to one if a common language is spoken in both destination and origin countries.

See Appendix 1 for data sources.

Table 3. Yearly regressions

Equation	1	2	3	4	5	6	7	8
	1981	1982	1983	1984	1985	1986	1987	1988
Dependent variable	Emigration rate							
log per worker gdp (destination)	0.0001	0.00079	0.00097	0.00038	0.00044	0.00052	0.00058	0.00062
	0.00039	0.00034*	0.00033**	0.00018*	0.00017*	0.00021*	0.00024*	0.00023**
log per worker gdp (origin)	-0.00008	-0.00006	-0.00004	-0.00002	-0.00002	-0.00003	-0.00003	-0.00003
	0.0001	0.00008	0.00006	0.00003	0.00003	0.00004	0.00004	0.00004
log distance	-0.00082	-0.00031	-0.00028	-0.00011	-0.00011	-0.00012	-0.00013	-0.00014
	0.00018**	0.00011**	0.00008**	0.00003**	0.00003**	0.00004**	0.00005**	0.00004**
common language	0.00058	0.00033	0.00024	0.00032	0.00033	0.0004	0.00048	0.00043
	0.00021**	0.00017+	0.00013+	0.00007**	0.00007**	0.00009**	0.00010**	0.00009**
share of young population (origin)	0.0056	0.00685	0.00476	0.00275	0.00265	0.00301	0.00342	0.00311
	0.00321+	0.00281*	0.00209*	0.00107*	0.00105*	0.00121*	0.00142*	0.00128*
constant	0.00532	-0.00682	-0.00855	-0.00359	-0.00419	-0.00494	-0.00561	-0.00599
	0.00497	0.00385+	0.00357*	0.00198+	0.00193*	0.00233*	0.00272*	0.00254*
number of obs	165	220	275	551	551	551	551	551
R-squared	0.19	0.11	0.11	0.07	0.07	0.08	0.08	0.08

OLS estimates. The standard errors are presented under each estimated coefficient. + significant at 10%; * significant at 5%; ** significant at 1%

per worker gdp is the level of per worker GDP, PPP-adjusted (constant 1996 international dollars), lagged by one year.

distance is the great-circle distance.

common language is a dummy variable equal to one if a common language is spoken in both destination and origin countries.

share of young population (origin) is the share of the population in the origin country aged 15-29.

See Appendix 1 for data sources.

Table 3. Yearly regressions (cont.)

Equation	1	2	3	4	5	6	7
	1989	1990	1991	1992	1993	1994	1995
Dependent variable	Emigration rate						
log per worker gdp (destination)	0.0011	0.00158	0.00171	0.0011	0.00128	0.0012	0.00093
	0.00032**	0.00034**	0.00032**	0.00025**	0.00022**	0.00021**	0.00020**
log per worker gdp (origin)	-0.00001	-0.00003	-0.00005	-0.00004	-0.00003	-0.00001	-0.00001
	0.00005	0.00005	0.00005	0.00003	0.00003	0.00003	0.00003
log distance	-0.00019	-0.00022	-0.00021	-0.00015	-0.00013	-0.00012	-0.00011
	0.00005**	0.00006**	0.00005**	0.00003**	0.00003**	0.00003**	0.00003**
common language	0.00043	0.00029	0.00018	0.00019	0.00021	0.00019	0.00021
	0.00011**	0.00011*	0.00010+	0.00007**	0.00006**	0.00006**	0.00006**
share of young population (origin)	0.00443	0.00513	0.00405	0.00172	0.00175	0.00176	0.00156
	0.00168**	0.00176**	0.00150**	0.00099+	0.00090+	0.00084*	0.00080+
constant	-0.01102	-0.01602	-0.01706	-0.01047	-0.01266	-0.01213	-0.00918
	0.00353**	0.00372**	0.00349**	0.00269**	0.00243**	0.00223**	0.00215**
number of obs	551	606	673	697	709	709	650
R-squared	0.08	0.08	0.08	0.07	0.08	0.08	0.08

OLS estimates. The standard errors are presented under each estimated coefficient.

+ significant at 10%; * significant at 5%; ** significant at 1%

per worker gdp is the level of per worker GDP, PPP-adjusted (constant 1996 international dollars), lagged by one year.

distance is the great-circle distance.

common language is a dummy variable equal to one if a common language is spoken in both destination and origin countries.

share of young population (origin) is the share of the population in the origin country aged 15-29.

See Appendix 1 for data sources.

Table 4. Panel data regressions: Robustness Checks

Equation	1	2	3	4
Method	Instrumental Variables Estimation	OLS	OLS	TOBIT
Dependent variable	Emigration rate			
log per worker gdp (destination)	0.00041	0.00037	0.00053	0.00086
	0.00015**	0.00011**	0.00019**	0.00036*
log per worker gdp (origin)	-0.00023	0.00006	-0.00018	-0.00009
	0.0002	0.0001	0.00013	0.00023
log distance	-0.00031	-0.00037	-0.00029	-0.00143
	0.00007**	0.00008**	0.00008**	0.00004**
common language	0.00007	0.0002	0.00015	0.0012
	0.00008	0.00013	0.00011	0.00009**
share of young population (origin)	0.00065	0.00218	-0.00014	0.00683
	0.001	0.00114+	0.00113	0.00300*
log yrs schooling (destination)		-0.00047	-0.00032	
		0.00015**	0.00018+	
log yrs schooling (origin)		0.00008	0.00048	
		0.00012	0.00037	
log capital per worker (destination)			-0.00028	
			0.00013*	
log capital per worker (origin)			0.00011	
			0.00014	
constant	-0.00001	-0.0009	0.00049	0.00063
	0.00243	0.00126	0.00094	0.00407
number of obs	7411	7313	4103	8010
R-squared	0.23	0.25	0.26	

Destination and origin countries' dummy variables are included in each specification. Standard errors, clustered by country pairs, are presented under each estimated coefficient. + significant at 10%; * significant at 5%; ** significant at 1%

In regression (1), I use terms of trade (lagged by one year) as an instrument for per worker GDP (lagged by one year) in both destination and origin country.

per worker gdp is the level of per worker GDP, PPP-adjusted (constant 1996 international dollars), lagged by one year.

distance is the great-circle distance. *common language* is a dummy variable equal to one if a common language is spoken in both destination and origin countries. *share of young population (origin)* is the share of the population in the origin country aged 15-29.

log yrs schooling is the log of the average schooling years in the total population over age 15, lagged by one year.

log capital per worker is non-residential capital stock per worker (1985 intl. prices), lagged by one year.

See Appendix 1 for data sources.

Table 5. Panel data regressions: Economic determinants more in detail

Equation	1	2	3	4
Dependent variable	Emigration rate			
log per worker gdp (destination)	0.00032		0.00038	0.00038
	0.00012**		0.00013**	0.00018*
log per worker gdp (origin)	0.00083		0.00005	0.00007
	0.00049+		0.00008	0.00017
square of log per worker gdp (origin)	-0.00004			
	0.00003			
unemployment rate (destination)		-0.00002		
		0.00001*		
unemployment rate (origin)		0.00001		
		0.00001		
multilateral pull			-0.00006	
			0.00003*	
origin country's relative inequality				0.00077
				0.00039*
square of relative inequality				-0.00028
				0.00013*
log distance	-0.00037	-0.00029	-0.00036	
	0.00008**	0.00008**	0.00007**	
common language	0.00019	0.00028	0.00019	
	0.00012	0.00014*	0.00012	
share of young population (origin)	0.00225	0.0007	0.0024	
	0.00111*	0.00083	0.00110*	
constant	-0.00433	0.00236	-0.0016	-0.00513
	0.00272	0.00083**	0.00132	0.00182**
number of observations	8010	5148	8010	4028
R-squared	0.24	0.19	0.24	0.18

OLS estimates. Destination and origin countries' dummy variables are included in each specification. Standard errors, clustered by country pairs, are presented under each estimated coefficient. + significant at 10%; * significant at 5%; ** significant at 1%. See the end of Table 1 for definitions of the main variables used.

multilateral pull gives, for each destination/origin country pair, the average of (log per worker gdp (destination)-log distance) over all the other destination countries.

origin country's relative inequality gives a measure of the inequality in the origin country relative to the destination country (it equals the gini coefficient in the origin country divided by the gini coefficient in the destination country).

See Appendix 1 for data sources.

Table 6: Dynamic regressions: network effects

Equation	1	2	3
	destination and origin countries d.v.	country pair d.v.	Arellano and Bond estimator
Dependent variable	Emigration rate		
emigration rate(t-1)	0.91536	0.60247	0.66532
	0.02641**	0.04166**	0.01485**
per worker gdp (destination)	-0.00003	0.00005	0.00018
	0.00003	0.00003+	0.00007*
per worker gdp (origin)	-0.00001	-0.00001	0.00003
	0.00002	0.00003	0.00005
log distance	-0.00003		
	0.00001**		
constant	0.00066	-0.00041	0.00000
	0.00027*	0.00025	0.00000
number of obs	8294	10101	6429
R-squared	0.88	0.9	

OLS estimates. The standard errors, clustered by country pair, are presented under each estimated coefficient.

+ significant at 10%; * significant at 5%; ** significant at 1%. *per worker gdp* is the level of per worker GDP, PPP-adjusted (constant 1996 international dollars), lagged by one year. *distance* is the great-circle distance. See Appendix 1 for data sources.

In equation (3), I include the emigration rate lagged by one, two, three and four years (the coefficients on the latter three lags are not shown). The reason is that, only by introducing all these lags, I don't reject the null of zero autocovariance in residuals of order 2 (which is one of the requirements of the Arellano and Bond estimator).

Column (3): Arellano-Bond test that average autocovariance in residuals of order 1 is 0: H0: no autocorrelation $z = -56.35$ $Pr > z = 0.0000$.

Column (3): Arellano-Bond test that average autocovariance in residuals of order 2 is 0: H0: no autocorrelation $z = -0.48$ $Pr > z = 0.6329$.

Appendix 1. Summary Statistics (1980-1995) and Data Sources

Variable	Obs	Mean	Std. Dev.	Min	Max
emigration rate	8010	0.0001324	0.0008154	0	0.0156894
per worker gdp (destination)	8010	40681.73	5895.097	25251.65	55361.18
per worker gdp (origin)	8010	20060.68	14106.18	1027.362	55361.18
distance	8010	8.171532	0.8693557	5.08715	9.383615
land border	8010	0.0268414	0.16163	0	1
common language	8010	0.170412	0.3760178	0	1
colony	8010	0.0384519	0.1922966	0	1
years schooling (destination)	8010	9.713563	1.3133	6.837	11.865
years schooling (origin)	7313	6.470704	2.521323	1.897	11.865
capital per worker (destination)	6651	36168.39	12168.11	16992	76733
capital per worker (origin)	4341	18267.41	13528.63	702	48135
share of young population (origin)	8010	0.2612177	0.0302848	0.1951073	0.3152469
unemployment rate (destination)	7667	6.730299	3.506799	0.5	14.1
unemployment rate (origin)	5263	8.137606	5.269609	0.08	27.6
origin country's relative inequality	1472	1.130724	0.3811316	0.3477012	2.59918

Data sources

The emigration rate (immigrant inflow from origin to destination country, divided by origin country's population) is from the IMS data set (OECD 1997).

Per worker GDP, PPP-adjusted (constant 1996 international dollars) is from the Penn World Tables, version 6.1.

Distance, land border, common language, and colony (countries ever in a colonial relationship) are from Glick and Rose (2001).

Years of schooling are from Barro and Lee (2000) data set.

Capital per worker (Nonresidential Capital Stock per Worker, 1985 intl. prices) is from the Penn World Tables, version 5.6.

The share of young population (origin) is based on data from the United Nations.

The unemployment rate is from the World Development Indicators (2001), World Bank.

The origin country's relative inequality is based on data on Gini coefficients from Deininger and Squire (1996) data set (only high-quality observations were used).

Data on terms of trade comes from the World Bank's Global Development Network Growth Database, Macro Time Series (2002).

The data set on immigration policy changes was constructed by Anna Maria Mayda and Krishna Patel (2004) - see Appendix to the paper.

All time-varying variables are lagged by one year.

Appendix 2. Average inflows into each destination country, by country of origin (1980-1995)

Australia (1983-1995)	
country of origin	inflow
UK	17095
New Zealand	11045
Vietnam	8048
Hong Kong	5739
Philippines	5379
Malaysia	3493
India	3069
China	2934
Former Yugoslavia	2790
South Africa	2441
Sri Lanka	2146
Lebanon	2089
USA	1724
Fiji	1682
Poland	1608
Ireland	1462
Chinese Taipei	1358
Germany	1303
Former USSR	1021
Portugal	767
<i>total (above inflows)</i>	77193
<i>percentage change</i>	-6.22%

Belgium (1984-1995)	
country of origin	inflow
France	6072
Netherlands	6014
USA	2930
Germany	2916
UK	2899
Morocco	2801
Italy	2495
Turkey	2239
Zaire	1966
Portugal	1435
Japan	833
Spain	833
Former Yugoslavia	829
Greece	759
Poland	655
China	589
Algeria	382
Tunisia	310
<i>total (above inflows)</i>	36957
<i>percentage change</i>	13.46%

Canada (1980-1995)	
country of origin	inflow
Hong Kong	19334
India	10437
Philippines	9441
UK	9034
Vietnam	8791
Poland	7550
USA	7459
China	6292
Lebanon	3917
Sri Lanka	3791
Portugal	3653
Jamaica	3543
Chinese Taipei	3255
Guyana	3108
El Salvador	2697
Haiti	2243
Iran	2193
France	2070
Former Yugoslavia	1933
South Korea	1584
Trinidad Tobago	1433
Romania	1241
Pakistan	1037
Former USSR	791
Somalia	195
<i>total (above inflows)</i>	117022
<i>percentage change</i>	48.29%

Denmark (1990-1994)	
country of origin	inflow
Somalia	1264
UK	1068
Turkey	1042
Germany	805
Iraq	789
Norway	699
Sweden	612
USA	606
Iran	570
Vietnam	549
Former Yugoslavia	481
Iceland	479
Poland	448
Thailand	366
Pakistan	356
Lebanon	335
Netherlands	304
France	269
Morocco	215
Italy	200
Finland	181
<i>total (above inflows)</i>	11638
<i>percentage change</i>	75.28%

total (above inflows) is the sum of the average immigrant inflows (1980-1995) by country of origin from the table.

percentage change is the percentage change of the *overall total* during the period between 1980 and 1995.

Appendix 2. Average inflows into each destination country, by country of origin (1980-1995) (cont.)

France (1984-1995)		Germany* (1984-1995)		Japan (1980-1995)		Luxembourg (1983-1995)	
country of origin	inflow	country of origin	inflow	country of origin	inflow	country of origin	inflow
Morocco	11892	Poland	117019	China	35425	Portugal	2170
Algeria	9187	Former Yugoslavia	92124	USA	35367	France	1272
Turkey	5777	Bosnia-Herzegovina	76836	Philippines	35121	Belgium	897
Tunisia	3083	Turkey	68791	South Korea	21052	Germany	662
Lebanon	2818	Romania	61910	Chinese Taipei	10882	Italy	441
USA	2403	Italy	39184	UK	9614	Netherlands	281
Haiti	2183	Croatia	24056	Brazil	6779	USA	256
Portugal	2050	Former USSR	23365	Hong Kong	6296	Spain	124
Vietnam	1761	Hungary	21835	Thailand	5913	<i>total (above inflows)</i>	6103
Zaire	1437	Greece	20372	Germany	5334	<i>percentage change</i>	29.73%
Poland	1422	Bulgaria	19245	Canada	3449		
Japan	1219	USA	17670	Peru	1008		
China	1084	Former CSFR	10692	<i>total (above inflows)</i>	176240		
Former Yugoslavia	1084	Portugal	9654	<i>percentage change</i>	-42.10%		
Sri Lanka	899	Spain	4705				
Romania	891	Morocco	4375				
Cambodia	860	Slovenia	2658				
Spain	400	Tunisia	2249				
<i>total (above inflows)</i>	50450	<i>total (above inflows)</i>	616740				
<i>percentage change</i>	-6.23%	<i>percentage change</i>	24.85%				

total (above inflows) is the sum of the average immigrant inflows (1980-1995) by country of origin from the table.

percentage change is the percentage change of the *overall total* during the period between 1980 and 1995.

* Figures for migrants from the former Yugoslavia to Germany do not include Croatia from 1992 and Bosnia-Herzegovina from 1993. Data from the former USSR to Germany does not include Russia from 1992.

Appendix 2. Average inflows into each destination country, by country of origin (1980-1995) (cont.)

Netherlands (1984-1995)	
country of origin	inflow
Turkey	8363
Former Yugoslavia	7392
Morocco	6537
Germany	5295
UK	4575
Suriname	4416
USA	2303
Belgium	2050
France	1517
Poland	1148
Italy	893
<i>total (above inflows)</i>	44489
<i>percentage change</i>	-16.04%

Norway* (1984-1995)	
country of origin	inflow
Bosnia-Herzegovina	3728
Denmark	2201
Sweden	1526
UK	1253
USA	987
Former Yugoslavia	868
Pakistan	682
Iran	669
Vietnam	612
Chile	537
Somalia	468
Sri Lanka	450
Germany	399
<i>total (above inflows)</i>	14380
<i>percentage change</i>	39.83%

Sweden (1984-1995)	
country of origin	inflow
Bosnia-Herzegovina	16972
Iran	4048
Finland	3880
Norway	3118
Former Yugoslavia	2840
Iraq	2051
Denmark	1877
Somalia	1724
Chile	1631
Poland	1484
Turkey	1214
Ethiopia	947
Russian Federation	910
Lebanon	896
USA	831
Croatia	784
Germany	761
Romania	746
UK	715
Thailand	603
India	369
Greece	311
<i>total (above inflows)</i>	48712
<i>percentage change</i>	61.88%

Switzerland (1984-1995)	
country of origin	inflow
Former Yugoslavia	18716
Portugal	9085
Germany	8333
Italy	8216
France	4655
Spain	4402
Turkey	4195
USA	2530
UK	2407
Austria	1728
Netherlands	1607
Canada	687
<i>total (above inflows)</i>	66561
<i>percentage change</i>	24.68%

total (above inflows) is the sum of the average immigrant inflows (1980-1995) by country of origin from the table.

percentage change is the percentage change of the *overall total* during the period between 1980 and 1995.

* Figures for migrants from the former Yugoslavia to Norway do not include Bosnia-Herzegovina from 1993.

Appendix 2. Average inflows into each destination country, by country of origin (1980-1995) (cont.)

United Kingdom (1982-1995)

country of origin	inflow
Pakistan	5817
India	5047
Bangladesh	3796
USA	3776
Australia	2659
New Zealand	1964
Nigeria	1556
Iran	1501
Japan	1474
Hong Kong	1287
Ghana	1093
Canada	1035
Sri Lanka	1021
Philippines	986
South Africa	926
Turkey	822
Jamaica	775
Malaysia	701
Iraq	500
Kenya	481
Poland	481
Thailand	444
Germany	419
Cyprus	402
Morocco	380
Spain	363
Sweden	355
France	345
Italy	340
Netherlands	289
Former Yugoslavia	276
Portugal	223

total (above inflows) 41534

percentage change -20.49%

United States (1980-1995)

country of origin	inflow
Mexico	199862
Philippines	51886
Vietnam	45041
China	32824
Dominican Republic	30471
India	29754
South Korea	29197
Former USSR	23231
El Salvador	21901
Jamaica	20219
Cuba	15174
Haiti	15168
UK	14939
Iran	14596
Poland	13534
Canada	12980
Chinese Taipei	12962
Colombia	12696
Laos	12165
Ireland	12054
Guatemala	9328
Guyana	9243
Cambodia	8108
Pakistan	7725
Peru	7637
Germany	7005
Hong Kong	6994
Thailand	6270
Ecuador	6189
Nicaragua	5626
Honduras	5507
Bangladesh	2684

total (above inflows) 702970

percentage change 35.79%

total (above inflows) is the sum of the average immigrant inflows (1980-1995) by country of origin from the table.
percentage change is the percentage change of the *overall total* during the period between 1980 and 1995.

Appendix 3. Total yearly inflows into each destination country, using data by country of origin (1980-1995)

Country of	1980	1981	1982	1983	1984	1985	1986	1987	1988	1989	1990	1991	1992	1993	1994	1995	TOTAL
Australia				69590	50729	57670	67543	83195	108277	113977	96460	98756	83289	57490	51623	64908	1E+06
Belgium					9200	8000	8800	10000	9500	11100	15142	18924	41762	39228	39472	40214	251342
Canada	100909	92308	85085	64669	64920	61187	71267	111124	116331	136398	162417	130550	141730	143184	132668	113164	2E+06
Denmark											8922	10064	9614	8063	10744		47407
France					25300	21000	22800	23300	25700	31600	60159	59608	50528	52160	37782	32275	442212
Germany					108400	127700	159100	173600	223100	233600	485500	583500	925518	680315	475267	470590	5E+06
Japan	266400	278100	103600	80900	100100	120100	122800	139300	191000	186600	176900	204500	214100	183245	186431	153948	3E+06
Luxembourg				4350	3911	4537	5042	4832	6535	6502	7086	7182	6101	6679	6617	6649	76023
Netherlands					23900	29000	33900	36300	37400	40800	48962	49868	50989	51268	35463	28738	466588
Norway					8998	9707	10459	14962	15378	11190	9122	9465	10615	16154	11596	9480	137126
Sweden					18300	19900	24800	27700	32900	44100	36100	25900	23100	43477	64099	26225	386601
Switzerland					45500	46500	52800	57700	60400	64200	80994	85545	88695	81181	69362	65845	798722
UK			43500	41540	40240	45020	38720	36980	39440	40170	42410	42310	41050	41950	42260	42650	578240
US	418341	485642	490427	459946	443291	466307	495255	493507	528355	951871	1E+06	2E+06	805446	761217	675347	573784	1E+07

Source: OECD (1997)

Appendix 3. Total yearly inflows into each destination country, using total inflows data (1980-1995)

Country of	1980	1981	1982	1983	1984	1985	1986	1987	1988	1989	1990	1991	1992	1993	1994	1995	TOTAL	%
Australia				93200	69800	78100	92400	113300	143500	145300	121200	121700	107400	76300	69800	87400	1E+06	76.06%
Belgium					37200	37500	39300	40100	38200	43500	50500	54100	55000	53000	56000	53100	557500	45.08%
Canada	143100	128600	121100	89200	88200	84300	99200	152100	161900	192000	214200	230800	252800	255800	223900	212200	3E+06	65.22%
Denmark											15100	17500	16900	15400	15600		80500	58.89%
France					51400	43400	38300	39000	44000	53200	102400	109900	116600	99200	69300	55700	822400	53.77%
Germany					331100	398200	478300	473300	648500	770800	842400	920500	1E+06	986900	774000	788300	9E+06	53.90%
Japan	362500	393100	134300	108100	131100	156500	157500	180300	234800	237400	223800	258400	267000	234500	237500	209900	4E+06	76.79%
Luxembourg				6200	6000	6600	7400	7200	8200	8400	9300	10000	9800	9200	9200	9600	107100	70.98%
Netherlands					37300	46200	52800	60900	58300	65400	81300	84300	83000	87600	68400	67000	792500	58.88%
Norway					12800	15000	16800	23800	23200	18500	15700	16100	17200	22300	17900	16500	215800	63.54%
Sweden					26100	27900	34000	37100	44500	58900	53200	43900	39500	54800	74800	36100	530800	72.83%
Switzerland					58600	59400	66800	71500	76100	80400	101400	109800	112100	104000	91700	87900	1E+06	78.33%
UK			53900	53500	51000	55400	47800	46000	49300	49600	53200	53900	52600	55600	55100	55500	732400	78.95%
US	530600	596600	594100	559800	543900	570000	601700	601500	643000	1E+06	2E+06	2E+06	974000	904300	804400	720500	1E+07	84.47%

Source: OECD (1997). % covered equals (TOTAL (disaggregate)/TOTAL (aggregate)) (in percentage terms).

Figure 1: The actual emigration rate as a function of mean income opportunities in the destination and origin country and of mean moving costs

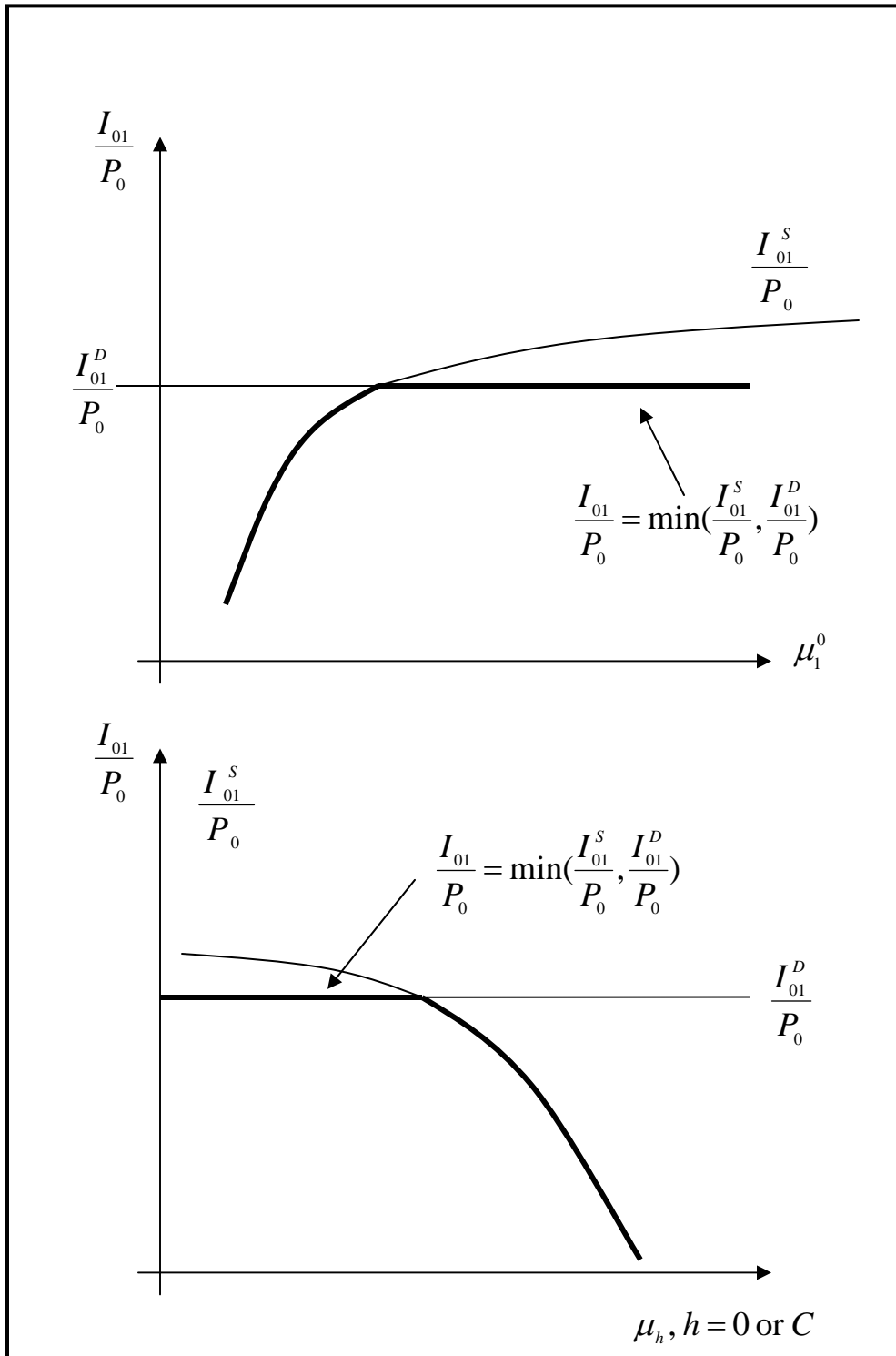
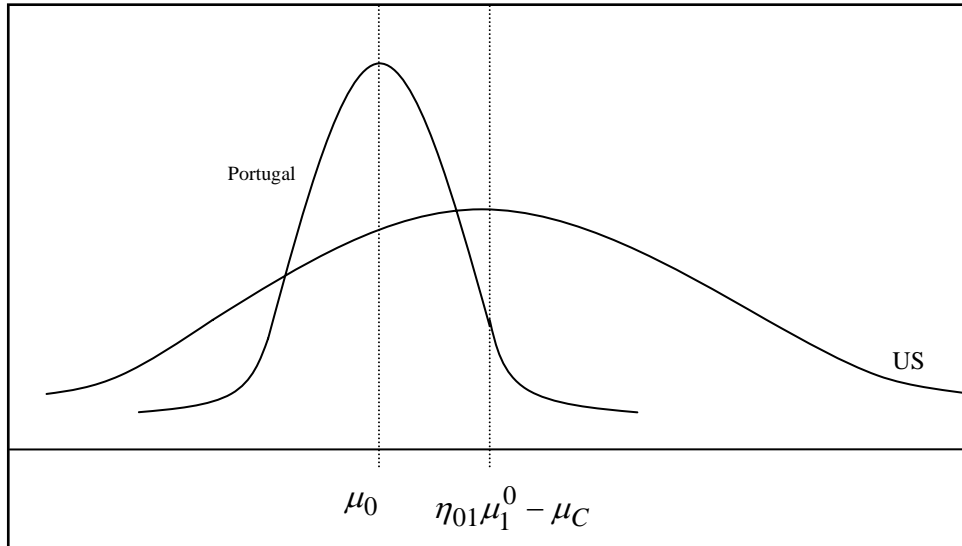
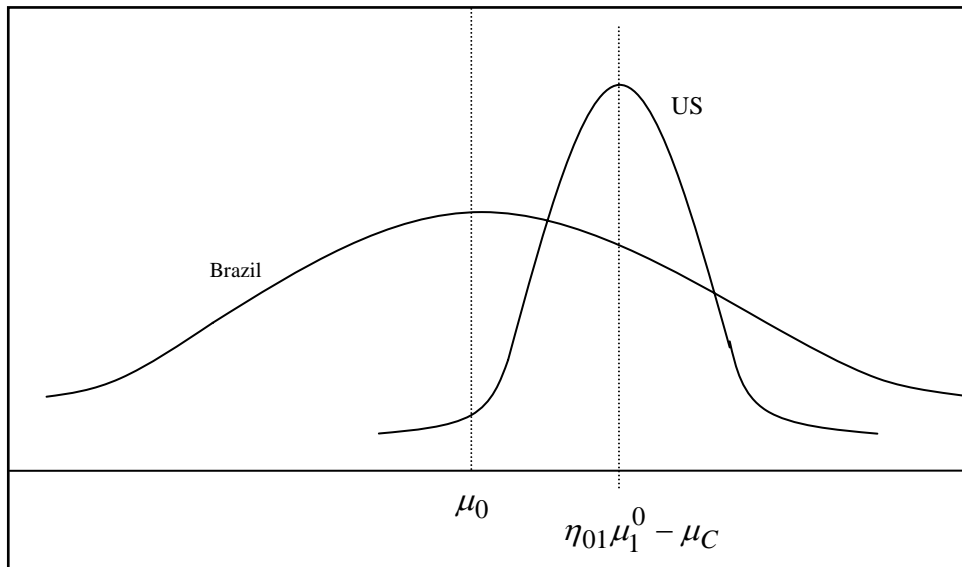


Figure 2: The impact of relative inequality on emigration rates according to Borjas (1987) selection model

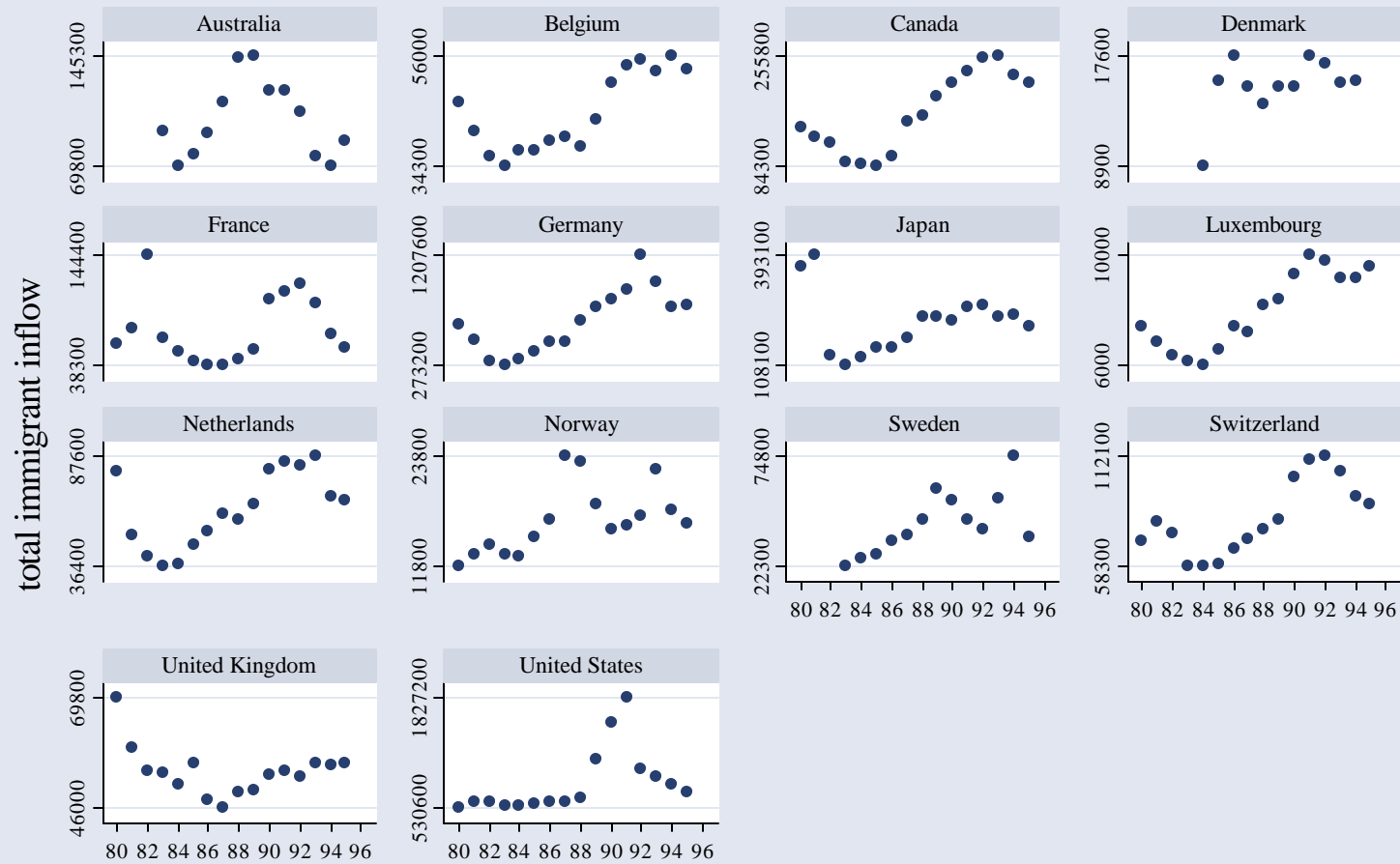


Each curve represents the net labor income in Portugal and in the U.S. respectively. Income inequality is assumed to be higher in the U.S. than in the Portugal, consistent with the data. The Gini coefficient for Portugal was 36.76 in 1990, while in the U.S. it was 37.8 (Deininger and Squire 1996).



Each curve represents the net labor income in Brazil and in the U.S. respectively. Income inequality is assumed to be higher in Brazil than in the U.S., consistent with the data. The Gini coefficient for Brazil was 61.76 in 1985, while in the U.S. it was 37.26 ((Deininger and Squire 1996).

Figure 3: Total immigrant inflow by destination country



Graphs by country of destination

year