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## Is the equalizing effect of retirement wealth wearing off?

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**Working Paper No. 420**

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**Is the Equalizing Effect of Retirement Wealth Wearing Off?**

by

Edward N. Wolff

The Levy Economics Institute, New York University, and NBER

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## **ABSTRACT**

Retirement wealth is often viewed as a great equalizer, offsetting the inequality in standard household net worth. One of the most dramatic changes in the retirement income system over the last two decades has been a decline in traditional Defined Benefit (DB) pension plans and a sharp rise in Defined Contribution (DC) pensions. Using data from the Federal Reserve Board's Survey of Consumer Finances, I find that retirement wealth (the sum of pension and Social Security wealth) had a considerably weaker offsetting effect on wealth inequality in 2001 than in 1983. Whereas standard net worth inequality increased modestly between 1983 and 2001, the inequality of augmented wealth (the sum of retirement wealth and net worth) surged from 1983 to 2001, very much in line with income inequality. Moreover, whereas median net worth climbed substantially from 1983 to 2001, median augmented wealth actually fell over this period.

Keywords: Pensions, Social Security, Wealth, Inequality

JEL codes: J32, D31, H55

## 1. INTRODUCTION

While almost all the attention of the media has been riveted on the Social Security system, the devolution of the private pension system has received surprisingly little attention. Indeed, one of the most dramatic changes in the retirement income system over the last two decades has been the replacement of many traditional Defined Benefit (DB) pension plans with Defined Contribution (DC) pensions. Moreover, retirement wealth is often viewed as the great equalizer, offsetting the inequality in standard household net worth. The main focus of the paper is to analyze the effects of this substitution on the overall distribution of household wealth.

The work of Poterba, Venti, and Wise (1998) suggests that the transition from DB to DC type plans increased pension wealth dramatically. My results, reported below, also confirm that *mean* pension wealth rose strongly between 1983 and 2001. However, *median* pension wealth grew much more slowly over this period and pension wealth inequality grew sharply as well.

A related issue that will be covered in this paper regards the distributional effects of Social Security wealth. Feldstein (1976), in a seminal paper on this subject found that the inclusion of Social Security wealth led to a sharp reduction in overall wealth inequality. I also find the same effect. However, my results show that median Social Security wealth grew more slowly than net worth over the period from 1983 to 2001 and its equalizing effect dissipated over these years.

The results of this paper also clear up a “puzzle” encountered in earlier work (Wolff, 2001). In particular, the share of total household wealth held by the richest one percent of households increased only a bit between 1989 and 1998, and the Gini coefficient for total household wealth actually declined slightly. I expected both the share of the top one percent and the Gini coefficient to have shown a considerable increase.

The reason is that in previous work (see Wolff, 2002) I identified two factors that seem to underlay much of the change in the share of wealth held by the top one percent. The first is the change in underlying income inequality and the second is the change in the ratio of stock prices to housing prices (see Figure A). In a simple regression of the share of the top one percent on these two factors, both variables proved statistically significant and the goodness of the fit of the equation was quite high.<sup>1</sup> Over the period from 1989 to 1998, income inequality, as

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<sup>1</sup> A regression of a wealth inequality index, measured by the share of marketable wealth held by the top 1 percent of households (WLTH) on income inequality, measured by the share of income received by the top 5 percent of

measured by the share of the top five percent increased by 2.8 percentage points, and the ratio of share prices to housing prices surged by a factor of 2.5. Extrapolating on the basis of the regression estimates, I would have expected a **9.9** percentage point increase in the share of the top one percent between 1989 and 1998, compared to its actual gain of 0.7 percentage points (see Figure B).

Wojciech Kopczuk and Emmanuel Saez (2004), using U.S. estate tax data from 1916 to 2000, also find very little change in the shares of wealth held by the top wealth groups in the 1990s. Indeed, they find very little change in the 1980s as well. The share of the top one percent was 21.1 percent in 1983 and 20.8 percent in 2000.

The next section of the paper (Section 2) provides a review of the pertinent literature on this subject. Section 3 describes the data sources and develops the accounting framework used in the analysis. Section 4 shows time trends in standard measures of household wealth over the 1983-2001 period. Sections 5 and 6 investigate changes in pension wealth and Social Security wealth, respectively, over this period. Section 7 presents summary measures on total (augmented) household wealth. Alternative pension wealth calculations are shown in Section 8. Concluding remarks are made in Section 9.

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families (INC), and the ratio of stock prices (the Standard and Poor index) to housing prices (RATIO), with twenty-one data readings between 1922 and 1998, yields:

$$\begin{array}{rcccc}
 \text{WLTH} = & 5.10 & + & 1.27 \text{ INC} & + & 0.26 \text{ RATIO}, & R^2 = 0.64, & N = 21 \\
 & (0.9) & & (4.2) & & (2.5) & & 
 \end{array}$$

with t-ratios shown in parentheses. Both variables are statistically significant (INC at the 1 percent level and RATIO at the 5 percent level) and with the expected (positive) sign. Also, the fit is quite good, even for this simple model.

Sources are (A) Share of income of the top 5 percent: The basic data source is the Current Population Report series on shares of income held by families that runs from 1947 to 1998. The data are available on internet at the address: <http://www.census.gov/hhes/income/histinc>. The earlier data, from 1922 to 1949, are from Kuznets's (1953) series on the percentage share of total income received by the top percentiles of tax units. This series is benchmarked against the census figure for 1949. (B): Standard and Poor 500 Composite Stock Index: (1) From 1922 to 1969: U.S. Bureau of the Census, *Historical Statistics of the United States, Colonial Times to 1970*, Bicentennial Edition, Part I, Washington, DC, U.S. Government Printing Office, 1975, p. 1004. (2) From 1970 to 1998: U.S. Council of Economic Advisers, *Economic Report of the President, 2000*, Washington, DC, U.S. Government Printing Office, 2001, Table B-93, p. 406. (C) Median House Prices: (1) From 1922 to 1969: U.S. Bureau of the Census, *Historical Statistics of the United States, Colonial Times to 1970*, Bicentennial Edition, Part I, Washington, DC, U.S. Government Printing Office, 1975, Series N 259 & 261, p. 647. (2) From 1970 to 1998: U.S. Bureau of the Census, *Statistical Abstract of the United States, 1999*, 119th edition, Washington, D.C., U.S. Government Printing Office, 1999, Table No. 1203, p. 725.

## 2. LITERATURE REVIEW

Previous work on the overall distributional effects of pension and Social Security wealth has been rather limited. The seminal paper on this subject is Martin Feldstein (1974), who introduced the concept of Social Security wealth and developed its methodology. His main interest was in the aggregate level of Social Security wealth and its effect on aggregate savings and retirement patterns. In a follow-up paper, Feldstein (1976) considered the effects of Social Security wealth on the overall distribution of wealth on the basis of the 1962 Survey of Financial Characteristics of Consumers (SFCC), a survey performed by the Federal Reserve Board of Washington. The paper found that the inclusion of Social Security wealth had a major effect on lowering the inequality of total household wealth (including Social Security wealth). The Gini coefficient for the sum of net worth and Social Security wealth among families in age class 35 to 64 was 0.51, compared to a Gini coefficient of 0.72 for net worth.

Edward Wolff followed up Feldstein (1976) by examining the distributional implications of both Social Security and private pension wealth. Wolff (1987), which used the 1969 Measurement of Economic and Social Performance (MESP) database, was one of the first papers to add estimates of private pension wealth and examine their effects on the overall distribution of wealth. The paper showed that while Social Security wealth has a pronounced equalizing effect on the distribution of augmented wealth (the sum of marketable wealth and retirement wealth), pension wealth has a disequalizing effect on augmented wealth. In particular, the addition of Social Security wealth to net worth reduced the overall Gini coefficient from 0.73 to 0.48 but the addition of pension wealth to the sum of net worth and Social Security wealth raised the Gini coefficient back to 0.66. The sum of Social Security and pension wealth has, on net, an equalizing effect on the distribution of augmented wealth but substantially less than Social Security wealth alone.<sup>2</sup>

Kathleen McGarry and Andrew Davenport (1997) used the 1992 wave of the Health and Retirement Survey (HRS) to estimate private pension wealth. They found that pension wealth is only slightly more equally distributed than net worth and that adding pension wealth to net worth has a modest effect on inequality (with the wealth share of the top decile declining from 52.9 to 44.8 percent with the addition of pension wealth). Arthur Kennickell and Annika Sunden

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<sup>2</sup> Also, see Wolff (1992) for a discussion of some of the methodological issues involved in estimating both Social Security and pension wealth.

(1999) used the 1989 and 1992 Survey of Consumer Finances (SCF) to analyze the effects of Social Security and pension wealth on the overall distribution of wealth. They also found a net equalizing effect from the inclusion of these two forms of retirement wealth. In particular, the inclusion of pension and net Social Security wealth<sup>3</sup> reduced the share of total wealth held by the top one percent with head less than the age of 65 in 1992 from 31.3 percent to 16.2 percent. Estimates are not provided separately for pension wealth and Social Security wealth. As far as we are aware, there are no estimates of Social Security or pension wealth that have been provided thus far from the 1998 or 2001 SCF.

In contrast, there are a host of studies that examine the intra-cohort redistributive effects of Social Security benefits relative to contributions into the Social Security system. They consider which groups are net gainers and which net losers from the Social Security System as a whole. These papers include Wolff (1993), who used the 1962 SFCC and the 1983 SCF; Coronado, Fullerton, and Glass (2000) who derived their estimates from the Panel Study of Income Dynamics (PSID); Smith, Toder, and Iams (2001), who based their work on Survey of Income and Program Participation (SIPP) data matched with Social Security administrative data and the micro-simulation MINT model; Jeffrey Liebman (2002), who matched Social Security Administration earnings and benefit records to the 1990 and 1991 panels of the SIPP; Dean Leimer (2003) and Leimer (2004), who based his analysis on Social Security administrative data.

There are several other topics of relevance to this paper. First, several papers have now used the Health and Retirement Survey (HRS) to examine what proportion of total (augmented) household wealth is comprised of pension and social security wealth. Alan Gustman, Olivia Mitchell, Andrew Samwick, and Thomas Steinmeier (1997) found that in 1992, pensions, Social Security, and health insurance accounted for half of the wealth held by all households aged 51 to 61 in the HRS; for 60% of total wealth of HRS households who are in wealth percentiles 45 to 55; and for 48% of health for those in the 90th to 95th wealth percentiles. In a follow-up study, Gustman and Steinmeier (1998) used data from the HRS to examine the composition and distribution of total wealth for a group of 51 to 61 year olds. They focused on the role of pensions in forming retirement wealth. They found that pension coverage is widespread, covering two thirds of households and accounting for one quarter of accumulated wealth on

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<sup>3</sup> Net social security wealth is defined as the discounted present value of future social security benefits less future taxes paid into the Social Security (OASI) system.

average. Social Security benefits accounted for another quarter of total wealth. They also reported that the ratio of wealth (excluding pensions) to lifetime earnings was the same for those individuals with pensions and for those without pensions. They concluded that pensions cause very limited displacement of other forms of wealth.

Second, several studies have documented changes in pension coverage in the United States, particularly the decline in DB pension coverage among workers over the last two decades. Before this, Laurence Kotlikoff and Daniel Smith (1983) provided one of the most comprehensive treatments of pension coverage and showed that the proportion of U.S. private-wage and salary workers covered by (traditional DB) pensions more than doubled between 1950 and 1979. David Bloom and Richard Freeman (1992), using Current Population Surveys (CPS) for 1979 and 1988, were among the first to call attention to the decline in DB pension coverage. They reported that the percentage of all workers in age group 25-64 covered by these plans fell from 63 to 57% over this period. Among male workers in this age group, the share covered dropped from 70 to 61%, while among females in the same age group, the share remained almost constant, at 53%.

Alan Gustman and Thomas Steinmeier (1992) documented the change-over from DB plans to DC plans between 1977 and 1985 on the basis of IRS 5500 filings. They estimated that only about half of the switch was due to a decline in DB coverage conditional on industry, size, and union status and the other half was due to a shift in employment mix toward firms with industry, size, and union status historically associated with low DB coverage rates. Other studies include William Even and David Macpherson (1994a, 1994b, 1994c, and 1994d). The 1994c study showed a particularly pronounced drop in DB pension coverage among workers with low levels of education; and Even and Macpherson (1994d), showed a convergence in pension coverage rates among female and male workers between 1979 and 1998.

Department of Labor (2000) found that a large proportion of workers, especially low wage, part-time, and minority workers, were not covered by private pensions. The coverage rate of all private sector wage and salary workers was 44 percent in 1997. Coverage of part-time, temporary and low-wage workers was especially low. This appears to be ascribable to the proliferation of 401(k) plans and the frequent requirement of employee contributions to such plans. It also found important racial differences, with 47 percent of white workers participating but only 27 percent of Hispanics. Another important finding is that 70 percent of unionized workers were covered by a pension plan, compared to only 41 percent of non-unionized



workers. Pension participation was found to be highly correlated with wages. While only 6 percent of workers earnings less than \$200 per week had a pension plan, 76 percent of workers earning \$1,000 per week participated.

Third, a related topic of interest is whether DC pension plans have substituted for DB-type plans. Leslie Popke (1999), using employer data (5500 filings) for 1992, found that, indeed, 401(k) and other DC plans have substituted for terminated DB plans and that offering a DC plan raises the chance of a termination in DB coverage.

Fourth, several papers looked at the issue of whether DC plans have substituted for other forms of wealth and whether there was any net savings derived from DC plans. James Poterba, Steven Venti, and David Wise (1992, 1993, 1995), using SIPP data for 1984 and 1991, Poterba, Venti, and Wise (1998), using HRS data for 1993, and Poterba, Venti, and Wise (2001), using both macro national accounting data and micro HRS data, concluded that the growth of IRAs and 401(k) plans did not substitute for other forms of household wealth and, in fact, raised household net worth relative to what it would have been without these plans. They found no substitution of DC wealth for either DB wealth or other components of household wealth.

In contrast, William Gale in a series of papers both by himself and with colleagues, found very little net savings emanating from DC plans. Gale (1995) concluded that when biases in estimation procedures in the previous literature on the subject are corrected, the offset of pension wealth on other forms of wealth can be very high. Using data from the 1984, 1987, and 1991 SIPP, Eric Engen and William Gale (1997) estimated that “at best” only a small proportion of 401(k) contributions represent net increments to household savings. In later work, Engen and Gale (2000) refined their analysis to look at the substitution effect by earnings groups. Using data from the 1987 and 1991 SIPP, they found that 401(k)s held by low earners may more likely represent additions to net worth than 401(k)s held by high earners, who hold the bulk of this asset. Overall, only between 0 and 30 percent of the value of 401(k)s represent net additions to private savings. Kennickell and Sunden (1999) found a significant negative effect of both defined benefit plan coverage and Social Security wealth on non-pension net worth but concluded that the effects of defined contribution plans, such as 401(k) plans, on other forms of wealth are statistically insignificant.

### 3. DATA SOURCES AND ACCOUNTING FRAMEWORK

The principal data sources used for this study are the 1983, 1989, 1998, and 2001 Survey of Consumer Finances (SCF) conducted by the Federal Reserve Board. Each survey consists of a core representative sample combined with a high-income supplement.<sup>4</sup> The SCF provides considerable detail on both pension plans and Social Security contributions. The SCF also gives detailed information on expected pension and Social Security benefits for both husband and wife. For 1983, the Federal Reserve Board has also made its own calculations of the wealth equivalent value of both expected pension benefits and Social Security benefits and made these available in its Public Use sample. However, this has not been done for the other years.

The principal wealth concept used here is marketable wealth (or net worth), which is defined as the current value of all marketable or fungible assets less the current value of debts. Total assets are the sum of: (1) the gross value of owner-occupied housing; (2) other real estate owned by the household; (3) cash and demand deposits; (4) time and savings deposits, certificates of deposit, and money market accounts; (5) government bonds, corporate bonds, foreign bonds, and other financial securities; (6) the cash surrender value of life insurance plans; (7) the current market value of Defined Contribution pension plans, including IRAs, Keogh, and 401(k) plans; (8) corporate stock and mutual funds; (9) net equity in unincorporated businesses; and (10) equity in trust funds. Total liabilities are the sum of: (1) mortgage debt, (2) consumer debt, including auto loans, and (3) other debt. I use the symbol NW to refer to standard net worth. It should be stressed that the standard definition of net worth includes the market value of DC pension plans. (We shall return to this point later on in the paper).<sup>5</sup>

A word should be said on why I use the SCF instead of the newer Health and Retirement Survey (HRS), which has much more complete data on earnings histories and has employer-provided information on individual DB pension plans of each employee covered by these plans. There are three reasons. First, the SCF provides much better data on the assets and liabilities

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<sup>4</sup> See, for example, Kennickell and Woodburn (1992), Kennickell, McManus, and Woodburn (1996), and Kennickell and Woodburn (1999) for details on the construction of the weights used in the SCF files.

<sup>5</sup> Only assets that can be readily converted to cash (that is, “fungible” ones) are included. As a result, consumer durables, such as automobiles, televisions, furniture, household appliances, and the like, are excluded here since these items are not easily marketed or their resale value typically far understates the value of their consumption services to the household.

that constitute marketable net worth. Second, the SCF data date from 1983, whereas the HRS data start in 1992. Since the most important transformation of the pension system occurred beginning in the late 1980s, the SCF data allow us to better track this change over the transition period. Third, the age coverage of the HRS is limited whereas the SCF covers the whole population.

The imputation of both pension and Social Security wealth involves a large number of steps, which is summarized below. Greater details can be found in the Appendix.

### **A. Pension Wealth**

For retirees (r) the procedure is straightforward. Let PB be the pension benefit currently being received by the retiree. The SCF questionnaire indicates how many pension plans each spouse is involved in and what the expected (or current) pension benefit is. The SCF questionnaire also indicates whether the pension benefits remain fixed in nominal terms over time for a particular beneficiary or is indexed for inflation. In the case of the former, the (gross) Defined Benefit pension wealth is given by:

$$(1a) \quad DB_r = \int_0 PB(1 - m_t)e^{-\delta t} dt$$

where  $m_t$  is the mortality rate at time  $t$  conditional on age, gender, and race;  $\delta$  the nominal discount rate, for which the (nominal) 10-year treasury bill rate is used; and the integration runs from the current year to age 109). In the latter case,

$$(1b) \quad DB_r = \int_0 PB(1 - m_t)e^{-\delta^* t} dt$$

and  $\delta^*$  is the real 10-year treasury bill rate, estimated as the current nominal rate less the Social Security Plan II-B assumption of 4.0 percent annual increase of the Consumer Price Index (CPI).

Among current workers (w) the procedure is somewhat more complex. The SCF provides detailed information on pension coverage among current workers, including the type of plan, the expected benefit at retirement or the formula used to determine the benefit amount (for example, a fixed percentage of the average of the last five year's earnings), the expected retirement age when the benefits are effective, the likely retirement age of the worker, and vesting requirements. Information is provided not only for the current job (or jobs) of each spouse but for up to five past jobs as well. On the basis of the information provided in the SCF

and on projected future earnings (see the Appendix for details), future expected pension benefits ( $EPB_w$ ) are then projected to the year of retirement or the first year of eligibility for the pension. Then the present value of pension wealth for current workers ( $w$ ) is given by:

$$(2) \quad DB_w = \int_{LR} EPB(1 - m_t)e^{-\delta t} dt$$

where  $RA$  is the expected age of retirement and  $LR = A - RA$  is the number of years to retirement. As above, and the integration runs from the expected age of retirement to age 109.<sup>6</sup>

It should be noted that the calculations of DB pension wealth for current workers are based on employee response, including his or her stated expected age of retirement (see Appendix D), *not* on employer-provided pension plans. A couple of studies have looked at the reliability of employee-provided estimates of pension wealth by comparing self-reported pension benefits with estimates based on provider data. Using data from the 1992 wave of the HRS, both Gustman and Steinmeier (1999) and Johnson, Sambamoorthi, and Crystal (2000) found that individual reports of pension benefits varied widely from those based on provider information. However, the latter also calculated that the median values of DB plans from the two sources were quite close (about a 6 percent difference).

## B. Social Security Wealth

For current Social Security beneficiaries ( $r$ ), the procedure is again straightforward. Let  $SSB$  be the Social Security benefit currently being received by the retiree. Again, the SCF provides information for both husband and wife. Since Social Security benefits are indexed for inflation, (gross) Social Security wealth is given by:

$$(3) \quad SSW_r = \int_0 SSB(1 - m_t)e^{-\delta^* t} dt$$

where it is assumed that the current social security rules remain in effect indefinitely.<sup>7</sup>

The imputation of Social Security wealth among current workers is based on the worker's projected earnings history estimated by regression equation (see the Appendix for details). The steps are briefly as follows. First, coverage is assigned based on whether the

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<sup>6</sup> Technically speaking, the mortality rate  $m_t$  associated with the year of retirement is the probability of surviving from the current age to the age of retirement.

<sup>7</sup> Separate imputations are performed for husband and wife and an adjustment in the Social Security benefit is made for the surviving spouse. See the Appendix for details.

individual expects to receive Social Security benefits and on whether the individual was salaried or self-employed. Second, on the basis of the person's earnings history, the person's Average Indexed Monthly Earnings (AIME) is computed. Third, on the basis of existing rules, the person's Primary Insurance Amount (PIA) is derived from AIME. Then,

$$(4) \quad SSW_w = \int_{LR} PIA(1 - m_t)e^{-\delta^*t} dt .$$

As with pension wealth, the integration runs from the expected age of retirement to age 109.<sup>8</sup>

Here, too, it should be noted that estimates of Social Security wealth are based on reported earnings at a single point in time. These estimates are likely to be inferior to those based on longitudinal work histories of individual workers (see, for example, Smith, Toder, and Iams, 2001, whose estimates are based on actual Social Security work histories.) In fact, actual work histories do show much more variance in earnings over time than one based on a human capital earnings function projection. Moreover, they also show many periods of work disruption that I cannot adequately capture here. However, I do have some *retrospective* information on work history provided by the respondent (see Appendix D for details). In particular, each individual is asked to provide data on the total number of years worked full-time since age 18, the number of years worked part-time since age 18, and the expected age of retirement (both from full-time and part-time work). On the basis of this information, it is possible to approximate the total number of full-time and part-time years worked over the individual's lifetime and use these figures in the estimate of the individual's AIME.

Nonetheless, since my estimates of SSW assume a continuous work life, I am likely to be overstating the value of SSW for many workers. This is likely to bias upward my estimates of mean and median SSW, as well as a downward bias in the variability of Social Security wealth. It may also lead to an understatement of the correlation between net worth and SSW. For all three reasons, this estimation procedure will likely lead to an *over-estimate* of the degree to which SSW equalizes total wealth inequality.

Estimates will be provided for the following components of household wealth:

$$(5) \quad NW = NWX + DC$$

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<sup>8</sup> As with pension wealth, the mortality rate  $m_t$  associated with the year of retirement is the probability of surviving from the current age to the age of retirement.

where DC is the current market value of Defined Contribution pension plans and HDWX is marketable household wealth excluding DC and NW corresponds to marketable wealth or net worth. Total pension wealth, PW, is given by:

$$(6) \quad PW = DC + DB$$

Private accumulations PA is then defined as the sum of NWX and total pension wealth:

$$(7) \quad PA = NWX + PW$$

The term “private accumulations” is used to distinguish contributions to wealth from private savings and employment contracts with both private and government employers from those of social insurance provided by the state — notably, Social Security (also see Section 8 below for more discussion of this concept).

Retirement wealth RW is then given as the sum of pension and social security wealth:

$$(8) \quad RW = PW + SSW$$

Finally, augmented household wealth, AW, is given by

$$(9) \quad AW = NWX + RW.$$

#### **4. TRENDS IN STANDARD MEASURES OF HOUSEHOLD WEALTH.**

Table 1 documents a robust growth in wealth during the 1990s. Median wealth was 16 percent greater in 2001 than in 1989. After rising by 7 percent between 1983 and 1989, median wealth fell by 17 percent from 1989 to 1995 and then rose by 39 percent from 1995 to 2001. As a result, median wealth grew slightly faster between 1989 and 2001, 1.32 percent per year, than between 1983 and 1989, at 1.13 percent per year.<sup>9</sup>

Mean net worth also showed a sharp increase from 1983 to 1989 followed by a rather precipitous decline from 1989 to 1995 and then, buoyed largely by rising stock prices, another surge in 2001. Overall, it was 65 percent higher in 2001 than in 1983 and 44 percent larger than in 1989. In fact mean wealth grew quite a bit faster between 1989 and 2001, at 3.02 percent per

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<sup>9</sup> The time trend is very similar when the unadjusted asset values are used instead of my adjusted values and when the value of vehicles is included in net worth. Similar results can also be derived from the estimates provided by Kennickell and Woodburn (1999) for 1989 and 1995.

year, than from 1983 to 1989, at 2.27 percent per year. Moreover, mean wealth grew almost three times as fast as the median, suggesting widening inequality of wealth over these years.

Median household income, based on Current Population Survey data, after gaining 11 percent between 1983 and 1989, grew by only 2.3 percent from 1989 to 2001—much slower than median net worth. The net change over the whole period was 14 percent. In contrast, mean income rose by 16 percent from 1983 to 1989 and by another 12 percent from 1989 to 2001, for a total change of 30 percent.

The robust performance of median net worth over the 1990s is particularly surprising in contrast to trends in median income. Indeed, between 1983 and 1998, while median income grew by 13.8 percent, median net worth gained 11.1 percent (and from 1989 to 1998, 2.3 percent and 3.8 percent, respectively). However, between 1998 and 2001, median net worth exploded by 11.5 percent while median income stagnated.

Looking at Panel B of Table 1, we see the important role played by DC pension wealth, which forms part of net worth NW. If DC pension wealth is excluded from net worth, then median wealth NWX actually declined over the 1990s, by 3.9 percent, and mean NWX grew more slowly than standard net worth NW. The rapid accumulation of DC pension wealth thus helped maintain household savings over the 1990s. We shall return to this point later in the paper.

Table 2 shows trends in both wealth and income inequality. It is most useful to begin with the income trends (Panel C). Household income inequality, based on CPS data, increased between 1983 and 1989, with the share of the top five percent rising by 2.5 percentage points, while the share of the next fifteen percent and that of the bottom four quintiles all fell.<sup>10</sup> The Gini coefficient rose from 0.414 to 0.431 over this period. Between 1989 and 2001, the share of the top five percent rose by another 3.5 percentage points while the next fifteen percent and the bottom four quintiles again lost ground, so that the Gini coefficient again increased, from 0.431 to 0.466. All told, according to the CPS figures, there was no abatement in the growth of inequality in the 1989-2001 period compared to 1983-1989.

The trends are different for wealth. As shown in Panel A, wealth inequality, after rising steeply between 1983 and 1989, remained virtually unchanged from 1989 to 2001. The share of wealth held by the top 1 percent rose by 3.6 percentage points from 1983 to 1989 and the Gini

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<sup>10</sup> The CPS tabulations provided by the U.S. Bureau of the Census do not include the income shares of the top percentile.

coefficient increased from 0.80 to 0.83. Between 1989 and 2001, the share of the top percentile actually declined sharply, from 37.4 to 33.4 percent, though this was almost exactly compensated for by an increase in the share of the next four percentiles. As a result, the share of the top five percent actually increased slightly, from 58.9 to 59.2 percent, as did the share of the top quintile, from 83.5 to 84.4 percent. The share of the fourth and middle quintiles also declined slightly, while that of the bottom 40 percent increased somewhat, so that overall, the Gini coefficient fell very slightly, from 0.832 to 0.826.

However, when we now exclude DC pension wealth from net worth, we find that inequality in wealth NWX actually rose over the 1990s (see Panel B). The share of the top one percent gained 1.3 percentage points, that of the top 5 percent grew by 3.7 percentage points, and the share of the top quintile increased by 2.6 percentage points between 1989 and 2001, while the Gini coefficient rose by 0.011 points. Here, too, the accumulation of DC pension wealth helped to moderate wealth inequality over the 1990s.

Table 3 shows changes in the overall portfolio composition of household wealth. In 2001, owner-occupied housing was the most important household asset in the breakdown shown in the table, accounting for 28 percent of total assets. However, net home equity amounted to only 19 percent of total assets. Real estate, other than owner-occupied housing, comprised 10 percent, and business equity another 17 percent. Liquid assets, including bank deposits, money market funds, CDs, and the cash surrender value of life insurance, made up 9 percent and (DC) pension accounts amounted to 12 percent. Financial securities made up 2 percent; corporate stock and mutual funds, 15 percent; and trust equity a little less than 5 percent. Debt as a proportion of gross assets was 13 percent, and the debt-equity ratio (the ratio of total household debt to net worth) was 0.14.

There have been some notable trends in the composition of household wealth over the period between 1983 and 2001. The most important from the standpoint of this paper is that DC pension accounts rose from 1.5 to 12.3 percent of total assets, with almost the entire gain occurring after 1989. This increase almost exactly offset the decline in liquid assets, from 17.4 to 8.8 percent—again, with almost all of the change occurring after 1989. Though there is no direct econometric evidence of substitution here, there is circumstantial evidence the explosion in the use of various types of pension accounts, like IRAs, 401(k) plans, and other thrift plans appears to have allowed households to substitute tax-free pension accounts for taxable savings deposits, rather than increasing overall family savings.



This tabulation provides a picture of the average holdings of all families in the economy, but there are marked class differences in how middle-class families and the rich invest their wealth. As shown in Table 4, the richest one percent of households (as ranked by wealth) invested almost 80 percent of their savings in investment real estate, businesses, corporate stock, and financial securities in 2001. Housing accounted for only 8 percent of their wealth, liquid assets 6 percent, and pension accounts 6 percent. Their debt-equity ratio was 2 percent and their ratio of debt to income was 34 percent. Among the next richest 19 percent, housing comprised 27 percent of their total assets, liquid assets another 9 percent, and pension assets 16 percent. Forty-six percent of their assets took the form of investment assets— real estate, business equity, stocks, and bonds. Debt amounted to 9 percent of their net worth and 77 percent of their income.

In contrast, almost 60 percent of the wealth of the middle three quintiles (60 percent) of households was invested in their own home in 2001. Another 25 percent went into monetary savings of one form or another and pension accounts. Together housing, liquid assets, and pension assets accounted for 84 percent of the total assets of the middle class. The remainder was about evenly split among non-home real estate, business equity, and various financial securities and corporate stock. The ratio of debt to net worth was 32 percent, much higher than for the richest 20 percent, and their ratio of debt to income was 100 percent, also higher than the top quintile.

There is remarkable stability in the composition of wealth by wealth class between 1983 and 2001. The most notable exception is a substitution of pension assets for liquid assets—a transition that occurred for all three wealth classes but that was particularly marked for percentiles 80-99 and for the middle three quintiles.

## **5. PENSION WEALTH**

Table 5 highlights trends in pension holdings over the 1983-2001 period. In this and the subsequent tables, it should be noted that the unit of observation is the household, not the individual worker. Moreover, I have divided households into three age groups: under 47, 47-64, and 65 and older. Data for the youngest group are the most problematic, since estimates of both DB pension wealth and social security wealth depend on projecting future work life and, in the case of the former, future job tenure with the same employer. Data for retirees are the most

secure since both pension and social security benefit levels are already determined. Estimates of both DB and social security wealth for the middle-aged group lie in between in terms of reliability. Individuals close to retirement have a fairly good idea of their expected age of retirement and have a high likelihood of remaining with their current employer (see Farber, 2001, for some evidence).

The share of all households with DC pension accounts skyrocketed over the 1983-2001 period, from 11 to 52 percent, or by 41 percentage points. The story is very similar for the three different age groups shown in Table 5—even among the elderly. The proportion holding pension accounts advanced by 40 percentage points in age group 46 and under, by 50 percentage points among households in age group 47-64, and by 33 percentage points among elderly households. In 2001, about 62 percent of households in the age range of 47 to 64 held some form of DC account, compared to 35 percent of elderly households and 54 percent of younger households. Most of the gains occurred after 1989.

Opposite trends are apparent for Defined Benefit (DB) pension wealth. The share of all households with DB pension wealth fell by 18 percentage points between 1983 and 2001, from 53 to 34 percent. Among households in age group 47-64, the decline was even more precipitous — by 24 percentage points, from 69 to 45 percent — while among elderly households the proportion fell by 20 percentage points and among young households, by 10 percentage points. In 2001, while 47 percent of elderly households held some form of DB pension wealth, only 45 percent of households in age group 47-64 and only 23 percent among young households recorded DB entitlements. Most of the loss in coverage again occurred after 1989.

The percentage of all households covered by either a DC or a DB plan increased from 54 to 66 percent between 1983 and 2001. Among the 47-64 age group, the proportion rose by 5.6 percentage points, to 76 percent in 2001, while among the elderly, the share fell by 4.3 percentage points, down to 63 percent in 2001. The biggest rise occurred among younger households, whose proportion surged from 36 to 61 percent. The share of households covered by pensions in 2001 was 76 percent among the middle-aged, compared to 63 percent among the elderly and 61 percent among the youngest age group.

As shown in Table 6, there were huge increases in the average holdings of DC pension accounts.<sup>11</sup> Among all households, the average value of these accounts increased ten-fold

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<sup>11</sup> Due to space constraints, I show results for only 1983 and 2001. Results for some other key variables are also available for 1989 and 1998.

between 1983 and 2001, from \$3,900 (in 2001 dollars) to \$53,700. Among age group 46 and under the increase was by a factor of 10.1 and among age group 47-64 the gain was by a factor of 11.1, while among elderly households, the rise was by a factor of 28.4. In 2001, mean DC pension wealth was greatest among age group 47-64, at \$96,600, second highest among elderly households, at \$53,600 and lowest among the youngest age group, at \$28,200.

Opposite trends are again evident for the average value of DB pension wealth. Among all households, the mean value fell by 26 percent between 1983 and 2001, from \$55,400 (in 2001 dollars) to \$41,200. Losses were also marked for age group 47-64, who saw their mean DB pension wealth decline by 21 percent between 1983 and 2001, and for younger households, whose average DB wealth fell by 27 percent. However, the valuation of pension rights among younger workers has to be interpreted cautiously, since these are based on projected benefits in 20 to 40 years hence.<sup>12</sup> However, the average value of DB plans among the elderly fell by somewhat less over this period, 14 percent. In contrast to DC pensions, the average value of DB pension wealth was highest among elderly households—\$105,600 in 2001—compared to \$74,300 among middle-aged households and only \$16,900 among young households.

We can now consider one of the issues raised in the Introduction: Has the spread of DC type pension plans adequately compensated for the decline in traditional DB pension coverage? The results indicate that the answer is both “yes” and “no.” Average pension wealth PW (the sum of DC and DB pensions) increased for all age groups between 1983 and 2001. Among all households, the mean value of total pension wealth PW climbed by 60 percent. Among those in age groups 46 and under and 47 to 64, the mean value increased by 76 and 66 percent, respectively, while among elderly households the mean value jumped by 70 percent.

The story is not quite as positive when we look at trends in median pension wealth. Among all households, median pension wealth PW more than doubled between 1983 and 2001, from \$4,900 to \$10,900 (result not shown). However, much of this gain is a result of increased pension coverage. When the sample is limited to those households with positive pension wealth, the median actually fell by 17 percent, from \$64,300 to \$53,700, over the period. Among young households who have pension wealth, median pension wealth increased by 14 percent, among middle-aged pension holders by 9 percent, and among elderly pension holders by 16 percent.

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<sup>12</sup> Moreover, there is a large number of missing values for this age group in the 1983 SCF data.

The bottom line is that while mean pension wealth grew robustly over this period, median pension wealth shows much more modest gains.

In Table 7, I show Gini coefficients for pension wealth. I again focus on dispersion among pension holders only within the age group. Among all households, the inequality of DC pension wealth fell over the period from 1983 to 2001. This was also true among young and middle-aged households but not among elderly DC pension owners.

However, the level of inequality in DC pension wealth was still very high in 2001, even among DC account holders alone. The Gini coefficient among all DC pension account holders was 0.741 in 2001. This compares to a Gini coefficient for net worth of 0.826. Inequality among DC account holders within age group was almost as great as among all DC account holders.<sup>13</sup>

In contrast, the dispersion in traditional DB pension wealth trended upward over the 1983-2001 period. Among all households who had plans, the Gini coefficient climbed from 0.555 to 0.638. Among middle-aged pension holders, DB pension wealth inequality also climbed sharply (by 0.088), as it did among elderly pension recipients (0.202). In contrast, DB wealth inequality declined by 0.078 points among young households with DB plans.

In terms of inequality levels, DB pension wealth inequality was considerably lower among DB pension wealth holders than among DC pension account holders. Not surprisingly, the switchover from DB pension plans to DC pension plans has resulted in an upsurge in pension wealth inequality.

This trend is also documented in Table 7. Among all households, the Gini coefficient for total pension wealth PW rose by 0.051 from 0.749 to 0.800, while among pension wealth holders, the increase was even more striking, by 0.157, from 0.537 to 0.695. The pension wealth inequality patterns are very similar among both middle-aged and elderly households. However, among the younger households, pension wealth inequality increased only slightly among all households in this group and declined sharply among young households with pension plans—a reflection of the large increase in the share of young households with such plans.

Figure 1 provides a look at the cumulative distribution of pension wealth among all households in 1983 and 2001. Because the percentage of households with positive pension wealth expanded over the period from 54.4 to 65.6, the 1983 cumulative distribution initially lies above the 1998 cumulative distribution. However, what is striking is that though the 1983

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<sup>13</sup> This result accords with media accounts of a large divide in the value of 401(k) plans between executives and staff workers in large corporations (see, for example, Leonhardt, 2002).

cumulative distribution lies above the 2001 cumulative distribution throughout, the two are very close from a pension wealth level of \$10,000 (in 2001 dollars) through \$65,000. In other words, despite the gain in the share of households with some pension wealth, between 1983 and 2001, a lot of them had very small amounts of pension wealth in 2001 (perhaps, an IRA of a few thousand dollars).

Figures 2a and 2b provide further details on the change in the distribution of pension wealth among all households over the 1983-2001 period. At lower percentile levels up to and including the median, there were large gains in pension wealth over the period, reflecting the increase in the share of households with a pension plan. Between percentiles 60 and 75, the percentage increase in pension wealth ran from 14. to 21.6 percent. However, the percentage increase in pension wealth was higher at the upper percentiles, ranging from 29.5 percent at the 80<sup>th</sup> percentile to a spectacular 82.4 percent at the 99<sup>th</sup> percentile. These results illustrate that the largest growth of pension wealth occurred at the top of the pension wealth distribution.

Results are even more dramatic for middle-aged households (Figures 2c and 2d) and elderly households (Figures 2e and 2f). Among the former, the share with positive pension wealth grew from 70.2 to 75.9 percent over the period, so that percentage gains in pension wealth were very high at lower percentiles. However, pension wealth at the 35<sup>th</sup> percentile fell by a substantial 14 percent and at the 40<sup>th</sup> percentile by 7 percent. Pension wealth at the 45<sup>th</sup> to 60<sup>th</sup> percentile gained less than 20 percent. In contrast, percentage gains in pension wealth increased from 53 percent at the 80<sup>th</sup> percentile to a remarkable 123 percent at the 99<sup>th</sup> percentile. The pattern is quite similar among elderly households. Pension wealth actually fell between the 45<sup>th</sup> and 65<sup>th</sup> percentile, from 100 to 9 percent. Percentage gains in pension wealth then rose monotonically by percentile, from 9 percent at the 70<sup>th</sup> percentile to a huge 146 percent at the 99<sup>th</sup> percentile. The results rather convincingly display that the size distribution of pension wealth, particularly among middle-aged and elderly households, substantially widened over the years from 1983 to 2001.

A decomposition of the inequality of pension wealth PW in 1983 and 2001 reveals the changing contribution of the two forms of pensions to overall pension wealth inequality. As derived in Wolff (1987), for any variable  $X = X_1 + X_2$ ,

$$CV^2(X) = p_1^2 CV^2(X_1) + p_2^2 CV^2(X_2) + 2CC(X_1, X_2)$$

where CV is the coefficient of variation (the ratio of the standard deviation to the mean), CC is the coefficient of covariation, defined as the ratio of the covariance to  $X^2$ ,  $p_1 = X_1/X$ , and  $p_2 = X_2/X$ .<sup>14</sup>

Results are shown in Table 8. As noted above, pension wealth inequality increased over the 1983-2001 period. In this case the coefficient of variation of PW among all households rose by 0.60 points, from 2.07 to 2.66. The coefficient of variation of DC wealth fell over the period by over half, from 8.5 to 3.7, while the coefficient of variation of DB pension wealth grew by a little over a half, from 2.0 to 3.2. Much of the change reflected the increasing share of households with DC pension wealth and the corresponding decline in the share of households with DB plans. Still, even in 2001, the inequality of DC pension wealth was greater than that of DB pension wealth.

Despite the decline in DC inequality, the largest contribution to the growth of PW inequality was the increasing share of DC pension wealth in total pension wealth (lines C1 and D1). Indeed, it more than accounted for the total increase in PW inequality over the 1983-2001 period. Moreover, despite the rise in DB inequality over the period, its declining share of total pension wealth helped to reduce overall pension wealth inequality (lines C2 and D2). The correlation between DC and DB pension wealth was quite small in both 1983 and 2001—0.17 and 0.14, respectively—and showed a decline over the period. The increase in the coefficient of covariation between the two components also contributed somewhat to the overall rise in PW inequality over the period. Results for age group 47 to 64 are quite similar.

I next look at trends in both mean and median private accumulations (PA) over the 1983 to 2001 period (see Table 9). We begin by recapitulating trends in net worth over the period. Among all households, mean NWX (net worth excluding DC) rose by 32 percent while the median barely changed. Mean net worth rose by 52 percent, while its median increased by only 14 percent. I next add pension wealth to NWX to obtain private accumulations. Its mean value was up by 38 percent, compared to 52 percent for net worth, while its median value was down by a very sizeable 14 percent, compared to a 14 percent increase in median net worth.

Among middle-aged households, the mean value of PA rose by 44 percent, compared to a 60 percent increase in net worth, whereas the median value of PA was down by 2 percent, compared to a 27 percent gain in median net worth. Among elderly households, the mean value of PA rose by 41 percent, compared to a 50 percent rise in net worth, while the median grew by

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<sup>14</sup> The Theil coefficient cannot be used here because of the large number of zero values.

30 percent, as opposed to a 48 percent increase in net worth. Among younger households, mean PA was up by 45 percent, compared to a 60 percent growth in net worth, while median PA remained completely unchanged, in contrast to a 21 percent decline in median net worth.<sup>15</sup>

Generally speaking, private accumulations fared worse than conventional net worth. Mean PA rose less than mean net worth, while median PA increased far less than median net worth (except for young households). A comparison of trends in PA with those in NWX suggests that households dipped into their private savings to finance their 401(k) and other DC plans.

In Table 9, I also consider the effects of pension wealth on overall wealth inequality. The Gini coefficient for net worth among all households is 0.826 in 2001. If we subtract DC pensions, then the Gini coefficient rises by 0.019 points to 0.845. Adding DB pension wealth to net worth results in a fall of the Gini coefficient of 0.026 points, to 0.800. This rather modest decline in the Gini coefficient is due to both the high level of pension wealth inequality in the population and the high correlation of pension wealth with marketable net worth (see below).

Looking over time, we also find that the equalizing effect of DB pension wealth has mitigated over the period from 1983 to 2001. Whereas the Gini coefficient for net worth among all households advanced by 0.028 points over the 1983-2001 period, the Gini coefficient for PA gained 0.054 points. Likewise, adding DB pension wealth to net worth results in a 0.053 decline in the Gini coefficient in 1983 but only a 0.026 decrease in 2001. However, the joint effect of adding total pension wealth to NWX is more similar in the two years: a 0.056 reduction in the Gini coefficient in 1983 and a 0.045 reduction in 2001.

The results are sharpest for middle-aged households. For this group, the Gini coefficient for net worth increased by 0.037, while that for PA ballooned by 0.081. Likewise, among the elderly, PA inequality increased by 0.020 points, whereas net worth inequality fell by 0.016 points. The exception to this pattern is young households, for whom PA inequality increased less than net worth inequality.

Table 10 provides a decomposition to provide an alternative assessment of the effects of adding private pension wealth to NWX (net worth less DC pension accounts) on the inequality of PA among all households and those in age group 47 to 64. As noted earlier, pension wealth inequality among all households as measured by the coefficient of variation (CV) increased by

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<sup>15</sup> Here, again, some caution should be exercised with regard to pension wealth figures for young households, particularly because of a large number of missing values in the 1983 data.

0.60 points over the 1983-2001 period. Inequality in NWX increased less over the period — by 0.25 points — and the coefficient of variation of PA grew by only 0.06 (Panel A). It is of note that the increase in the inequality of PA based on CV is considerably smaller than that based on the Gini coefficient. This reflects the fact the CV is much more sensitive to changes in the upper tail of the distribution than is the Gini coefficient, and, as we shall see below, a large portion of the change in the distribution of PA occurs in the middle deciles.

The correlation between NWX and PW increased over the period, from 0.21 to 0.26, and in this decomposition the rising correlation accounted for 110 percent of the rise in the coefficient of variation of PA. The growth in pension wealth inequality together with its rising share in PA also made a positive contribution to the rise in the CV of PA (33 percent), while the declining share of NWX in PA made a negative contribution.

Results are rather different for the 47 to 64 age group. Here the largest contribution is made by the rising inequality in NWX (62 percent), followed by the rising correlation between NWX and PW (29 percent) and then by the rise in PW inequality (9 percent). The results suggest that for this age group at least lower wealth households substituted DC pension wealth for other components of net worth, accounting for the sharp jump in the inequality of NWX.<sup>16</sup>

Figures 3a and 3b provide a closer look at the size distribution of PA among all households in 1983 and 2001. Here it becomes quite clear that the major gains over the 1983-2001 period were made by households at the high end of the wealth distribution. Indeed, comparing the size distributions among all households in the two years at different percentile levels, we find an almost monotonic relation between percentile level and percentage change in PA over the period. The percentage growth in PA surges from -73 percent at the tenth percentile to 79 percent at the ninety-ninth percentile. The crossover point in the two distributions occurs just about at the sixtieth percentile. Results are quite similar for households in age group 47 to 64, with the percentage growth in PA ballooning from -98 percent at the tenth percentile to 94 percent at the ninety-ninth percentile and the two distributions crossing slightly below the median (see Figures 3c and 3d). In contrast, among elderly households, PA shows positive growth at and above the 15<sup>th</sup> percentile, though here again the percentage increase in PA tends to rise with the percentile level.

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<sup>16</sup> Here it should be stressed that a decomposition is not a behavioral model—that is, we do not take into account how other wealth components of NWX change in response to changes in PW.



## 6. SOCIAL SECURITY AND TOTAL RETIREMENT WEALTH

In contrast to the strong growth in pension wealth, mean social security wealth among all households rose by only 10 percent between 1983 and 2001 (see Table 11). The average value of social security wealth gained 21 percent among middle-aged households between 1983 and 2001 and 65 percent among young households, though it did remain largely unchanged among elderly households. Likewise, median social security wealth increased by a very modest 7 percent among all households. Households in the 47-64 age bracket saw their median social security wealth grow by 22 percent, while elderly households experienced a 4 percent decline. In contrast, young households saw their median social security wealth climb by 68 percent.

Social security wealth averaged \$139,500 in 2001. This compares to a mean net worth (NW) figure of \$380,300 and mean pension wealth of \$94,800. Median social security wealth in 2001 was \$120,700 — close to that of mean social security wealth. This suggests a normal or close to normal distribution of social security wealth. Moreover, median social security wealth was almost double median net worth (\$73,500) and more than ten times higher than median pension wealth (\$10,900). Mean and median social security wealth among middle-aged households was about 40 percent greater than that among elderly households and about 70 percent greater than that among young households. The smaller social security wealth among the elderly relative to the middle-aged largely reflects the higher lifetime earnings of the latter, whereas the smaller social security wealth among the young relative to the middle-aged largely reflects the greater discount factor applied to future social security benefits (from the larger number of years left before retirement).

The inequality of social security wealth is much lower than that of net worth or pensions (see Table 11). In 2001, the Gini coefficient for social security wealth among all households was 0.34, compared to 0.83 for net worth and 0.80 for pension wealth. In this case, inequality in social security wealth plummeted over the 1983-2001 period, with the Gini coefficient falling by 0.11 points. The Gini coefficient for social security wealth also fell considerably among young, middle-aged, and elderly households over the period.

More details on the change in the distribution of social security wealth between 1983 and 2001 are provided in Figures 4a-4f. Among all households, there was strong growth in social security wealth over the period, reflecting the increase in social security coverage from 82

to 98 percent. Moreover, in contrast to pension wealth, the largest gains were made by the lowest percentiles. Indeed, the percentage increase in SSW was almost inversely related to the percentile level, ranging from 160 percent at the 20<sup>th</sup> percentile to -6 percent at the 99<sup>th</sup> percentile. The pattern is similar for middle-aged households, though here all the gains were positive, ranging from 163 percent at the 10<sup>th</sup> percentile to 6 percent at the 99<sup>th</sup> percentile. Among elderly households, the same inverse relation can be observed, though in this case the gains are positive up to the 45<sup>th</sup> percentile and negative over the remainder of the distribution.

Total retirement wealth is the sum of DC pensions, DB pension wealth, and social security wealth. Among all households, mean retirement wealth grew by 17 percent from 1983 to 2001 (see Table 11). Moreover, mean retirement wealth was up among each of the three age groups over the 1983-2001 period: by 41 percent among young households, by 40 percent among middle-aged ones, and by 20 percent among the elderly. In contrast, median retirement wealth among all households *fell* by 6 percent between 1983 and 2001. It declined by 3 percent among elderly households, though among young households it rose by 29 percent over the 1983-1998 period and among middle-aged ones it was up by 15 percent.

The inequality of retirement wealth lies between that of social security wealth on the one hand and both pension wealth and net worth on the other hand (Table 11). In 2001, the Gini coefficient for retirement wealth among all households was 0.49, compared to 0.34 for social security wealth, 0.80 for pension wealth, and 0.83 for net worth. The main news here is a sharp *rise* in the inequality of retirement wealth over the 1983 to 2001 period among all households, among whom the Gini coefficient increased by 0.061 points, among middle-aged households by 0.076 points, and among the elderly by an even more sizeable 0.111 points. However, among young households, the Gini coefficient declined by 0.077 points, reflecting the increase in private pension coverage reported earlier.<sup>17</sup>

Figures 5a-5f portray changes in the distribution of retirement wealth RW between 1983 and 2001. Among all households, percentage gains in RW form a U-shaped pattern, falling from 52 percent at the 10<sup>th</sup> percentile to -9 percent at the 45<sup>th</sup> percentile and then climbing to 55 percent at the 99<sup>th</sup> percentile. The strong growth in RW among the lowest percentiles reflects the large increases in SSW at the bottom, while the sharp gains among the top percentiles is due to the huge gains of pension wealth at the top. The decline in the middle is a reflection of the

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<sup>17</sup> Here, again, some caution should be exercised with regard to retirement wealth figures for young households, particularly because of a large number of missing values in the 1983 data.

losses in PW in this part of the distribution. The pattern of percentage gains is also U-shaped among middle-aged households, though in this case all the gains are positive. In contrast, among the elderly, the percentage increases rise almost monotonically with the percentile level, from -18 percent at the 10<sup>th</sup> percentile to 105 percent at the 99<sup>th</sup> percentile. This pattern reflects the slight decline in SSW at the middle and the large decline of PW at the bottom, as well as the substantial increase of PW at the top.

Table 12 provides a decomposition to analyze the sources of the large gains in the inequality of RW over the 1983 to 2001 period. Here, it is clear that the main contributor to this increase is the substantial growth in pension wealth inequality. The latter accounted for 85 percent of the total growth in RW inequality among all households and for 89 percent of the growth among households in age group 47 to 64. The correlation between PW and SSW also intensified over the period, from 0.11 to 0.39, and in this decomposition the rising correlation accounted for 29 percent of the rise in the coefficient of variation of RW among all households and 22 percent among middle-aged households. Inequality in SSW lessened over the period and offset somewhat the rise in PW inequality.

## **7. AUGMENTED WEALTH**

Table 13 provides figures on the trend in both mean and median augmented wealth over the 1983 to 2001 period. We begin by recapitulating trends in net worth over the period. Among all households, mean net worth rose by 52 percent, while median net worth increased by 14 percent. If we next add pension wealth to NW, then the mean value of PA was up by 38 percent, compared to 52 percent for net worth, and the median value was down sharply, by 14 percent. The mean value of augmented wealth AW, the sum of net worth, PW, and SSW, rose even less, by only 25 percent, over the 1983-2001 period, and median AW plummeted by 21 percent.

The patterns are similar by age group. Among young households, mean AW increased by 39 percent, compared to 60 percent for NW, and median AW plunged by 18 percent, compared to a 21 percent drop in NW. Among middle-aged households, mean AW grew by 38 percent, compared to 60 percent for NW, and median AW increased by a tiny 3 percent, compared to 27 percent for NW. The elderly experienced a 30 percent gain in mean AW,

compared to 50 percent in mean NW, and median AW advanced by 13 percent, much less than the 50 percent gain in median NW.

Table 14 considers the effects of adding retirement wealth on overall wealth inequality. In 2001 the Gini coefficient for net worth among all households was 0.826. Adding pension wealth to net worth (PA) results in a modest decline of the Gini coefficient of only 0.026 points to 0.800. This is due to both the high level of pension wealth inequality in the population and the high correlation of pension wealth with marketable net worth. However, adding social security wealth to net worth results in a sizeable reduction in the Gini coefficient of 0.161 points, from 0.826 to 0.665 (result not shown). This reflects both the much lower level of inequality in SSW than in marketable wealth, as well as its relatively low (though positive) correlation with net worth. Finally, adding both pension wealth and social security wealth to net worth results in an only very modest further reduction of the Gini coefficient to 0.663. The main equalizing effect of retirement wealth comes from social security, not private pensions. Results are very similar for the three age groups.

Looking over time, we find as our main result that the equalizing effect of retirement wealth has mitigated over the 1983-2001 period. Whereas the Gini coefficient for net worth among all households advanced by 0.028 points over the 1983-2001 period, the Gini coefficient for PA gained 0.054 points and that for the sum of NWX and SSW grew by 0.034 points, while the Gini coefficient for augmented wealth ballooned by 0.072. In other words, while the addition of retirement wealth to net worth reduced the Gini coefficient by 0.208 points in 1983, the difference was only 0.164 in 2001. Thus, the inequality reducing effects of adding retirement wealth to net worth has fallen over the years from 1983 to 2001.

The results are very similar for middle-aged and elderly households. Among the former, the Gini coefficient for net worth increased by 0.037 while that for augmented wealth climbed by 0.082. Among the latter, the Gini coefficient for net worth actually fell by 0.016 while that for AW grew by 0.047. The exception is young households, for whom the net worth Gini coefficient advanced by 0.062 while that for AW remained unchanged.

Figures 6a-6f give a graphical picture of changes in the distribution of AW between 1983 and 2001. Among all households, percentage gains in AW, like those of RW, form a U-shaped pattern, falling from -10 percent at the 10<sup>th</sup> percentile to a low of -28 percent at the 40<sup>th</sup> percentile and then advancing to 48 percent at the 99<sup>th</sup> percentile. The pattern seems to mirror rather closely the pattern of percentage gains in RW. It is of note that for AW the crossover

point occurs at the 80<sup>th</sup> percentile, with households at or above this point enjoying positive gains and those below suffering losses.

The pattern of percentage gains is quite similar among middle-aged households, with percentage gains falling from -1 percent at the 10<sup>th</sup> percentile to a low of -7 percent at the 25<sup>th</sup> percentile and then climbing to 90 percent at the 99<sup>th</sup> percentile. In this case, the crossover point occurs at the 35<sup>th</sup> percentile. In contrast, among the elderly, the percentage increases remain relatively flat between the 10<sup>th</sup> and 60<sup>th</sup> percentile, at around 10 percent, and then rise to 58 percent at the 99<sup>th</sup> percentile.

Figures 7a and 7b show retirement wealth by net worth percentile among all households. It is apparent that there is a strong correlation between net worth and retirement wealth. In 2001, RW ranged from a low of \$116 thousand for the lowest wealth class to \$739 thousand for the top wealth class in 2001 — more than a sixfold difference. These results also indicate a relatively strong positive correlation between NW and RW. It also appears that the correlation became stronger between 1983 and 2001 (as we shall see in Table 15). It is also evident from Figure 7b that between 1983 and 2001 RW grew generally faster at higher wealth percentiles, particularly at the 80<sup>th</sup> percentile and above.

A decomposition, shown in Table 15, provides more details on the effects of adding retirement wealth RW to NWX (net worth less DC pension accounts) on the change in wealth inequality among all households and among middle-aged ones alone over the 1983-2001 period. The results indicate that the main source — about three quarters — of the inequality gain is due to rising inequality in NWX. As noted above, the correlation between NWX and RW is positive and also increased over the period, from 0.22 to 0.27. The rising covariance of RW and NWX accounted for another fifth or so of the increase in AW inequality. Finally, the rise in the inequality of RW itself explained the other 6 percent.<sup>18</sup>

## **8. ALTERNATIVE PENSION WEALTH CALCULATIONS**

So far I have treated DB and DC pension wealth as well as SSW as comparable concepts. However, there are important differences in their estimation. Most notably, the calculation of DB wealth is estimated on the basis of the future stream of pension benefits on the assumption

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<sup>18</sup> It should be stressed once again that a decomposition is not a behavioral model. In particular, in this case, we do not take into account how other wealth components like NWX change in response to changes in PW or in SSW.

that the employee remains at his or her firm of employment until the person's expected retirement date. The computation of SSW is also based on the assumption that the worker remains at work until the person's expected retirement date. On the other hand, the DC valuation is based on the current market value of DC plans.

There are two ways of putting DB (and SSW) and DC on an "equal footing." First, one can compute DB and SSW on a "year-to-date" (YTD) basis, assuming that the worker stopped working as of the year of the survey, say 2001. It is possible to do this for SSW by computing the SS benefits that would be received if the worker stopped working in 2001. For DB, we can do this for responses that indicate DB benefits as a percentage of last year's salary by using current earnings (as opposed to projected earnings) to calculate the expected DB benefit. However, this method is a bit problematic since we do not know how the worker's DB benefit depends on his or her years of service in the company or when the worker is vested. Moreover, most respondents simply indicate what they expect their DB benefit to be at year of retirement. Since we do not know the actual formula used, it is impossible to compute the expected benefit on the basis of current earnings. Moreover, in the case of the 1983 SCF, we do not have the underlying data to perform the necessary calculations for YTD SSW or DB wealth.<sup>19</sup>

Second, one can project forward DC wealth to year of retirement (YOR). The 2001 SCF does provide information on the employer contribution to DC plans, like 401(k)s (see Part E of the Appendix). If we assume, as in the case of DB pensions, that workers remain at their company until retirement and that the terms of their DC contract with their employer stays the same, then it is possible to do this. In most cases, the employer contribution is a fixed percentage of the employee's salary. On the basis of the estimated human capital earnings functions for each worker, I can calculate the annual stream of future employer contributions to the DC plan until retirement (which I call DCEMP).<sup>20</sup>

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<sup>19</sup> This comparison might actually overstate the case. Some future DB entitlements are based on past jobs (as opposed to current jobs), which, of course, do not require any additional work time with the company to secure future DB benefits. However, the portion of future DB benefits that accrue from past jobs is relatively small—I estimate only 9.65 percent of total DB benefits in 2001.

<sup>20</sup> Moreover, the 1983 data do not present a problem, since DC wealth was a trivial amount, so that we can again safely ignore this in the wealth comparison between 1983 and 2001.

The issue remains how to treat future employee contribution to DC plan (as well as future wealth accumulation).<sup>21</sup> Insofar as DC contributions add to net savings over time, it would be necessary to add in this net amount of the employee portion to the employer contribution in future projections. However, as discussed in Section 2 above, the evidence is rather mixed, with Poterba, Venti, and Wise (2001) concluding that the growth of IRAs and 401(k) plans did not substitute for other forms of household wealth while the work of Engen and Gale (2000) finding that DC plans do *not* add to net savings but rather substitute for other forms of savings. It seems that including only future employer contributions (DCEMP) might be a prudent compromise. Moreover, this would likely put the treatment of DC pension wealth on a par with that of DB pension wealth.<sup>22</sup>

Nonetheless, even this “compromise” solution might overstate the degree of comparability between DB and the sum of DC and DCEMP. The reason is that in the case of current jobs with DB benefits, it is not possible to tell how much of the value of DB wealth is already “sewn up” as of the time of the survey for the reasons stated above. Among older workers, in their fifties and sixties, the incremental gain in DB from continuing to work until expected retirement age is probably relatively small, since much of DB wealth from the current job is already guaranteed. A similar argument applies to SSW, since only the highest 35 years of earnings are used to calculate the worker’s AIME. It is quite likely in the case of DB wealth (and SSW), the incremental gain from continuing to work is substantially less in both absolute and relative terms than the future contributions to DC pension plans.

As a result, adding the total discounted value of future employer contributions to DC plans to DC wealth might overstate the net addition to the household’s wealth relative to the net gains in DB wealth from continuing to work. It is quite likely that the true value of private accumulations (and augmented wealth) lies between PA and the sum of PA and DCEMP (likewise, between AW and the sum of AW and DCEMP). PA and the sum PA+DCEMP (as well as AW and the sum AW+DCEMP) might be thought of as lower and upper bounds on the true value of private accumulations (and augmented wealth).<sup>23</sup>

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<sup>21</sup> It is beyond the scope of the paper to project forward total wealth accumulation over time. This process would require a household microsimulation model such as the MINT model that the Urban Institute and the Social Security Administration use (see, for example, Smith, Toder, and Iams, 2001).

<sup>22</sup> Moreover, it is also possible that participation in a DB plan might augment future savings though here, too, the evidence is generally negative (see, for example, Munnell, 1976, and Kennickell and Annika Sunden, 1999).

<sup>23</sup> Another difference between DB wealth and DC wealth is the greater risk associated with the latter. The benefit

### A. DC Employer Contribution Pension Wealth.

The SCF questionnaire indicates how many DC pension plans each spouse has (up to three per spouse).<sup>24</sup> Information on the employer contribution to DC pensions plans is recorded in two ways. First, in some cases, the contribution is given as a flat dollar amount. Though it is not indicated in the survey data whether the dollar contribution is indexed to inflation over time, I assume that it is indexed to the CPI, which seems the more likely arrangement.<sup>25</sup> Let EMPAMT be the dollar amount of the employer contribution to the DC plan. Then, the present value of the stream of future employer contributions, DCEMP<sub>a</sub>, is given by:

$$(10) \quad \text{DCEMP}_a = \int_0^{\text{LR}} \text{EMPAMT} (1 - m_t) e^{-\delta^* t} dt$$

where  $m_t$  is the mortality rate at time  $t$  conditional on age, gender, and race;  $\delta^*$  is the real 10-year treasury bill rate, estimated as the current nominal rate less the Social Security Plan II-B assumption of 4.0 percent annual increase of the CPI.<sup>26</sup> The integration runs from the current year to LR, where RA is the expected age of retirement and  $\text{LR} = A - \text{RA}$  is the number of years to retirement.

Second, in most cases, the employer contribution is given as a percent of earnings. If we assume that the proportion, EMPPER, is fixed over time, then DCEMP<sub>b</sub>, is given by:

$$(11) \quad \text{DCEMP}_b = \int_0^{\text{LR}} \text{EMPPER} \cdot E^*_t (1 - m_t) e^{-\delta^* t} dt$$

where  $E^*_t$  is the predicted earnings of the worker at time  $t$  in constant dollars (see Appendix C for details).

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levels in DB plans are already set by the terms of the plans – that is why these are called *defined benefits*. DB wealth depends only on future labor force participation in the company and future earnings. The establishment of the Pension Benefit Guarantee Corporation in 1974 does, at least, insure the pension benefits (up to a fixed amount) in the event of the bankruptcy of a company. In comparison, DC wealth depends not only on future labor force participation and future earnings but also on future employee contributions, future employer contributions, and future rates of return. Indeed, the stock market experience of the 2000 to 2003 period shows how difficult it would have been to project the future value of DC wealth even over this short three-year period. DB benefits are more certain than DC benefits. Indeed, the shifting of the risk from employer to employee is one of the reasons behind the rise of DC plans (see Wolff, forthcoming, for a discussion of this issue).

<sup>24</sup> The SCF records DC plans only for the main job of each respondent. No information on DC plans is provided for secondary employment. This does not appear to be a significant problem because in 2001, 99.4 percent of the total labor earnings of the head and 98.8 percent of that of the spouse came from the person's primary job.

<sup>25</sup> This will, if anything, bias upward the estimated employer contribution to the DC pension plan

<sup>26</sup> It should be noted that past employer contributions to DC plans are already included in the current market value of DC wealth.



I then modify the basic accounting variables as follows:

$$DC^* = DC + DCEMP$$

$$PW^* = DB + DC^*$$

$$PA^* = NWX + PW^*$$

$$RW^* = PW^* + SSW$$

$$AW^* = NWX + RW^*$$

Results are shown in Table 16. It is first of interest to note that in 2001, the average value of DCEMP among all households was \$29,300, or 55 percent of mean DC. This represents a substantial addition to the value of DC wealth and puts total DC wealth (DC\*) almost on a par with DB wealth. The addition of DCEMP in 2001 increases the mean value of PW by 31 percent, mean PA by 7 percent, the mean value of RW by 12 percent, and mean AW by only 5 percent. Among age group 47 to 64, the net effect of adding DCEMP is smaller. It increases mean DC wealth by 32 percent, mean PW by 18 percent, mean PA by 5, mean RW by 9 percent, and mean AW by only 4 percent. The inclusion of DCEMP actually has a larger effect on median values. In 2001, adding DCEMP increased the median value of PA by 24 percent, the median value of RW by 12 percent, and the median value of AW by 11 percent among all households, and median PA by 12 percent, median RW by 6 percent, and median AW by 6 percent among age group 47 to 64.

The addition of DCEMP, not surprisingly, greatly enhances the rate of growth of the mean value of these wealth components between 1983 and 2001. Among all households, mean PW\* increased by 109 percent (compared to 60 percent for PW), mean PA\* by 47 percent (compared to 38 percent for PA), mean RW\* by 31 percent (compared to 17 percent for RW), and mean AW\* by 32 percent (compared to 25 percent for AW). The pattern is similar for age group 47 to 64, though the differences are somewhat smaller.

Median values also show higher growth rates when DCEMP is added, though the differences are notably smaller than in the case of mean values. Among all households, median PA\* increased by 8 percent from 1983 to 2001 (compared to a 14 percent drop in median PA), median RW\* by 6 percent (compared to a 6 percent decline in median RW), and median AW fell by 12 percent (compared to a 21 percent drop in median AW). Results are similar for middle-aged households. Among this age group, median PA\* was up by 10 percent (compared to a 2 percent decline in median PA), median RW\* by 22 percent (compared to a 15 percent

increase in median RW), and median AW by 10 percent (compared to 3 percent gain in median AW).

Three points are noteworthy. First, median augmented wealth among all households still contracted over the 1983-2001 period, even after the inclusion of employer contributions to DC plans. Second, when DCEMP is included, median private accumulations among all households grew by only 7.6 percent over the 18-year period, or by 0.41 percent per year, a little more than half the annual rate of growth of net worth (NW). This gain is also slower than the 13.7 percent rise in median household income over the period. Third, among age group 47 to 64, both median PA\* and AW\* grew by 0.5 percent per year, much slower than the 1.3 percent per year gains in median NW. Thus, the main results derived in Sections 5 through 7 of the paper above without the inclusion of DCEMP remain largely unaltered.

The story is very similar when we consider new measures of wealth inequality (Table 16). It is first of interest that adding DCEMP reduces overall pension wealth inequality (from a Gini coefficient of 0.80 to 0.76 among all households and from 0.74 to 0.72 among age group 47-64). This is true despite the fact that the correlation of DCEMP and DC in 2001 is positive (0.24 among all households and 0.37 among middle-aged households), as is the correlation of DCEMP and PW (0.21 among all households and 0.36 among middle-aged households). Likewise, the inclusion of DCEMP lowers the inequality of private accumulations and augmented wealth (the correlation of DCEMP and NW is 0.11 and that between DCEMP and AW is 0.12 among all households and 0.21 and 0.22, respectively among middle-aged households), though the differences are smaller than for pension wealth, but increases somewhat the inequality of retirement wealth.

Nonetheless, the main conclusion, that the increase of inequality over the 1983-2001 period is greater when retirement wealth is included in the wealth definition, holds up. While the Gini coefficient for net worth among all households increased by 0.028 over the period, the Gini coefficient for PA\* rose by 0.031 (instead of by 0.054 in the case of PA) and that for AW\* climbed by 0.061 (instead of 0.072 in the case of AW). Among middle-aged households, the Gini coefficient for PA\* grew by 0.072 and that for AW\* by 0.079, in comparison to a gain of 0.037 in the case of net worth. Thus, it remains true that the mitigating effect of retirement wealth on overall wealth inequality mitigated over the years 1983 to 2001.

## **B. Future Tax Liability on Pension Wealth.**

A related issue concerns the tax liability attached to pension wealth. There are three types of tax liability associated with wealth holdings: (1) income taxes on the receipt of a pension benefit or a lump-sum pension distribution; (2) capital gains tax on the sale of an asset; and (3) estate tax liability on inheritances.

First, DC contributions are tax sheltered when they are made and subject to income tax on withdrawal. As a result, their current post-tax value is lower than their current market value. In principle, the post-tax value of DC plans should be used when computing current net worth, PA, and AW. DB pension benefits (and lump-sum distributions) are likewise taxable on receipt, so that the value of DB pension wealth should also be discounted when computing PA or AW. Second, marketable assets included in NWX are subject to capital gains taxes on the sale of the asset. The value of DC wealth or DB wealth would not be subject to capital gains taxes. Third, estate tax liability would apply to all asset components of net worth, including DC wealth, though not generally to DB wealth.<sup>27</sup>

It is beyond the scope of the paper to adjust wealth holdings for capital gains or estate taxes. However, it is possible to make a relatively crude adjustment to the values of DB and DC wealth for future taxes on receipt.<sup>28</sup> In principle, to make a proper calculation we would have to predict future income (and its composition), future tax deductions and exemptions, and the future tax schedule as well at retirement. For simplicity, I assume that for current workers income at retirement equals 80 percent of the pre-retirement income.<sup>29</sup> In the case of current beneficiaries, I assume that their (post-retirement) income remains fixed over their remaining life. I assume that marital status remains unchanged and that couples file joint returns. I assume that the tax schedule remains fixed over the remaining lifetime of the individual. I also treat the taxation of social security benefits according to the tax code current at the time of the survey.<sup>30</sup>

Results for 1983 and 2001 are shown in Table 17. It is first of note that for 2001, the

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<sup>27</sup> The exception would be lump-sum distributions from DB plans. This is a relatively small amount. In 2001, lump-sum distributions from DB plans comprised only 2.7 percent of total DB wealth.

<sup>28</sup> See Poterba (2004) for further discussion of the tax treatment of retirement savings.

<sup>29</sup> The 80 percent figure is a typical replacement rate. I also use Adjusted Gross Income for the income concept, which is provided in the SCF data.

<sup>30</sup> In 2001, social security benefits are “tax-free” if Adjusted Gross Income excluding Social Security benefits was less than \$32,000 for a married couple filing jointly and \$25,000 for singles or couples filing separately.

marginal tax rate on pension wealth averaged 13.4 percent and that on Social Security wealth averaged 9.8 percent (the corresponding figures for 1983 are 12.4 percent and 4.5 percent). The average net value of retirement wealth (that is, net of expected taxes on receipt and excluding DCEMP) was 89 percent of its gross average value in 1983 and 83 percent in 2001. Thus, federal taxes took a bigger bite out of expected retirement benefits in 2001 than in 1983.

As a result, while the mean value of gross RW (excluding DCEMP) among all households grew by 16.5 percent between 1983 and 2001, its net value increased by only 8.2 percent (the corresponding figures are 31.1 and 20.7 percent including DCEMP). Likewise, while the median value of gross RW among all households declined by 5.9 percent (increased by 5.8 percent with DCEMP), the median value of net RW dipped by 8.5 percent (rose by 3.2 percent with DCEMP). Correspondingly, the mean value of gross AW climbed by 25 percent (32 percent with DCEMP) over the period but the net value gained only 4 percent (9 percent with DCEMP). Finally, while the median value of gross AW dipped by 21 percent (12 percent with DCEMP), the median value of net AW plummeted by 24 percent (15 percent with DCEMP).

The results are quite similar for age group 47 to 64. The mean value of gross RW grew by 40 percent (52 percent with DCEMP), while the mean value of net RW gained only 31 Percent (41 percent with DCEMP). However, in this case the median value of both gross and net RW increased by about the same degree — 15 percent without DCEMP and 22 percent with DCEMP. The mean value of net AW also rose less than that of net AW among age group 47 to 64, and while the median value of gross AW gained 3.4 percent (9.7 percent with DCEMP), the median value of net AW increased by only 2.5 percent (6.9 percent with DCEMP).

As shown, in Table 17, netting out the implicit taxes in retirement wealth has a small effect on overall pension wealth inequality among all households (the Gini coefficient for PW without DCEMP in 2001 is reduced by 0.011) but has a larger effect on retirement wealth inequality (the Gini coefficient is lessened by 0.036). The latter result makes particular sense since the implicit tax rates on Social Security benefits are notably higher for high earners and netting out taxes from SSW is strongly redistributive. However, netting out implicit taxes on pension and Social Security wealth barely affects either PA or AW inequality. As a result, inequality trends in both PA and AW between 1983 and 2001 remain almost unchanged after subtracting implicit taxes from retirement wealth. For middle-aged households, netting out taxes does not have much effect on inequality trends for PA and AW (without DCEMP) but does

augment the degree to which inequality grows for PA\* and AW\* (including DCEMP) between 1983 and 2001 (the Gini coefficient grows by 0.098 for net PA\* compared to 0.072 for gross PA\* and by 0.093 for net AW\* compared to 0.079 for gross AW\*).

Thus, the basic story line is strengthened when net values are used instead of gross values. Between 1983 and 2001, the median net value of AW among all households plunged by 24 percent excluding DCEMP and by 12 percent when DCEMP is included. Among middle-aged households, median net AW budged upward by only 2.5 percent without DCEMP and by 6.9 percent with DCEMP included. The Gini coefficient for net AW rose by 0.072 and that for net AW\* by 0.061 among all households and by 0.083 for net AW and by 0.093 for net AW\* among age group 47 to 64.

## **9. CONCLUSION**

The 1980s and 1990s have witnessed the devolution of the traditional Defined Benefit pension system in favor of Defined Contribution pension coverage. In general, middle-aged and elderly Americans have seen marked improvements in both the mean and median levels of their marketable net worth over the period 1983 to 2001. On the other hand, traditional Defined Benefit pension coverage declined over the period from 1983 to 1998. The share of households covered by a DB plan fell by 18 percentage points, from 53 percent in 1983 to 34 percent in 2001. Mean DB pension wealth also decreased, from \$55,400 to \$41,200 (both in 2001 dollars). Average DC pension wealth, on the other hand, skyrocketed. By 2001, over half of all households held some form of DC type pension plan, compared to 11 percent in 1983.

The rise of DC pension plans more than fully compensated for the loss of DB type pension plans over the 1983-2001 period in terms of average values. Mean total pension wealth (the sum of DB plus DC wealth) increased by 60 percent in real terms over the period. The share of households covered by either a DB or a DC pension plan also grew over the period, from 54 to 66 percent. Moreover, mean private accumulations (the sum of net worth and pension wealth) soared by 37.5 percent.

However, mean social security wealth gained only 9.7 percent over the period, from \$127,200 to \$139,500 (both in 2001 dollars). Still, mean retirement wealth increased by 16.5 percent over these years. All told, mean augmented wealth grew by 25 percent, from \$448,200 in 1983 to \$561,000 in 2001.

However, the story looks quite different when we look at trends in median values. The median value of private accumulations tumbled by 13.5 percent. Median retirement wealth showed a 5.9 percent decline, from \$162,400 in 1983 to \$152,800 in 2001 (both in 2001 dollars). Altogether, median augmented wealth in 2001 dollars plummeted by 21 percent, from \$291,700 to \$257,600 (see Table 18).

Social security wealth continues to have a strong mitigating effect on overall wealth inequality, while private pension wealth has a very modest equalizing effect. However, retirement wealth had a weaker offsetting effect on wealth inequality in 2001 than in 1983. The addition of retirement wealth to net worth reduced the overall Gini coefficient in 1983 by 0.208, from 0.799 to 0.590, but by only 0.164 in 2001, from 0.826 to 0.663. As a result, while the Gini coefficient for net worth gained 0.028 over the 1983-2001 period, that for augmented wealth surged by 0.072.

Results for age group 47 to 64, those households most affected by the transition of the pension system, are similar. While median net worth climbed by 27.0 percent, median augmented wealth increased at a snail's pace, by only 3.4 percent, and median private accumulations actually fell by 2.2 percent. While the Gini ratio for standard wealth grew by 0.037, that for augmented wealth jumped by 0.082.

When future expected employer contributions to DC plans are added to pension wealth, the story line remains basically the same, though the results are softened a bit. On the other hand, when expected taxes on receipt are netted out of both pension and social security wealth, the basic findings are strengthened. When employer contributions are added, median augmented wealth among all households shrinks by 12 percent instead of 21 percent from 1983 to 2001, and when taxes are then netted out, it tumbles by 15 percent. Likewise, when adding employer DC contributions, the Gini coefficient for augmented wealth rises by 0.061 instead of 0.072 and when then subtracting expected taxes, it still rises by 0.061. A similar pattern of results exists for middle-aged households. When employer contributions are added and taxes are then netted out, median AW still grows very modestly by 6.9 percent instead of 3.4 percent and the Gini coefficient soars by 0.093 instead of 0.082.

The inequality “puzzle” discussed in the “Introduction” is now resolved when retirement wealth (and, in particular, the replacement of DB plans with DC ones) is included in the definition of household wealth. In particular, overall wealth inequality now climbs much more

strongly between 1983 and 2001 and much in line with the surge in income inequality and the steep gains in the stock market.

In sum, I find that despite the proliferation of defined contribution plans at a time when the stock market experienced one of its longest bull runs in history, general well-being did not improve. Indeed, median retirement wealth and median "augmented wealth" actually deteriorated over the 1983-2001 period. This is true because of a large shift in the composition of private retirement wealth away from defined benefit (DB) plans toward defined contribution (DC) plans.

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## **APPENDIX: ESTIMATION OF PENSION AND SOCIAL SECURITY WEALTH**

I follow the methodology (with a few modifications indicated below for subsequent years) laid out in the 1983 Survey of Consumer Finances codebook. This allows consistency with the estimates of both pension and social security wealth already provided in the 1983 Survey of Consumer Finances. The computations of retirement wealth in 1983 followed the following steps:

### **A. Pension Wealth**

Total gross pension wealth consists of two main components.<sup>31</sup>

1. (Gross) Present value of pensions from past jobs: The sum of the present value of past job pensions for head and spouse.
2. Gross present value of pensions from current jobs: The sum of the gross present value of current job non-thrift benefits for head and spouse. Expectations data are used for calculations.

The procedure is as follows. Pension coverage is first ascertained for current jobs. There are five possible categories: (1) covered and vested, anticipates benefits; (2) covered but not vested yet, anticipates benefits; (3) covered but not vested yet, does not anticipate benefits; (4) not covered, anticipates will be. Age when expected to be covered is ascertained; and (5) not covered, never will be.

For those who are covered by a pension plan or expect coverage, the person is asked how many distinct pensions plans he or she is covered by. For each plan, the age at which the pension benefits are expected to be given is then asked.

The actual expected annual retirement benefit is then determined by the following steps. First, the age at which the respondent will be vested in each plan is determined. Second, the age at which the respondent could retire with full benefits is ascertained. Third, the respondent was asked the nature of the formula used to determine the retirement benefits. There are six possibilities: (1) retirement formula based on age; (2) retirement formula based on years of service.; (3) retirement formula based on meeting both age and years of service criteria; (4) retirement formula based on the sum or age and years of service; (5) retirement formula based on meeting either age or years of service criteria; and (6) other combinations or formulas.

Fourth, the age at which the respondent could retire with some benefits was asked. The same six choices of the formula used was then given. Fifth, the age at which the respondent expected benefits to start was then asked.

Sixth, the expected retirement benefit was computed depending on the type of formula. This consists of three possibilities. (1) the annual pay in the final year of the job was computed. This variable, used in pension benefit calculations, is computed by projecting current pay to the year respondents say he/she will leave the job or retire. Wage growth is assumed to have three components: (i) occupation specific (adjusted for age) taken from the slopes in the CPS log-

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<sup>31</sup> A third though minor component is also provided: pensions from other non-specified sources.

wage regressions (for high-income observations this is assumed to be zero); (ii) a Social Security Plan II-B assumption of 1.5% annual economy wide real wage growth; and (iii) a Social Security Plan II-B assumption of 4.0% inflation.

(2) In some cases, the respondent reported expected retirement benefits. This variable is the expected dollar retirement benefits in the first year of eligibility as answered by the respondent. For some observations the dollar amount was reported directly, but for others it was computed by multiplying reported benefits as a percentage times the calculated projected final wage. The variable is given as an annual amount except when a lump sum is expected (in which case the lump sum amount is given).

(3) In some cases, the respondent reported expected retirement benefits as a percent of final pay. This variable is the expected retirement benefits in the first year of eligibility as answered by the respondent, expressed as a percent of their projected wages in their final year of work. For some observations the percent was reported directly, but for others it was computed by dividing the reported dollar benefit by the calculated projected final wage.

Seventh, on the basis of the responses above, the present value of pension benefits from each current and past plan applicable to both head and spouse was then computed. This variable is measured assuming an annual (or lump sum) pension benefit as given above, starting in the year of first benefits. Benefits for that and each succeeding year are adjusted for the probability of death and are discounted back to 1983. Sex-based Social Security mortality tables are used to compute the probabilities of death (standard for each year). These are capped at 109 years. Spousal survival benefits are assumed to be opted for 75 percent of the time and are randomly assigned when appropriate. Spousal survival benefits are also adjusted for death probabilities. Benefits are discounted at the 1983 long-term U.S. government bond rate of 10.85 percent.

Eighth, pension wealth was also computed for those individuals currently receiving pension benefits from past jobs. This was based on the following responses: (1) number of years receiving benefits and (2) amount of pension benefit pay received in 1982. For pensions already being received, the nominal value of the pension is assumed to be fixed, and is indexed to the year it started by the actual price changes observed as measured by the CPI. The present value of pension benefits from each job is then measured assuming an annual pension benefit as given starting in the year of first benefits (or 1983). Benefits for that and each succeeding year (adjusted for probability of receipt) are discounted back to 1983. Sex-based Social Security mortality tables are used to compute the probabilities of dying each year and/or living to receive any benefits. These are capped at 109 years. Spousal survival benefits are assumed to be opted for 75 percent of the time and are randomly assigned when appropriate. Spouse mortality tables are also used. Benefits are discounted at the 1983 long-term U.S. Government bond rate of 10.85 percent.

## **B. Social Security Wealth**

The gross present value of social security benefits is defined as: The sum of the gross present value of Social Security benefits for head and spouse. Social Security formula and current receipts used for calculations.

Among current Social Security benefit recipients, the steps are as follows: First, it was determined the kind of Social Security benefit received. The possibilities are: (1) retirement; (2) disability; (3) both retirement and disability; and (4) other kind.

Second, the respondent was asked the number of years receiving Social Security benefits. Third, both head and spouse were asked the amount received in 1982.

Among future recipients, the steps are as follows. First, both head and spouse were asked to report the age at which they expected to receive Social Security benefits (zero if he or she does not expect benefits). Second, the age at which Social Security benefits were expected to start was asked. Third, the number of years until the start of Social Security benefits was determined. Fourth, the respondent was asked the total number of years on Social Security jobs to current date. If this was not answered, then an estimate of Social Security coverage was used, summing over current and the three possible past jobs. Fifth, an estimate of future years on Social Security jobs was computed from retirement years indicated by head and spouse.

Sixth, data on number of years on Social Security jobs, wage rates for each known job, estimates of retirement dates, and dates of starting benefits were used as inputs to Social Security formulas to compute benefits. Seventh, estimates of Social Security benefits were provided. A calculated value was based on current job wage. All persons were assumed to work continuously until their stated age of full-time retirement, and then part-time until their stated age of final retirement. All persons were assumed to retire no later than 72 or age + 1 if currently over 72. Persons not currently working and over 50 were assumed not to work again. Wages were calculated by projecting current wages by the same method used to calculate final wages. Wage growth was assumed to have three components: (1) occupation specific (adjusted for age) taken from the slopes in the CPS log-wage spline regressions; (2) a Social Security plan II-B assumption of 1.5 percent annual economy wide real wage growth; and (3) a Social Security plan II-B assumption of 4.0 percent inflation. Part-time years (if currently working full-time) were assigned wages equal to 1/2 the projected full-time wages or the maximum amount allowable for full benefit receipt allowed by Social Security, whichever was smaller.

Eighth, the Social Security AIME (Average Indexed Monthly Earnings) used as the basis of computing the Social Security benefit base. The variable is the average covered Social Security earnings per month (including zeros) for all years from 1951 or age 22 (which ever is later) to age 60. These are indexed by a Social Security wage index to the year respondent is 60. Years after 60 can be substituted at nominal value. The five lowest years are dropped before an average AIME is computed. These procedures are mimicked using the SCF data on job earnings and future retirement plans to estimate an AIME value. Past and current job wages are projected back (and forward) to estimate earnings for each known year of work. These projections assume within-occupation real wage adjustments as taken from the CPS regressions (see past/current job), and economy-wide productivity growth and inflation as occurred or is projected to occur under the Social Security plan II-B. Other years of unknown jobs are filled in with terms from the closest known job to fill in the total number of Social Security covered years. Wages are then capped at the actual or projected Social Security maximum and minimum coverage amounts. The AIME was then computed using actual or projected Social Security wage indices. The variable is currently estimated for all persons projected to have future Social Security benefits.

Ninth, the Social Security PIA (Primary Insurance Amount) on an annual basis is the basis of the calculation of Social Security benefits. It is computed from the AIME. In 1982 the monthly PIA was computed as 90 percent of the first \$254 of AIME plus 32 percent of the next \$1274 plus 15 percent of the amount above. Calculations here take account of legislatively planned changes in this formula. The PIA is currently computed for all non-receivers projected to have future Social Security benefits.

Tenth, the present value of Social Security benefits is then computed assuming an annual benefit as given by the PIA estimate and starting in the year of first benefits (or 1983). Benefits for that and each succeeding year (adjusted for probability of receipt) are discounted back to 1983. Sex-based Social Security mortality tables are used to compute the probabilities of dying each year and/or living to receive any benefits. These are capped at 109 years. Benefits are discounted at the 1983 long-term U.S. Government bond rate of 10.85 percent.

Eleventh, spousal benefits are also assumed at 50 percent of the primary benefit if a spouse is present. However, this variable will be zero if no spousal benefits are expected (such as when the individual's own benefits are larger than their spousal benefits). The age at which spousal benefits begin is estimated. Spouse mortality tables are also used for these calculations. The age at which widow's benefits first could be drawn is also estimated. It is an estimate of the age at which the individual could start to receive Social Security widow's benefits upon the death of their spouse. This variable will be zero if widow's benefits could never be drawn. An adjustment is also made if it appeared that the recipient's benefits had been reduced because of work. Benefits are discounted at the 1983 long-term U.S. Government bond rate of 10.85 percent.

### C. Modifications for years after 1983

A few changes were made in the procedures for computing both pension and social security wealth. First, the regression equations used to compute future earnings was modified as follows:

Human capital earnings functions are estimated by gender, race, and schooling level. In particular, the sample is divided into 16 groups by the following characteristics: (i) white and Asian versus African-American and Hispanic; (ii) male and female; and (iii) less than 12 years of schooling, 12 years of schooling, 13 to 15 years of schooling, and 16 or more years. For each group, an earnings equation is estimated as follows:

$$\text{Log}(E_i) = b_0 + b_1 \text{Log}(H_i) + b_2 X_i + b_3 X_i^2 + b_4 \text{SE}_i + \sum_j b_j \text{OCCUP}_{ij} + b_{10} \text{MAR}_i + b_{11} \text{AS}_i + \epsilon_i,$$

where log is the natural logarithm;  $E_i$  is the current earnings of individual  $I$ ;  $H_i$  is annual hours worked in the current year;  $X_i$  is years of experience at current age (estimated as age minus years of schooling minus 5);  $\text{SE}_i$  is a dummy variable indicating whether the person is self-employed or working for someone else,  $\text{OCCUP}$  is a set of five dummy variables indicating occupation of employment ((a) professional and managerial; (b) technical, sales, or administrative support; (c) service; (d) craft, and (e) other blue-collar, with farming the omitted category);  $\text{MAR}$  is a dummy variable indicating whether the person is married or not married;  $\text{AS}$  is a dummy variable indicating whether the person is Asian or not (used only for regressions on the first racial

category); and  $\varepsilon$  is a stochastic error term. Future earnings are projected on the basis of the regression coefficients.<sup>32</sup>

Second, the ten-year treasury bond rate prevailing for each individual year (1989, 1992, 1995, and 1998) was used as the discount factor.

Third, I have used mortality rates by age, gender, and race instead of by age and gender alone in the computation of the present value of both pensions and social security wealth.

Fourth, for consistency with 1983, I have continued to employ the Social Security Plan II-B assumption of 1.5 percent annual economy wide real wage growth, even though this seems too high in comparison with the actual post-1973 growth in annual earnings (which has averaged about 0.2 percent per year). I have also used the Social Security Plan II-B assumption of 4.0 percent annual inflation, even though this seems too high.

#### **D. Questions on Work History**

Following is a sample of questions on work history drawn from the 1989 SCF codebook that is used to calculate the earnings profile of both head and spouse and to calculate the AIME for each:

1. Including any periods of self-employment, the military, and your current job, since you were 18, how many years have you worked full-time for all or most of the year?
2. Not counting your current job, have you ever had a full-time job that lasted for three years or more?
3. I want to know about the longest such job you had. Did you work for someone else, were you self-employed, or what?
4. When did you start working at that job?
5. When did you stop working at that job?
6. Since you were 18, have there been years when you only worked part-time for all or most of the year?
7. About how many years in total did you work part-time for all or most of the year?
8. Thinking now of the future, when do you expect to stop working full-time?
9. Do you expect to work part-time after that?
10. When do you expect to stop working altogether?

#### **E. Questions on DC plans.**

1. Does your employer make contributions to this [Defined Contribution] plan? Does the business make contributions to this plan?
2. What percent of pay or amount of money per month or year does your employer currently contribute?

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<sup>32</sup> This implicitly assumes that deviations from the regression line in the current year are a result of a transitory component to current income only. This procedure follows the conventions of the 1983 SCF codebook. .



**Table 1: Mean and Median Wealth and Income, 1983-2001**  
(In thousands, 2001 dollars)

	1983	1989	1992	1995	1998	2001	Percent Change		
							1983-89	1989-01	1983-01
<b>A. Net Worth (NW)</b>									
1. Median	59.3	63.5	54.2	53.0	65.9	73.5	7.0	15.8	23.9
2. Mean	231.0	264.6	257.3	237.7	293.6	380.1	14.6	43.7	64.6
<b>B. Net Worth excluding DC Pension Accounts (NWX)</b>									
1. Median	57.6	60.4	49.3	42.8	50.4	58.0	4.8	-3.9	0.7
2. Mean	247.1	278.5	243.7	208.6	278.8	326.6	12.7	17.3	32.2
<b>C. Income (CPS)<sup>a</sup></b>									
1. Median	37.1	41.3	38.7	39.6	42.2	42.2	11.2	2.3	13.7
2. Mean	45.2	52.2	49.0	52.2	56.3	58.2	15.5	11.6	28.9

Source: own computations from the 1983, 1989, 1992, 1995, 1998, and 2001 Surveys of Consumer Finances.  
a. Source for household income data: U.S. Bureau of the Census, Current Population Surveys, available on the internet: <http://www.census.gov/hhes/income/histinc/ineqtoc.html>.

**Table 2. The Size Distribution of Wealth and Income, 1983-2001**

Year	Percentage Share of Wealth or Income Held by:									Gini Coefficient
	Top 1.0%	Next 4.0%	Next 5.0%	Next 10.0%	Top 20.0%	Next 20.0%	3rd 20.0%	Bottom 40.0%	All	
<b>A. Net Worth (NW)</b>										
1983	33.8	22.3	12.1	13.1	81.3	12.6	5.2	0.9	100.0	0.799
1989	37.4	21.6	11.6	13.0	83.5	12.3	4.8	-0.7	100.0	0.832
1992	37.2	22.8	11.8	12.0	83.8	11.5	4.4	0.4	100.0	0.823
1995	38.5	21.8	11.5	12.1	83.9	11.4	4.5	0.2	100.0	0.828
1998	38.1	21.3	11.5	12.5	83.4	11.9	4.5	0.2	100.0	0.822
2001	33.4	25.8	12.3	12.9	84.4	11.3	3.9	0.3	100.0	0.826
<b>B. Net Worth excluding DC Pension Accounts (NWX)</b>										
1983	33.9	22.8	12.0	12.9	81.7	12.3	5.2	0.8	100.0	0.802
1989	35.6	23.3	11.9	12.9	83.7	12.2	4.8	-0.7	100.0	0.835
1998	38.1	23.7	11.2	12.1	85.1	11.1	4.0	-0.3	100.0	0.842
2001	36.9	25.7	11.7	12.0	86.2	10.2	3.6	-0.1	100.0	0.845
<b>C. Income (CPS)<sup>a</sup></b>										
	[Top 5 Percent]		[Next 15 Percent]							
1983	---	16.4---	---	28.3---	44.7	24.7	16.5	14.1	100.0	0.414
1989	---	18.9---	---	27.9---	46.8	24.0	15.8	13.3	100.0	0.431
1992	---	18.6---	---	28.3---	46.9	24.2	15.8	13.2	100.0	0.434
1995	---	21.0---	---	27.7---	48.7	23.3	15.2	12.8	100.0	0.450
1998	---	21.7---	---	27.5---	49.2	23.2	15.0	12.5	100.0	0.456
2001	---	22.4---	---	27.7---	50.1	23.0	14.6	12.2	100.0	0.466

Note: own computations from the 1983, 1989, 1992, 1995, 1998, and 2001 SCF.

a. Source for household income data: U.S. Bureau of the Census, Current Population Surveys, available on the internet: <http://www.census.gov/hhes/income/histinc/ineqtoc.html>.

**Table 3. Composition of Total Household Wealth, 1983, 1989 1998, and 2001  
(Percent of gross assets)**

<b>Component</b>	<b>1983</b>	<b>1989</b>	<b>1998</b>	<b>2001</b>
<b>Principal residence (gross value)</b>	<b>30.1</b>	<b>30.2</b>	<b>29.0</b>	<b>28.2</b>
<b>Other real estate (gross value)</b>	<b>14.9</b>	<b>14.0</b>	<b>10.0</b>	<b>9.8</b>
<b>Unincorporated business equity<sup>a</sup></b>	<b>18.8</b>	<b>17.2</b>	<b>17.7</b>	<b>17.2</b>
<b>Liquid assets<sup>b</sup></b>	<b>17.4</b>	<b>17.5</b>	<b>9.6</b>	<b>8.8</b>
<b>Pension accounts<sup>c</sup></b>	<b>1.5</b>	<b>2.9</b>	<b>11.6</b>	<b>12.3</b>
<b>Financial securities<sup>d</sup></b>	<b>4.2</b>	<b>3.4</b>	<b>1.8</b>	<b>2.3</b>
<b>Corporate stock and mutual funds</b>	<b>9.0</b>	<b>6.9</b>	<b>14.8</b>	<b>14.8</b>
<b>Net equity in personal trusts</b>	<b>2.6</b>	<b>3.1</b>	<b>3.8</b>	<b>4.8</b>
<b>Miscellaneous assets<sup>e</sup></b>	<b>1.3</b>	<b>4.9</b>	<b>1.8</b>	<b>1.8</b>
<b>Total</b>	<b>100.0</b>	<b>100.0</b>	<b>100.0</b>	<b>100.0</b>
<b>Debt on principal residence</b>	<b>6.3</b>	<b>8.6</b>	<b>10.7</b>	<b>9.4</b>
<b>All other debt<sup>f</sup></b>	<b>6.8</b>	<b>6.4</b>	<b>4.2</b>	<b>3.1</b>
<b>Total debt</b>	<b>13.1</b>	<b>15.0</b>	<b>15.0</b>	<b>12.5</b>
<b><u>Memo:</u></b>				
<b>Debt / equity ratio</b>	<b>15.1</b>	<b>17.6</b>	<b>17.6</b>	<b>12.5</b>
<b>Net home equity / total assets</b>	<b>23.8</b>	<b>21.6</b>	<b>18.2</b>	<b>18.8</b>

**Note:** own computations from the 1983, 1989, 1998 and 2001 Survey of Consumer Finances.

**a.** Net equity in unincorporated farm and non-farm businesses and closely-held corporations.

**b.** Checking accounts, savings accounts, time deposits, money market funds, certificates of deposits, and the cash surrender value of life insurance.

**c.** IRAs, Keogh plans, 401(k) plans, the accumulated value of defined contribution pension plans, and other retirement accounts.

**d.** Corporate bonds, government bonds, open-market paper, and notes.

**e.** Gold and other precious metals, royalties, jewelry, antiques, furs, loans to friends and relatives, future contracts, and miscellaneous assets.

**f.** Mortgage debt on all real property except principal residence; credit card, installment, and other consumer debt.

**Table 4. Composition of Household Wealth by Wealth Class, 1983 and 2001  
(Percent of gross assets)**

Component	Top One Percent		Next 19 Percent		Middle 3 Quintiles	
	1983	2001	1983	2001	1983	2001
Principal residence	8.1	8.4	29.1	26.8	61.6	59.2
Liquid assets (bank deposits, money market funds, and cash surrender value of life insurance)	8.5	5.7	21.4	9.3	21.4	12.1
Pension accounts	0.9	5.5	2.0	16.5	1.2	12.7
Corporate stock, financial securities, mutual funds, and personal trusts	29.5	33.6	13.0	21.4	3.1	6.2
Unincorporated business equity other real estate	52.0	44.3	32.8	24.4	11.4	8.5
Miscellaneous assets	1.0	2.6	1.6	1.7	1.3	1.2
<b>Total assets</b>	<b>100.0</b>	<b>100.0</b>	<b>100.0</b>	<b>100.0</b>	<b>100.0</b>	<b>100.0</b>
<b>Memo:</b>						
Debt / equity ratio	5.9	2.4	10.9	8.9	37.4	31.7
Debt / income ratio	86.8	33.6	72.8	76.9	66.9	100.3

**Note:** own computations from the 1983 and 2001 Surveys of Consumer Finances.

**Table 5. Percent of Households with Pension Wealth, 1983-2001**

	1983	1989	1998	2001	Percentage Point Change		
					1983-89	1989-01	1983-01
<b><u>A. All Households</u></b>							
1. DC Accounts	0.9	24.0	48.8	52.2	13.0	28.2	41.3
2. DB Plans	2.6	45.6	37.1	34.4	(7.0)	(11.2)	(18.2)
3. Pension Wealth	54.4	56.0	64.7	65.6	1.5	9.6	11.2
<b><u>B. Ages 46 and under</u></b>							
1. DC Accounts	13.7	31.2	52.0	53.8	17.5	22.6	40.1
2. DB Plans	32.6	37.9	27.4	22.8	5.3	(15.1)	(9.8)
3. Pension Wealth	35.6	52.2	60.4	60.7	16.6	8.6	25.1
<b><u>C. Ages 47-64</u></b>							
1. DC Accounts	11.9	28.3	59.7	62.0	16.4	33.7	50.1
2. DB Plans	68.9	56.8	45.9	45.3	(12.1)	(11.5)	(23.7)
3. Pension Wealth	70.2	67.5	74.7	75.9	(2.8)	8.4	5.6
<b><u>D. Ages 65 and over</u></b>							
1. DC Accounts	2.1	1.3	32.3	35.0	(0.8)	33.7	32.9
2. DB Plans	66.2	1.3	48.9	46.5	(14.9)	(4.7)	(19.6)
3. Pension Wealth	66.9	1.8	62.2	62.6	(15.1)	10.8	(4.3)

Note: own computations from the 1983, 1989, 1998, and 2001 Surveys of Consumer Finances. Households are classified into age groups by the age of the head of household.

Key: Pension Wealth PW = DB + DC

**Table 6. Mean and Median Household Pension Wealth, 1983 and 2001**  
(In thousands, 2001 dollars)

	1983	2001	% Change, 1983-2001
<b><u>A. All Households</u></b>			
1. Mean DC Pension Wealth	3.9	53.7	1273.7
2. Mean DB Pension Wealth	55.4	41.2	-25.7
3. Mean Pension Wealth	59.3	94.8	59.8
Memo: Median Pension Wealth among PW Holders Only	64.3	53.7	-16.5
<b><u>B. Ages 46 and under</u></b>			
1. Mean DC Pension Wealth	2.6	28.2	1005.6
2. Mean DB Pension Wealth	23.0	16.9	-26.8
3. Mean Pension Wealth	25.6	45.1	76.1
Memo: Median Pension Wealth among PW Holders Only	26.8	30.4	13.6
<b><u>C. Ages 47-64</u></b>			
1. Mean DC Pension Wealth	8.0	96.6	1105.8
2. Mean DB Pension Wealth	94.5	74.3	-21.4
3. Mean Pension Wealth	102.5	170.8	66.6
Memo: Median Pension Wealth among PW Holders Only	89.9	98.4	9.4
<b><u>D. Ages 65 and over</u></b>			
1. Mean DC Pension Wealth	1.8	53.6	2835.1
2. Mean DB Pension Wealth	60.5	52.1	-14.0
3. Mean Pension Wealth	62.3	105.6	69.5
Memo: Median Pension Wealth among PW Holders Only	66.0	76.7	16.3

Note: own computations from the 1983 and 2001 Surveys of Consumer Finances.  
Households are classified into age groups by the age of the head of household.

Key: Pension Wealth PW = DB + DC

**Table 7. Inequality of Pension Wealth among Pension Holders, 1983 and 2001  
(Gini coefficients)**

	1983	2001	Change, 1983-2001
<b><u>A. All Pension Holders</u></b>			
1. DC Accounts	0.792	0.741	-0.051
2. DB Plans	0.555	0.638	0.084
3. Pension Wealth	0.537	0.695	0.157
Memo: PW among all households in group	0.749	0.800	0.051
<b><u>B. Pension Holders: Ages 46 and under</u></b>			
1. DC Accounts	0.778	0.719	-0.059
2. DB Plans	0.646	0.568	-0.078
3. Pension Wealth	0.664	0.677	0.013
Memo: PW among all households in group	0.881	0.804	-0.077
<b><u>C. Pension Holders: Ages 47-64</u></b>			
1. DC Accounts	0.732	0.714	-0.018
2. DB Plans	0.530	0.617	0.088
3. Pension Wealth	0.539	0.654	0.115
Memo: PW among all households in group	0.677	0.738	0.061
<b><u>D. Pension Holders: Ages 65 and over</u></b>			
1. DC Accounts	0.687	0.703	0.015
2. DB Plans	0.463	0.665	0.202
3. Pension Wealth	0.472	0.666	0.194
Memo: PW among all households in group	0.646	0.791	0.144

**Note:** own computations from the 1983 and 2001 Surveys of Consumer Finances.  
Households are classified into age groups by the age of the head of household.

**Key:** Pension Wealth PW = DB + DC

**Table 8. Contribution of DC and DB Pensions to Overall Pension Wealth (PW) Inequality, 1983 and 2001**

Category	All Households				Households, Ages 47-64			
	Change		Contribution to Change	Change		Contribution to Change		
	1983	2001		1983-2001	1983		2001	
<b>A. Coefficient of Variation (CV)</b>								
1) PW	2.07	2.66	0.60	1.63	2.15	0.51		
2) DC	8.52	3.69	-4.83	6.65	2.96	-3.69		
3) DB	2.02	3.20	1.18	1.58	2.67	1.09		
<b>B. Correlation (DC,DB)</b>								
	0.17	0.14	-0.03	0.19	0.12	-0.07		
<b>C. Decomposition of CV<sup>2</sup>(PW)</b>								
1) p <sub>1</sub> <sup>2</sup> CV <sup>2</sup> (DC)	0.34	4.35	4.01	0.25	2.80	2.55		
2) p <sub>2</sub> <sup>2</sup> CV <sup>2</sup> (DB)	3.55	1.93	-1.62	2.15	1.35	-0.80		
3) 2CC(DC,DB)	0.37	0.80	0.43	0.28	0.46	0.18		
4) CV <sup>2</sup> (PW)	4.27	7.08	2.82	2.67	4.61	1.94		
<b>D. Percentage Decomposition of CV<sup>2</sup>(PW)</b>								
1) p <sub>1</sub> <sup>2</sup> CV <sup>2</sup> (DC)	8.1	61.4		9.3	60.8		131.9	
2) p <sub>2</sub> <sup>2</sup> CV <sup>2</sup> (DB)	83.2	27.3		80.4	29.3		-41.2	
3) 2CC(DC,DB)	8.7	11.2		10.4	9.9		9.3	
4) CV <sup>2</sup> (PW)	100.0	100.0		100.0	100.0		100.0	

Note: own computations from the 1983 and 2001 Survey of Consumer Finances. Key:

- CV = Coefficient of variation
- CC = Coefficient of covariation
- p<sub>1</sub> = DC / PW
- p<sub>2</sub> = DB / PW

Note that the 1983 observation is excluded from the calculation if a missing value exists for either DB or DC.

a. Defined as the change in the component over the 1983-2001 period divided by the change in the CV Squared of PW over the period.

**Table 9. Mean, Median, and the Inequality of Private Accumulations, 1983 and 2001**  
(In thousands, 2001 dollars)

	Mean Values		% Change, 1983-2001	Median Values		% Change, 1983-2001	Gini Coefficients		Change, 1983-2001
	1983	2001		1983	2001		1983	2001	
<b>A. All Households</b>									
1. Net Worth minus DC (NWX)	247.1	326.6	32.2	57.6	58.0	0.7	0.802	0.845	0.043
2. Net Worth (NW)	251.0	380.3	51.5	64.5	73.5	14.0	0.799	0.826	0.028
3. Private Accumulations (PA)	306.4	421.4	37.5	111.6	96.5	-13.5	0.746	0.800	0.054
<b>B. Ages 46 and under</b>									
1. Net Worth minus DC (NWX)	106.8	146.8	37.4	23.1	12.8	-44.8	0.801	0.892	0.091
2. Net Worth (NW)	109.4	175.1	60.0	26.3	20.7	-21.3	0.797	0.859	0.062
3. Private Accumulations (PA)	132.4	191.9	44.9	31.8	31.8	0.0	0.793	0.832	0.039
<b>C. Ages 47-64</b>									
1. Net Worth minus DC (NWX)	365.1	501.7	37.4	108.1	103.5	-4.3	0.762	0.823	0.061
2. Net Worth (NW)	373.1	598.3	60.4	108.3	137.6	27.0	0.761	0.798	0.037
3. Private Accumulations (PA)	467.6	672.6	43.8	204.4	199.9	-2.2	0.681	0.762	0.081
<b>D. Ages 65 and over</b>									
1. Net Worth minus DC (NWX)	370.8	504.0	35.9	101.4	142.0	40.0	0.777	0.766	-0.011
2. Net Worth (NW)	372.7	557.6	49.6	101.9	150.8	47.9	0.778	0.762	-0.016
3. Private Accumulations (PA)	433.2	609.7	40.7	158.5	206.0	30.0	0.715	0.735	0.020

Note: own computations from the 1983 and 2001 Surveys of Consumer Finances. Households are classified into age groups by the age of the head of household.

Key: Private Accumulations PA = NWX+PW.



Table 10. Contribution of Pensions to Inequality in Private Accumulations (PA), 1983 and 2001

Category	All Households			Households, Ages 47-64		
	1983	2001	Change 1983-2001	1983	2001	Change 1983-2001
<b>A. Coefficient of Variation (CV)</b>						
1) PA	4.89	4.95	0.06	3.89	4.13	0.24
2) NWX	5.90	6.15	0.25	4.87	5.30	0.43
3) PW	2.07	2.66	0.60	1.63	2.15	0.51
<b>B. Correlation (NWX,PW)</b>						
	0.21	0.26	0.05	0.21	0.26	0.05
<b>C. Decomposition of CV<sup>2</sup>(PA)</b>						
1) p <sub>1</sub> <sup>2</sup> CV <sup>2</sup> (NWX)	22.97	22.69	-0.27	14.43	15.63	1.20
2) p <sub>2</sub> <sup>2</sup> CV <sup>2</sup> (PW)	0.15	0.36	0.21	0.13	0.30	0.17
3) 2CC(NWX,PW)	0.80	1.50	0.70	0.57	1.13	0.56
4) CV <sup>2</sup> (PA)	23.91	24.55	0.63	15.12	17.05	1.93
<b>D. Percentage Decomposition of CV<sup>2</sup>(PA)</b>						
1) p <sub>1</sub> <sup>2</sup> CV <sup>2</sup> (NWX)	96.0	92.4		95.4	91.7	62.3
2) p <sub>2</sub> <sup>2</sup> CV <sup>2</sup> (PW)	0.6	1.5		0.9	1.7	8.7
3) 2CC(NWX,PW)	3.3	6.1		3.7	6.6	29.0
4) CV <sup>2</sup> (PA)	100.0	100.0		100.0	100.0	100.0

Note: own computations from the 1983 and 2001 Surveys of Consumer Finances. Key:

CV = Coefficient of variation

CC = Coefficient of covariation

p<sub>1</sub> = NWX / PA

p<sub>2</sub> = PW / PA

Note that the 1983 observation is excluded from the calculation if a missing value exists for either DB or DC.

a. Defined as the change in the component over the 1983-2001 period divided by the change in the CV Squared of PA over the period.

Table 11. Mean, Median and the Inequality of Retirement Wealth, 1983 and 2001  
(In thousands, 2001 dollars)

	Mean Values		% Change,		Median Values		% Change,		Gini Coefficients		Change, 1983-2001
	1983	2001	1983-2001		1983	2001	1983-2001		1983	2001	
<b>A. All Households</b>											
1. Pension Wealth PW	59.3	94.8	59.8		4.9	10.9	122.3		0.749	0.800	0.051
2. Social Security Wealth SSW	127.2	139.5	9.7		112.5	120.7	7.2		0.455	0.344	-0.111
3. Retirement Wealth RW	201.2	234.4	16.5		162.4	152.8	-5.9		0.430	0.492	0.061
<b>B. Ages 46 and under</b>											
1. Pension Wealth PW	25.6	45.1	76.1		0.0	3.5	--		0.881	0.804	-0.077
2. Social Security Wealth SSW	66.5	109.8	65.2		61.1	102.4	67.7		0.488	0.320	-0.168
3. Retirement Wealth RW	109.8	154.9	41.1		89.7	115.7	29.1		0.507	0.430	-0.077
<b>C. Ages 47-64</b>											
1. Pension Wealth PW	102.5	170.8	66.6		42.4	50.0	18.0		0.677	0.738	0.061
2. Social Security Wealth SSW	152.7	184.5	20.8		144.7	176.3	21.9		0.366	0.297	-0.069
3. Retirement Wealth RW	254.8	355.3	39.5		212.7	244.6	15.0		0.390	0.466	0.076
<b>D. Ages 65 and over</b>											
1. Pension Wealth PW	62.3	105.6	69.5		34.8	10.7	-69.3		0.646	0.791	0.144
2. Social Security Wealth SSW	147.3	146.6	-0.4		132.6	127.0	-4.3		0.461	0.356	-0.105
3. Retirement Wealth RW	209.6	252.2	20.4		177.4	171.4	-3.4		0.377	0.488	0.111

Note: own computations from the 1983 and 2001 Surveys of Consumer Finances. Households are classified into age groups by the age of the head of household. Computations are for all households in the group. Key:

$$PW = DC + DB. \quad RW = PW + SSW.$$

Table 12. Contribution of Pension and Social Security Wealth to Inequality in Retirement Wealth, 1983 and 2001

Category	All Households				Households, Ages 47-64			
	Change		Contribution		Change		Contribution	
	1983	2001	1983-2001	1983-2001 <sup>a</sup>	1983	2001	1983-2001	1983-2001 <sup>a</sup>
<b>A. Coefficient of Variation (CV)</b>								
1) RW	0.86	1.27	0.41		0.79	1.16	0.37	
2) PW	1.75	2.66	0.91		1.63	2.15	0.51	
3) SSW	0.81	0.63	-0.18		0.65	0.53	-0.12	
<b>B. Correlation (PW,SSW)</b>								
	0.11	0.39	0.29		0.09	0.35	0.26	
<b>C. Decomposition of CV<sup>2</sup>(RW)</b>								
1) p <sub>1</sub> <sup>2</sup> CV <sup>2</sup> (PW)	0.42	1.16	0.74		0.43	1.06	0.63	
2) p <sub>2</sub> <sup>2</sup> CV <sup>2</sup> (SSW)	0.26	0.14	-0.12		0.15	0.08	-0.08	
3) 2CC(PW,SSW)	0.07	0.32	0.25		0.05	0.20	0.15	
4) CV <sup>2</sup> (RW)	0.75	1.62	0.87		0.63	1.34	0.71	
<b>D. Percentage Decomposition of CV<sup>2</sup>(RW)</b>								
1) p <sub>1</sub> <sup>2</sup> CV <sup>2</sup> (PW)	56.2	71.7		85.0	68.6	79.3		88.9
2) p <sub>2</sub> <sup>2</sup> CV <sup>2</sup> (SSW)	34.5	8.7		-13.5	24.0	5.6		-10.6
3) 2CC(PW,SSW)	9.3	19.6		28.5	7.4	15.0		21.7
4) CV <sup>2</sup> (RW)	100.0	100.0		100.0	100.0	100.0		100.0

Note: own computations from the 1983 and 2001 Surveys of Consumer Finances. Key:

CV = Coefficient of variation

CC = Coefficient of covariation

p<sub>1</sub> = PW / RW

p<sub>2</sub> = SSW / RW

Note that the 1983 observation is excluded from the calculation if a missing value exists for either DB or DC.

a. Defined as the change in the component over the 1983-2001 period divided by the change in the CV Squared of RW over the period.

	Mean Values		% Change,		Median Values		% Change,	
	1983	2001	1983-2001	1983-2001	1983	2001	1983-2001	1983-2001
<b>A. All Households</b>								
1. Net Worth (NW)	251.0	380.3	51.5		64.5	73.5	14.0	
2. Augmented Wealth (AW)	448.2	561.0	25.2		291.7	231.5	-20.6	
<b>B. Ages 46 and under</b>								
1. Net Worth (NW)	109.4	175.1	60.0		26.3	20.7	-21.3	
2. Augmented Wealth (AW)	216.6	301.7	39.3		172.4	140.9	-18.3	
<b>C. Ages 47-64</b>								
1. Net Worth (NW)	373.1	598.3	60.4		108.3	137.6	27.0	
2. Augmented Wealth (AW)	619.9	857.0	38.3		376.4	389.2	3.4	
<b>D. Ages 65 and over</b>								
1. Net Worth (NW)	372.7	557.6	49.6		101.9	150.8	47.9	
2. Augmented Wealth (AW)	580.4	756.3	30.3		312.1	352.6	13.0	
<p>Note: own computations from the 1983 and 2001 Surveys of Consumer Finances. Households are classified into age groups by the age of the head of household. Key:</p> <p>Augmented Wealth AW = NWX + PW + SSW.</p>								

Table 14. Inequality of Augmented Wealth, 1983 and 2001 (Gini coefficients)		1983	2001	Change, 1983-2001
<b><u>A. All Households</u></b>				
1.	Net Worth (NW)	0.799	0.826	0.028
2.	Private Accumulations (PA)	0.746	0.800	0.054
3.	NWX + SSW	0.631	0.665	0.034
4.	Augmented Wealth (AW)	0.590	0.663	0.072
<b><u>B. Ages 46 and under</u></b>				
1.	Net Worth (NW)	0.797	0.859	0.062
2.	Private Accumulations (PA)	0.793	0.832	0.039
3.	NWX + SSW	0.644	0.612	-0.032
4.	Augmented Wealth (AW)	0.617	0.617	0.000
<b><u>C. Ages 47-64</u></b>				
1.	Net Worth (NW)	0.761	0.798	0.037
2.	Private Accumulations (PA)	0.681	0.762	0.081
3.	NWX + SSW	0.600	0.655	0.055
4.	Augmented Wealth (AW)	0.558	0.640	0.082
<b><u>D. Ages 65 and over</u></b>				
1.	Net Worth (NW)	0.778	0.762	-0.016
2.	Private Accumulations (PA)	0.715	0.735	0.020
3.	NWX + SSW	0.627	0.637	0.010
4.	Augmented Wealth (AW)	0.586	0.632	0.047
<p><b>Note:</b> own computations from the 1983 and 2001 Surveys of Consumer Finances.  Households are classified into age groups by the age of the head of household. Key:  Private Accumulations PA = NWX+PW.  Augmented Wealth AW = NWX + PW + SSW.</p>				

Table 15. Contribution of Net Worth and Retirement Wealth to Inequality in Augmented Wealth, 1983 and 2001

Category	All Households				Households, Ages 47-64			
	Change		Contribution		Change		Contribution	
	1983	2001	1983-2001	1983-2001 <sup>a</sup>	1983	2001	1983-2001	1983-2001 <sup>a</sup>
<b>A. Coefficient of Variation (CV)</b>								
1) AW	3.36	3.76	0.40		2.95	3.27	0.32	
2) NWX	5.38	6.15	0.77		4.87	5.30	0.43	
3) RW	0.86	1.27	0.41		0.79	1.16	0.37	
<b>B. Correlation (NWX,RW)</b>								
	0.22	0.27	0.06		0.21	0.27	0.06	
<b>C. Decomposition of CV<sup>2</sup>(AW)</b>								
1) p <sub>1</sub> <sup>2</sup> CV <sup>2</sup> (NWX)	10.70	12.81	2.11		8.18	9.63	1.45	
2) p <sub>2</sub> <sup>2</sup> CV <sup>2</sup> (RW)	0.11	0.28	0.17		0.11	0.23	0.12	
3) 2CC(NWX,RW)	0.48	1.04	0.56		0.39	0.81	0.42	
4) CV <sup>2</sup> (AW)	11.30	14.13	2.83		8.67	10.67	1.99	
<b>D. Percentage Decomposition of CV<sup>2</sup>(AW)</b>								
1) p <sub>1</sub> <sup>2</sup> CV <sup>2</sup> (NWX)	94.7	90.6		74.4	94.3	90.2		72.7
2) p <sub>2</sub> <sup>2</sup> CV <sup>2</sup> (RW)	1.0	2.0		5.9	1.2	2.2		6.2
3) 2CC(NWX,RW)	4.3	7.4		19.7	4.5	7.6		21.2
4) CV <sup>2</sup> (AW)	100.0	100.0		100.0	100.0	100.0		100.0

Note: own computations from the 1983 and 2001 Survey of Consumer Finances. Key:

CV = Coefficient of variation

CC = Coefficient of covariation

p<sub>1</sub> = NWX / AW

p<sub>2</sub> = RW / AW

Note that the 1983 observation is excluded from the calculation if a missing value exists for either DB or DC.

a. Defined as the change in the component over the 1983-2001 period divided by the change in the CV Squared of AW over the period.

**Table 16. Effect of including Employer Contribution to DC Plans (DCEMP) on Net Worth, Private Accumulations, and Augmented Wealth, 1983 and 2001**  
(In thousands, 2001 dollars)

	1983	2001		Change, 1983-2001	
		W/out DCEMP	With DCEMP	W/out DCEMP	With DCEMP
<b>I. All Households</b>					
<b>A. Mean Values</b>					
1. DC Pension Wealth	3.9	53.7	83.0	1273.7%	2023.3%
2. Pension Wealth (PW)	59.3	94.8	124.1	59.8%	109.2%
3. Net Worth (NW)	251.0	380.3	380.3	51.5%	51.5%
4. Private Accumulations (PA)	306.4	421.4	450.7	37.5%	47.1%
5. Retirement Wealth (RW)	201.2	234.4	263.6	16.5%	31.1%
6. Augmented Wealth (AW)	448.2	561.0	590.2	25.2%	31.7%
<b>B. Median Values</b>					
1. DC Pension Wealth	0.0	0.6	2.0	--	--
2. Pension Wealth (PW)	4.9	10.9	27.7	122.3%	464.5%
3. Net Worth (NW)	64.5	73.5	73.5	14.0%	14.0%
4. Private Accumulations (PA)	111.6	96.5	120.0	-13.5%	7.6%
5. Retirement Wealth (RW)	162.4	152.8	171.8	-5.9%	5.8%
6. Augmented Wealth (AW)	291.7	231.5	257.6	-20.6%	-11.7%
<b>C. Gini Coefficients</b>					
1. Pension Wealth (PW)	0.749	0.800	0.761	0.051	0.012
2. Net Worth (NW)	0.799	0.826	0.826	0.028	0.028
3. Private Accumulations (PA)	0.746	0.800	0.778	0.054	0.031
4. Retirement Wealth (RW)	0.430	0.492	0.498	0.061	0.068
5. Augmented Wealth (AW)	0.590	0.663	0.652	0.072	0.061
<b>II. Ages 47-64</b>					
<b>A. Mean Values</b>					
1. DC Pension Wealth	8.0	96.6	127.3	1105.8%	1490.1%
2. Pension Wealth (PW)	102.5	170.8	201.6	66.6%	96.6%
3. Net Worth (NW)	373.1	598.3	598.3	60.4%	60.4%
4. Private Accumulations (PA)	467.6	672.6	703.4	43.8%	50.4%
5. Retirement Wealth (RW)	254.8	355.3	386.1	39.5%	51.6%
6. Augmented Wealth (AW)	619.9	857.0	887.8	38.3%	43.2%
<b>B. Median Values</b>					
1. DC Pension Wealth	0.0	9.6	19.0	--	--
2. Pension Wealth (PW)	42.4	50.0	69.0	18.0%	62.8%
3. Net Worth (NW)	108.3	137.6	137.6	27.0%	27.0%
4. Private Accumulations (PA)	204.4	199.9	224.1	-2.2%	9.6%
5. Retirement Wealth (RW)	212.7	244.6	259.6	15.0%	22.0%
6. Augmented Wealth (AW)	376.4	389.2	413.1	3.4%	9.7%
<b>C. Gini Coefficients</b>					
1. Pension Wealth (PW)	0.677	0.738	0.722	0.061	0.045
2. Net Worth (NW)	0.761	0.798	0.798	0.037	0.037
3. Private Accumulations (PA)	0.681	0.762	0.752	0.081	0.072
4. Retirement Wealth (RW)	0.390	0.466	0.477	0.076	0.087
5. Augmented Wealth (AW)	0.558	0.640	0.637	0.082	0.079

Note: own computations from the 1983 and 2001 Surveys of Consumer Finances. Key:

Private Accumulations PA = NWX+PW. Retirement Wealth RW = PW + SSW

Augmented Wealth AW = NWX + PW + SSW.

**Table 17. Augmented Wealth, Net of Federal Income Taxes on Receipt, both excluding and including Employer Contribution to DC Plans, 1983 and 2001**  
(In thousands, 2001 dollars)

	1983	2001		Change, 1983-2001	
		W/out DCEMP	With DCEMP	W/out DCEMP	With DCEMP
<b><u>I. All Households</u></b>					
<b><u>A. Mean Values</u></b>					
1. Pension Wealth (PW)	48.6	73.0	95.5	50.3%	96.5%
2. Net Worth (NW)	251.0	380.3	380.3	51.5%	51.5%
3. Private Accumulations (PA)	308.1	399.6	422.1	29.7%	37.0%
4. Retirement Wealth (RW)	179.8	194.6	217.0	8.2%	20.7%
5. Augmented Wealth (AW)	499.4	521.2	543.6	4.4%	8.9%
<b><u>B. Median Values</u></b>					
1. Pension Wealth (PW)	4.2	9.2	23.5	--	--
2. Net Worth (NW)	64.5	73.5	73.5	14.0%	14.0%
3. Private Accumulations (PA)	105.9	91.2	113.7	-13.9%	7.4%
4. Retirement Wealth (RW)	148.7	136.0	153.4	-8.5%	3.2%
5. Augmented Wealth (AW)	280.4	213.7	237.3	-23.8%	-15.4%
<b><u>C. Gini Coefficients</u></b>					
1. Pension Wealth (PW)	0.751	0.789	0.746	0.037	-0.006
2. Net Worth (NW)	0.799	0.826	0.826	0.028	0.028
3. Private Accumulations (PA)	0.748	0.802	0.781	0.054	0.033
4. Retirement Wealth (RW)	0.413	0.456	0.460	0.043	0.047
5. Augmented Wealth (AW)	0.591	0.664	0.652	0.073	0.061
<b><u>II. Ages 47-64</u></b>					
<b><u>A. Mean Values</u></b>					
1. Pension Wealth (PW)	82.5	129.8	152.5	57.3%	84.8%
2. Net Worth (NW)	373.1	598.3	598.3	60.4%	60.4%
3. Private Accumulations (PA)	444.5	631.6	654.3	42.1%	47.2%
4. Retirement Wealth (RW)	220.1	287.9	310.7	30.8%	41.1%
5. Augmented Wealth (AW)	582.1	789.7	812.4	35.7%	39.6%
<b><u>B. Median Values</u></b>					
1. Pension Wealth (PW)	38.5	41.6	58.0	8.2%	50.7%
2. Net Worth (NW)	108.3	137.6	137.6	27.0%	27.0%
3. Private Accumulations (PA)	192.1	182.3	208.0	-5.1%	8.3%
4. Retirement Wealth (RW)	185.3	214.2	227.9	15.6%	22.9%
5. Augmented Wealth (AW)	349.9	358.6	374.1	2.5%	6.9%
<b><u>C. Gini Coefficients</u></b>					
1. Pension Wealth (PW)	0.656	0.722	0.703	0.066	0.047
2. Net Worth (NW)	0.761	0.798	0.798	0.037	0.037
3. Private Accumulations (PA)	0.683	0.765	0.781	0.082	0.098
4. Retirement Wealth (RW)	0.366	0.425	0.460	0.060	0.094
5. Augmented Wealth (AW)	0.558	0.642	0.652	0.083	0.093

Note: own computations from the 1983 and 2001 Surveys of Consumer Finances.

Households are classified into age groups by the age of the head of household. Key:

Private Accumulations PA = NWX+PW.

Retirement Wealth RW = PW + SSW.

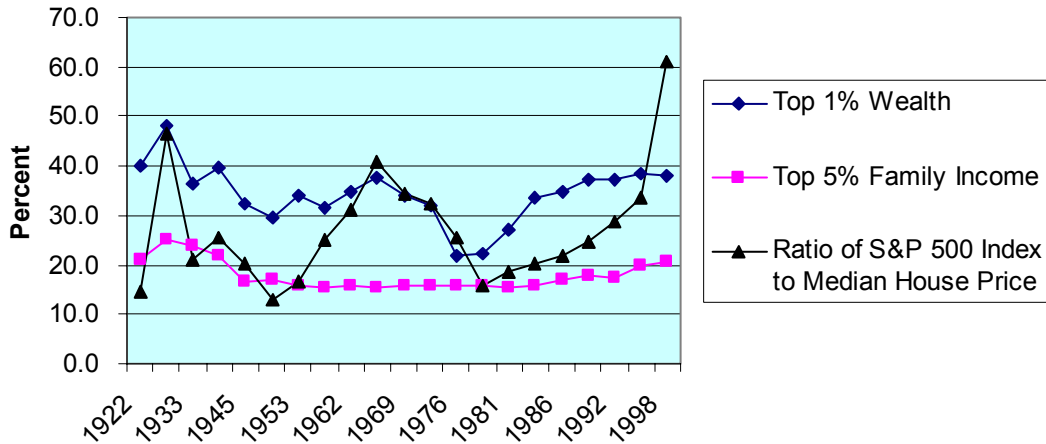
Augmented Wealth AW = NWX + PW + SSW.



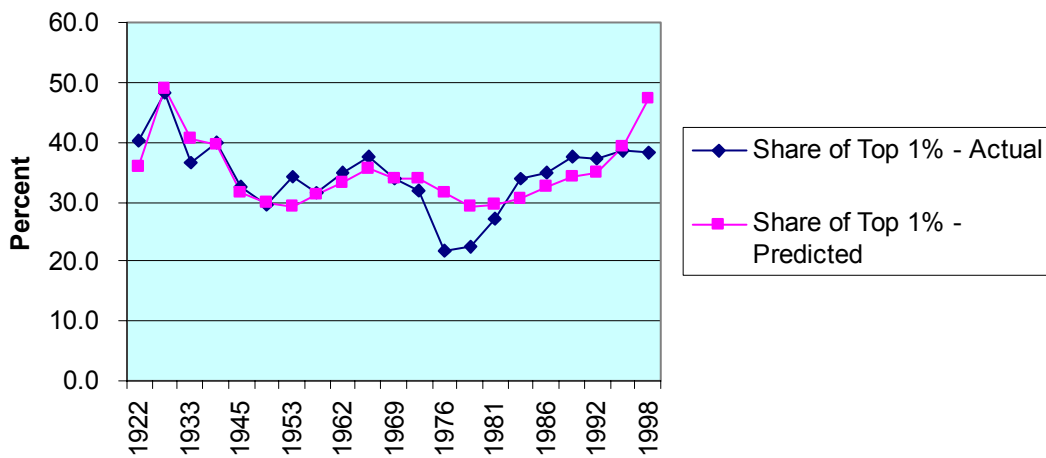
Table 18. Summary Table of Wealth Trends, 1983 and 2001  
(In thousands, 2001 dollars)

	Median Values		% Change, 1983-2001	Gini Coefficients		Change, 1983-2001
	1983	2001		1983	2001	
<b>I. All Households</b>						
1. Income (CPS) <sup>a</sup>	37.1	42.2	13.7	0.414	0.466	0.052
2. Net Worth (NW)	64.5	73.5	14.0	0.799	0.826	0.028
3. Private Accumulations (PA)	111.6	96.5	-13.5	0.746	0.800	0.054
3a. Private Accumulations with DCEMP (PA*)	111.6	120.0	7.6	0.746	0.778	0.031
3b. Net Private Accumulations (PA)	105.9	91.2	-13.9	0.748	0.802	0.054
3c. Net Private Accumulations with DCEMP (PA*)	105.9	113.7	7.4	0.748	0.781	0.033
4. Augmented Wealth (AW)	291.7	231.5	-20.6	0.590	0.663	0.072
4a. Augmented Wealth with DCEMP (AW*)	291.7	257.6	-11.7	0.590	0.652	0.061
4b. Net Augmented Wealth (AW)	280.4	213.7	-23.8	0.591	0.664	0.073
4c. Net Augmented Wealth with DCEMP (AW*)	280.4	237.3	-15.4	0.591	0.652	0.061
<b>C. Ages 47-64</b>						
1. Net Worth (NW)	108.3	137.6	27.0	0.761	0.798	0.037
2. Private Accumulations (PA)	204.4	199.9	-2.2	0.681	0.762	0.081
2a. Private Accumulations with DCEMP (PA*)	204.4	224.1	9.6	0.681	0.752	0.072
2b. Net Private Accumulations (PA)	192.1	182.3	-5.1	0.683	0.765	0.082
2c. Net Private Accumulations with DCEMP (PA*)	192.1	208.0	8.3	0.683	0.781	0.098
3. Augmented Wealth (AW)	376.4	389.2	3.4	0.558	0.640	0.082
3a. Augmented Wealth with DCEMP (AW*)	376.4	413.1	9.7	0.558	0.637	0.079
3b. Net Augmented Wealth (AW)	349.9	358.6	2.5	0.558	0.642	0.083
3c. Net Augmented Wealth with DCEMP (AW*)	349.9	374.1	6.9	0.558	0.652	0.093
Note: own computations from the 1983, 1989, and 2001 Surveys of Consumer Finances.						
Households are classified into age groups by the age of the head of household. Key:						
Private Accumulations PA = NWX+PW.				Retirement Wealth RW = PW + SSW.		
Augmented Wealth AW = NWX + PW + SSW.						
a. Source for household income data: U.S. Census of the Bureau, Current Population Surveys, available on the internet: <a href="http://www.census.gov/hhes/income/histinc/ineqtoc.html">http://www.census.gov/hhes/income/histinc/ineqtoc.html</a> .						

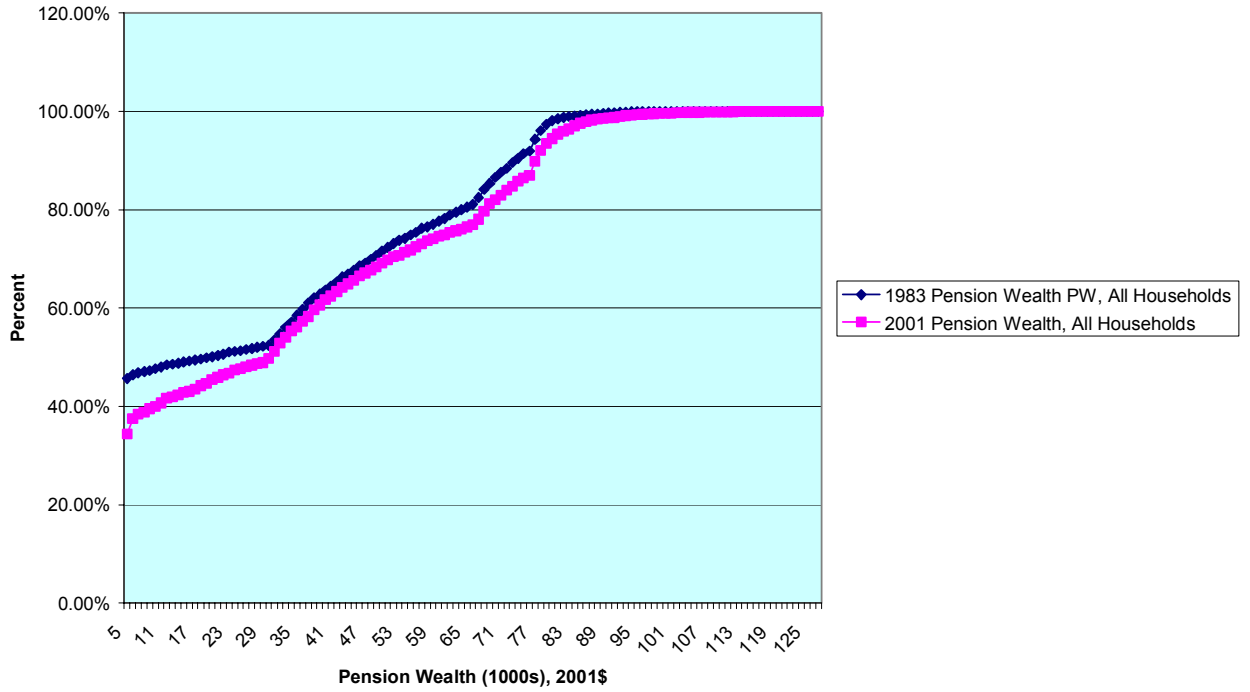
**Figure A. Wealth Inequality, Income Inequality, and the Ratio of Stock Prices to House Prices, 1922-1998**



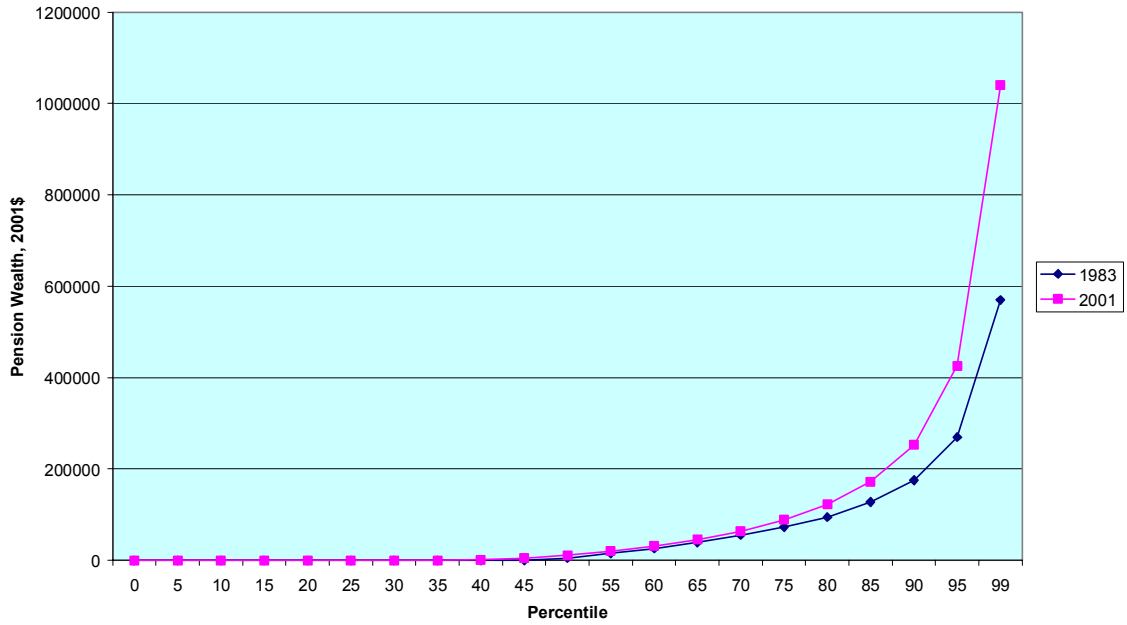
**Figure B. Predicted versus Actual Wealth Inequality, 1922-1998**



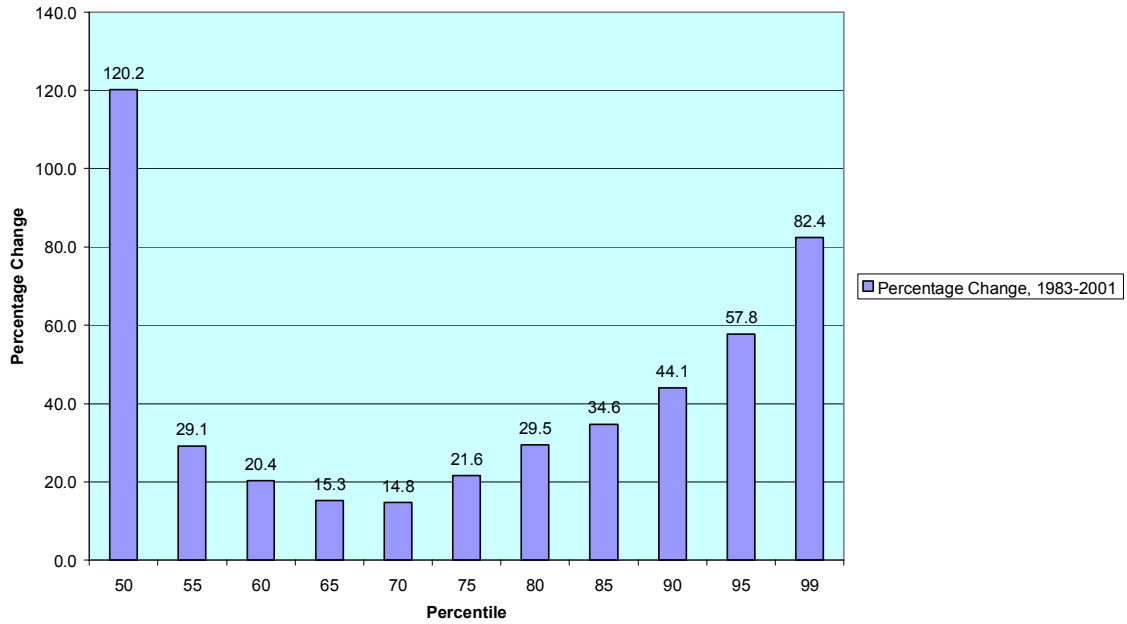
**Figure 1. The cumulative Distribution of Households by Pension Wealth Level, 1983 and 2001**



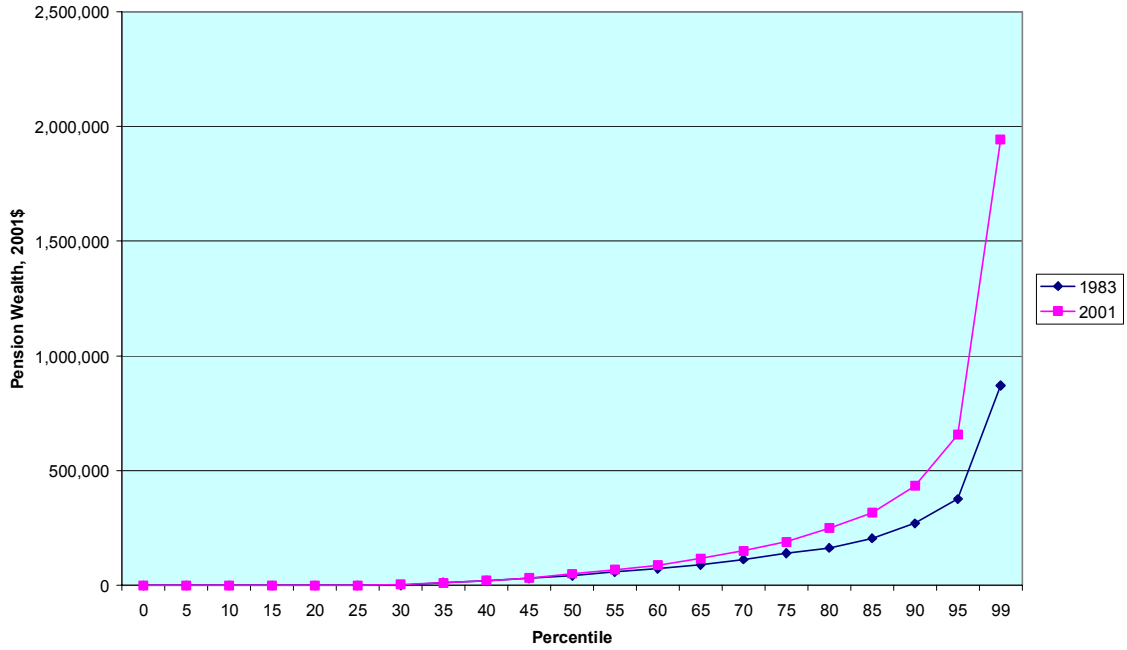
**Figure 2a. Pension Wealth PW in 2001 Dollars by Pension Wealth Percentile, All Households, 1983 and 2001**



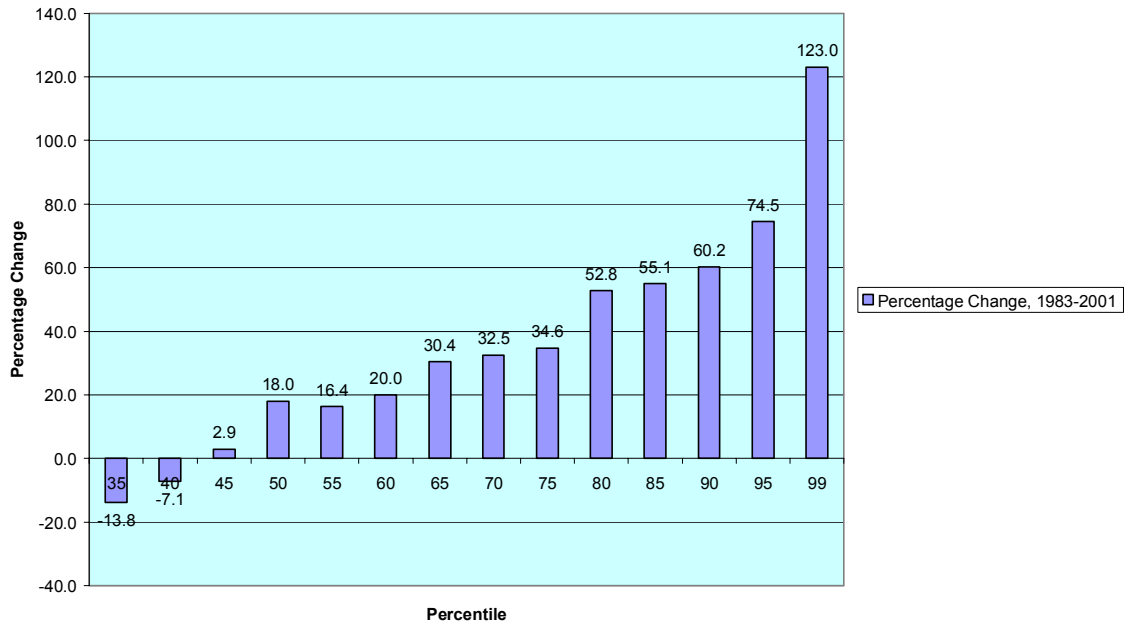
**Figure 2b. Percentage Change in Pension Wealth PW in 2001 Dollars by Pension Wealth Percentile, All Households, 1983 and 2001**



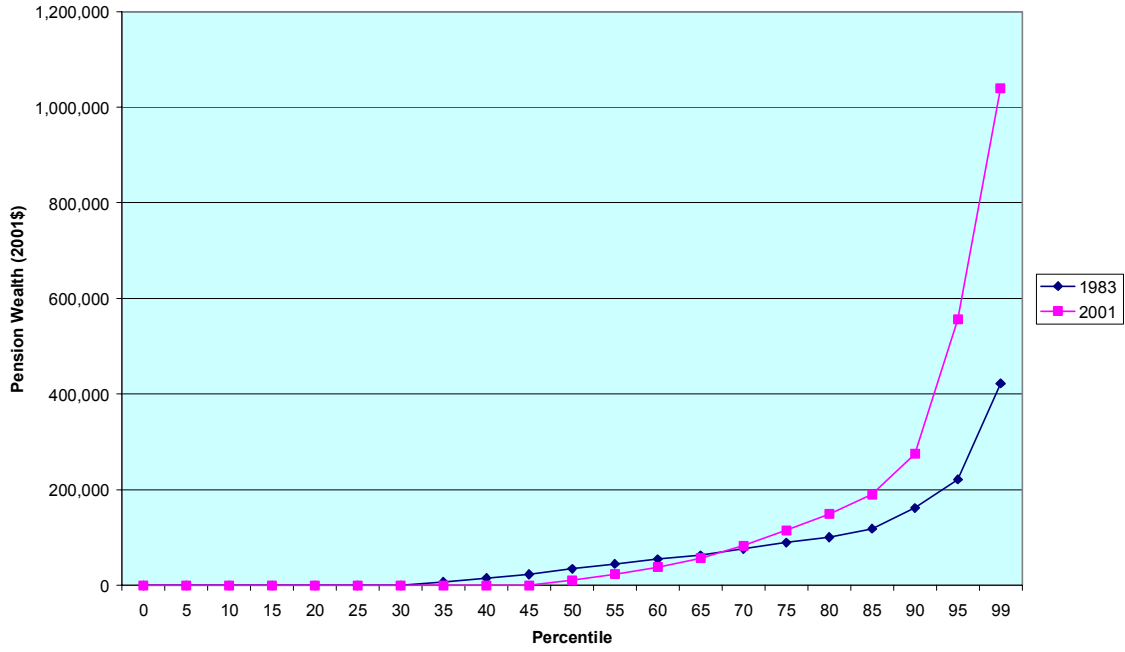
**Figure 2c. Pension Wealth PW in 2001 Dollars by Pension Wealth Percentile, Ages 47-64, 1983 and 2001**



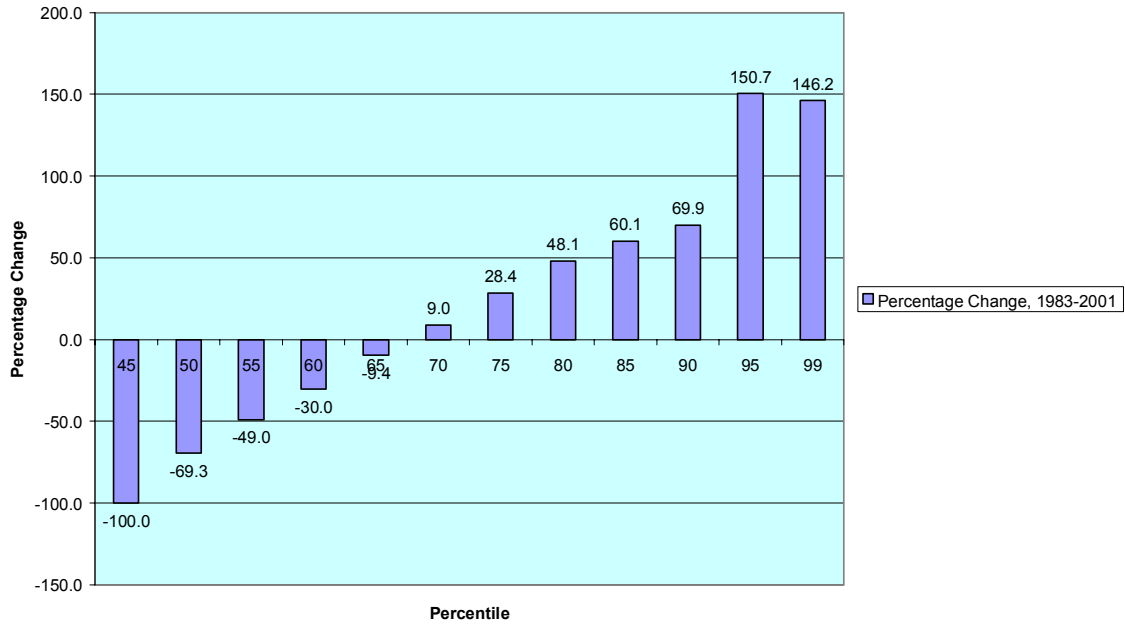
**Figure 2d. Percentage Change in Pension Wealth PW in 2001 Dollars by Pension Wealth Percentile, Ages 47-64, 1983 and 2001**



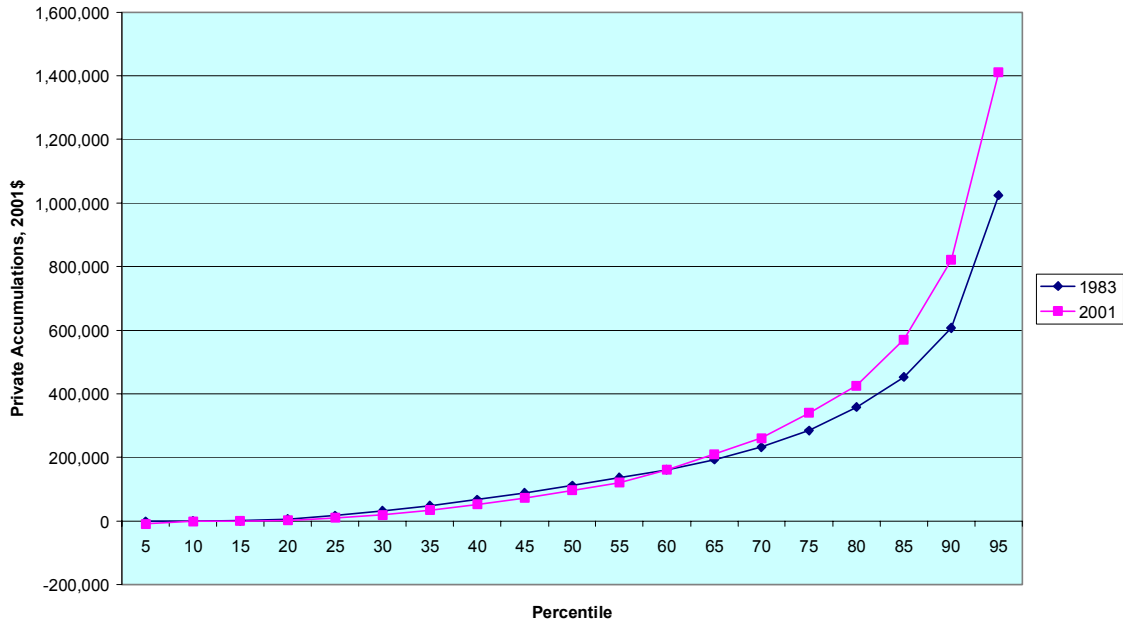
**Figure 2e. Pension Wealth PW in 2001 Dollars by Pension Wealth Percentile, Ages 65 and Over, 1983 and 2001**



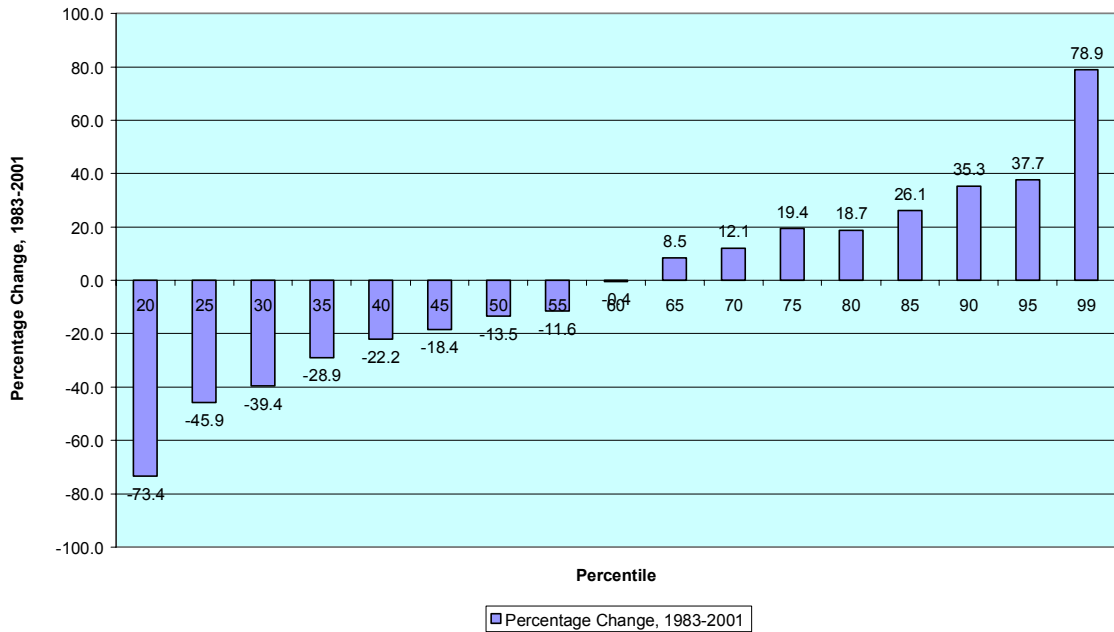
**Figure 2f. Percentage Change in Pension Wealth PW in 2001 Dollars by Pension Wealth Percentile, Ages 65 and over, 1983 and 2001**



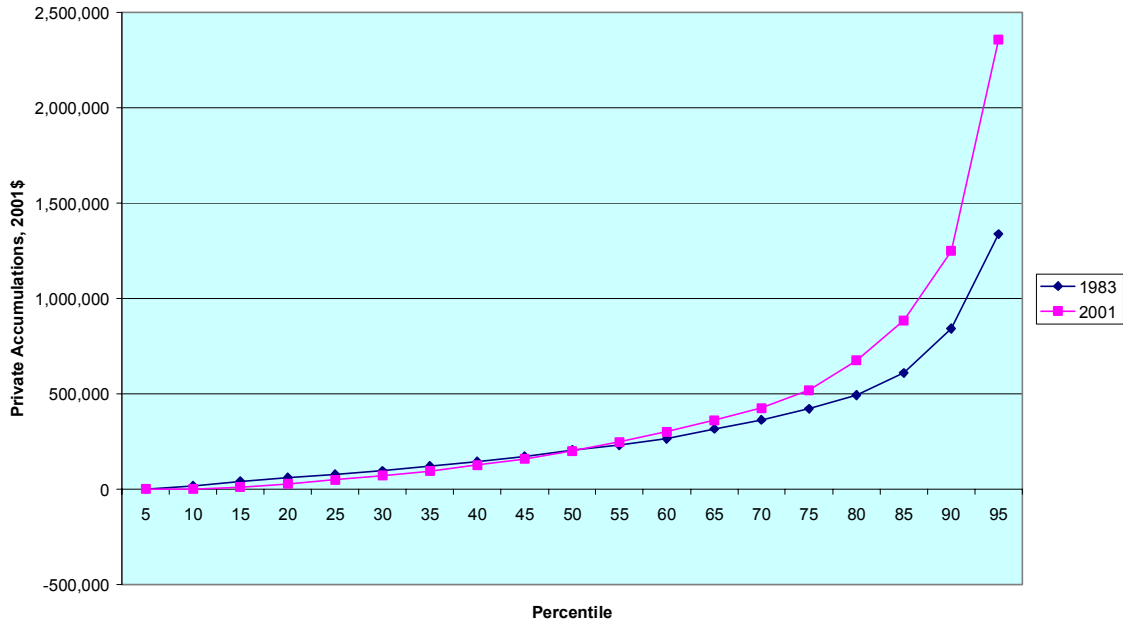
**Figure 3a. Private Accumulations (PA) in 2001 Dollars by PA Percentile, All Households, 1983 and 2001**



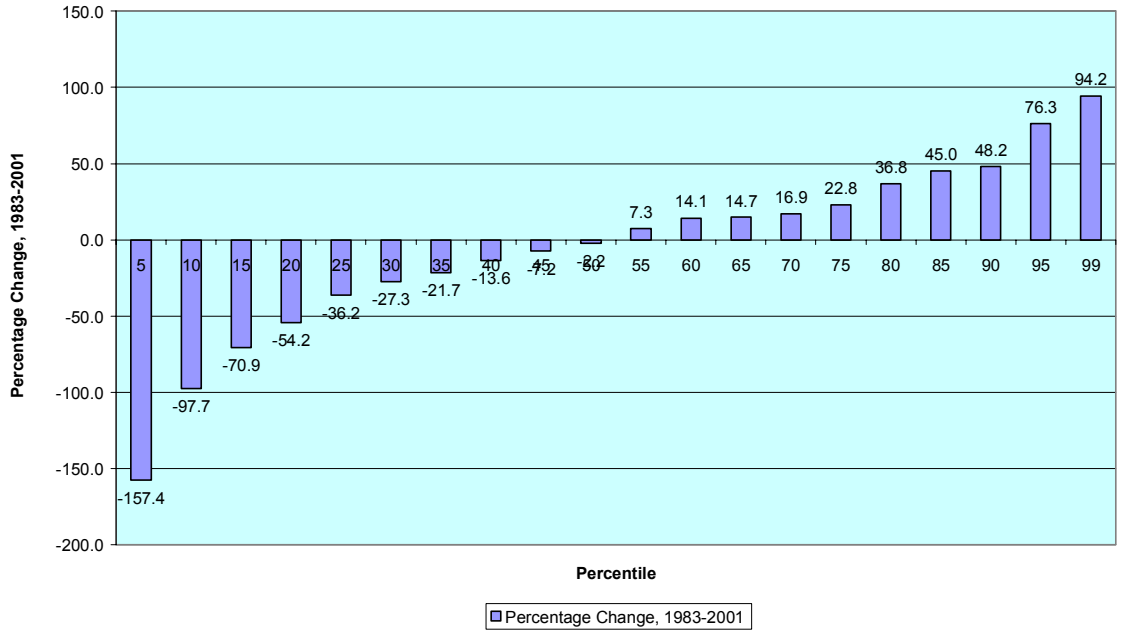
**Figure 3b. Percentage Growth of Private Accumulations (in 2001 Dollars), All Households, by Percentile, 1983 to 2001**



**Figure 3c. Private Accumulations (PA) in 2001 Dollars by PA Percentile, Ages 47-64, 1983 and 2001**

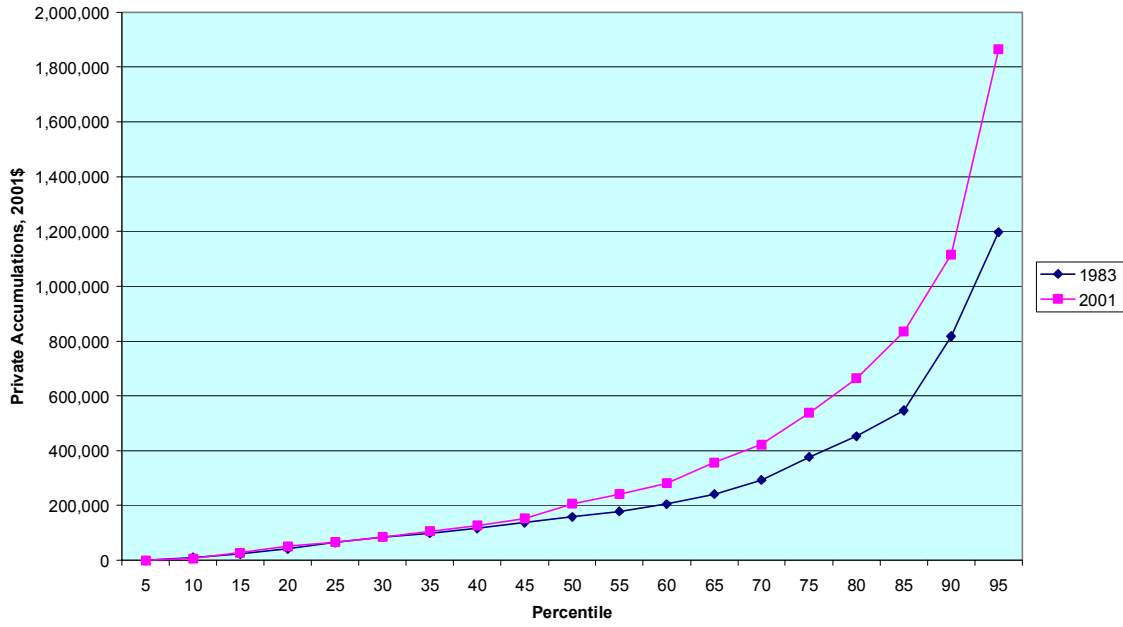


**Figure 3d. Percentage Growth of Private Accumulations (in 2001 Dollars), Ages 47-64, By Percentile, 1983 to 2001**

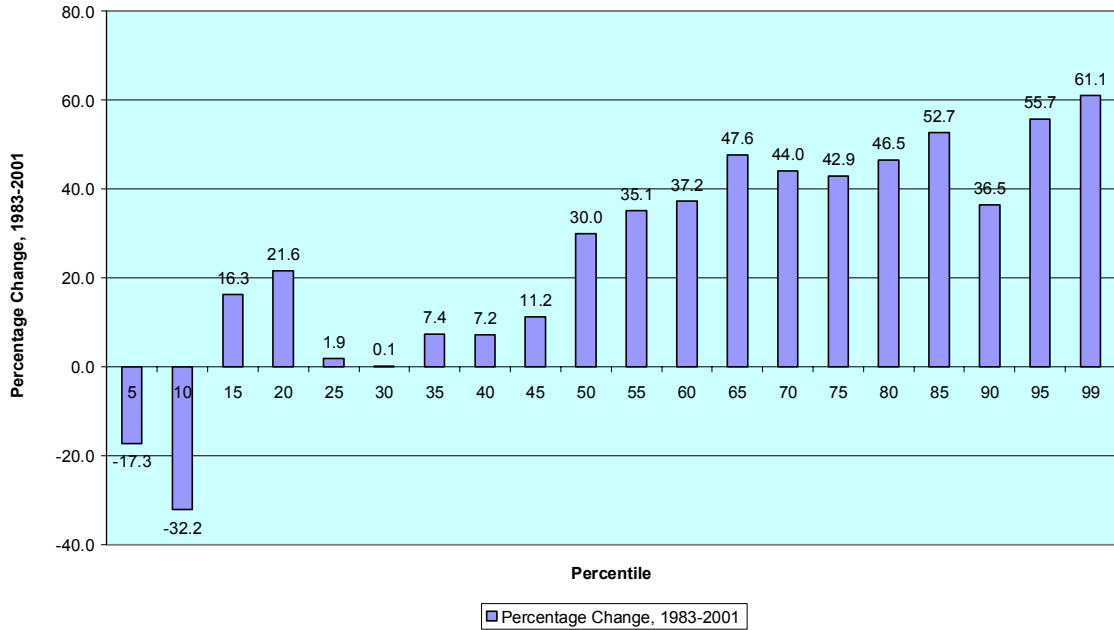




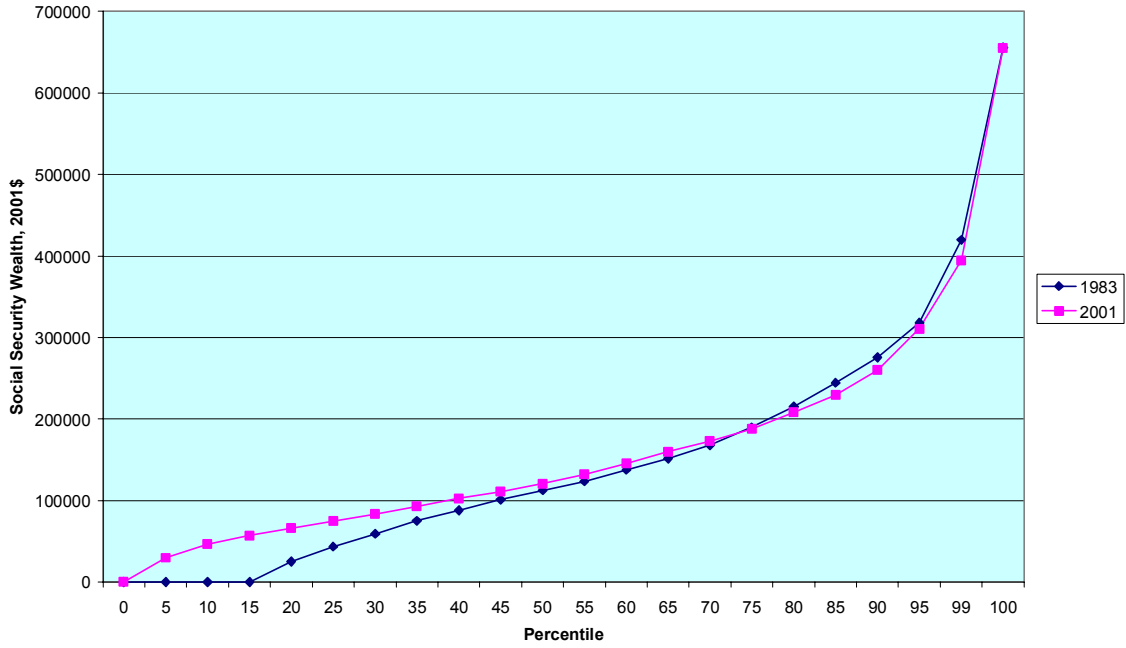
**Figure 3e. Private Accumulations (PA) in 2001 Dollars by PA Percentile, Ages 65 and Older, 1983 and 2001**



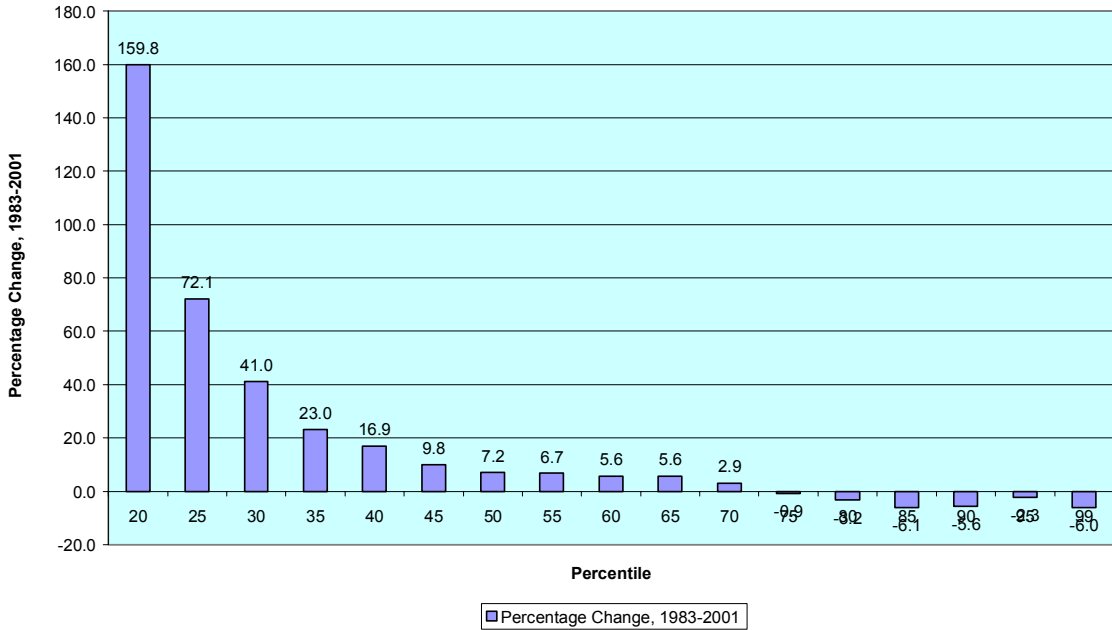
**Figure 3f. Percentage Growth of Private Accumulations (in 2001 Dollars), Ages 65 and Older, By Percentile, 1983 to 2001**



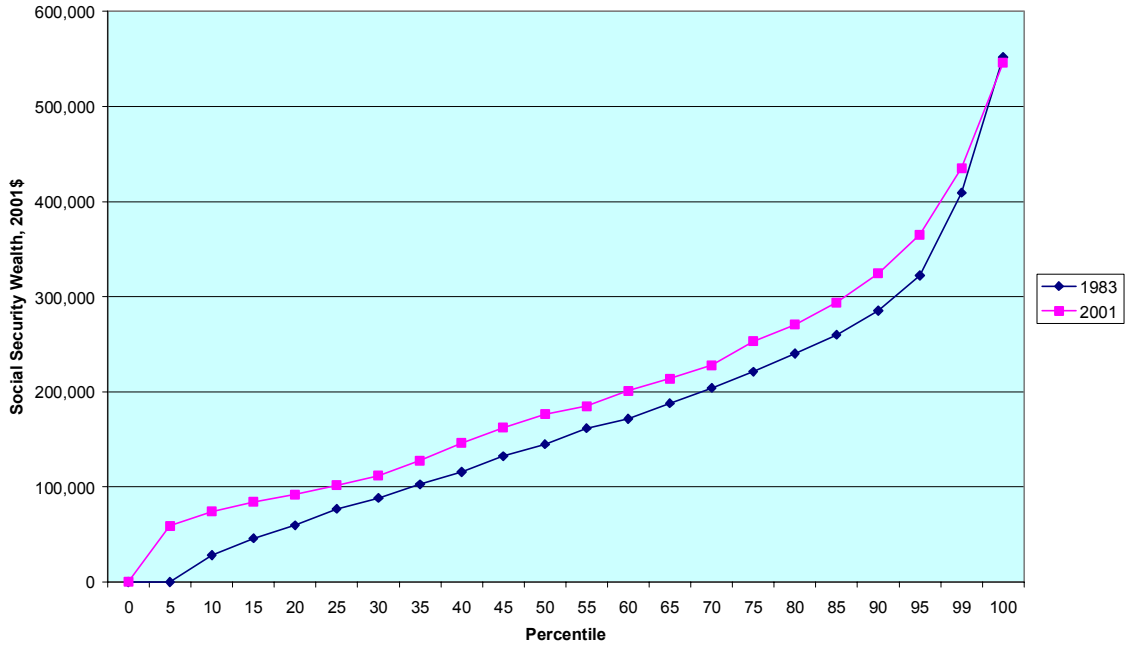
**Figure 4a. Social Security Wealth SSW in 2001 Dollars by Social Security Wealth Percentile, All Households, 1983 and 2001**



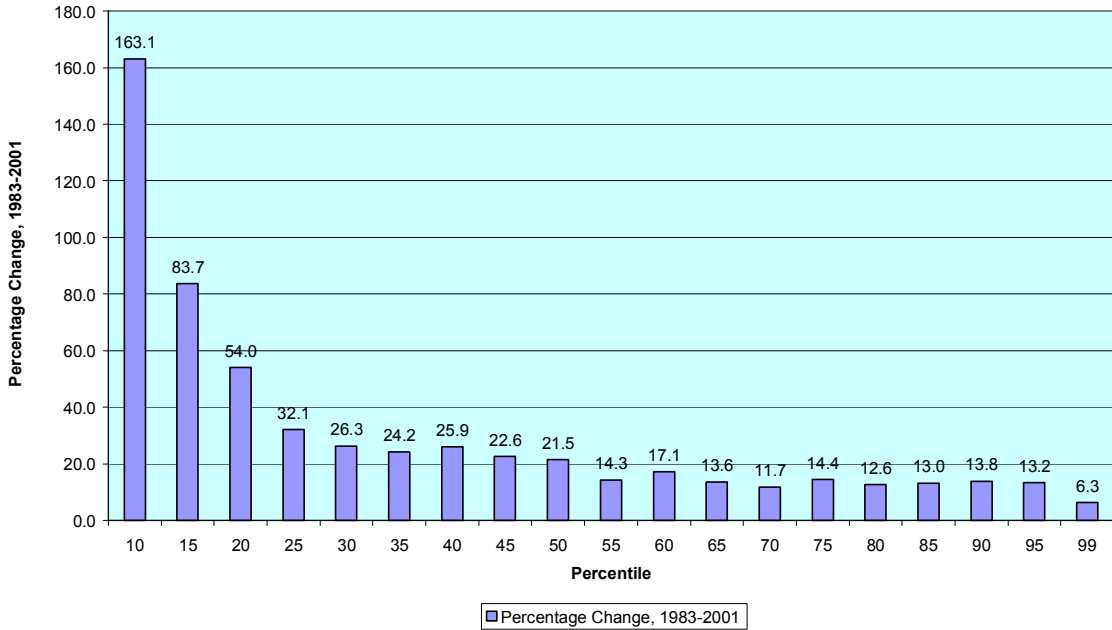
**Figure 4b. Percentage Growth of Social Security Wealth (in 2001 Dollars), All Households, by Percentile, 1983 to 2001**



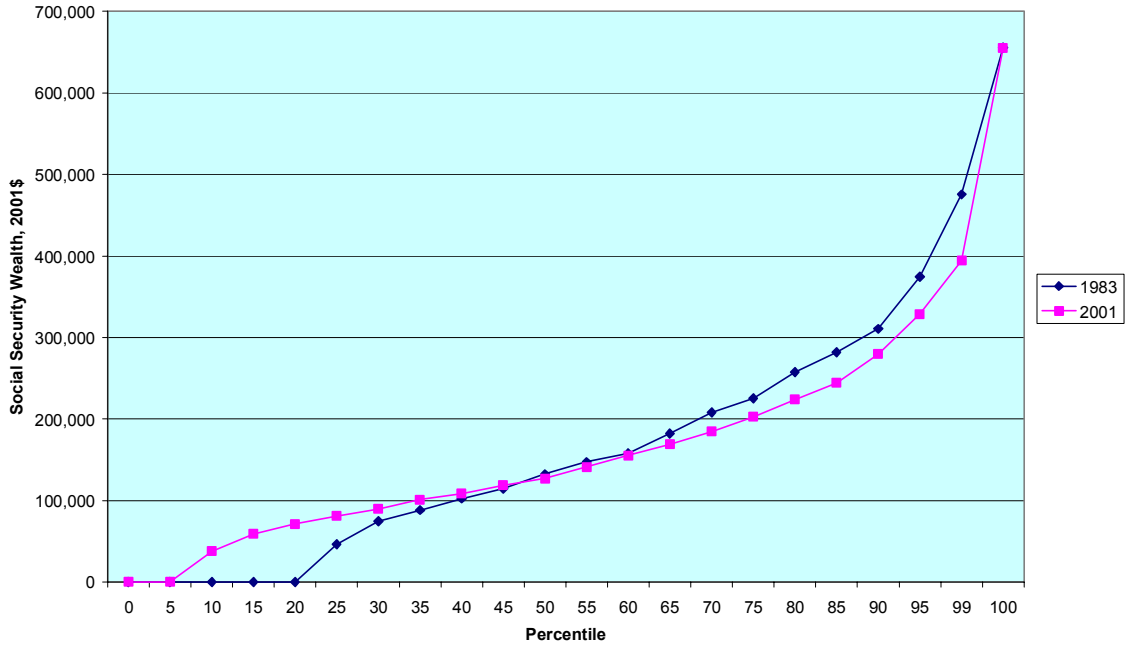
**Figure 4c. Social Security Wealth SSW in 2001 Dollars by Social Security Wealth Percentile, Ages 47-64, 1983 and 2001**



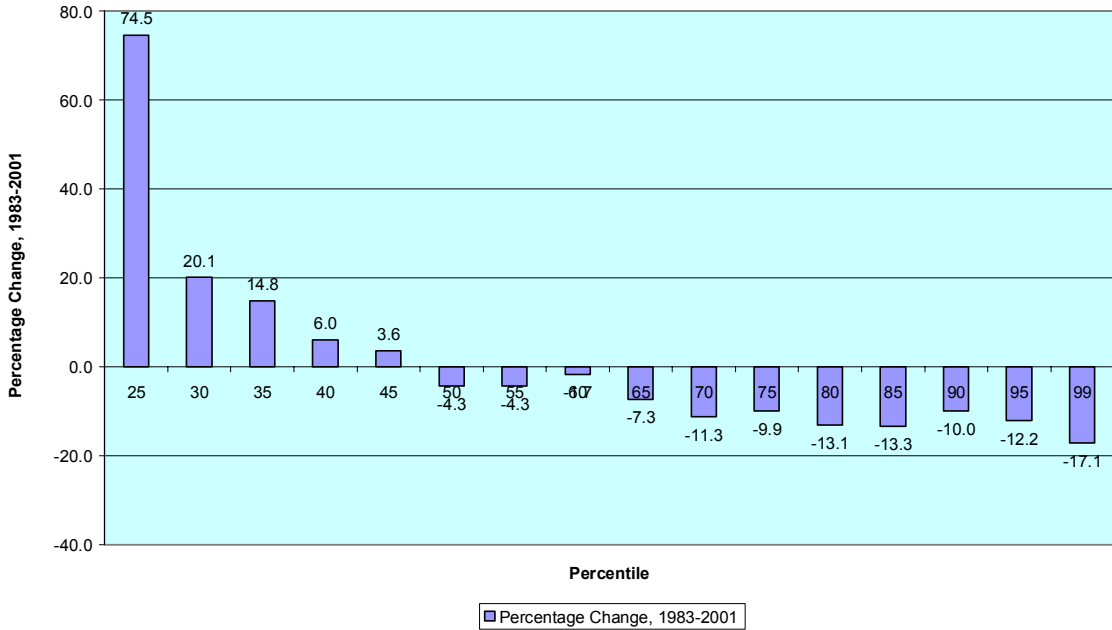
**Figure 4d. Percentage Growth of Social Security Wealth (in 2001 Dollars), Ages 47 to 64, by Percentile, 1983 to 2001**



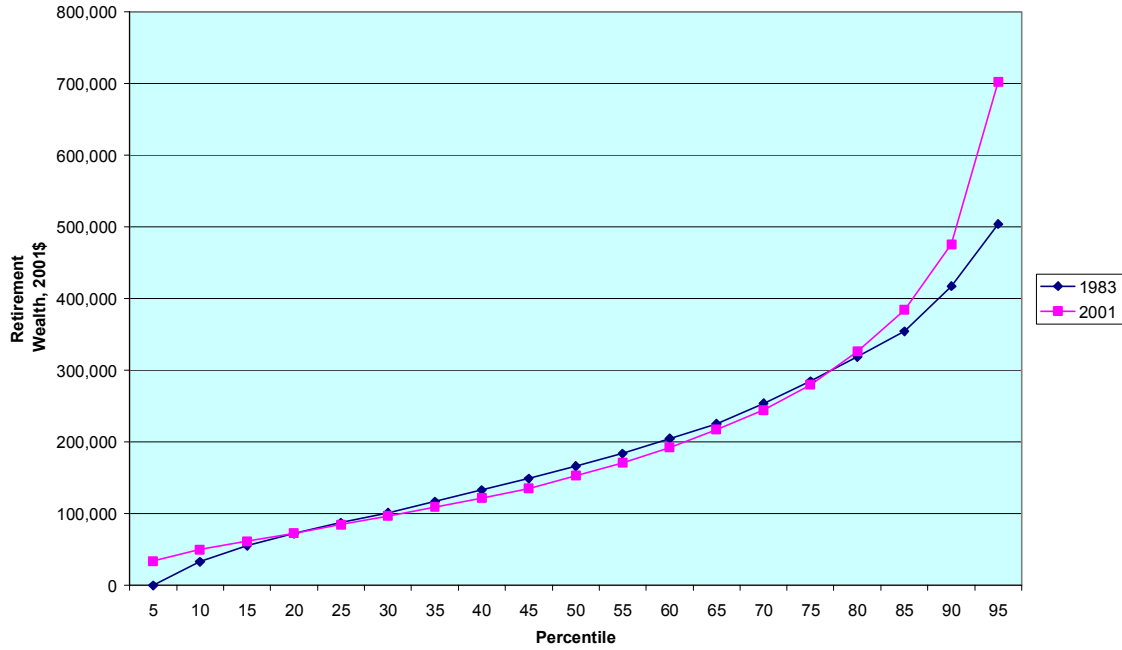
**Figure 4e. Social Security Wealth SSW in 2001 Dollars by Social Security Wealth Percentile, Ages 65 and over, 1983 and 2001**



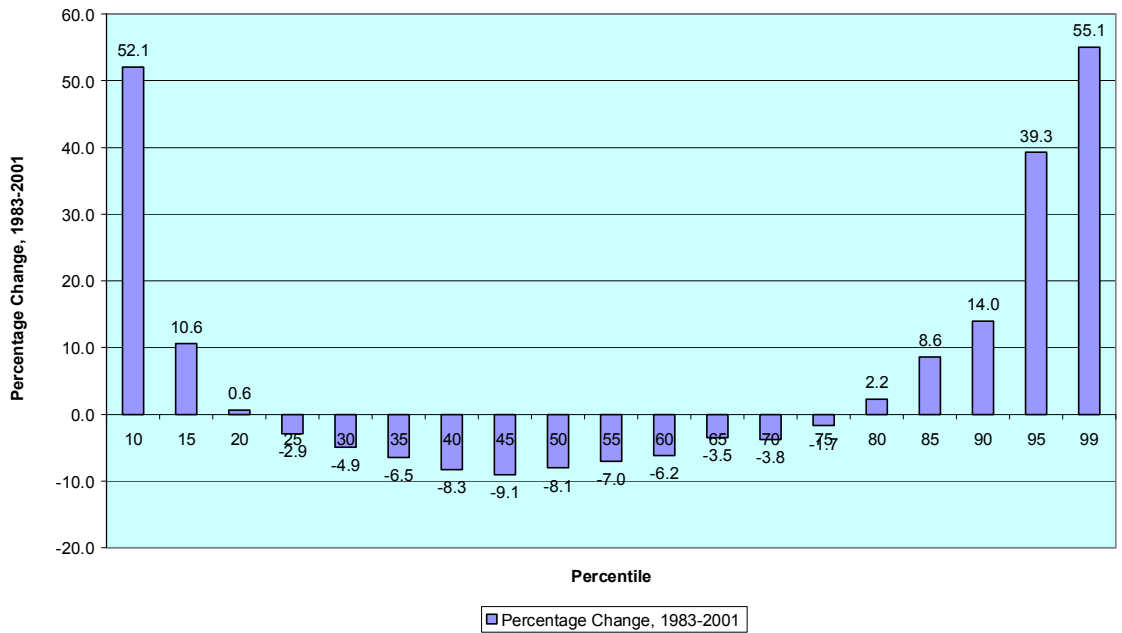
**Figure 4f. Percentage Growth of Social Security Wealth (in 2001 Dollars), Ages 65 and older, by Percentile, 1983 to 2001**



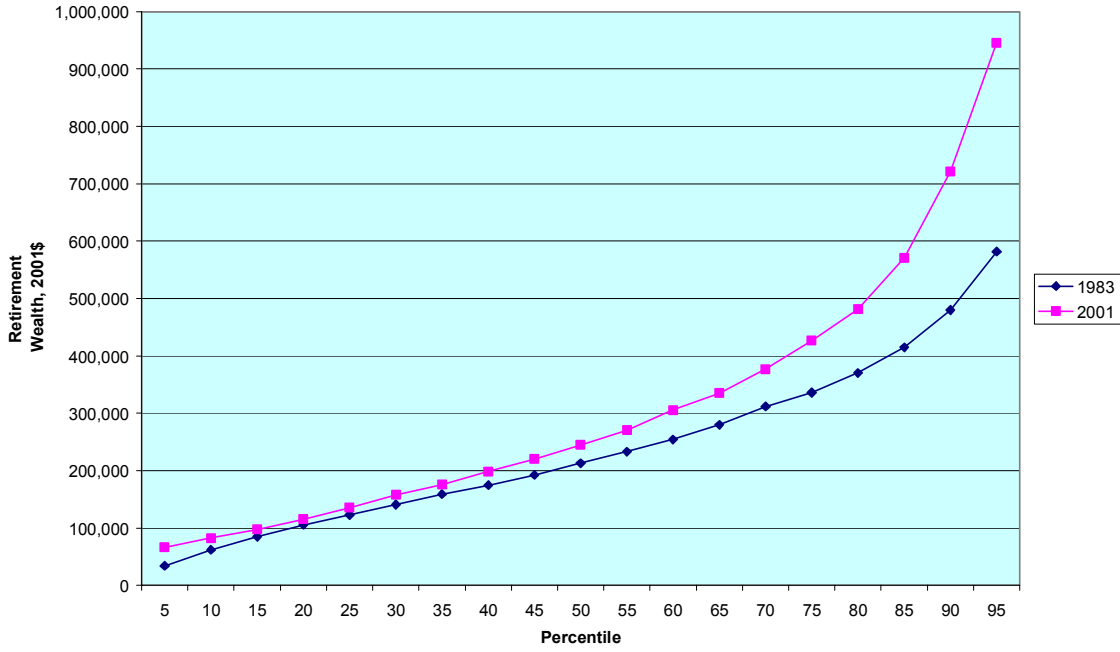
**Figure 5a. Retirement Wealth in 2001 Dollars by Retirement Wealth Percentile, All Households, 1983 and 2001**



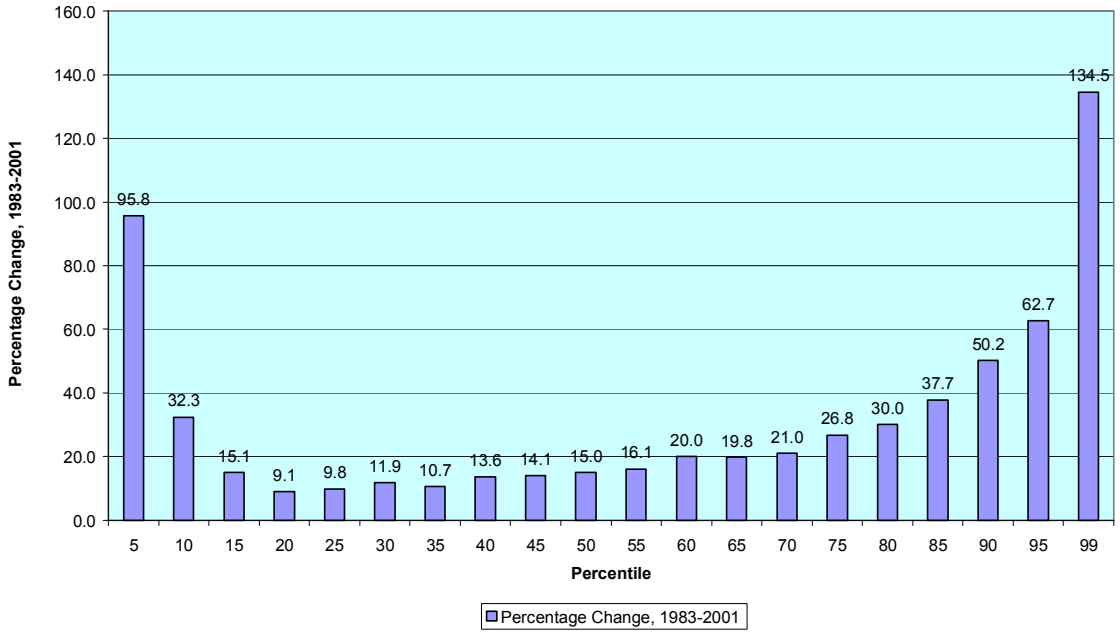
**Figure 5b. Percentage Growth of Retirement (in 2001 Dollars), All Households, by Percentile, 1983 to 2001**



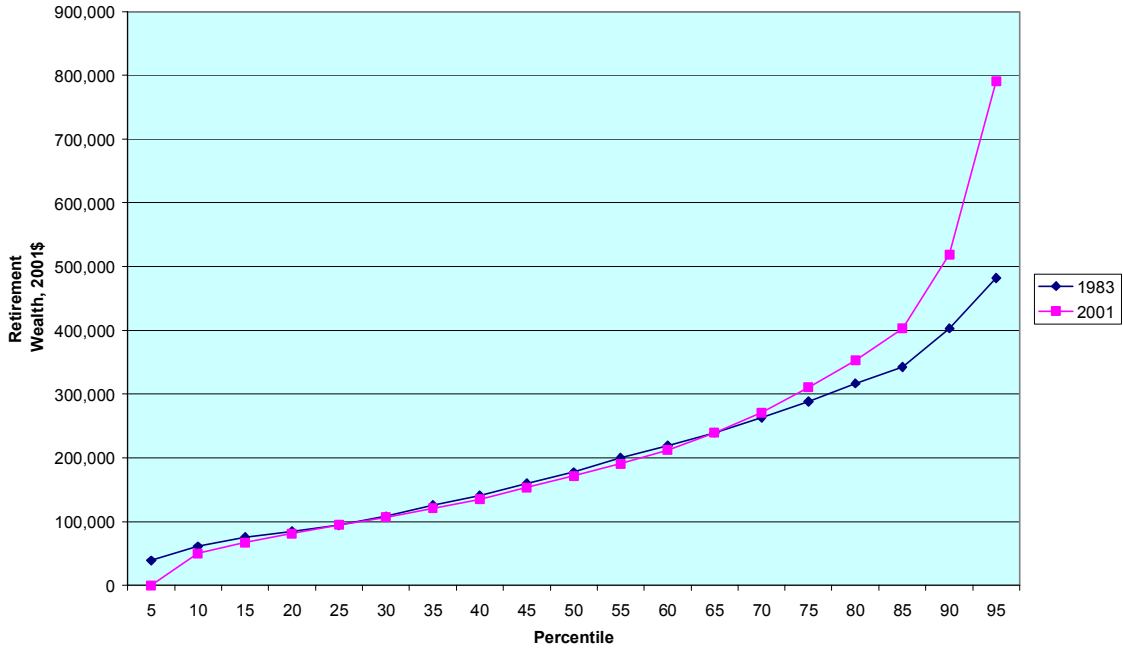
**Figure 5c. Retirement Wealth in 2001 Dollars by Retirement Wealth Percentile, Ages 47-64, 1983 and 2001**



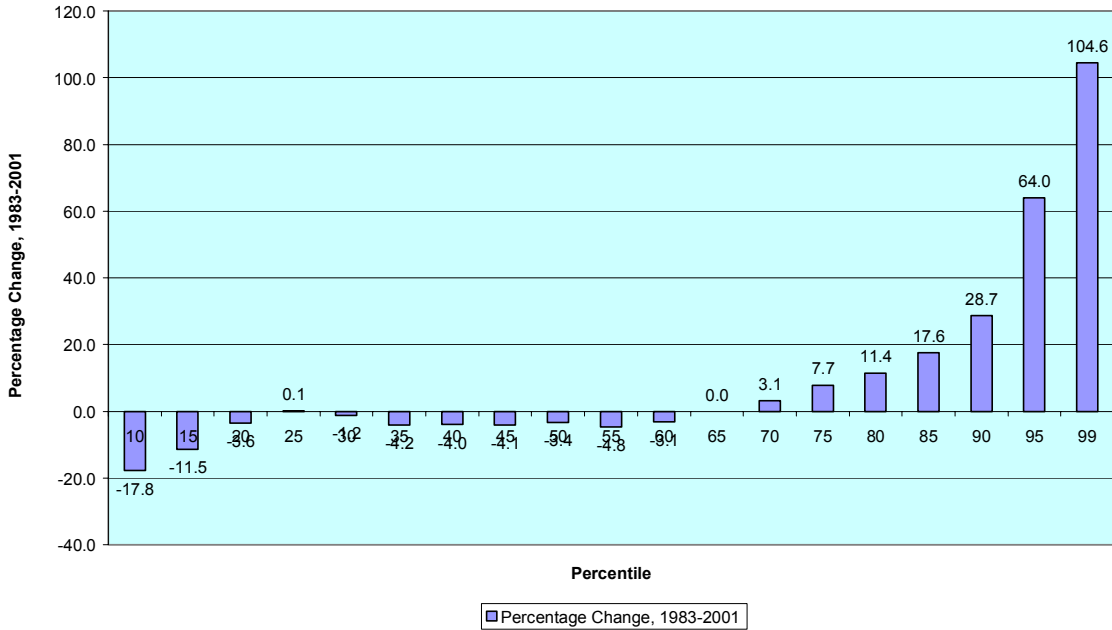
**Figure 5d. Percentage Growth of Retirement (in 2001 Dollars), Ages 47 to 64, by Percentile, 1983 to 2001**



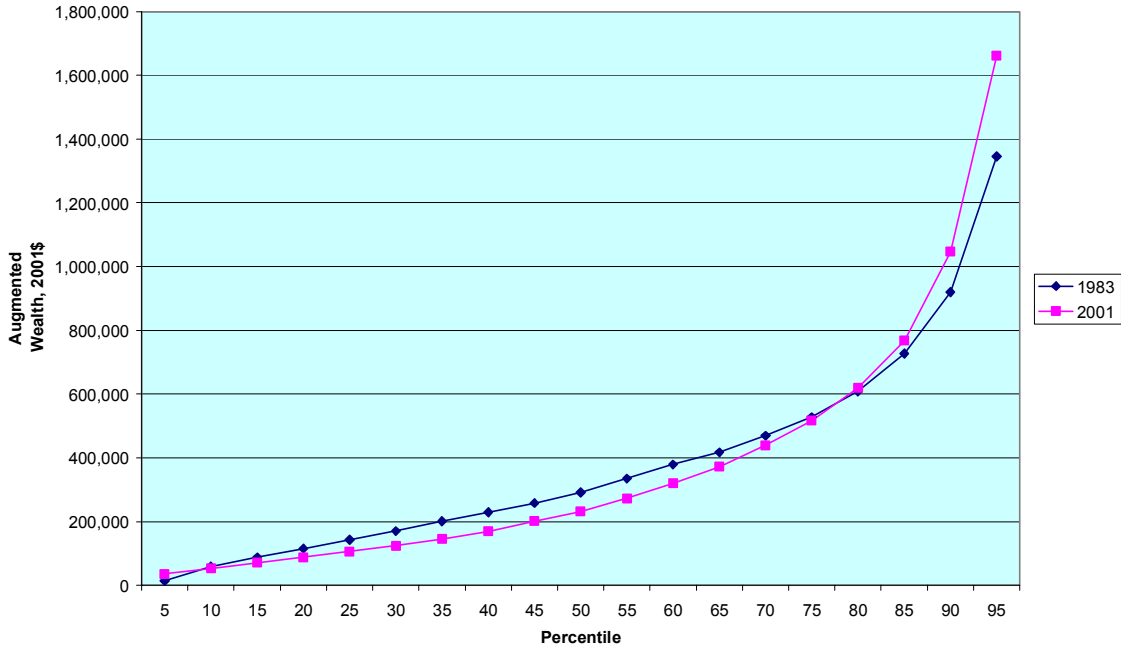
**Figure 5e. Retirement Wealth in 2001 Dollars by Retirement Wealth Percentile, Ages 65 and over, 1983 and 2001**



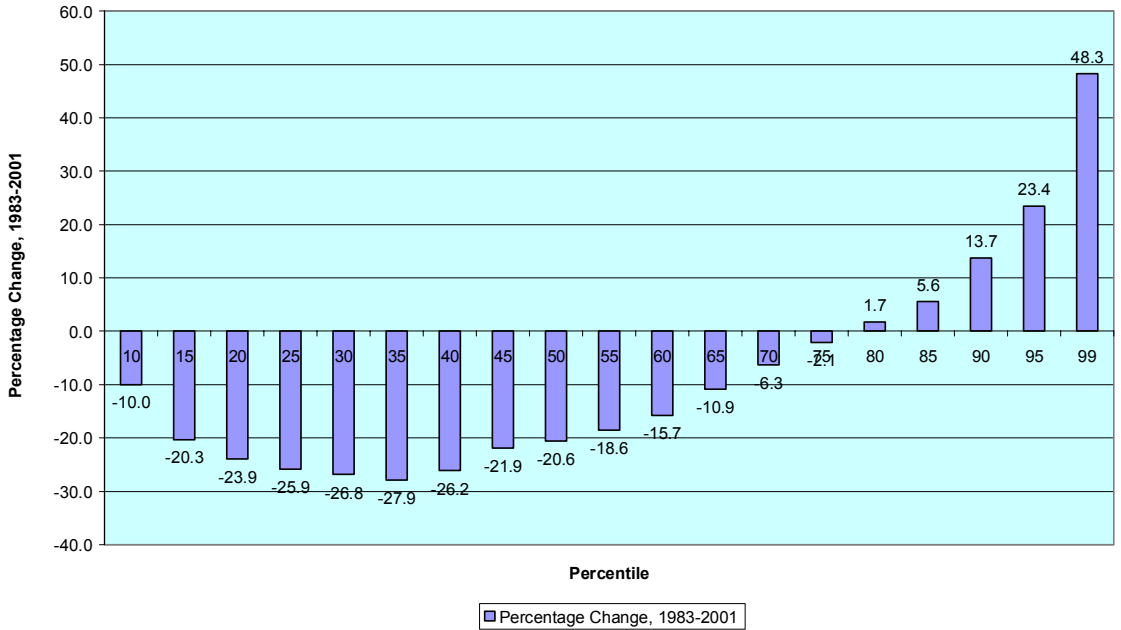
**Figure 5f. Percentage Growth of Retirement (in 2001 Dollars), Ages 65 and over, by Percentile, 1983 to 2001**



**Figure 6a. Augmented Wealth in 2001 Dollars by Augmented Wealth Percentile, All Households, 1983 and 2001**

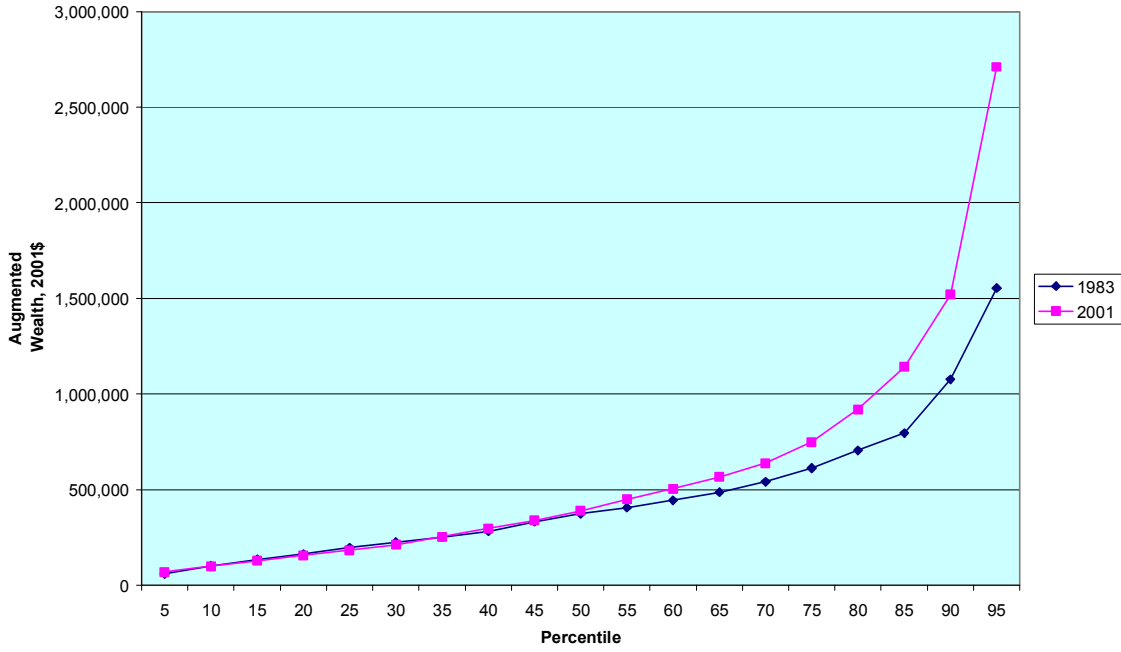


**Figure 6b. Percentage Growth of Augmented Wealth (in 2001 Dollars), All Households, by Percentile, 1983 to 2001**

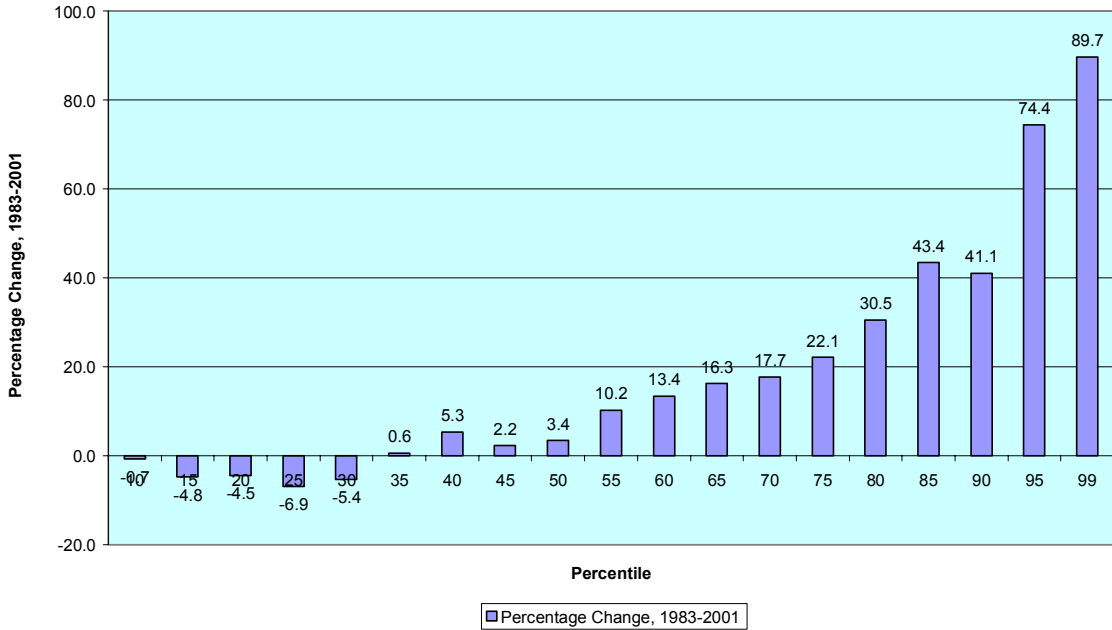




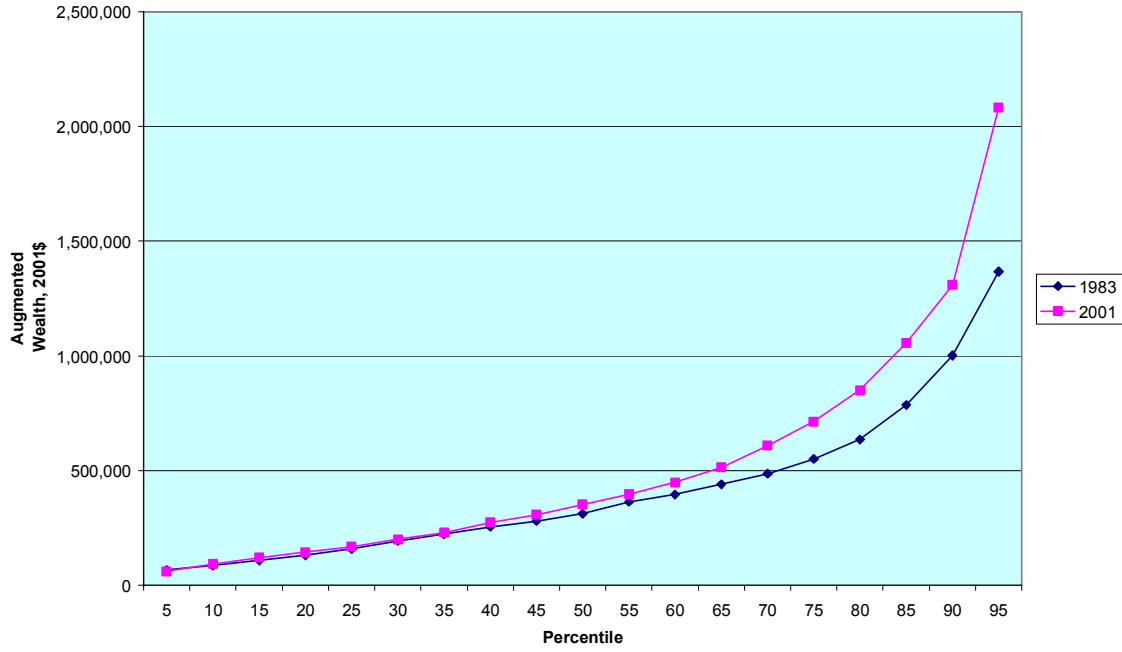
**Figure 6c. Augmented Wealth in 2001 Dollars by Augmented Wealth Percentile, Ages 47-64, 1983 and 2001**



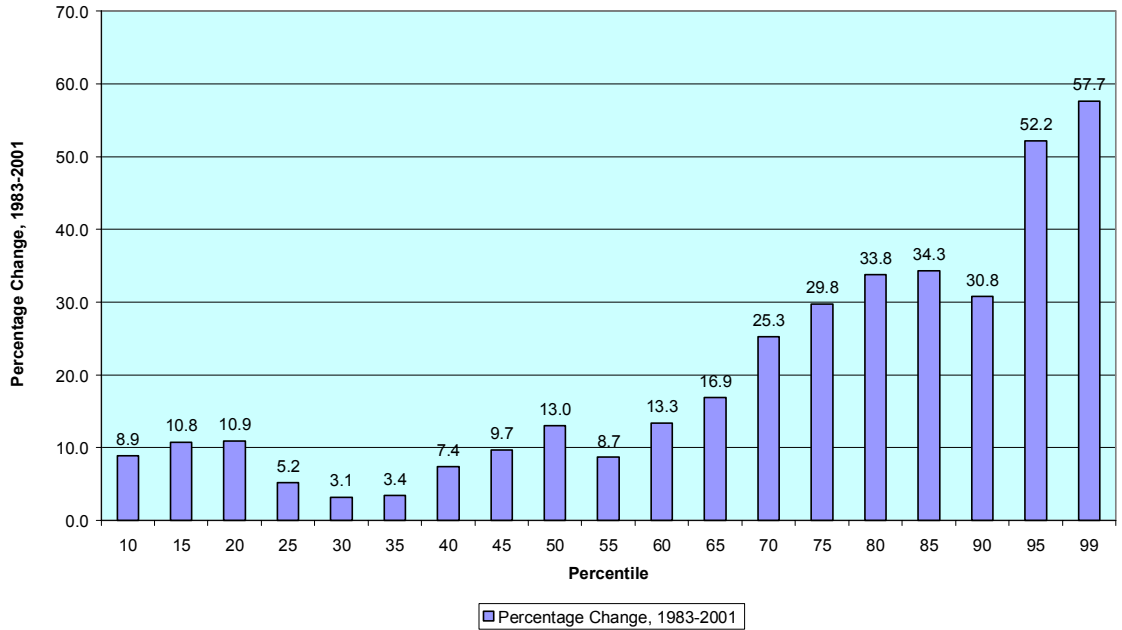
**Figure 6d. Percentage Growth of Augmented Wealth (in 2001 Dollars), Ages 47-64, by Percentile, 1983 to 2001**



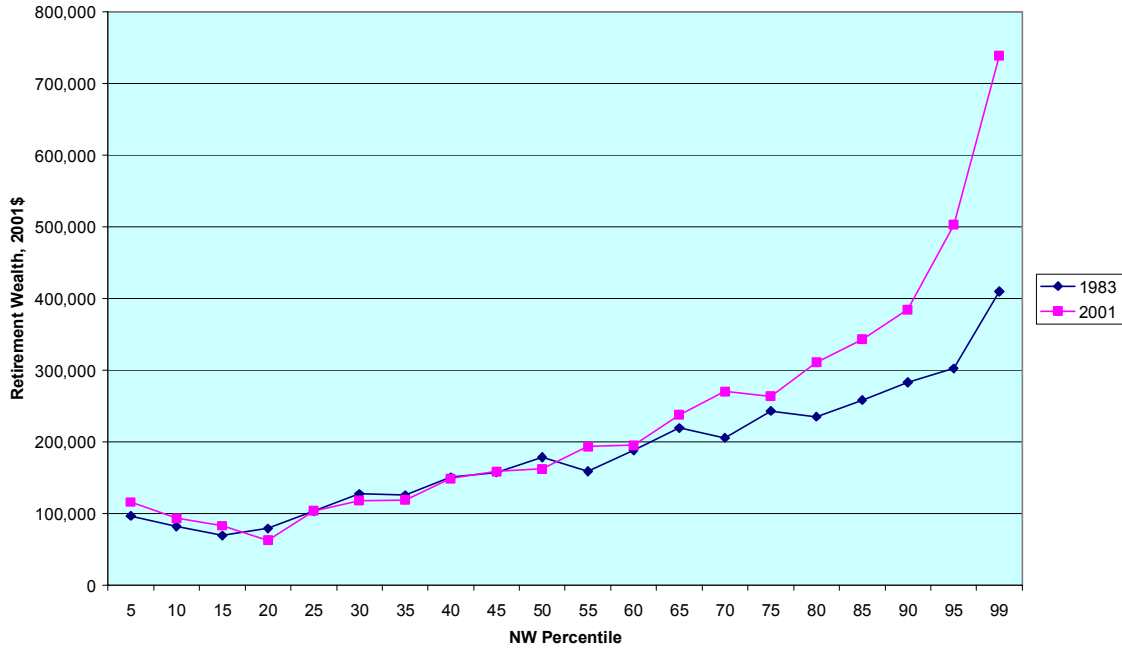
**Figure 6e. Augmented Wealth in 2001 Dollars by Augmented Wealth Percentile, Ages 65 and over, 1983 and 2001**



**Figure 6f. Percentage Growth of Augmented Wealth (in 2001 Dollars), Ages 65 and over, by Percentile, 1983 to 2001**



**Figure 7a. Retirement Wealth RW in 2001 Dollars by Net Worth (NW) Percentile, All Households, 1983 and 2001**



**Figure 7b. Percentage Growth of Retirement Wealth (in 2001 Dollars), All Households, by Net Worth Percentile, 1983 to 2001**

